# Unifying occupancy-detection and local frequency scaling (Frescalo) models

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#### Abstract 4

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Frescalo's "local frequency scaling" and classical occupancy-detection models both seek to recover true species-occurrence signals from imperfect data. In this paper, we show that the two approaches rest on the same underlying detection mathematics. Occupancy models treat each site's repeat visits as independent detection trials and separately estimate occupancy probability and per-visit detectability. Frescalo, by contrast, pools data across ecologically defined neighbourhoods and infers a single combined detection rate and a temporal "time-factor" to capture trends. We demonstrate that the Bernoulli-trial formulation of occupancy-detection converges to Frescalo's Poisson-process framework, with occupancy and detectability collapsing into a single rate parameter. This equivalence clarifies how Frescalo's neighbourhood and time corrections function as a coarser-scale analogue of repeat-visit models. By casting Frescalo in occupancy modelling terms, we hope to promote further investigation into the adoption of occupancy-model diagnostics, extensions and covariate tests within Frescalo analyses, improving transparency and rigour when working with less structured biodiversity data.

Keywords: occupancy models, sampling effort, effort correction, citizen science, unstructured data, Frescalo

### 1. Introduction

Occupancy-detection models [9] and the Frescalo "local frequency scaling" method [6] both aim to correct raw biological records (i.e. species occurrence) data for imperfect sampling. Classical occupancy models do this at the scale 10 of repeated visits to individual sites, explicitly estimating true presence prob-11 12 abilities  $(\psi)$  and detectability (p) via a hierarchical likelihood. Frescalo was designed to work at larger spatio-temporal scales, exploiting emergent patterns 13 of relative frequency in "neighbourhoods" to derive Poisson-process-based scaling 14 factors ( $\alpha$ ) and species' relative "time factors" indexing true fluctuations in 15 site occupancy. Given that many datasets lack repeat-visit structure, and/or 16 may exhibit variation in the detection process that is unmodellable due to a 17

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lack of knowledge of its determinants [11], understanding how Frescalo recovers
effort-adjusted trends from aggregated data can broaden the toolkit of ecologists.

Although the two models can appear quite different, Pescott et al. [11] 20 informally suggested that Frescalo could be seen as a type of occupancy-detection 21 model "where an adjustment for overlooked species is made in relation to spatial 22 rather than temporal replication, whilst simultaneously adjusting for variable 23 regional effort". We here show that this suggestion can be formalised due to 24 the two model types' reliance on the same core mathematics of Bernoulli versus 25 Poisson detections. Below we (1) recall each framework, (2) write down their 26 key equations, and (3) algebraically map one onto the other, demonstrating that 27 Frescalo time trends are based on an implicit occupancy-detection model whose 28 "visits" and "occupancy" are folded into a single site/species discoverability rate 29 parameter  $\lambda$  and standardised neighbourhood effort index. 30

# 31 2. Occupancy-detection models

# 32 2.1. Basic single-season model

Following MacKenzie et al. [9], at each site i for species j assume a latent occupancy indicator

$$z_{ij} \sim Bernoulli(\psi_{ij})$$

<sup>35</sup> Conditional on presence, v total survey visits indexed by k produce

$$y_{ij1}, \dots, y_{ijv} \mid z_{ij} = 1 \sim Bernoulli(p_{ij})$$

where p is detectability. If  $z_{ij} = 0$  (i.e. species absent), then all  $y_{ijk} = 0$ .

<sup>37</sup> Marginalising out  $z_{ij}$ , it is well-known that the probability of at least one <sup>38</sup> detection across v visits is

$$P(\max_{k} y_{ijk} = 1) = \psi_{ij} [1 - (1 - p_{ij})^{v}].$$

<sup>39</sup> Thus the model simultaneously estimates

 $\psi_{ij} = \Pr(\text{occupied}), \quad p_{ij} = \Pr(\text{detect}|\text{occupied}),$ 

40 and inference proceeds via the full likelihood over all sites and detection histories.

# 41 3. Frequency scaling using local occupancy (Frescalo)

42 3.1. Neighbourhood frequencies

Frescalo [6] pools presence-only data across a neighbourhood around target site *i*. We denote the observed proportion of neighbourhood sites in which species *j* was recorded by  $f_{ij}$  (in practice this frequency may relate to a weighted neighbourhood as per Hill [6], but this detail is not crucial for what follows). Under a Poisson-process model of species discovery with rate  $\lambda_{ij}$  and unknown total neighbourhood-level sampling effort  $s_{i(N)}$ , one has

$$f_{ij} = 1 - \exp(-\lambda_{ij} s_{i(N)}).$$

<sup>49</sup> Subsequently, a frequency-weighted neighbourhood index

$$\phi_i = \frac{\sum_j f_{ij}^2}{\sum_j f_{ij}}$$

is then "standardised" to a target value  $\Phi$  by solving for a site-specific effort multiplier  $\alpha_i$  such that

$$\phi_i(\alpha_i) = \frac{\sum_j [1 - (1 - f_{ij})^{\alpha_i}]^2}{\sum_j [1 - (1 - f_{ij})^{\alpha_i}]} = \Phi.$$

<sup>52</sup> Mathematically,  $\Phi$  is chosen so that every neighbourhood's weighted-mean <sup>53</sup> frequency  $\phi_i = \sum_j f_{ij}^2 / \sum_j f_{ij}$  equals  $\Phi$ . Hill [6] showed that  $\phi_i$  is the ratio of <sup>54</sup> the mean species richness to the 'effective number of common species' (often <sup>55</sup> called  $N_2$ , the reciprocal of Simpson's index; Hill [5]), which means that  $\phi_i$ <sup>56</sup> isolates sampling intensity from true differences in richness and evenness. By <sup>57</sup> fixing  $\phi_i = \Phi$ , we therefore align all neighbourhoods to the same effort scale <sup>58</sup> without erasing real ecological differences.

<sup>59</sup> This process yields the standardised neighbourhood frequencies

$$\tilde{f}_{ij} = 1 - (1 - f_{ij})^{\alpha_i}$$

which are independent of time (i.e. they are calculated with respect to the entire time period under consideration, rather than any subdivisions of this used for trend calculations), and serve as a proxy for the "true" discoverability- or effort-standardised neighbourhood species rank-frequency curve.

# 64 3.2. Temporal correction

Within each time period t, one chooses a set of "benchmark" species [8] and 65 computes the proportion recorded per site and time period (Hill's  $s_{it}$ ) as an index 66 of site-level recording effort. (Note that there are potentially many ways to choose 67 ones' neighbourhood benchmarks, but Hill [6] proposed a fixed proportion  $R^*$  of 68 the standardised species rank-frequency curve after an additional normalisation 69 step involving the division of species' ranks by the expected species count  $\sum_{i} f_{ij}$ ; 70 however, the precise method of choosing benchmarks does not affect what follows). 71 For each species j in period t, Hill then defines a Poisson-link intensity 72

$$Q_{ijt} = -\ln[1 - s_{it}\tilde{f}_{ij}],$$

<sup>73</sup> The modelled "discovery" probability is then

$$P_{ijt}(x_{jt}) = 1 - \exp(-Q_{ijt}x_{jt}).$$

Hill [6] estimates the time-factor  $x_{jt}$  by matching the *total* modelled to *total* observed presences  $y_{ijt}$ :

$$\sum_{i} y_{ijt} = \sum_{i} P_{ijt}(x_{jt}).$$

In practice one iterates  $x_{jt}$  in the exact Poisson form above until those sums coincide (e.g. see the *R* code of Pescott [10]). The difference between the (summed) observed presences  $y_{ijt}$  and the model's baseline expectation after standardising time-independent neighbourhood effort  $\alpha_i$  and adjusting for site/time specific effort  $s_{it}$  is therefore captured by the time factor  $x_{jt}$ . Frescalo can thus deliver detection-corrected trends from unstructured data when its core assumptions are met.

# 4. Bridging the gap

# 84 4.1. Static occupancy and detection

We can compare the static (i.e. single season) single-species occupancydetection model probability of at least one detection in v visits

$$\psi [1 - (1 - p)^v]$$

with the Poisson-process discovery rate (conditional on the all-time frequency
curve) used in Frescalo

$$1 - e^{-\lambda s_{i(N)}}.$$

For small pv,  $(1-p)^v \approx e^{-pv}$ , hence  $\psi[1-\exp(-pv)] \approx 1-\exp(-\psi pv)$ ; now identifying  $\lambda = \psi p$  and  $v = s_{i(N)}$  recovers the approximate Frescalo detection probability  $1-\exp(-\lambda s_{i(N)})$ . For any value of p, exact equivalence can be found by solving

$$1 - e^{-\lambda v} = \psi [1 - (1 - p)^v] \tag{1}$$

93 for

$$\lambda = -\frac{1}{v} \ln[1 - \psi(1 - (1 - p)^v)], \qquad (2)$$

but this only reduces to  $\psi p$  in the limit  $pv \to 0$ . Frescalo's Poisson rate  $\lambda$  is 94 therefore exactly the function of occupancy, detectability and (latent) visit count 95 that makes equation (1) true. Whilst in Frescalo we never observe v directly, 96 we can infer it via the continuous neighbourhood effort index  $s_{i(N)}$ , which is 97 aligned across all neighbourhoods by the spatial scaler  $\alpha_i$ . Frescalo can therefore 98 be interpreted as an occupancy-detection analogue at the neighbourhood scale: 99 it replaces the two-parameters  $(\psi, p)$  and discrete v with a Poisson rate  $\lambda$  and 100 a continuous effort-multiplier  $\alpha$  equalising variable survey effort  $s_i$  (inferred by 101 the neighbourhood level  $s_{i(N)}$  across sites. 102

A key step in recognising the equivalent elements of these models is to appreciate that Frescalo applies its discoverability standardisation at a large scale: not only is the adjustment done with respect to the multi-site neighbourhood and across all species, but it is also calculated across all time periods in the analysis. The standardised neighbourhood frequencies  $\tilde{f}_{ij}$  and the species rank-frequency curve they form is estimated once, independently of time, before temporal change is examined.

### 110 4.2. Time trend interpretation

A time trend in occupancy derived from a classical occupancy-detection 111 model is modelled simply by letting  $\psi_{ij}$  vary linearly or non-linearly over time, 112 conditional on both model-specific [16, 17] and other standard survey sampling 113 assumptions [3] being reasonable. Frescalo, by contrast, posits a single time-114 independent set of discoverability-adjusted baseline frequencies  $f_{ij}$ , and then uses 115 benchmarks and the site/period effort index  $s_{it}$  to compute expected frequencies 116 under an assumption of stasis, subsequently letting the time factors  $x_{it}$  absorb 117 any residual differences as true ecological change. 118

This underscores a key difference in how effort-adjustment processes func-119 tion in each model type. Occupancy-detection models assume that true site 120 occupancies, and so trends in these, are directly recoverable from visit-level 121 information. Frescalo assumes that fine-scale visit data is generally unavailable 122 and/or uninformative for all or part of the time series of interest, and so models 123 species' discoverability at a much larger scale. The main aim of this adjustment 124 is to ensure a common scale across which neighbourhoods, and therefore sites, 125 can be compared: without the harmonisation of effort across neighbourhoods, 126 the time factors estimated for each site for a species would not be comparable. 127 making average time factors and trends in these meaningless. 128

Another fundamental difference is the meaning of the site occupancy values 129 produced. As noted,  $\psi_i$  has the simple meaning of predicted site occupancy under 130 the classical model (notwithstanding debates around usage versus occupancy 131 when these types of models are applied at different scales; [14]). The Frescalo 132 time factor  $x_{it}$  is, however, defined by the benchmark average, and values >1 or 133 <1 indicate that a species is at a higher or lower average frequency relative to the 134 common species where it occurs, rather than in absolute occupancy probability. 135 This may be an important limitation to inferring effort via observable recording 136 outcomes, rather than having knowledge of those factors that directly map onto 137 effort, such as the actual number of visits and covariates that are known to explain 138 an important portion of observed variance in species' visit-level detectability 139 [6, 15, 7].140

One way around this issue is the observation of Bijlsma [1] that site occupancy 141 probabilities can actually be derived from Frescalo via the combination of the 142 standardised species' frequencies  $f_{ij}$  and the time-factors  $x_{jt}$ , and this has been 143 exploited in at least one published analysis [4]. However, this requires a note of 144 caution: whilst sensitivity analyses published in Hill [6] suggested that the trends 145 in time-factors estimated by Frescalo can be relatively insensitive to the choice of 146  $R^*$ , the benchmark threshold (variation in this parameter changing the intercept 147 of estimated trends but not their slope), the same is not true of back-calculated 148 site occupancy probabilities. Because the relationship between time-factors and 149 species frequencies is non-linear, the shifts in time trend intercept seen using 150 151 different values of  $R^*$  will not translate into the same proportional changes in predicted site occupancies over time (Pescott, pers. obs.) This may be 152 particularly important when these trends are used to classify species' into risk 153 categories, as for example happens in Red Listing exercises [e.g. 12]. 154

### 155 **5.** Conclusions

Unstructured species occurrence data are too valuable to ignore, especially 156 for historical periods where no information about the visit-level data collection 157 process survives [11]. Hill's "frequency scaling using local occupancy" or Frescalo 158 method allows the careful analyst to infer a large-scale detectability or effort 159 metric that can subsequently be used to place neighbourhoods on a common 160 footing for the estimation of distribution trends. The large-scale formulation of 161 this approach not only allows for the potential inclusion of more data sources 162 (e.g. records extracted from Atlases or museums), but may also act to reduce 163 the actual error in species' trends intrinsically [2, 13]. 164

<sup>165</sup> By demonstrating how Frescalo collapses the classical occupancy-detection <sup>166</sup> model's  $\psi$  and p into  $\lambda$ , and how it infers visit-related effort via an emergent <sup>167</sup> community-level mean rate, the approach performs an occupancy-detection-<sup>168</sup> type correction even when explicit or informative temporal repeat-visit data <sup>169</sup> are lacking. By highlighting this link, we hope to promote the development of <sup>170</sup> additional diagnostics, extensions and more rigorous uncertainty quantification <sup>171</sup> for the frequency scaling using local occupancy method.

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# 175 **References**

- [1] Bijlsma, R.J., 2013. The Estimation of Species Richness of Dutch Bryophytes
   between 1900 and 2011. Technical Report 15. Dutch Bryological and Licheno logical Society. Gouda, the Netherlands.
- [2] Boyd, R.J., Bowler, D.E., Isaac, N.J.B., Pescott, O.L., 2024. On the
   trade-off between accuracy and spatial resolution when estimating species
   occupancy from geographically biased samples. Ecological Modelling 493,
   110739. doi:10.1016/j.ecolmodel.2024.110739.
- [3] Boyd, R.J., Powney, G.D., Pescott, O.L., 2023. We need to talk about
   nonprobability samples. Trends in Ecology & Evolution doi:10.1016/j.
   tree.2023.01.001.
- [4] Eichenberg, D., Bowler, D.E., Bonn, A., Bruelheide, H., Grescho, V., Harter,
   D., Jandt, U., May, R., Winter, M., Jansen, F., 2021. Widespread decline in
   Central European plant diversity across six decades. Global Change Biology
   27, 1097–1110. doi:10.1111/gcb.15447.
- [5] Hill, M.O., 1973. Diversity and evenness: A unifying notation and its consequences. Ecology 54, 427–432.

- [6] Hill, M.O., 2012. Local frequency as a key to interpreting species occurrence
   data when recording effort is not known. Methods in Ecology and Evolution
   3, 195–205.
- [7] Kéry, M., Royle, J.A., Schmid, H., Schaub, M., Volet, B., Häfliger, G.,
   Zbinden, N., 2010. Site-Occupancy Distribution Modeling to Correct
   Population-Trend Estimates Derived from Opportunistic Observations. Con servation Biology 24, 1388–1397. doi:10.1111/j.1523-1739.2010.01479.
   x.
- [8] Latour, J., van Swaay, C., 1992. Dagvlinders als indicatoren voor de regionale milieukwaliteit. De Levende Natuur 93, 19–22.
- [9] MacKenzie, D.I., Nichols, J.D., Lachman, G.B., Droege, S., Royle, J.A.,
   Langtimm, C.A., 2002. Estimating Site Occupancy Rates When Detection
   Probabilities Are Less Than One. Ecology 83, 2248–2255. doi:10.1890/
   0012-9658(2002)083[2248:ESORWD]2.0.C0;2.
- [10] Pescott, O., 2025. An R translation of Hill's fortran code for Frescalo.
   doi:10.5281/zenodo.15305437.
- [11] Pescott, O.L., Humphrey, T.A., Stroh, P.A., Walker, K.J., 2019. Temporal changes in distributions and the species atlas: How can British and Irish plant data shoulder the inferential burden? British & Irish Botany 1, 250–282. doi:10.33928/bib.2019.01.250.
- [12] Stroh, P., Leach, S., August, T., Walker, K., Pearman, D., Rumsey, F.,
  Harrower, C., Fay, M., Martin, J., Pankhurst, T., Preston, C., Taylor, I.,
  2014. A Vascular Plant Red List for England. Botanical Society of Britain
  and Ireland, Bristol.
- [13] Stroh, P., Walker, K., Humphrey, T., Pescott, O., Burkmar, R. (Eds.), 2023.
  Plant Atlas 2020. Mapping Changes in the Distribution of the British and Irish Flora. Botanical Society of Britain and Ireland & Princeton University Press, Princeton.
- [14] Valente, J.J., Jirinec, V., Leu, M., 2024. Thinking beyond the closure assumption: Designing surveys for estimating biological truth with
  occupancy models. Methods in Ecology and Evolution 15, 2289–2300.
  doi:10.1111/2041-210X.14439.
- [15] van Strien, A.J., van Swaay, C.A., Termaat, T., 2013. Opportunistic citizen
   science data of animal species produce reliable estimates of distribution
   trends if analysed with occupancy models. Journal of Applied Ecology 50,
   1450–1458. doi:10.1111/1365-2664.12158.
- [16] Warton, D.I., Stoklosa, J., Guillera-Arroita, G., MacKenzie, D.I., Welsh,
   A.H., 2017. Graphical diagnostics for occupancy models with imperfect
   detection. Methods in Ecology and Evolution 8, 408–419. doi:10.1111/2041 210X.12761.

- <sup>232</sup> [17] Wright, W.J., Irvine, K.M., Rodhouse, T.J., 2016. A goodness-of-fit test
- <sup>233</sup> for occupancy models with correlated within-season revisits. Ecology and
- Evolution 6, 5404–5415. doi:10.1002/ece3.2292.