1	Weaning age and supersuckling in Galápagos Sea Lions (Zalophus
2	wollebaeki): Maternal, offspring and environmental drivers of life-
3	history strategies
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5	Authors: Childs, Alexandra* ^{†1} ; Feldmann, Carlina ² ; Dyck, Julia ² ; Stoehr, Svenja ¹ ; Twiss, Sean
6	D. ³ ; Krüger, Oliver ^{1,4} .
7	
8	Affiliations:
9	1 Department of Animal Behaviour, Faculty of Biology, Bielefeld University, Konsequenz 45,
10	33501, Germany.
11	2 Department of Business Administration and Economics, Bielefeld University,
12	Universitätsstraße 25, 33615 Bielefeld, Germany.
13	3 Department of Biosciences, Durham University, Durham DH1 3LE, UK.
14	4 Joint Institute for Individualisation in a Changing Environment (JICE), Bielefeld University
15	and University of Münster, Konsequenz 45, Bielefeld 33615, Germany.
16	
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20	*Corresponding Author: Alexandra Childs; alexandra.childs@uni-bielefeld.de
21	†Current Address: Department of Animal Behaviour, Faculty of Biology, Bielefeld University,
22	Konsequenz 45, 33615 Bielefeld, Germany
23	
24	

25 Abstract

26 Weaning marks a critical life-history transition in mammals, shaped by maternal investment 27 strategies, offspring condition, and environmental constraints. In Galápagos sea lions (Zalophus 28 wollebaeki), weaning age is highly variable, with some individuals continuing to suckle into 29 adulthood (supersucklers). Using 20 years of mark-recapture data from 1890 individuals, we 30 applied multi-state capture-recapture models to investigate the probability of weaning age, and 31 logistic regression models to explore the drivers of supersuckling. We examined maternal effects 32 (age, birth interval, and pup condition at birth), offspring traits (sex, early growth), and 33 environmental variability (Oceanic Niño Index [ONI]). Weaning probability peaked at four years 34 of age, with 95% of individuals weaned by this time. Longer birth intervals were associated with 35 later weaning. Offspring sex had no influence on weaning probability. However, ONI affected 36 juvenile survival, indirectly shaping weaning outcomes. No intrinsic predictors of supersuckling 37 were identified, although low resource availability during an individual's second or third year of 38 life increased the likelihood of becoming a supersuckler. While weaning is typically driven by 39 escalating mother-offspring conflict, the persistence of supersuckling challenges this paradigm. 40 We propose that the regular occurrence of supersuckling in this species may therefore be socially 41 driven, reflecting the maintenance of long-term mother-offspring relationships that contribute to a 42 complex adult social structure. This study highlights the interplay between environmental 43 unpredictability and social behaviours in shaping life-history strategies.

44

45 **1. Introduction**

46 Weaning in mammals marks the end of maternal nutritional investment and is a fundamental life-47 history transition for offspring. The timing of weaning is shaped by a complex interplay of maternal investment strategies, offspring condition, and environmental constraints (Lee 1996; 48 49 Schulz & Bowen 2005; Avery & Zinn 2023). As an energetically costly form of parental care, 50 lactation must be balanced against the mother's future reproductive success, leading to an inherent 51 mother-offspring conflict over the duration of care (Trivers 1974). Offspring who benefit from 52 protracted maternal investment in the form of nutrition may have increased chances of survival 53 (Gerber et al. 2021; Dittus & Baker 2023; Vitikainen et al. 2023). However, gluttonous offspring, 54 once weaned, may be unable to maintain the same nutritional input and subsequently perish (Reiter 55 et al 1978). Mothers, on the other hand, may incur an immediate cost to body condition and reproductive fitness but be rewarded for their extended investment through delayed generational
success (Péron et al. 2019; Zipple et al. 2020; Bleu et al. 2021; Chapman et al. 2021; Nielsen et al

58 2023).

59 Amongst pinnipeds the duration of lactation is particularly variable due to their terrestrial or floe 60 ice breeding and aquatic foraging (Pomeroy 2011; Avery & Zinn 2023). Phocids generally exhibit 61 abrupt weaning after an intensive lactation period, whereas otariids (eared seals) experience 62 prolonged lactation and a gradual transition to nutritional independence that can last up to 36 63 months in some species (Sepúlveda & Harcourt 2021). This lengthily dependence increases the 64 potential for variability in weaning age and may intensify mother-offspring conflict (Trillmich & 65 Wolf 2008; Sepúlveda & Harcourt 2021; Dittus & Baker 2023). Maternal investment may be determined by pup traits such as sex, mass and demand (Goldsworthy 2006; McDonald et al. 2012; 66 Kraus et al. 2013). However, lactation duration, and therefore weaning age, is frequently dictated 67 68 by maternal factors such as parity, maternal resources and reproductive experience (Schulz & 69 Bowen 2004; Trillmich & Wolf 2008; Costa & Maresh 2022). Environmental factors such as 70 predation (Ferguson 2006) and climatic variability (Burkanov et al. 2011; Costa & Valenzuela-71 Toro 2021; Schwarz et al. 2022) further influence weaning by affecting maternal foraging success 72 and offspring condition. During periods of resource scarcity, females may prioritize their own 73 survival and future reproductive opportunities, leading to earlier weaning or decreased investment 74 in current offspring (Soto et al. 2004). Conversely, in times of higher resource availability, 75 prolonged lactation may be more feasible, allowing offspring to remain nutritionally dependent 76 for extended periods.

77 Galápagos sea lions (GSL) Zalophus wollebaeki are an endangered otariid endemic to the 78 Galápagos archipelago. As an equatorial species GSL are exposed to extreme stochastic 79 environmental fluctuations in the form of El Niño events which directly impact resource 80 availability for foraging mothers (Soto et al. 2004; Schwarz et al. 2022). The reproductive period 81 is protracted starting in September and ending in January (Pörschmann et al. 2010) although 82 females show high asynchronosity and low seasonality in their breeding cycle (Villegas-Amtmann 83 et al. 2009). There are high amounts of individual variation in age at primiparity (4-9 years) and 84 the time until the birth of the next sibling (birth interval). However, females produce on average, 85 a single pup every two years from the age of six (Kalberer et al. 2018). The expected lifespan for

86 GSL is 21 years (Krüger et al. 2021) and the oldest recorded age at parturition is 18 years (personal 87 *communication*). Adoption and alloparenting are extremely rare in this species (Childs et al. 2025). 88 Reproductive overlap is not uncommon and, whilst some females have been observed supporting 89 multiple dependent offspring, more often than not, the situation results in sibling competition and 90 the rejection or death of one offspring (Trillmich & Wolf 2008). Recent research has shown that 91 GSL exhibit a prolonged and flexible lactation period lasting between one and three years (Childs 92 et al. 2025), with some individuals, 'supersucklers' continuing to suckle into adulthood. However, 93 robust statistical analysis of age at weaning has yet to be conducted.

94 Accurately determining weaning age is crucial for understanding both the physiological and 95 ecological determinants of maternal investment strategies in otariids (Costa & Valenzuela-Toro 96 2021; McHuron et al. 2023). However, direct observations of weaning are often logistically 97 challenging, particularly in long-lived, wide-ranging marine species (Trillmich et al. 2016; 98 Osiecka et al. 2020; Hastings et al. 2021). Multi-state hidden Markov models (HMMs) provide a 99 powerful tool for inferring weaning transitions from mark-recapture data, allowing estimates of 100 state probabilities and the role life-history traits and environmental fluctuation may play in 101 variability (Hastings et al. 2021).

102 Using multi-state capture-recapture models (MSCRM) framed as HMMs, we investigate weaning 103 age in GSL in relation to life-history traits and environmental conditions. We predicted that low 104 resource availability, represented by the Oceanic Niño Index (ONI) anomaly, would delay 105 weaning. Moreover, we hypothesised that male offspring would wean later due to male-dominant 106 size dimorphism and that lower maternal investment and inexperience would prolong the lactation 107 period. Based on the MSCRM results we further explored the possible drivers of supersuckling 108 using logistic regression models (LRM). We theorise that supersucklers exhibit slower growth than 109 their non-supersuckling conspecifics and that poor environmental conditions during the year of 110 expected weaning may lead to delayed weaning. A better understanding of weaning patterns and 111 the mother-offspring relationship in this species is particularly relevant in the context of increasing 112 environmental variability, as shifts in resource availability may influence maternal investment 113 strategies.

114

115 **2. Methods**

117 2.1. Data collection & Study Site

118 Caamaño islet (0°45'S, 90°16'W) is located off the southern coast of Santa Cruz, Galápagos, and 119 is home to a breeding colony of GSL that has been intensively studied since 2003 (Wolf & 120 Trillmich, 2007). Data for this study were collected between February 2003 and April 2024 during 121 two annual field seasons (October-December and February-April) and grouped by reproductive 122 season (August–July), referred to by the calendar year in which each season began. All individuals 123 that are born on, or regularly use the islet are given a pit-tag, and marked using Allflex tags, 124 resulting in a database of extensive individual life history (for details on capture methodology see 125 Wolf & Trillmich 2007; Meise et al. 2013). Identification rounds were conducted daily (2-5 times), 126 during which the ID, location, and suckling events of all individuals present on the islet were 127 recorded (for details on identification rounds see Trillmich et al. 2016; and for suckling events 128 Childs et al. 2025). Missing tags were recorded whenever tag-loss was observed and used as means 129 of tracking an individual's absence from the database. Considerable effort was made to mark 130 unidentified mothers and replace missing tags to maintain life history records.

131

132 Mother-offspring pairs were identified through direct observation of birth, perinatal bonding, or 133 repeat suckling events within a month (Childs et al. 2025). Each individuals' date of birth (DoB) 134 was categorised: where 'exact' is accurate to the day, 'weak' is ± 2 weeks, and 'unknown' is 135 greater than 2 weeks uncertainty. Field-based age estimates for juveniles have proven to be highly accurate (Jeglinski et al. 2010). We therefore recategorised all individuals originally labelled as 136 137 'unknown' DOB and first marked at, or under, 6 months of age as 'roughly known' so that they 138 could be reassigned to their appropriate reproductive year of birth. All individuals born before 139 2003, when data collection first began, were categorised as DOB 'unknown'. Individuals with 140 unknown DoB were not included in any statistical modelling or testing involving age.

141

Sex verification and morphological measurements were taken every time an individual was captured. Individuals were weighed with an electric scale, and length was measured from the tip of the snout to the tip of the tail (see Mueller et al. 2011). All individuals were captured 1. for marking, 2. to observe pup growth during their first 2 months, 3. annually to measure juvenile morphological differences and cohort response to environmental variations (Wolf & Trillmich

147 2007). To avoid disturbance to the mother-offspring relationship, offspring are only captured when 148 their mother is absent. Offspring body condition was calculated using scaled mass index (SMI) 149 (Peig & Green 2009) for both birth (10-day-old or less) and early life (90-day-old mean) as 150 indicators of offspring health and maternal investment (Mueller et al. 2011). Which we take as 151 representative of maternal investment due to the offspring's nutritional dependence on milk at this 152 life stage (Jeglinski et al. 2012; Schwarz et al. 2022). For investigating growth differences between 153 supersucklers and non-supersucklers, SMI was calculated for each time point (age) where both 154 length and weight measurements were available.

155

156 Birth interval, measured as the number of years until the birth of the next sibling, was used to gain 157 insight into the effect the presence of a younger sibling may have on weaning age as sibling conflict 158 is not uncommon in this species and females are known to reject older offspring upon the birth of 159 a new pup (Trillmich & Wolf 2008). Due to the possibility that longer than average intervals (2 160 years, Kalberer et al. 2018) might be due to tag-loss, unrecorded miscarriages or stillbirths, females not observed on the islet for more than 2 years were excluded from the sample. We further cross-161 162 referenced all individuals with known multiple-year tag-loss to check for mothers who may have 163 given birth, but no offspring was assigned due to an inability to positively identify the pair. 164

165 Throughout the study, we define pups as individuals younger than 6 months, after which all GSL 166 are referred to as juveniles until they reach 5 years of age and are assumed to be adults (Trillmich 167 et al. 2016). Due to high site fidelity and limited exploration, deaths of individuals under the age 168 of 1 year were accurately recorded either by recovering the body, disappearance from the site of 169 birth and entire islet for very young individuals (< 3 months), or repeated observations of the 170 mother without her offspring over the course of a field season. Deaths of individuals older than 1 171 year are only recorded if a body is recovered and a positive ID obtained. We used correlated ONI 172 anomaly (NOAA Climate Prediction Centre 2024) and sea surface temperature ([SST] Galápagos 173 Climatology Database 2024), aggregated and averaged by reproductive year, to represent food 174 availability (Soto et al. 2004; Mueller et al. 2011).

175

176 2.2. Multi-state capture-recapture model

177 For this study weaning is defined as the point in time when an offspring stops suckling. We employed MSCRM which we frame as an HMM (see supplementary material) and used the 178 179 package LaMa (2025) in R (version 4.3.3) for building the HMM likelihood. We assumed three 180 possible latent states for offspring: unweaned, weaned, and absent (Figure 1). All animals' 181 trajectories begin in their reproductive year of birth when they are certainly unweaned. From there, 182 they may transition to the weaned state after a certain number of seasons, followed by the absent 183 state after additional season(s). Alternatively, they may transition from the unweaned state directly 184 to the absent state. We assumed irreversible state transitions such that an offspring, once considered weaned, cannot return to unweaned, and once considered absent, will stay absent. A 185 186 potential trajectory with one state per season could be U, U, U, U, W, W, A. As our primary interest 187 lies in age at weaning, we included the animal's age in years as a covariate in the state process. 188 Finding the appropriate functional form of this covariate was addressed during model selection 189 (see sp. *Methods*).



Figure 1. Both processes of the MSCRM for the observations suckling (S), observed (O), not observed (N), dead (D) generated given the states unweaned (U), weaned (W), and absent (A). Solid lines show possible state transitions. Dashed lines indicate possible observations given each state. Lines corresponding to probabilities of zero and index are omitted for visual clarity.

192 To describe the state dynamics over time, we modelled the survival rate ϕ_t for $t = 1, \dots, T$ and 193 the weaning rate v_t as functions of an animal's age. It should be noted that the survival rate refers 194 only to apparent survival from the age of one year onwards, as the absent state is comprised of 195 dead animals as well as those who were no longer seen on the islet. Given the latent state, different 196 observations could have arisen: an unweaned animal can be observed suckling with probability π , observed but not suckling with probability ρ_t , or not observed $(1 - \pi_t - \rho_t)$. In the weaned state, 197 an animal can either be observed (but not suckling) with probability λ_t , or not observed $(1 - \lambda_t)$. 198 199 Absent animals can either be recovered dead with probability ε_t or not observed $(1 - \varepsilon_t)$. We 200 accounted for observation effort by writing π_t , ρ_t , λ_t , ε_t as functions of the number of observation 201 rounds in each season t.

202

203 The variables offspring sex, ONI, offspring birth SMI (as a proxy for maternal fitness), birth 204 interval and maternal age in years (at the time of the birth of her offspring) were investigated and their effect assessed by comparing models using the specific subset both with and without the 205 206 variable included. SST was investigated during early exploration. GSL are known to be entirely 207 nutritionally dependent on milk until 12 months (Jeglinski et al. 2012; Piedrahita 2015). 208 Individuals who died before reaching 1 year of age were therefore unable to provide us with information to estimate the probability of any of the possible state transitions. Although the 209 210 MSCRM may have been flexible enough that the inference on weaning would not change, we were 211 concerned that the flexibility would become absorbed as the model attempted to explain early age 212 survival. As this study's central focus was to investigate age-at-weaning in GSL, we decided to 213 exclude all individuals registered as deceased before the age of 1 year (n = 156).

214

We used the Viterbi algorithm (Zucchini, MacDonald & Langrock 2016, Chapter 5) to obtain the most likely state sequence for each animal given the data and the model parameters. Model selection for the functional form of the age covariate and inclusion of other covariates in the state process was guided by the AIC whilst a visual inspection of results served as a final criterion for judging the relevance of added complexity.

220

221 2.3. Logistic regression model

222 LRMs were used to investigate the probability of an offspring becoming a supersuckler. We 223 classed supersucklers as offspring deemed unweaned by the selected MSCRM i.e. after the 95% 224 weaning threshold, which conforms with standard statistical practices (Wei et al. 2006; Waldmann 225 2018) and previous work on this behaviour (Childs et al. 2025). Supersucklers are therefore 4 years 226 of age or older. In order to understand the drivers behind the likelihood of individuals becoming 227 supersucklers we analysed both time-invariant and time-variant variables. The former consisted of 228 offspring age, age at the birth of a sibling, sex, SMI, and maternal age at parturition. The latter 229 were ONI anomaly average and observation effort for each reproductive year. Additionally, we 230 incorporated the decoded state of the offspring in each season based on the MSCRM. For an 231 unbiased comparison between supersucklers and non-supersucklers, we only included non-232 supersucklers who according to the MSCRM had not transitioned to the absent state at 4 years old. 233 This resulted in a dataset of 766 individuals: 650 non-supersucklers and 116 supersucklers.

234

235 Based on this data an LRM was fitted including offspring sex and ONI anomaly. A series of LRMs 236 were then built using subsets of data to investigate whether maternal fitness and experience affects 237 the probability of their offspring becoming a supersuckler: 1. maternal age and birth, 2. birth SMI, 238 90-day average SMI, and both SMI measurements, 3. birth interval, and the data for this latter 239 LRM. To control for any possible external ONI or cohort effects, we included year of birth as a 240 random effect in all models. The effect of a variable was informed by significance results and BIC, 241 which also guided model selection and parameter inclusion for all LRMs. Missings were checked 242 for informativeness by comparing the proportion of supersucklers in the data group with missing 243 values, with the data group with non-missing values for that variable.

244

245 To understand the potential reproductive cost of mothers supporting a supersuckler we visually 246 inspected the relationship between time until the birth of the next sibling and whether that sibling 247 survived to yearling age for both supersucklers and non-supersucklers. We further investigated the 248 number of female supersucklers who were reproductively active at the time of supersuckling. Here, 249 reproductively active was defined as females who either gave birth within the same reproductive 250 year as having been observed supersuckling - and were therefore pregnant at the time - or had 251 already given birth to an offspring prior to their first observed supersuckling event. Male 252 reproductive data were not available for this study. To assess whether supersucklers grow slower during early life than their non-supersuckling counterparts we plotted and visually inspected
length, weight and SMI curves of supersucklers and non-supersucklers, separated by sex and fitted
with a spline model.

256

257 **3. Results**

258 3.1. Multi-state capture-recapture model

All MSCRM model candidates and their fit can be found in Table 1. During early exploration no effect of SST was found nor was there any improvement to the fit of the MSCRMs, it was therefore excluded from all models. We found that accounting for observation effort (n = rounds per year) increased the model fit. Effort has a direct positive effect on the amount of suckling observed in the unweaned state: between 56.9% and 86.3% for min. and max. number of observation rounds, as well as the probability of observing an individual at all (33.1% - 48.2% in the weaned state). Dead individuals are also more likely to be recovered (0.2% - 2.7% in the absent state) with higher effort. **Table 1.** All fitted MSCRM described by the covariates included in the observation process via Z and the state process via Y. Number of parameters (np) and267number of individuals (nind) for which the corresponding variables were available are used as a basis for comparison via the AIC. Variable notation age_{dec} refers268to the exponential decay function of age $age^2 * exp^{-age}$, *SMI* to the SMI of an individual's first 10 days, *sibl* to time until sibling, m_{age} to the age of the mother.269Index refers to the steps in model selection. The index in bold indicates the model that was preferred at each stage of model selection.

Index	Covariates in	covariates in Y	np	nind	AIC
	Ζ				
0A	-	-	6	1890	38038.0
0B	rounds	-	10	1890	37586.4
1A	rounds	age	12	1890	37510.8
1B	rounds	age, age^2	14	1890	37091.2
1C	rounds	age, age^2, age^3	16	1890	37046.4
1D	rounds	age, age _{dec}	14	1890	37032.0
1E	rounds	$I(age_t = 1),, I(age_t = 14))$	38	1890	36821.0
2	rounds	age, age_{dec} , sex, sex \cdot $age,$ sex \cdot age_{dec}	20	1890	37032.2
3	rounds	age, age_{dec} , ONI, ONI \cdot $age, ONI \cdot$ age_{dec}	20	1890	36969.6
4A	rounds	$age, age_{dec}, ONI, ONI \cdot age, ONI \cdot age_{dec}, SMI, SMI \cdot age, SMI \cdot age_{dec}$	26	1266	24529.4
4B	rounds	age, age _{dec} , ONI, ONI · age, ONI · age _{dec}	20	1266	24526.2
5A	rounds	$age, age_{dec}, ONI, ONI \cdot age, ONI \cdot age_{dec}, sibl, sibl \cdot age, sibl \cdot age_{dec}$	26	317	5961.2
5B	rounds	$age, age^2, ONI, ONI \cdot age, ONI \cdot age^2$	20	317	5993.4
6A	rounds	age, age_{dec} , ONI, ONI · age, ONI · age_{dec} , m_{age}, m_{age} · age, m_{age} · age_{dec}	26	328	no convergence
6B	rounds	age, age _{dec} , ONI, ONI · age, ONI · age _{dec}	20	328	no convergence

271 Offspring effects

272 When including age as a functional form (MSCRM 1D), the preferred approach was 273 age²·exp(-age). Although estimating an effect separately for each age category was supported by 274 AIC, visual inspection showed no meaningful differences to model 1D. As less data was available 275 for older ages additional complexity was deemed unnecessary. Exponential decay improved the 276 overall fit compared to simple polynomials without increasing the number of parameters. We 277 therefore favoured the model in its simplest form, using only exponential decay and the polynomial 278 effect as parameters. This ensured the flexibility of the polynomial function is prioritised for 279 younger ages whilst assuming the effect in older ages is captured sufficiently by the linear term. 280 Viterbi decoding of model 1D provided age-at-weaning estimates of 50% by 1.5 years, 90% by 3 281 years, and 95% by 4 years (Figure 2). The highest number of state switches was observed between 282 1 - 2 years of age and rapidly decreased after this point (Figure S1). The highest number of switches to the weaned state was observed at 1.5 years of age. Switches to the absent state were extremely 283 284 elevated throughout the first 2.5 years, peaking at the yearling stage and only levelling out towards 285 the end of the 3rd year.



Figure 2. Viterbi state proportions from Model 1D showing the most likely state (unweaned [U], weaned [W], absent [A]) an individual is in at any given time point given their observational history and age. White-boxed numbers within the figure indicate the total number of possible observable animals at 5-year intervals.

287 Due to increasing exploration and time spent at sea, the likelihood of recovering a body and 288 obtaining a positive ID for death registration dwindles dramatically with age. Over the entire data 289 collection period 247 deaths have been recorded, of which 63% (n = 156) were individuals under 290 1 year of age. Only 40 adult deaths have been recorded. That these numbers are so low is likely due 291 to most individuals within this population dying at sea or temporary haul-out locations. Without 292 positive confirmation of these deaths' individuals remain in the dataset even after multiple years of 293 absence. It should therefore be noted that the sample sizes (Figure 2 & 3) indicate the total number 294 of possible observable individuals.

295

The selection of additional covariates was guided by a combination of AIC, BIC, and contextual considerations. Among animals that could be observed, both sexes were relatively evenly represented. We found no apparent effect of sex on weaning age in MSCRM 2, and both visual inspection and AIC indicated that including sex as a covariate was unnecessary. However, Viterbi decoding revealed a slight effect on the probability of females transitioning to the absent state later in life, with males more likely to transition at younger ages (Figure 3).





Figure 3. Viterbi state proportions from Model 2 showing the most likely state (unweaned [U], weaned [W], absent [A]) of an individual at any given time point is in according to sex. White-boxed numbers within the figure indicate the total number of possible observable animals at 5-year intervals.

304 Environmental effects

317

305 Visual inspection and AIC results of MSCRM 3 indicated that ONI should be retained as a 306 covariate and a visualisation of the relationship (Figure 4) showed that there is a strong effect on 307 juvenile survival between the ages of 1 - 4 years. This can be seen in the high variation in the 308 number of individuals in both the unweaned (U) and absent (A) states between El Niño and La 309 Niña conditions. During poor El Niño conditions individuals are much less likely to be in the 310 unweaned state and more likely to transition to the absent state, indicating a reduced survival rate 311 during these events compared to more favourable La Niña periods. No probability on survival prior 312 to yearling age can be inferred due to the exclusion of these deaths from the model. The probability 313 of individuals being observed in the weaned (W) state is also affected by ONI between the ages of 314 2 - 4, where during El Niño individuals between these ages are more likely to be observed in this 315 state than during La Niña conditions. The effect of ONI on all states reduces with age showing 316 only a slight influence on weaned and absent between the ages of 7-16.



Figure 4. ONI category effect on state behaviours unweaned (U), weaned (W) and absent (A). Where the interaction during an Intermediate ONI event is represented by a solid line and El Niño (pale) and La Niña (dark) as points along this line.

319 Maternal effects

When investigating maternal life-history effects, information on mothers was unavailable or incomplete for most offspring, resulting in different subsets when incorporating variables related to an individual's mother. This provided the base for model fitting.

323

324 Including birth SMI as a covariate is neither supported by AIC nor did the weaning rate in the first 325 few years of an individual differ given different SMI values. However, SMI did appear to have an 326 effect on offspring survival (Figure S2) with survival rates estimated between 78% and 96% for 327 offspring with minimum and maximum SMI values in the data, respectively. When birth interval 328 was included as a covariate (MSCRM 5A & 5B) AIC results indicated a slight preference for its 329 inclusion. The MSCRM suggests that a longer interval before the birth of the next sibling is 330 associated with later weaning (Figure S3) with about 98% of offspring with an early sibling (born 331 in the same year) weaned by age, but only 61% of offspring with a late sibling (born 7 years later) 332 based on a simulated state distribution given the model.

333

334 The inclusion of maternal age in MSCRM 6A and 6B produced unstable results, preventing the 335 identification of a global maximum. This was likely due to the small sample size (n = 328) and 336 specificity of the subset. Due to the colony's decline a large proportion of the females who 337 recruited and successfully reproduced on Caamaño were first identified as juveniles or adults 338 during the early years of data collection and therefore have unknown DOB (personal 339 *communication*). Simplification of the models failed to yield more stable results; we were therefore 340 unable to draw any meaningful conclusions of the effect this life-history trait may have on weaning 341 age.

342

343 **3.2.** Logistic regression model

All LRM model candidates and their fit can be found in Table 2. Supersucklers and nonsupersucklers were classified according to the offspring decoded states of MSCRM 3. Due to the lack of effect of SST in the MSCRM at an annual scale, SST was not included in any LRM. LRM 0B confirmed that sex has no significant effect on the probability of becoming a supersuckler. Averaged ONI anomalies were included for the year of birth and the first four years of their life as a covariate, in order to understand whether ONI during offspring's early life or the year they become a supersuckler are more important. We found that an increase in ONI anomaly during an individual's 2nd and 3rd year of life has a negative impact on their probability of becoming a supersuckler. Yet the inclusion of ONI anomalies did not greatly improve the model, with BIC indicating that no covariate inclusion was a better fit. In all other LRMs we therefore included the year of birth as a random effect to control for any cohort effect, and only ONI anomalies of the 2nd and 3rd year as covariates.

356

Table 2. All fitted LRMs predicting the probability for an individual to become a supersuckler described by the predictors and random effect included. The number of parameters (np) and number of individuals (nind) for which the corresponding variables were available are used as a basis for comparison via the BIC. Variable notation ONI_0 – ONI_4 refers to the oscillation index at the given age of an individual in, SMI_{10} , SMI_{90} to the SMI of an individual's first 10 or 90 days, *sibl* to time until sibling, m_{age} to the age of the mother. Index refers to the steps in model selection. The index in bold indicates the model that was preferred at each stage of model selection.

Index	random	predictors	BIC	nind	np
	effect				
0A		-	658.04	766	1
0B		sex, ONI ₀ , ONI ₁ , ONI ₂ , ONI ₃ , ONI ₄	677.24	766	7
1A	cohort	m_{age}	102.91	98	3
1B	cohort	-	98.41	98	2
2A	cohort	SMI ₉₀	425.74	492	3
2B	cohort	<i>SMI</i> ₁₀ , <i>SMI</i> ₉₀	430.38	492	4
2C	cohort	SMI ₁₀	426.44	492	3
2D	cohort	-	422.27	492	2
3A	cohort	sibl	99.87	140	3
3B	cohort	-	95.97	140	2

365 The results of LRM 1A indicate that maternal age at birth does not play an important role in the 366 probability of offspring becoming a supersuckler. The proportion of offspring without a known-367 age mother was similar for supersucklers and non-supersucklers, however, we still assume some 368 structure in the missingness that might not have been addressed in this model setup. Within this 369 LRM only a few offspring with older mothers (+12 years) were included; since the data collection 370 period did not span more than a single lifetime, the younger the mother, the more likely her age 371 was to be in the data set. We therefore concluded that if there was an effect, we might not be able 372 to find it, and that this variable should be excluded from further analysis.

373

374 We found no relationship between becoming a supersuckler and SMI at birth or early life SMI and 375 so excluded this variable from further analysis. It should be noted that the proportion of missing 376 values for SMI was slightly higher for supersucklers than non-supersucklers, likely a reflection of 377 colony proportional differences. There was also no indication of a relationship between birth 378 interval to the next sibling and the probability of an offspring becoming a supersuckler. Due to the 379 exclusion of mothers with recorded absence (2+ years) from the database and the previously stated 380 data imbalance, the proportion of missings for supersucklers was again much higher than for non-381 supersucklers. Additional visualisation of birth interval and survival of the next sibling to yearling 382 age (Figure S4), suggested that there may be some relationship between sibling first-year survival 383 and older offspring becoming a supersuckler but further highlighted the high number of missing 384 values. We concluded that, due to the small and imbalanced sample size, it was unlikely that we 385 might be able to find the effect if there was one and that this variable could be excluded from 386 further analysis.

387

Visual inspection of weight, length and SMI curves (Figure S5) showed high overlap between the predicted means for pups who would become supersucklers and those who would not, indicating no difference in growth between supersucklers and non-supersucklers. The widening of the confidence intervals (95%) from 700 days (1.9 years) onwards indicates that the curve becomes implausible at this point due to the reduced sample size.

- 393
- **394 Conflict of Interest**
- 395 None declared.

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- 407 Supplementary material
- 408

409 **1. Methods: Model formulation of the MSCRM**

The multi-state capture-recapture model (MSCRM) belongs to the class of hidden Markov models
(HMMs) comprising two stochastic processes: the state process and the observation process.

- 412 HMMs assume that (partially) hidden states determine the distribution of the observations at each
- 413 point in time.
- 414

415 <u>State Process:</u>

The state process $\{S_t\}_{t=1,\dots,T}$ is defined as a Markov chain with s_t the state at season t. It aims to 416 describe the weaning process by distinguishing the unweaned state "U" and the weaned state "W". 417 418 Additionally, the absent state "A" accounts for events of death and emigration. This absorbing state ensures that dead or absent animals which cannot be observed anymore are not considered 419 420 weaned and the weaning process is therefore only informed by animals that can be observed. The 421 upper half of the directed graph in Figure 1 shows the possible state transitions. Notably, once the 422 unweaned state is left, an animal cannot transition back to it. The same applies for the weaned 423 matrix $\Gamma =$ state. Therefore. the following transition probability $(\gamma_{UU} \gamma_{UW} \gamma_{UA} 0 \gamma_{WW} \gamma_{WA} 0 0 1) = (\phi \cdot (1 - \nu) \phi \cdot \nu 1 - \phi 0 \phi 1 - \phi 0 0 1)$ 424

is assumed with γ_{kl} denoting the probability of switching from state *k* to *l* for all $k, l \in \{U, W, A\}$. 426

In defining a survival rate that reflects the probability that an animal continues to be observable, and the assumption that survival rate is independent of the current state, we could then rewrite the entries in Γ in terms of this survival rate ϕ and a weaning rate ν . In practical terms, by estimating only γ_{UU} and γ_{UW} and considering that $\gamma_{UU} + \gamma_{UW} + \gamma_{UA} = 1$, we can derive $\phi = \gamma_{UU} + \gamma_{UW}$ and $\nu = \frac{\gamma_{UU}}{\gamma_{UU} + \gamma_{UW}}$.

432

433 Covariates that were assumed to affect the weaning and survival rate (ν, ϕ) could be included in 434 the state process leading to an inhomogeneous Markov chain $\{S_t\}$ with Γ_t and therefore (ν_t, ϕ_t) 435 varying over time. In order to do so, the transition probabilities to be estimated are linked to linear 436 predictors $\tau_{j,t} = Y_t \beta_j$ for j = 1,2 and t = 1, ..., T#(1) given the covariate matrix Y_t and the

437	regression	coefficients	β_j	such	that	$\gamma_{UU,t} = \frac{expexp(\tau_{1,t})}{1 + expexp(\tau_{1,t}) + expexp(\tau_{2,t})}$	and	$\gamma_{UW,t} =$
438	$\frac{expe}{1+expexp(\tau_1$	$exp(\tau_{2,t})$ $exp(\tau_{2,t}) + expexp(\tau_{2,t})$).					

440 <u>Observation process:</u>

The observation process $\{X_t\}_{t=1,...,T}$ depends on the underlying state process $\{S_t\}$ as depicted in the bottom half of Figure 1. The observations $x_{t,q}$ correspond to being observed suckling (S), being observed but not suckling (O), not being observed (N), or recovered dead (D) for each field season t and each animal q.

445

446 We constructed three state-dependent probability distributions as follows:

447
$$X|j \sim Categorical(q_{\cdot|j}), \text{ for } j = U, W, A$$

448
$$q_{\cdot|U} = (\pi, \rho, 1 - \pi - \rho)$$

449
$$q_{\cdot|W} = (0, \lambda, 1 - \lambda, 0)$$

$$q_{\cdot|A} = (0,0,1-\epsilon,\epsilon)$$

451

Given the unweaned state U, we obtained a categorical distribution where individuals who were observed suckling had probability π , observed doing something else had probability ρ and unobserved with probability $1 - \pi - \rho$. Given the weaned state W and the absent state A, the distributions are essentially Bernoulli with λ the probability of being observed given the weaned state and ϵ the probability of being recovered dead.

- 457
- 458 We collect variables that are assumed to affect the resighting probability and thereby the 459 distribution of observations in a covariate matrix Z_t and build a linear $\eta_{i,t} = Z_t \alpha_i$ for

460 i = 1,2,3,4 and $t = 1, \dots, T$ with $\alpha_{i,t}$ the regression coefficients.

461

462 The four parameters $\rho_t, \pi_t, \lambda_t, \epsilon_t$ are then time varying by being linked to the linear predictors as 463 follows:

464
$$\rho_t = \frac{exp \exp(\eta_{1,t})}{1 + exp \exp(\eta_{1,t}) + exp \exp(\eta_{2,t})}$$

465
$$\pi_t = \frac{exp \ exp \ (\eta_{2,t})}{1 + exp \ exp \ (\eta_{1,t})} + exp \ exp \ (\eta_{2,t})}$$

466
$$\lambda_t = \frac{1}{1 + exp \ exp \ (\eta_{3,t})}$$

467
$$\epsilon_t = \frac{1}{1 + \exp \exp(\eta_{4,t})}$$

469 <u>Parameter estimation and state decoding:</u>

By framing the model as a hidden Markov model, we leveraged inferential tools like the forward algorithm to estimate the latent weaning process and the Viterbi algorithm to decode the most likely state sequence. Note that we constructed a partially hidden Markov model in the sense that a suckling observation "S" and a death recovery "D" inform directly about the state having generated this observation. We used the package LaMa in R for conveniently building the HMM likelihood and maximized the likelihoods directly by using the numerical optimizer nlm.

476

477 **2. Results**



Figure S1. Total number of switches between each state according to offspring age: unweaned-weaned (UW, blue), unweaned-absent (UA, light green), weaned-absent (WA, dark green).



Figure S2. Apparent survival probability (A) of offspring according to their birth SMI, indicated by red (low birth SMI), blue (mean birth SMI) and green (high birth SMI) lines, and their age. As well as their probable age at weaning if they survive (B).



Figure S3. Birth interval effect - calculated as time until the birth of the next sibling - on state behaviours unweaned (U), weaned (W) and absent (A) according to offspring age. Indicated by solid lines where red is a short interval (early sibling = 1 year), blue is a mean interval and green a long interval (late sibling = 7 years) they survive (B).



Figure S4. A visualisation of the possible relationship between supersuckling and younger sibling survival to yearling age, where all younger siblings are plotted as jitter points according to the birth interval between them and the elder offspring and grouped according to the elder offsprings classification as a supersuckler or non-supersuckler. Younger sibling survival to yearling age is denoted by the colour of their point where blue is 'survived' and red is 'did not survive'.



Figure S5. The thick lines show the predicted average **a**) body weight, **b**) length and **c**) SMI given the age and sex for individuals later becoming supersucklers (orange) and non-supersucklers (purple). The dashed lines indicate the 95% confidence intervals of the predicted mean.

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