

1 Ocean warming drives abrupt declines in fish productivity at 2 global scale

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10 **Abstract**

11 Marine life is under multiple pressures, including climate change and overfishing. Environmental change and
12 variability can threaten fishery sustainability, especially when it results in large, abrupt and persistent shifts in
13 productivity of fish stocks. Reports of abrupt shifts in marine systems are not uncommon, but a global
14 assessment of their occurrence and drivers is seriously lacking. Here, we systematically classified the
15 temporal dynamics of fish stock productivity from agency-assessed fisheries worldwide. Among the 315 fish
16 stocks with time series available, we detected at least one productivity abrupt shift for more than a quarter of
17 the stocks. Using an integrative modeling approach including life history, environmental conditions, and
18 fishing intensity variables, we showed that abrupt declines are over-represented in stocks where sea surface
19 temperature increases have been larger during the period covered by fish stock monitoring, while abrupt
20 increases are more likely under lower fishing intensity. We investigated the link between productivity abrupt
21 shifts and stock collapses. We found that abrupt declines in productivity preceded stock collapses by ten to
22 twenty years in 25% of the cases, suggesting that some major stock collapses could be anticipated if abrupt
23 shift were more systematically detected and examined. Overall, our results highlight the importance of
24 considering productivity abrupt shifts to prevent a pervasive risk of fish population collapse in warming
25 oceans.

26 **Significance Statement**

27 Using the largest global stock assessment dataset for marine fisheries, we found a large proportion of abrupt
28 shifts in the trajectories of productivity time series for 315 fish stocks. We evidenced that abrupt productivity
29 declines were over-represented in marine regions with highest warming rates. We also demonstrate that
30 abrupt productivity declines preceded stock collapses by about a decade in a quarter of the stocks that
31 shifted. Our results shed light on a likely warming-related timeline to fisheries collapse and call for more
32 systematic examination and early detection of abrupt shifts. This paper contributes to setting priorities in
33 managing stocks that would be more likely to undergo strong and persistent shifts in their productivity.

34 **Keywords:** abrupt shift, fisheries, nonlinear dynamics, stock assessment, warming temperatures

35 **Introduction**

36 Marine ecosystems are facing intense anthropogenic pressure (Georgian et al., 2022; Halpern et al., 2008),
37 putting the survival of marine species along with the food security and economies of many coastal human
38 populations at risk (FAO, 2022). Although the capture of wild fish from the oceans has been maintained at
39 high levels since its peak in the 1990s (Pauly and Zeller, 2016), it does not imply that all stocks are
40 sustainably exploited (FAO, 2022). Indeed, despite the overall success of some management strategies
41 (Frank and Oremus, 2023; Melnychuk et al., 2021), overexploitation and stock collapses are still major threats
42 (Georgian et al., 2022) with no guarantee that recent improvements in fishery sustainability (Hilborn et al.,
43 2020) can be maintained in the face of climate change.

44 One particular challenge for fisheries management is the ability to avoid abrupt, strong, and persistent
45 declines of exploited stocks, which undermine sustainability goals (King et al., 2015; Levin and Möllmann,
46 2015). Even though the existence of such so-called “regime shifts” in fisheries has long been documented
47 (deYoung et al., 2004; Hare and Mantua, 2000; Steele, 1998), such events have yet to be broadly integrated
48 within stock assessment and management (Conversi et al., 2015; Sguotti et al., 2022). Regime shifts are
49 usually considered at the ecosystem level involving the synchronous change in multiple variables from fish to
50 phytoplankton (deYoung et al., 2008). But it has been argued that abrupt shifts at the level of fish populations
51 could be indicative of putative regime shifts (Daskalov et al., 2007; Pedersen et al., 2020). In fisheries,
52 examples of once plentiful stocks that crashed to very low levels are not uncommon, notably with the case of
53 the Peruvian anchoveta (Idyll, 1973), North Atlantic cod (Myers et al., 1997), or Western Atlantic bluefin tuna
54 (Safina and Klinger, 2008). However, beyond these emblematic examples the extent to which regime shifts in
55 fisheries occurred in recent decades might be underestimated.

56 Accounting for regime shifts within fisheries management can be hindered by two major knowledge gaps.
57 First, they are hard to detect. The search for regime shifts in time series often relies on the sole use of
58 breakpoint detection algorithms without relevant alternative non-abrupt models (Spake et al., 2022), which
59 can limit the confidence in the conclusions (Rudnick and Davis, 2003). Second, regime shifts are overlooked
60 in most stock assessments, despite evidence that regime shift models are frequently a better descriptor of
61 stock dynamics (Vert-pre et al., 2013). Attempts to account for regime shifts and non-stationarity remain

63 scarce and insufficient to improve current practices (Levin and Möllmann, 2015; Sguotti et al., 2022). Instead,
64 the usual standard for evaluating the sustainability of management strategies relies on fixed thresholds –
65 related to biological reference points like minimal biomass or maximal fishing mortality – to trigger actions
66 (Walters and Martell, 2004), despite recent efforts to introduce more dynamic reference points (Berger, 2019;
67 Hodgdon et al., 2022).

68 Although the notions of regime shifts and stock collapses appear strongly connected, they are not equivalent.
69 Regime shifts are usually characterized by an abrupt temporal trajectory with a final state potentially
70 persistent in time (Bestelmeyer et al., 2011). In contrast, stock collapses relate to stock depletion based on
71 fixed arbitrary thresholds related to stock size or catch (Yletyinen et al., 2018) that give no indication about the
72 circumstances preceding the collapse nor the persistence of the collapse state. For instance, (Vert-pre et al.,
73 2013) showed that abrupt shifts from high to low surplus production levels – a proxy for stock productivity
74 corresponding to the change in abundance in the absence of fishing (Walters et al., 2008) – can happen quite
75 regularly and were unrelated to abundance levels for a substantial proportion of stocks. Although this finding
76 implies that abrupt shifts in productivity do not necessarily lead to stock collapse, the relationship between the
77 two phenomena has not been systematically explored. In addition to not fully understanding the
78 consequences of abrupt shifts in productivity, we also poorly understand what causes them. The relative
79 contributions of climate and exploitation to stock collapse has been well investigated (Möllmann and
80 Diekmann, 2012), but usually focused on a few data-rich stocks (Beaugrand et al., 2022; Pershing et al.,
81 2015). Broader syntheses of stock collapse mostly focus on a single pressure at a time, either fishing intensity
82 (Essington et al., 2015), climate change (Free et al., 2019), or life history (Pinsky et al., 2011) but rarely all
83 pressures together (Pinsky and Byler, 2015). Overall, we lack a global overview of the prevalence of abrupt
84 shifts in fisheries productivity, along with information on their drivers, and potential link with stock collapses.

85 The aim of this study is to identify potential regime shifts in fish stocks globally, by looking at abrupt shifts in
86 productivity. Here, we systematically classify the dynamics of 315 productivity time series, estimated from
87 assessments of marine fish stocks from around the world, to address the following questions: (i) How
88 prevalent are productivity abrupt shifts (PAS) and how are they distributed in space and time? (ii) Are PAS

89 related to species life history, environmental conditions or fishing pressure? (iii) Are PAS associated with
90 stock collapses?

91 **Results**

92 **Productivity Abrupt Shifts (PAS) prevalence and distribution**

93 We classified 315 fish stock productivity time series into basic trajectory types based on shape and trend and
94 found all types of trajectories over the period 1950–2020 (Fig. 1). Quality scores indicated that selected
95 models, especially for abrupt and quadratic trajectories, performed well and were robust to the removal of
96 individual data points (Fig. 1, Fig. S1). Still, the variance explained by the models remained low given the
97 intrinsically high variability of productivity time series (Fig. S1). The full list of trajectories is available in the
98 supplement (Table S1).

99 Globally, PAS were found for more than a quarter of stocks (25.7%, N=81), with roughly equal numbers of
100 negative (13.3%, N=42) and positive (12.4%, N=39) PAS (Fig. 2A). The other types of trajectories (quadratic,
101 linear, and no change) were found in similar proportions. Balanced proportions between different trajectory
102 types were also found at the scale of FAO major fishing areas (Fig. 2B), with no significant difference across
103 areas (Chi-square test, $p = 0.32$). The direction of shifts was however unevenly distributed in space (Chi-
104 square test, $p = 0.02$), with regions like the North-West Atlantic and North-West Pacific comprising a higher
105 proportion of negative PAS (30% and 29% respectively), whereas the region with most positive PAS (32%)
106 was the South-West Pacific. Similar patterns were found using large marine ecosystems (LMEs) as grouping
107 areas (Fig. S2). Considering taxonomic groups, we found significant differences in the proportions of PAS
108 against non-abrupt trajectories across the five most numerous orders in number of stocks (Chi-square test,
109 $p = 0.01$), with notably a higher proportion of positive PAS in Perciformes and of negative PAS in
110 Clupeiformes (Fig. S3).

111 Over time, the occurrence of positive PAS were spread from the 1960's to the 2010's and the distribution did
112 not differ significantly from the coverage of time series (Fig. 2C, Kolmogorov-Smirnov test, $p = 0.32$).
113 However, the distribution of negative PAS differed from coverage (Kolmogorov-Smirnov test, $p = 0.04$) and
114 tended to cluster during the 1980's (Fig. 2D), before the maximum of time series available was reached

115 between 1994 and 2000. The observation of PAS being less frequent after 2010 could arise from a lack of
116 available data along with the difficulty of detecting very recent shifts.

117 **Drivers of Productivity Abrupt Shifts (PAS)**

118 We tested which drivers among life history, environment (sea surface temperature, SST), and fishing intensity
119 (exploitation rate, ER) were more related to the occurrence of PAS using generalized additive mixed models
120 to account for phylogeny and spatial location. We found that the main drivers of PAS differed between
121 negative and positive PAS. Negative PAS were mostly influenced by environmental conditions like positive
122 trends in SST occurring before the shift ($p = 0.005$, Fig. 3A) and marginally by lower average SST ($p = 0.054$,
123 Fig. 3A) compared to other trajectories. Positive PAS were influenced by fishing intensity and life history and
124 were associated with higher trends in ER ($p = 0.005$, Fig. 3B), higher age at maturity ($p = 0.048$, Fig. 3B), and
125 marginally lower trends in SST ($p = 0.054$, Fig. 3B).

126 Clearly, as both models for negative and positive productivity abrupt shifts explained between 13% and 14%
127 of the total variance, the occurrence of PAS was only partly explained by the variables selected. But the
128 identity and relative importance of those variable is not a matter of model structure. Indeed, analysis of PAS
129 using an entirely different method (hierarchical partitioning) identified the same significant predictors (Table
130 S2–S3).

131 **Productivity Abrupt Shifts PAS and stock collapse**

132 We also tested the extent to which negative PAS were associated with stock collapse defined as being below
133 25% of the average stock biomass recorded to date following (Essington et al., 2015). While not all negative
134 PAS led to stock collapse (only 11 out of 42 stocks did, 26%, Fig. 4A), among the stocks that did collapse, we
135 found a higher proportion of negative PAS (23%, Fig. 4A) compared to those that did not collapse (12%, Fig.
136 4A). Collapsed stocks also tended to have more decreasing and fewer increasing productivity trajectories
137 (Fig. 4A). On average, collapsed stocks experienced a stronger magnitude of negative PAS than stocks that
138 did not collapse (t-test $p = 0.06$, Fig. 4B). Most of the negative PAS occurred between 4 to 12 years before
139 the stock collapsed, while no such temporal lags were found for positive PAS (Fig. 4C).

140 The influence of SST on PAS and collapse was further investigated at the scale of large marine ecosystems
141 (LMEs). We evaluated the relationship between warming rate (SST change) between 1950 and 2020 and the
142 proportion of stocks that underwent a PAS or a collapse. We found a significant positive linear relationship
143 between negative PAS and warming rate ($p = 0.006$, Fig. 4D), meaning that LMEs with the most rapid
144 warming rate were also those with the highest proportion of stocks that underwent negative PAS. However,
145 we only found a not significant positive linear relationship between warming rate and the proportion of
146 collapsed stocks ($p = 0.205$, Fig. 4F), even when using different definitions of collapse ($0.25 < p < 0.44$, Fig.
147 S4). That is, the LMEs warming most rapidly had more frequent negative PAS, but this did not translate into
148 more frequent collapse of stocks in these LMEs.

149 **Discussion**

150 In this work we distinguished productivity abrupt shifts (PAS) in fisheries time series from gradual productivity
151 trajectories using a systematic classification of trajectory types based on shape and trend (no change, linear,
152 quadratic, abrupt). We found that PAS were detected in more than 25% of stocks worldwide and that PAS
153 occurrence varied in space and time. We provide evidence that large negative PAS frequently preceded stock
154 collapses and were associated with a higher warming rate. Those findings could have critical implications for
155 fisheries management in warming oceans.

156 Our results expand on the well-documented examples of regime shifts (e.g., (Blöcker et al., 2023; Möllmann
157 et al., 2021)), giving a more complete picture of the prevalence of such shifts. Our classification aptly
158 identifies stocks like Newfoundland cod (Myers et al., 1997), Baltic sea cod (Möllmann et al., 2021), or
159 Japanese sardine (Watanabe et al., 1995), which underwent among the most prominent and rapid abrupt
160 collapses previously characterized (full list available in Table S1). More importantly, we also identify others
161 that were surprisingly not extensively treated in the literature (e.g., Greenland halibut off Labrador Shelf -
162 Grand Banks in the 1990s). Assuming that retrospective analyses of productivity trajectories can give good
163 insights into how stocks are likely to react in the future, our approach enables the identification of stocks that
164 could be more prone to abrupt decline and thus require more careful management.

165 The prevalence of abrupt shifts we found (25%) is somewhat below those from previous analyses (46% for
166 (Sellinger et al., 2024), 39% for (Vert-pre et al., 2013)). The differences in prevalence could be explained by
167 the set of stocks analyzed, the different relationships considered (stock-recruitment in (Sellinger et al., 2024)),
168 the different models used (productivity-abundance relationship in (Vert-pre et al., 2013)), and perhaps most
169 importantly because our classification involved the congruence of two independent models to confirm a
170 trajectory as abrupt, making the attribution of abruptness more strict but also probably more reliable. Our
171 trajectory classification without the confirmation step doubles the prevalence of abrupt shifts (51%, Fig. S5).

172 Among the life history, climate change, fishing intensity related factors that have been proposed to affect
173 stock collapses either regionally or globally, we found that trend in SST to be the most significant variable
174 related to the occurrence of negative PAS. Climate change, alone or in conjunction with other factors, had
175 already been stated as one of the most prominent drivers of marine regime shifts (Rocha et al., 2015) and
176 stock collapses (Pinsky and Byler, 2015). Interestingly, the large marine ecosystems that underwent the
177 largest SST increases between 1982 and 2006 ($> 1^{\circ}\text{C}$, (Belkin, 2009)) are also those for which we found the
178 largest proportions of stocks with negative PAS during the same period, namely the Baltic and North Sea, the
179 East China Sea and Sea of Japan, and Newfoundland - Labrador Shelf. As fishing intensity alone did not
180 explain negative PAS, we did not explore potential interactions between fishing and temperature that have
181 previously been examined for global (Pinsky and Byler, 2015) and regional (Rouyer et al., 2014) stock
182 collapses. However, we found an effect of fishing intensity related to positive PAS. This effect might in fact
183 correspond to the early stages of a fishery when the productivity of unfished stocks is usually low because the
184 stock is assumed to be near carrying capacity. It corresponds to the basic principle of maximum sustainable
185 yield, that productivity is maximized for intermediate levels of fish abundance corresponding to fishing
186 mortality rates near those associated with maximum sustainable yield (F_{MSY} , (Walters and Martell, 2004)). No
187 sign of abrupt recovery after collapse was found in the data, except for the herring in the North Sea, which
188 recovered rapidly following reduction in fishing pressure - a result already documented by (Dickey-Collas et
189 al., 2010).

190 We accounted for life history essentially through maturity related metrics and principal habitat, but in contrast
191 with previous studies (Pinsky et al., 2011; Pinsky and Byler, 2015) we found no significant effect of those life

192 history traits on negative shifts. Only the age at maturity (which is negatively correlated with somatic growth
193 rate, Fig. S6A) tended to be positively associated with positive shifts, meaning that rapid productivity
194 increases were more often found in slower growing, later maturing species. This pattern is consistent with the
195 periodic life history strategy (Winemiller and Rose, 1992) where delayed maturation and consequent large
196 numbers of eggs can result in occasional very large recruitment events when these eggs encounter favorable
197 conditions for early life survival. These periodic strategists often exhibit years of minimal or even negative
198 productivity punctuated by occasional large year classes that create a period of high productivity as the year
199 class ages and grows.

200 We also found support for a link between negative PAS and subsequent stock collapse for a quarter of the
201 stocks that shifted negatively. This sequence of events could be expected with productivity decline impacting
202 stock abundance and biomass. However, collapse following negative PAS was far from inevitable, a fact likely
203 explained by the high prevalence of harvest control rules, which reduced fishing pressure as abundance
204 declines (Punt, 2010). Our results also suggest that gradual stock collapses were associated with lower levels
205 of warming, while areas that underwent rapid warming were more likely to experience negative abrupt shifts.
206 There is evidence for such warming-driven declines notably in the case of insufficient management adaptation
207 (Pershing et al., 2015) but a more extensive coverage of stocks with lower warming rate would be necessary
208 to confirm this. These results are all robust to different definitions of stock collapse (Fig. S7).

209 Additional analyses would be needed to ascertain whether the abrupt shifts in productivity of individual fish
210 stocks that we detected here correspond with ecosystem-wide regime shifts. , Such analysis would require
211 time series of other ecosystem components, many of which (e.g., zooplankton abundance) are unavailable at
212 a global scale (deYoung et al., 2008; Lees et al., 2006). Still, the fact that abrupt shifts occur is meaningful,
213 even in the absence of a broader regime shift, as they have serious implications for management regardless
214 of the mechanisms driving such shifts (Beaugrand et al., 2022; Möllmann and Diekmann, 2012).

215 While the RAM Legacy database, on which this analysis was based, has broad geographic coverage, it does
216 not reflect the actual distribution of fisheries worldwide and is biased towards intensively monitored stocks

217 mostly from wealthy nations at temperate latitudes (Ricard et al., 2012). Fisheries from tropical latitudes are
218 more often data limited or monitored for a lower amount of time, which limits reliable shift detection. Yet, the
219 correlation we found between warming rate and negative PAS suggests that marine areas like the North-West
220 and South Atlantic, or South Indian Oceans that are projected to be hotspots of rapid temperature increase in
221 the next decades (Cheng et al., 2022) require careful monitoring and management. This is a challenging task
222 since it has been suggested that regional fisheries agencies are still struggling to account effectively for
223 climate change in their practices (Sumbly et al., 2021). As our results contribute to the understanding of fish
224 stock long-term dynamics, they could pave the way for advancing the anticipation of productivity abrupt shifts
225 to improve effective fishing regulation in the face of rising global change pressures.

226 **Materials and Methods**

227 **Fisheries data**

228 Time series of catch and stock biomass were download from the freely available RAM Legacy Stock
229 Assessment database (RAMLDB v4.61, (Ricard et al., 2012)). From those, we estimated stock productivity (or
230 surplus production) time series, which is a biological variable particularly relevant to investigate putative
231 regime shifts (Vert-pre et al., 2013). Abrupt shifts in productivity time series have been used as hallmark of
232 regime shift in some particular fish stocks (Blöcker et al., 2023; Möllmann et al., 2021). Surplus production
233 can be assimilated to stock productivity if it is independent to biomass, which has been suggested to be the
234 case for most of the stocks (Vert-pre et al., 2013). We therefore refer to surplus production as productivity.
235 The productivity $S(t)$ in the year t was estimated from catch and total biomass time series according to the
236 following formula:

$$237 \quad S(t) = B(t+1) - B(t) + C(t)$$

238 with $B(t)$ stock total biomass and $C(t)$ catch in the year t . While catch values are raw data, often assumed to
239 be measured without error, biomass estimates are model outputs and concerns have been raised when
240 considering model outputs as input data (Brooks and Deroba, 2015). Here, the greatest concern is that
241 interannual smoothing of biomass time series by the assessment model could cause us to miss true abrupt

243 shifts. This concern is minimized by the fact that most biomass time series in the database are from highly
244 flexible statistical catch at age models that allow for substantial process error (e.g., interannual variation in
245 reproductive output or recruitment, (Ricard et al., 2012; Thorson et al., 2014)). Our data filtering criteria
246 (described below) removed a small number of stocks where smooth biomass trajectories were clearly a model
247 artifact..

248 The RAMLDB includes stock assessments from several phyla including mollusks, crustaceans, and
249 vertebrates (fishes). For consistency, we only focused on marine fish stocks including both ray-finned and
250 cartilaginous fishes, and ignored data before 1950 that sometimes corresponded to model extrapolations.
251 Fish productivity time series with no missing data points were estimated for 397 stocks but a total of 82 were
252 discarded either because they had time series length below 25 years (68 stocks) or because biomass
253 estimates were apparently generated from deterministic models which could only produce smooth changes
254 (14 stocks). The classification was thus performed on 315 stocks corresponding to 161 taxa, among which
255 158 were at the species and 3 at the genus level. Median time series length was 41 years and the longest
256 time series was 71 years. To allow for inter-stock comparisons, productivity was normalized by average stock
257 biomass following (Essington et al., 2015).

258 **Time series classification**

259 Productivity time series were fitted with four different types of model – intercept-only (no change), linear trend,
260 quadratic trend, and abrupt change – and Akaike Information Criterion corrected for small sample size (AICc)
261 were computed for each model fit. The best trajectory was considered to be the model with lowest AICc and
262 validated following (Pélissié et al., 2024). To validate abrupt trajectories, an independent breakpoint detection
263 method (*asdetect*, (Boulton and Lenton, 2019)) was run and we checked whether both methods agreed on
264 the shift date with a tolerance of five years. For quadratic and linear cases, we tested the significance of the
265 higher order coefficient. We used the classification with default parameters for *asdetect* method (anomalous
266 rate of change equal to three medians absolute deviations and detection threshold set to 0.15) that have been
267 tested to be most reliable for time series of at least 25 time points (Pélissié et al., 2024). To avoid uncertain
268 shifts, shifts less than five year from the start or the end of the time series were not considered. If the abrupt

269 shift model and the *asdetect* method agreed that an abrupt shift occurred but disagreed on the year of the
270 shift, the year identified by the abrupt shift model was used. We also computed three classification quality
271 score to assess different aspects of model reliability, namely the relative support for each fitted model (AICc
272 weight), model choice robustness to individual data point removal (Leave-One-Out cross validation), and the
273 ratio between variation not explained by the model and the overall variation (normalized RMSE). This
274 framework permits the identification of at most one shift in each time series, which is assumed to be the
275 largest one if several might be present. The standardization by stock average biomass changes the absolute
276 AICc values but does not affect model ranking or best model choice.

277 **Spatial data**

278 A stock represents a consistent population unit that is spatially bounded. Most stock polygons were extracted
279 from the dataset used by (Free et al., 2019). An addition of 64 stock polygons to the dataset was made
280 following Free's methodology either by combining fishing subareas or divisions, approximating polygons to
281 already existing ones, or digitizing fishing areas based on individual assessment data. All added polygons
282 were processed using QGIS v3.16. Stocks were grouped by FAO Major Fishing Areas as directly available
283 from RAMLDB dataset. We also grouped stocks by Large Marine Ecosystems (LMEs) that represent
284 biogeographically more relevant units. We assigned each stock to the LME with which stock polygons
285 overlapped most, except for high seas fisheries for which they were assigned to main ocean areas.

286 **Explanatory variables**

287 Potential explanatory variables were gathered for each taxa or stock, spanning life history, environmental, and
288 fishing-related pressure. To avoid missing data, we used imputed life history traits from the FishLife database
289 (Thorson, 2020) and selected two relevant variables among the least correlated ones, namely age and length
290 at maturity (Fig. S6A). Primary habitat was retrieved from FishBase (Froese and Pauly, 2023) with classes
291 grouped as pelagic (pelagic-neritic, pelagic-oceanic, bathypelagic) and demersal (bathydemersal,
292 benthopelagic, demersal, reef-associated). Sea surface temperature (SST) was used as the most relevant
293 environmental variable using data from the Met Office HadISST1 dataset (Rayner et al., 2003) from 1870 to
294 2020 on a monthly basis (then averaged annually) and with one degree spatial resolution. Annual SST maps

295 were clipped based on stock or LME boundaries and averaged to obtain a single SST time series per stock
296 (Free et al., 2019). Exploitation rate (ER) was used as a proxy for fishing intensity and corresponds to the
297 ratio of catch to total biomass in the same year. Exploitation rate relative to that at maximum sustainable yield
298 U/U_{msy} may have been a better proxy since it accounts for intrinsic differences in productivity among stocks
299 but was not available for 15% of the stocks. SST and ER means and trends were estimated from the start of
300 productivity time series up to the shift if any was detected or to the latest year available otherwise. The
301 absence of strong multicollinearity of explanatory variables was determined by performing pairwise Pearson
302 correlation tests (Fig. S6B).

303 To test the association between stock collapse and PAS, we defined a stock as collapsed if biomass in a
304 given year was below 25% of the average stock biomass recorded to date following the definition by
305 (Essington et al., 2015) and if such threshold was crossed for at least two consecutive years to limit artifacts.
306 We focused on whether a stock ever collapsed and the first year of collapse. We repeated the analyses that
307 involved collapsed status with alternative definitions used by (Yletyinen et al., 2018), namely 10% of
308 maximum biomass and 15%, and 50% of average biomass.

309 **Statistical analyses**

310 We assessed the homogeneity of trajectories found across regions and taxonomic orders separately using
311 Chi-square test and computed p-values by Monte Carlo simulation using 10^5 replicates to deal with cases of
312 low expected values.

313 We used logistic generalized additive mixed models (GAMM) to assess the effect of independent variables on
314 the occurrence of productivity abrupt shifts using the *mgcv* package (Wood, 2011). Positive and negative
315 shifts were considered separately as dependent variable against all other trajectories to contrast conditions of
316 abrupt shifts against all conditions available. The effect of life history, environmental conditions, and fishing
317 intensity were estimated as fixed effects by considering age at maturity, length at maturity, main habitat,
318 average and trend in SST, and average and trend in ER altogether. Numerical variables were standardized to
319 allow comparison. To account for space, stock centroid coordinates were considered as smooth terms.
320 Taxonomic orders and families were treated as random effect to account for broad phylogenetic relationships.

321 Two stocks from the Japanese Seto inland sea lacked SST data due to limited spatial resolution and were
322 thus discarded for these analyses.

323 An alternative approach was also performed to assess the contribution of each predictor variable
324 independently and in conjunction with the other predictors on the occurrence of abrupt shifts. This approach
325 controls for the possible influence of unchecked multicollinearity among variables. We used the *hier.part*
326 package (Nally and Walsh, 2004) to run hierarchical partitioning and randomization tests with 999 repetitions
327 on positive and negative shift occurrence as binomial variables with a logit link function. Phylogeny was not
328 used for these analyses because of excessive number of groups in each taxonomic level.

329 **Sensitivity analyses**

330 We also ran the classification without the step for abrupt shift confirmation to determine how frequently this
331 step was eliminating potential abrupt shift classifications. Without the confirmation step, the proportion of
332 stocks classified as PAS reached 51.4% with similar proportions of negative (26%) and positive (25.4%) shifts
333 (Fig. S5). We also repeated the GAMM analysis of predictors of PAS with this less strict classification for
334 PAS. In that case, we still found the effect of SST change ($p = 0.049$, Fig. S8A) for negative PAS, with
335 average SST and ER also found as significant predictor ($p = 0.006$ and $p = 0.008$ respectively, Fig. S8A). For
336 positive PAS, the effect of both average ($p < 0.001$) and trend in ER ($p = 0.013$) were found as well as a
337 positive effect of average SST ($p = 0.034$, Fig. S8B).

338 We repeated the GAMM with fishing intensity relative to that at maximum sustainable yield (U/Umsy) instead
339 of exploitation rate (ER). U/Umsy accounts for intrinsic species differences in withstanding fishing pressure
340 but was available only for 269 stocks. Similar effects to the initial models in Fig. 3 were found in estimates
341 although not significant for SST change ($p = 0.084$), probably due to the loss of many stocks with negative
342 PAS (14 out of 42, among which 6 also collapsed), with in addition a significant positive effect of trend in
343 U/Umsy ($p = 0.002$, Fig. S8C). Similar effects to the initial models in Fig. 3 were also found for positive PAS
344 with age at maturity as the only significant variable ($p = 0.049$, Fig. S8D).

345 We also tested the robustness of the results related to stock collapse by using different thresholds of collapse
346 (10% of maximum stock biomass, 15% and 50% of the average stock biomass), which gave similar results in
347 terms of relationship with negative PAS (Fig. S7) and warming rate (Fig. S4).

348 **Author Contributions**

349 All authors designed research; M.P. performed empirical analyses; M.P. analyzed the data and all authors
350 contributed to writing the paper.

351 **Data availability**

352 All analyses were conducted using R software (version 4.3.3). The R script to replicate all analyses is
353 available on Github (https://anonymous.4open.science/r/ocean_warming_fisheries).

354 **Acknowledgments**

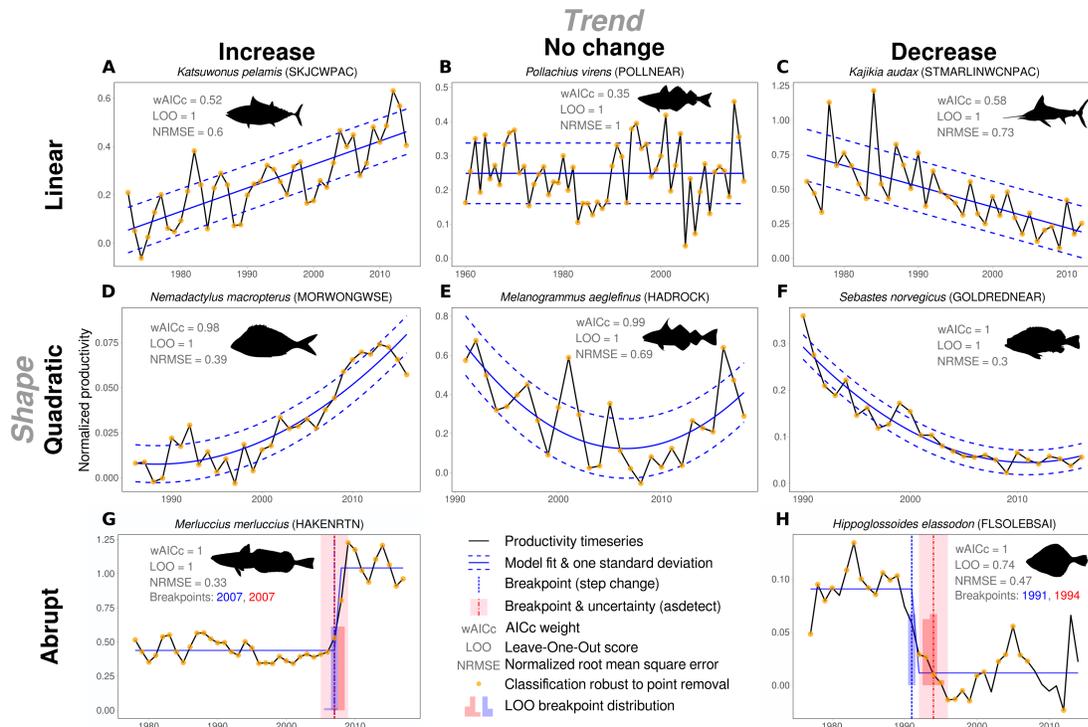
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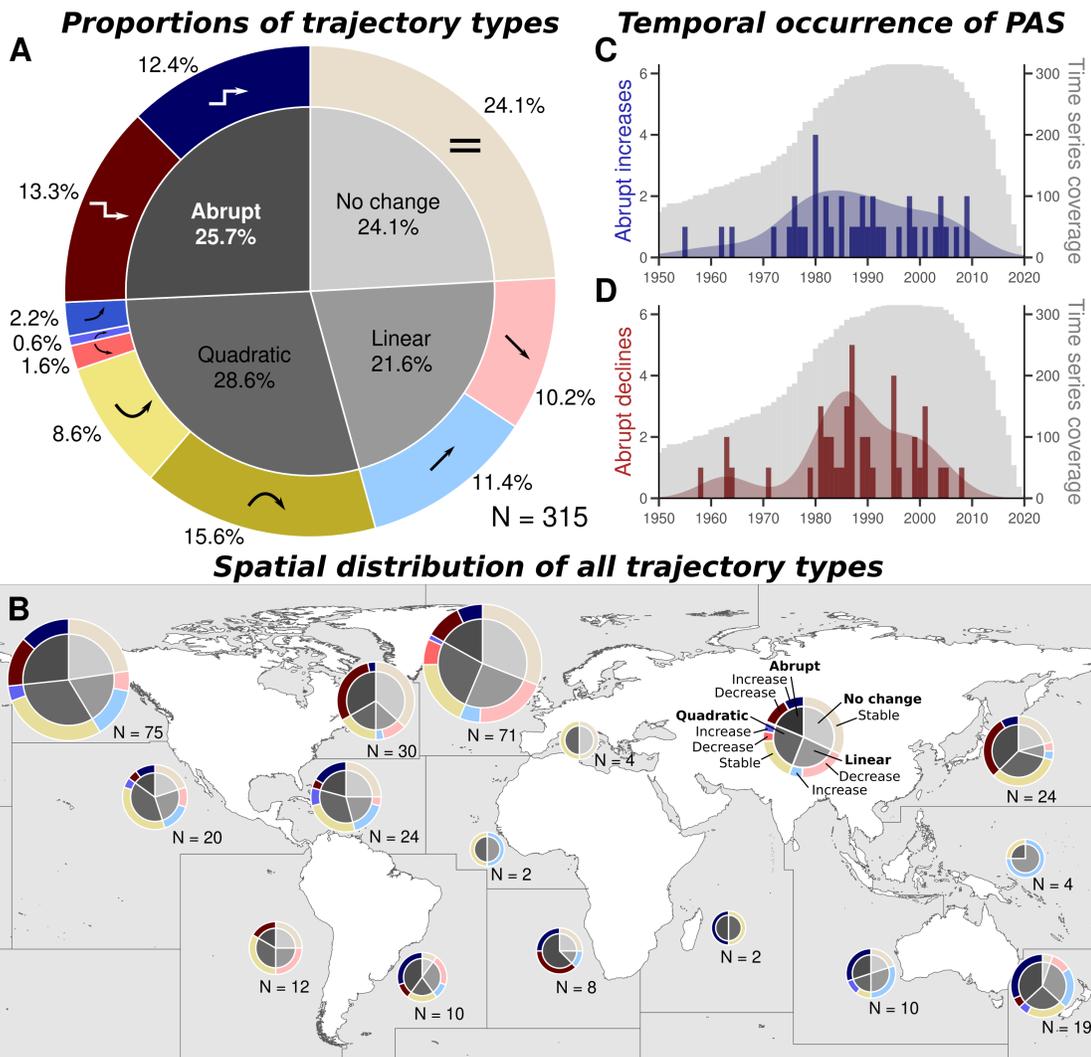
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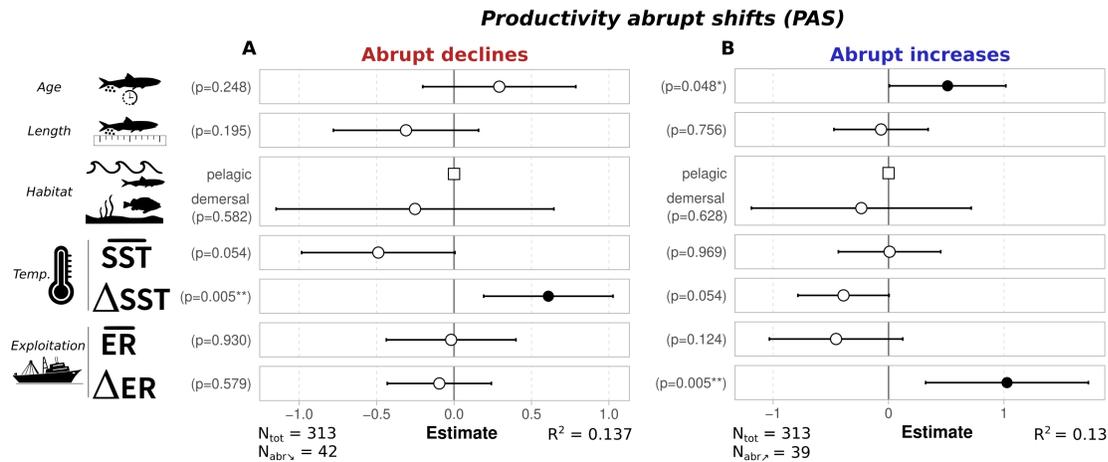
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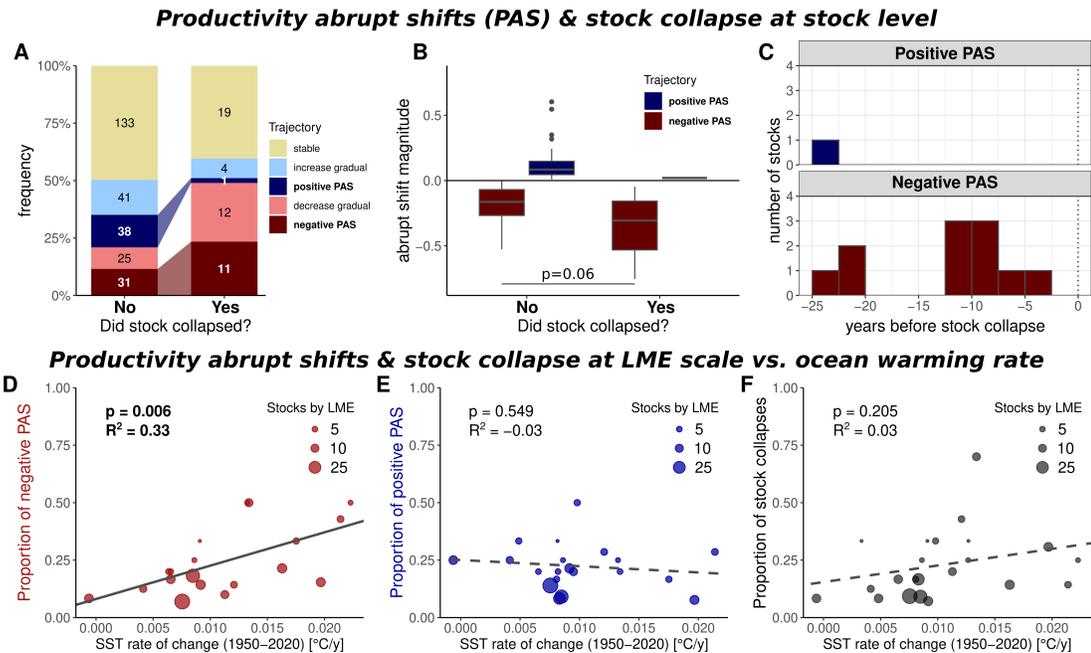
360 **Figure 1.** Examples of productivity time series classified into trajectory types based on shape and trend.
 361 Some time series were best described by either linear (A-C), quadratic (D-F), or abrupt trajectories (G, H),
 362 and overall trend either positive (A, D, G), or negative (C, F, H). No change is a types of shape that
 363 corresponds to a trajectory without trend (B, E). Each panel shows fish stock productivity time series
 364 normalized by the average stock biomass (black line) with the best model fit (solid blue line) and standard
 365 deviation (dashed lines) following the classification procedure detailed in the methods. Three classification
 366 quality scores are specified for each time series: AICc weight (wAICc), leave-one-out score (LOO), and
 367 normalized root mean square error (NRMSE). Timepoints that if removed in the LOO process would result in
 368 the same shape are highlighted by orange dots. For abrupt trajectories (G, H), the location of breakpoints is
 369 indicated by vertical lines, the pink background corresponds to the breakpoint uncertainty for the *asdetect*
 370 method, and the distribution of breakpoint locations from LOO time series are represented by color bars. Note
 371 that scales are not the same for the different panels. Species name and (stock ID) are also displayed. Fish
 372 silhouettes come from <https://www.phylopic.org/>.



373 **Figure 2.** Overview and spatial distribution of all productivity trajectory types alongside the temporal
 374 distribution of productivity abrupt shifts (PAS). (A) Overall proportion of all productivity trajectory types,
 375 with the inner pie chart indicating trajectory shape and the outer layer specifying the trend. (B) Spatial
 376 distribution of all trajectory types by FAO major fishing areas. The size of the charts is indicative of the number of stocks,
 377 which are also displayed. (C-D) Temporal distribution of positive (C) and negative PAS (D) (color bars and
 378 density) with in gray the coverage (i.e., number of stocks with data available in each year) of the time series
 379 classified.



380 **Figure 3.** Generalized additive mixed model (GAMM) coefficient estimates for models accounting for negative
 381 (A) and positive (B) PAS against all other trajectories. Positive estimates mean that PAS are more often found
 382 for larger values of the variable and conversely. Dependent variables are arranged from top to bottom with
 383 age at maturity, length at maturity, main habitat (pelagic as reference category), mean (SSTbar) and linear
 384 trend (deltaSST) in sea surface temperature, mean (ERbar) and linear trend (deltaER) in exploitation rate.
 385 Mean and trend in SST and ER were computed from the first year of productivity available up to the shift if
 386 any was detected. Model estimates (circles) with confidence intervals at 95% (horizontal bars) are presented
 387 and significant estimates are indicated with filled circles. All numeric variables were standardized.



388 **Figure 4.** Relationship between PAS and stock collapse at stock level (A-C) and at the scale of Large Marine
 389 Ecosystems (LMEs) related to ocean warming rate (D-F). (A) Proportion of trajectories depending on whether
 390 stocks ever reached collapsed state (defined as 25% of the mean biomass) with actual number of stocks
 391 indicated. (B) Standardized PAS magnitude for negative (dark red) and positive (dark blue) shifts and
 392 collapsed state. (C) Distribution of PAS location relative to the first year as collapsed (vertical dotted line), with
 393 positive PAS (top panel in blue) and negative PAS (bottom panel in blue). (D-F) Relationship between the
 394 proportion of PAS (D, E) or stock collapse (F) within LMEs and Sea surface temperature (SST) rate of change
 395 between 1950 and 2020. LMEs with less than two stocks were not included. Significant linear regression are
 396 drawn with a solid line, and dashed line otherwise.