

1 Measuring natural selection on the transcriptome

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32 **Summary**

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34 The level and pattern of gene expression is increasingly recognized as a principal
35 determinant of plant phenotypes and thus of fitness. The estimation of natural selection
36 on the transcriptome is an emerging research discipline. We here review recent
37 progress and consider the challenges posed by the high dimensionality of the
38 transcriptome for the multiple regression methods routinely used to characterize
39 selection in field experiments. We consider several different methods, including
40 classical multivariate statistical approaches, regularized regression, latent factor
41 models, and machine learning, that address the fact that the number of traits potentially
42 affecting fitness (each expressed gene) can greatly exceed the number of plants that
43 researchers can reasonably monitor in a field study. While such studies are currently
44 few in number, extant data is sufficient to illustrate several of these approaches. With
45 additional methodological development coupled with applications to a broader range of
46 species, we believe prospects are favorable for directly characterizing selection on gene
47 expression within natural plant populations.

48 **Introduction**

49

50 One of the fundamental goals of evolutionary biology is to understand how natural
51 selection acts on phenotypes. Understanding the form, strength, and direction of
52 selection is crucial to making predictions about the evolutionary trajectory of traits,
53 understanding adaptation, and quantitatively testing alternative hypotheses about the
54 extent to which organismal features evolve by adaptive or non-adaptive mechanisms.
55 For this reason, evolutionary biologists have devoted considerable effort to measuring
56 natural selection in field, experimental, and common garden environments (Kingsolver
57 et al. 2001, 2012; Siepielski et al. 2013). While the rapid progress in molecular biology
58 and genomics continually offers the promise of characterizing the genetic basis of
59 complex traits (Hill 2010), there is a growing realization that these techniques and
60 approaches yield a suite of molecular phenotypes that are themselves amenable to
61 evolutionary (and genetic) analysis. Here we outline the prospects and challenges for
62 characterizing natural selection on one particularly relevant– and increasingly
63 attainable– set of molecular phenotypes, gene expression.

64

65 Several lines of evidence suggest that gene expression is an important determinant of
66 organismal fitness, and thus likely to experience selection. Early experimental results,
67 from mutation accumulation experiments in which the strength of selection has been
68 minimized or reduced, suggested that stabilizing selection was acting on gene
69 expression (Rifkin et al. 2005, Gilad et al. 2006). Likewise, observations from the
70 microarray-era indicated that populations experiencing different environmental
71 conditions can diverge in gene expression, even in the face of substantial gene flow
72 (Oleksiak et al. 2002), potentially indicating the past action of selection. More recent
73 studies have demonstrated changes in gene expression in response to severe weather
74 events like cold snaps (Campbell-Staton et al. 2017), between ancestors and surviving
75 descendants of droughts (Hamann et al. 2021), and in response 3-4 generations of
76 experimental evolution in the field under new ecological conditions (Ghalambor et al.
77 2015). Collectively, these and other studies reveal that gene expression can and does
78 evolve on a wide array of time-scales, including in the laboratory (Rifkin et al. 2005),
79 between adjacent populations of the same species (Oleksiak et al. 2002), in response to
80 severe weather events (Campbell-Staton et al. 2017; Hamann et al. 2021), and in
81 ecologically realistic, complex communities within a handful of generations (Ghalambor
82 et al. 2015, 2018).

83

84 Despite prominent examples of gene expression evolution on microevolutionary
85 timescales, as well as theorizing on its relevance on macroevolutionary time scales
86 (e.g., King and Wilson 1975), we have few direct estimates of natural selection on gene
87 expression. In contemporary populations, is gene expression subject to stabilizing
88 selection as first predicted, or is it frequently subject to directional selection as might be

89 deduced from these studies of evolutionary divergence on short time scales? How does
90 the strength of selection on gene expression compare to that on 'macroscopic' traits
91 such as life history, morphology, or behavior? Are the levels of transcription among
92 multiple genes in the transcriptome sufficiently correlated as to require distinguishing
93 between direct and indirect selection? Is there a relationship between the level of
94 expression and the strength of phenotypic selection, analogous to the relationship
95 between the level of expression and rates of molecular evolution (Wright et al. 2004,
96 Slotte et al. 2011)? These and a host of other questions require extending the Lande-
97 Arnold revolution (Lande and Arnold 1983; Svensson 2023) from traditional
98 macroscopic phenotypes to include gene expression.

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100 **Transcriptomes as Quantitative Traits**

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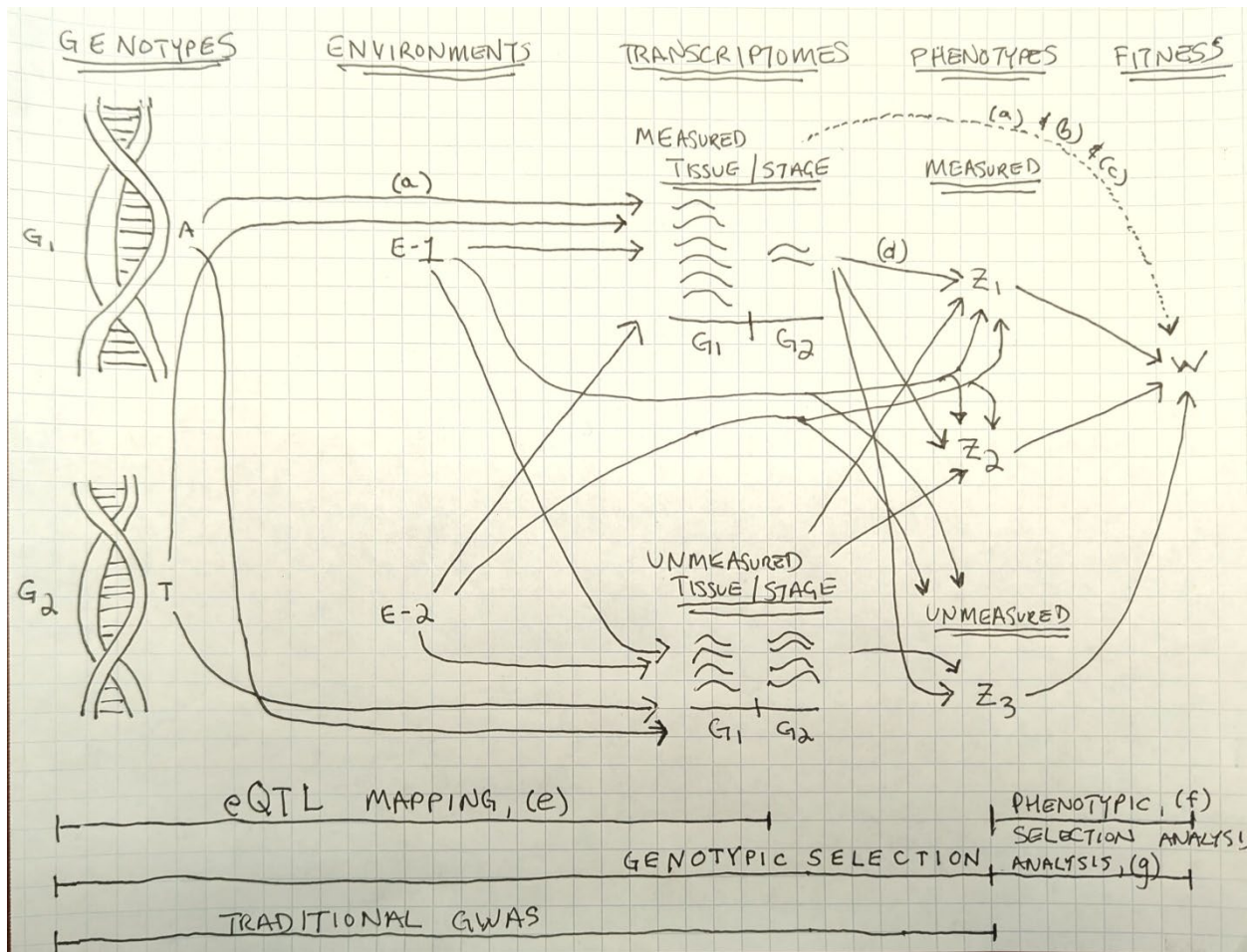
102 Progress on these questions starts with the recognition that gene expression is itself a
103 quantitative trait. The expression levels of genes across the genome are quantitative
104 traits with strong environmental influences combined with multi-locus genetic effects
105 (Liu et al. 2019). In fact, given that modern RNA-seq experiments often obtain
106 expression estimates for many genes simultaneously (N in the 1,000s), the
107 transcriptome is really a collection of vectors (or a matrix), as expression levels of a
108 gene change with tissue, lifestage, and the expression of other genes. Considering the
109 transcriptome as a set of correlated characters with a quantitative genetic perspective
110 offers several insights. First, the transcriptome is hugely multivariate and thus offers
111 investigators a chance to measure a large number of phenotypes simultaneously. In
112 addition, by characterizing the phenotype in the broad sense, the traits measured are
113 less prone to bias about the types of traits that might be important or under strong
114 selection (e.g., size versus floral traits, those involved in mating versus anti-herbivore
115 defense), although important choices must still be made about the time and tissue
116 sampled. Perhaps most importantly, there is a well-developed machinery to analyze
117 selection on correlated quantitative traits (Lande & Arnold, 1983; Rausher, 1992).
118 However, the application of quantitative genetics to the transcriptome must confront the
119 serious challenge of scale.

120

121 Most studies of phenotypic selection utilize a regression framework. In the simplest
122 implementation of this approach, an estimate of relative fitness (e.g., individual seed set
123 divided by mean seed set for the population) is regressed on a single phenotype in a
124 univariate regression. In the context of gene expression, this would involve regressing
125 an estimate of relative fitness on the expression of an individual gene, for all the
126 individuals in the experimental population or sample. If expression has been
127 standardized (i.e., $\bar{x} = 0$ and $\sigma = 1$), the resulting parameter estimate is the standardized
128 selection differential for the expression of that gene; positive values would indicate that

129 greater expression of the gene was associated with increased relative fitness. Groen et
130 al. (2020) applied this approach to populations of rice growing under field and drought
131 environments. They found that selection differentials for gene expression were generally
132 weak, but stronger under drought than well-watered conditions.
133

134 The scale of the transcriptome introduces two key problems with the univariate
135 approach. First, RNA-seq experiments estimate the expression of thousands of genes
136 at a time. Simply repeating a univariate analysis for all the genes for which one has data
137 introduces severe multiple testing problems: A large number of genes associated with
138 relative fitness will undoubtedly be false positives. Addressing the number of tests thus
139 requires multiple testing or false discovery rate corrections. Second, selection
140 differentials measure total selection on a phenotype, which is the sum of direct selection
141 on the trait and indirect selection through correlated traits (Lande and Arnold 1983).
142 Because the expression of any individual gene is likely to be correlated with the
143 expression of other genes (and other traits), a selection differential alone cannot tell
144 whether it is the expression of a focal gene that is directly important for relative fitness,
145 or whether the expression of that gene is simply correlated with other traits that are
146 under selection.
147



148
 149 **Figure 1. A schematic depicting the mapping from genotype to fitness.** From left to right
 150 we present a hypothetical case of two genotypes (G_1 and G_2) and two
 151 environments (E_1 and E_2) to illustrate how genetic variation and environmental variation
 152 affect transcriptomes, phenotypes, and fitness. In the middle column, we highlight that
 153 genetic and environmental variation may lead to differences in expression in some
 154 tissues and stages (measured tissue/stage) but not others (in this case, an unmeasured
 155 tissue/stage). Expression and environmental variation, in turn, both affect macroscopic
 156 phenotypes (z_1 , z_2 , z_3). In this case, we highlight that while z_1 and z_2 have been
 157 measured, it is likely that unmeasured phenotypes (z_3) are affected by expression and
 158 also affect fitness. In the arrows leading to fitness, we note that expression can affect
 159 fitness directly (dotted arrow) and via phenotypes (z_1 and z_2). Bars across the bottom
 160 are labeled with common analytical approaches to understanding expression, the
 161 genetic basis of traits, and selection. Key paths: (a) Groen et al. (2020), (b)
 162 Stinchcombe and Henry (2025), (c) Figure 2, this paper, (d) Brown and Kelly (2022), (e)
 163 Josephs et al. (2015, 2020), (f) Lande and Arnold (1983), (g) Rausher (1992).

164
 165

166 The standard approach for measuring selection on (potentially) correlated traits involves
 167 multiple regression (Lande and Arnold 1983; Figure 1 path (f)). In this context, a
 168 regression of relative fitness on the expression of all the genes in the transcriptome

169 would yield selection gradients for gene expression. These gradients measure the direct
170 effect of expression on relative fitness, accounting for the effects of the other traits (i.e.,
171 expression of other genes) included in the model. While promising in the abstract, with
172 real data and sample sizes, such a model quickly runs into the N - p problem: there are
173 far more parameters to estimate (p) than there are total samples (N) in even the most
174 heroic of experiments. Consequently, one of the primary advantages of the Lande-
175 Arnold (1983) approach—its ability to distinguish direct and indirect selection on
176 correlated traits—is lost. In the remaining sections, we outline a handful of promising
177 statistical and experimental approaches that can be used to address the N - p problem
178 of measuring selection gradients for transcriptomes.

179

180 **Selection Gradients for the Transcriptome: Statistical Approaches**

181

182 There are several statistical approaches for measuring selection gradients for gene
183 expression, and here we comment on some variants that appear to be emerging in the
184 literature. Our expectation is that there will be continued work in this area, and that
185 future developments are likely. At their core, these methods share one fundamental
186 feature: dimensionality reduction, the compression of the data so as to estimate fewer
187 parameters than the sample size. To use a hypothetical example, if an investigator has
188 estimates of fitness for 500 individuals, and estimates of expression for >500 genes in
189 those same 500 individuals, the goal of these approaches is to reduce the problem to
190 estimating selection from far fewer than 500 parameters (so that N is greater than p).

191

192 *PCA and Gene Coexpression Modules*

193

194 The most straightforward approach is likely familiar to many users of selection gradient
195 analysis— principal component analysis (PCA). Because PCA is a widely used
196 technique and familiar to many biologists, we do not consider the mathematical or
197 technical details of its implementation; Jolliffe (2002) provides an extensive coverage. In
198 short, after a PCA, an investigator obtains independent axes capturing variation in the
199 original traits. In many cases, far fewer axes (PCs) are required to describe the data
200 than there were original traits. In the context of gene expression, these PC axes can be
201 used as independent variables to predict relative fitness. An important point is that
202 fewer— ideally far fewer— PC axes have to be used than there were original traits,
203 otherwise nothing is gained. Groen et al. (2020; Figure 1, path (a)) used this approach
204 with PC axes, and were able to detect significant selection on several PC axes. They
205 used these findings to detect selection on the expression of genes related to
206 photosynthesis and growth.

207

208 One downside of this approach is that a PC axis reflects— simultaneously— all the
209 individual traits included in the study. A PC score is a weighted average of the
210 expression of all measured genes with the magnitude and direction of the weights
211 differing among principal components. This can make PC scores difficult to interpret.
212 Chong et al. (2018) illustrate a method to ‘back-transform’ selection estimates for PC
213 scores into selection estimates on the original traits. They argued these are much easier
214 to interpret and suggested the technique would be useful for studies of selection on
215 gene expression, metabolomics, and other high dimensional traits. In brief, one
216 performs some matrix algebra computations involving selection estimates for PC scores
217 and the eigenvectors of the original PCA. This rotation yields an estimate of a selection
218 gradient on individual gene expression traits, accounting for the patterns of correlation
219 among the traits, but only within the portion of multivariate space described by the PC
220 axes included (Chong et al. 2018). Similar calculations can be performed to estimate
221 standard errors for these reconstituted estimates of selection gradients for gene
222 expression.

223

224 Henry and Stinchcombe (2025, Figure 1, path (b)) also used PCA to understand
225 selection on gene expression. Like Groen et al. (2020), they regressed relative fitness
226 on PC axes of gene expression. However, rather than using the PC axes as objects of
227 study in themselves, they used the methods described by Chong et al. (2018) to back-
228 transform selection on PC scores into selection gradient estimates for individual genes.
229 In their study of *Ipomoea hederacea* (Ivyleaf morning glory), they had estimates of
230 relative fitness for 96 individuals, and estimates of gene expression for 2,753 genes
231 throughout the genome. The best model used 61 PCs to describe patterns of variation
232 in gene expression, which collectively explained 55% of variation in relative fitness.
233 Turning these back into selection gradients for the expression of individual genes
234 suggest several important, if tentative, findings about selection on gene expression.
235 First, they found a very strong positive relationship between selection differentials and
236 selection gradients for gene expression, suggesting that overall in their study much of
237 the selection on gene expression was direct, rather than indirect due to the expression
238 of other genes. Second, they found a wide distribution of selection gradients for
239 expression, approximately symmetrical around zero: some genes were under selection
240 for increased expression, and a similar number for decreased expression. Finally, they
241 observed that selection gradients for gene expression were substantially smaller than
242 their past findings of selection on size and life history traits in the same population
243 (Henry and Stinchcombe 2023).

244

245 An alternative approach to dimensionality reduction is to first identify gene co-
246 expression modules using programs like WGCNA (Langfelder and Horvath 2008).
247 These modules are constructed by identifying sets of genes whose expression is more

248 tightly correlated with other genes in the module than genes in other modules. The
249 expression of the genes within a module can be summarized with PCA– so-called
250 eigen-genes– and the PC1 score for the first eigenvector of a module can be estimated
251 for each individual in a data set. These PC scores represent a weighted sum of gene
252 expression of the genes within the module. As before, PC scores for a module’s
253 expression– which might summarize the expression of dozens to hundreds of genes–
254 can be used as ‘traits’ in Lande-Arnold style analyses.

255
256 Several investigators have applied this approach, relating gene coexpression module
257 PC scores to aspects of plant performance, size, or life history traits that are likely to be
258 under strong selection (e.g., Palakurty et al. 2018, Josephs et al. 2020, Brown and Kelly
259 2022). For example, Brown and Kelly (2022; Figure 1 path (d)) found that PC1 scores
260 from twenty gene coexpression modules could explain 47% of variation in flower size in
261 *Mimulus guttatus*. They used permutation testing to verify that these modules indeed
262 significantly predicted flower size, and that the observed co-expression modules
263 performed significantly better than random groupings of genes of the same size. In
264 other words, the coexpression modules contain biological signal in predicting traits (in
265 this case, flower size) beyond a random grouping of genes in the transcriptome. Flower
266 size is not itself a fitness component, but is under strong selection in *Mimulus guttatus*
267 (Mojica and Kelly 2010), suggesting that transcriptomic variation affecting flower size
268 can also potentially affect fitness. Interestingly, while several studies have related eigen-
269 gene expression from coexpression modules to performance and fitness traits, to our
270 knowledge none have used the PC rotations of Chong et al. (2018) to estimate selection
271 gradients for expression of the individual genes within the module.

272
273 The use of gene co-expression modules entails a few benefits and drawbacks that are
274 worth considering. Co-expression modules have the benefit that individual genes
275 appear in one and only one module. As a consequence, the interpretation of the
276 expression of the entire module is more straightforward than the output of a PCA, where
277 the expression of each gene will load onto all the PC axes. Discrete, non-overlapping
278 modules, in our view, might offer greater biological interpretation of the types of genes
279 (or GO categories) that are associated with any given module. One drawback of
280 coexpression modules, or PC scores summarizing the expression levels of genes within
281 a module (eigen-gene expression), is that the scores summarizing multiple modules are
282 not guaranteed to be uncorrelated across a sample, in contrast to a PCA using all of the
283 data. As a consequence, understanding selection on multiple modules simultaneously
284 requires a multiple regression and the estimation of selection gradients.

285
286 *Machine Learning*

287

288 There is great enthusiasm for machine learning approaches in evolutionary biology
289 (Schridder and Kern 2018). While this field is moving quickly and a full review is beyond
290 our scope here (see Schridder and Kern 2018 for an entry point), there are several
291 features of these algorithms that suggest promise in the context of measuring selection
292 on gene expression. Machine learning approaches often focus on overall prediction
293 rather than individual parameter estimation. In this context, it would be to predict relative
294 fitness from expression of the set of genes for which investigators have expression,
295 rather than hypothesis testing about the individual contribution of any one gene's
296 expression. Several features of the mechanics of how the algorithms work aid this. First,
297 data are often split into "training" and "testing" sets, which can prevent overfitting and
298 noise being fit to the model, and allow an evaluation of the overall performance of the
299 model. Second, many of the approaches identify features (gene expression in this case)
300 in a way that reduces the overall number of parameters that are estimated, which is a
301 start towards addressing the issue of the scale of the transcriptome. Third, in many
302 cases the output of a machine learning algorithm is a measure of importance, rather
303 than a parameter estimate like a selection differential or gradient— for example, the
304 expression of these genes are important in determining whether an individual survives
305 or dies before reproduction.

306
307 Assuming that as an evolutionary biologist one has managed to implement one of the
308 many machine learning algorithms available, and obtained a list of genes (features)
309 whose expression is related to a fitness component, how does one make that
310 information compatible and conversant with traditional measures of selection like
311 differentials or gradients? One potential way forward is to use this reduced set of
312 genes— that having survived cross-validation, evaluation in the testing data set, and
313 acceptable performance metrics— appear to have expression that predicts relative
314 fitness to estimate selection differentials and gradients the traditional way. In other
315 words, one can use machine learning algorithms to identify an important subset of
316 genes to focus on, and then traditional selection analysis to estimate selection
317 differentials and gradients.

318
319 In Henry and Stinchcombe's (2025) study, they used machine learning classification
320 algorithms to determine which genes' expression were important for determining
321 whether an individual set seed versus failed to set seed. After model fitting, they
322 identified 278 genes whose expression was identified as important for determining
323 whether an individual set seed or failed to set seed; 29 of these genes were also
324 identified with PCA, having strong selection gradients for their expression. Interestingly,
325 the distribution of selection differentials and gradients for the expression of these 29
326 genes was bimodal, with few instances of weak (near-zero) selection. In other words,
327 the machine learning classifier identified genes whose expression was important for

328 successfully setting seed and these genes showed the strongest patterns of phenotypic
329 selection.

330

331 *Regularized regression*

332

333 Many evolutionary biologists (including ourselves!) find aspects of machine learning to
334 be a bit of a black box: its hard to fully visualize the functions and models being fit by
335 the algorithms. This is especially in the case of neural networks where the output of one
336 function is used as the input for another, in a series of layers. Fortunately, there's a set
337 of statistical techniques closely related to machine learning– and indeed used by some
338 machine learning algorithms– that is closer to the typical statistical toolkits of practicing
339 evolutionary biologists While to our knowledge regularized regression has not been
340 used to estimate selection on gene expression, several features suggest that it could be
341 useful.

342

343 Regularized regression is a useful analytical tool for fitting regressions with many
344 predictors, varying degrees of multicollinearity between the predictors, and limited data
345 (Morrissey 2014; Sztepanacz and Houle 2024). In contrast to ordinary least squares
346 (OLS) univariate or multivariate regressions which estimate parameters by minimizing
347 the sum of squared errors, regularized regressions minimize functions which include a
348 penalty (Morrissey 2014; Sztepanacz and Houle 2024). A consequence of this is that
349 individual parameter estimates are shrunk towards zero (i.e., regularized), which also
350 reduces their variance. Parameter estimates obtained from regularized regression are
351 thus biased compared to least-squares estimates, but the overall model predictive
352 accuracy can be improved, in the presence of a bias-variance trade-off. For these
353 reasons, regularized regression approaches are likely to be of use in the case of
354 multicollinearity (Chong et al. 2018; Sztepanacz and Houle 2024).

355

356 Sztepanacz and Houle (2024) performed a simulation study that illustrates the potential
357 utility of regularized regression for measuring selection on multiple, potentially highly
358 correlated traits. While their focus was not on gene expression, the lessons likely apply.
359 They showed that in the face of limited data, and multicollinearity between predictors
360 (as might be expected with the expression of thousands of genes as traits), regularized
361 estimates provided more accurate estimates of the total strength of selection and the
362 overall multivariate direction of selection. The frequentist implementation of regularized
363 regression, however, does not yield traditional measures of uncertainty like standard
364 errors and statistical significance for the individual predictors (Morrissey 2014;
365 Sztepanacz and Houle 2024). While this is a potential limitation for future meta-analyses
366 which require estimates of uncertainty for parameters, it is important to note that the
367 importance of a gene's expression in predicting relative fitness can be judged from the

368 magnitude of the estimated parameters, especially because regularized regression
369 approaches require the predictor data to be scaled to $\bar{x} = 0$ and $\sigma = 1$. In this manner,
370 gene's whose expression leads to large estimated coefficients are worthy of further
371 investigation and follow up. To our knowledge, no one has used regularized regression
372 to estimate natural selection on gene expression and transcriptomes.

373

374 **Measuring Selection Gradients at the Genotypic scale**

375

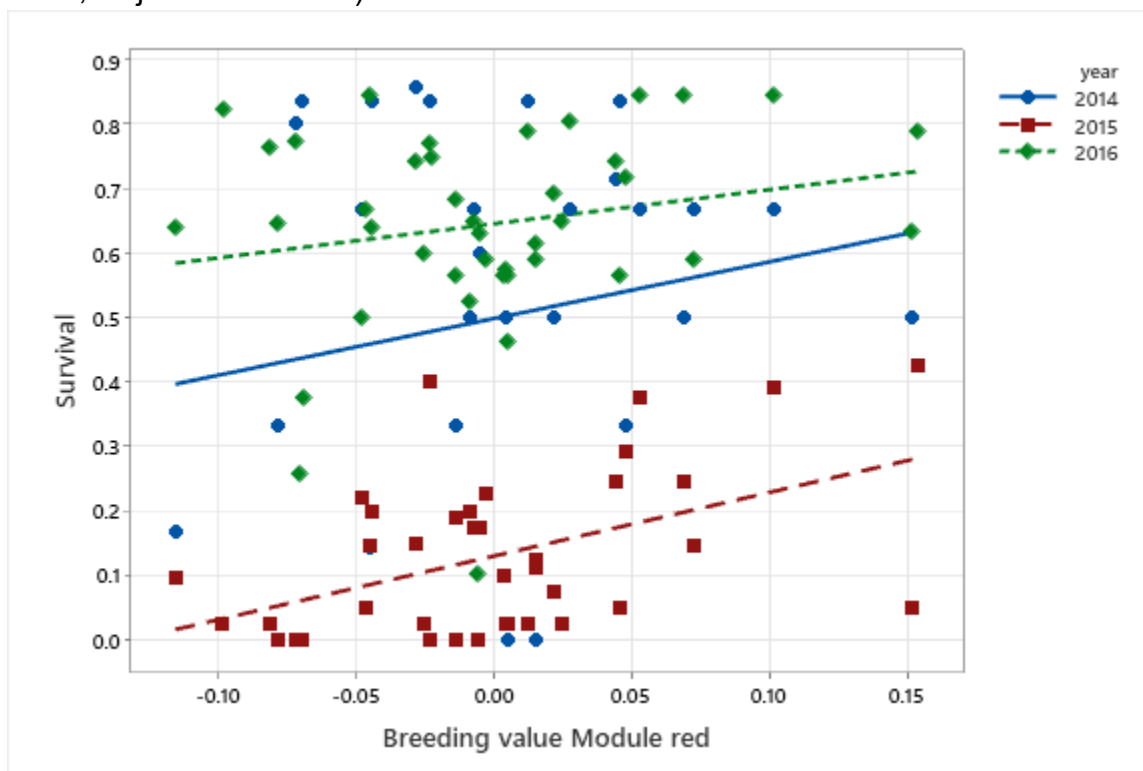
376 In evolutionary quantitative genetics, it is common to distinguish phenotypic selection
377 from response to selection. The former is the relationship between the multivariate
378 phenotype and fitness while the latter is determined by the mapping from genotype to
379 phenotype and requires an additional generation to measure. Separating selection from
380 response enables an operational division of labor. Field studies without a genetic
381 component can characterize selection, usually employing the Lande-Arnold regression
382 framework. Given phenotypic selection estimates, an evolutionary response can be
383 predicted using estimates of additive genetic variances and covariances from genetic
384 experiments. Genetic statistics can be estimated from classical breeding designs or
385 pedigrees or from genomic genotyping of individuals (Lynch and Walsh 1998).

386

387 The separation of selection from response is certainly convenient, but it is encumbered
388 with serious assumptions (Morrissey et al. 2010). There are many situations in which it
389 is advantageous to predict fitness from genetic statistics. One downfall of predicting
390 fitness directly from phenotypic traits (of any variety) is the possibility that the
391 relationship may be environmentally induced (Mitchell-Olds and Shaw 1987; Price et al.
392 1988; Rausher 1992). For plant systems, it is easy to envision that individuals growing
393 in high resource soils (e.g., high N, P, or K) both have higher fitness and also larger
394 values of traits requiring N and P— for example, size, branching, or plant defense traits.
395 In this instance, a naive application of the Lande-Arnold approach would detect
396 selection on these traits even if size, branching, and plant defense have no effect on
397 fitness at all. In this scenario, both fitness and the other phenotypic traits are
398 responding, independently, to soil resource variation, and investigators observe an
399 environmentally induced relationship rather than a causal one. In regression terms, one
400 has omitted a 'trait' (in this case, soil NPK concentrations) that is correlated with both
401 the predictors and the response variable, leading to inaccurate parameter estimates.
402 Importantly, such relationships will not lead to responses to selection and evolutionary
403 change (Rausher 1992). It is highly likely that gene expression, as a trait, will be
404 environmentally sensitive to aspects of soils, temperature, weather, abiotic and biotic
405 conditions, and a multitude of other influences. *A priori*, this suggests that the potential
406 for environmental covariances to bias estimates of phenotypic selection on gene
407 expression is high.

408

409 Fortunately, Rausher (1992; Figure 1, path (g)) provided a solution to this problem:
 410 estimating selection using either breeding values or estimates of genotypic values for
 411 both phenotypes and fitness. While this approach comes at a cost of sample size and
 412 statistical power (Stinchcombe et al. 2002), covariances between fitness and
 413 phenotypes estimated with breeding values reflect genetic relationships, rather than
 414 environmentally induced ones. As a consequence, not only are parameter estimates of
 415 selection more accurate, but they also reflect relationships that have the potential to
 416 produce evolutionary change (Stinchcombe et al 2002). While formal studies remain
 417 rare (but see Stinchcombe et al. 2002 and Hadfield 2008), existing evidence suggests
 418 that many estimates of phenotypic selection on macroscopic traits are highly biased by
 419 environmental covariances (Stinchcombe et al. 2002; Kruuk 2002; Morrissey et al.
 420 2012; Hajduk et al. 2020).



421 **Figure 2. Selection gradients for expression module ‘Red’ are positive for survival**
 422 **in all three years. The overall effect of Red on survival to flower (all years**
 423 **included) is significantly positive ($F_{1,107} = 6.43, p = 0.013$), although only the 2015**
 424 **regression, where survival was generally low, is significant when considered in**
 425 **isolation ($F_{1,38} = 11.07, p < 0.002$).**

427
 428 The breeding value regression approach is as applicable to gene expression as it is to
 429 macro-phenotypes. This is also true for the compression methods discussed in the
 430 previous section (e.g. PCA and WGCNA). They can be applied as readily to breeding
 431 value regressions as to phenotypic regressions. To illustrate, we revisit the gene
 432 expression modules of Brown and Kelly (2022) which were obtained for homozygous

433 lines of *Mimulus*. Many of these same lines were intercrossed to make F₁ plants and
434 then measured for survival and reproductive success in the natural habitat by Troth *et*
435 *al.* (2018). The breeding values (a.k.a. additive genetic values) for the F₁ plants are the
436 average of the values obtained from the parental lines – this is how breeding values
437 work (Lynch and Walsh 1998). As a consequence, we can use gene expression
438 estimates obtained in the greenhouse to predict fitness in nature. Over three successive
439 generations, one expression module (Red) was a consistent predictor of survival to
440 flower in each of the three field seasons (**Fig 2**; note that Fig. 2 is path (c) in Fig. 1). In
441 the greenhouse, genotypes with high values for ‘Red’ are associated with earlier
442 germination (Brown and Kelly 2022).

443
444 Unlike phenotypes, which are unique features of individual plants, breeding values are
445 ‘portable’; indeed, this is one of their features at the heart of their success in agricultural
446 applications. They can be carried across experiments whenever genotypes can be
447 replicated. Portability enables highly powered experiments because gene expression
448 can be studied on large samples of plants grown under controlled conditions (e.g., the
449 greenhouse environment). Also, because breeding values of expression are the
450 predictors of field fitness, we avoid the serious difficulty of environmental factors
451 inducing spurious correlations between phenotype and fitness. The downside is that
452 expression levels in the greenhouse could prove to be the “wrong traits.” The same
453 genes could be expressed in different ways under field conditions than under those
454 used to obtain breeding values, or different genes could be expressed in response to
455 different environments. This would be an example of genotype by environment
456 interaction, where the amount of expression, or which genes are expressed, depends
457 on the environmental context (greenhouse or field). Of course, this is always a concern
458 with RNAseq studies, whether the transcripts sampled from a particular tissue at a
459 particular life stage are the most relevant determinants of phenotype and/or fitness.

460
461 To date, there has been a great deal more work on the genetic basis of transcriptional
462 variation than on how this variation affects fitness. Research on the genetics of gene
463 expression has also been confronting the issue of scale. Above, we discussed PCA and
464 WGCNA based on the ‘P matrix’, the variances and covariances among plants in the
465 expression level of each gene (the phenotype). An alternative approach is to partition
466 the phenotypic variance into genetic and environmental components and then apply the
467 compression to these underlying components. For example, Blows *et al* (2015) show
468 that the genetic component of variation in expression of 8750 genes of *Drosophila*
469 *serrata* could be distilled into the contributions of a much smaller number of underlying
470 variables using matrix completion methods.

471

472 A qualitatively different approach to understanding gene expression evolution is to apply
473 factor analysis or latent factor modelling. While these approaches are more common in
474 psychology and other disciplines, they have received less adoption in evolutionary
475 biology (for exceptions, see McGuigan and Bows 2010; Frichot et al. 2013). In the
476 context of gene expression, the idea is that the expression of each gene is influenced
477 by a limited number of underlying ‘factors.’ These factors are not directly observed but
478 can be modeled and estimated from data. Variation in factors can be partitioned into
479 genetic and environmental components, and through the mapping from the factors to
480 the expression levels of genes, one can characterize the variances and covariances for
481 the entire transcriptome. The problem thus shifts from analyzing the genetic variances
482 and covariances in the expression of thousands of individual genes to understanding
483 the variances and covariances of a much more limited set of inferred factors. Two
484 implementation methods– Bayesian Sparse Factor Analysis (BSFA; Runcie and
485 Mukherjee 2013) and MegaLMM (Runcie et al. 2021)– have been developed that are
486 suited to predicting the high dimensional structure of genetic variances and covariances
487 of the transcriptome from a more limited set of variables. These approaches provide a
488 natural means to reduce the dimensionality of the determination of gene expression
489 levels from genetic and environmental influences. Correlations between expression
490 levels emerge when different genes share a common factor.

491
492 Factor analysis could be applied to estimate selection on the transcriptome in either of
493 two distinct ways. The first would be to apply BSFA or MegaLMM strictly to the
494 partitioning of transcriptome variation into genetic and environmental components,
495 without including fitness variables in the model. Given estimated breeding values for
496 factors, one could predict field fitness in a way analogous to the *Mimulus* example of Fig
497 2 (except using factors instead of module eigenvalues). This approach addresses the
498 scale issue because factors are uncorrelated with each other. Moreover, given the
499 mapping from factors to expression levels, one can extrapolate from selection gradients
500 on factors to gradients on individual genes. The second way would be to apply factor
501 analysis to transcriptomes and fitness measurements simultaneously. This is essentially
502 adding fitness measures to the list of phenotypes (transcript levels). One then estimates
503 genetic and environmental covariances among the expression levels of genes
504 simultaneously with their covariances with fitness. Estimated factors with strong
505 contributions from fitness would be identified as under selection. Genes whose
506 expression loaded heavily on those factors are thus under selection.

507
508 The simultaneous approach has the advantage that the sparse factor model directly
509 estimates the genetic covariance between fitness and gene expression. This is the
510 predicted change in the mean expression level into the next generation (Robertson
511 1966, Price 1970, 1972). The two-part method is more consistent with the traditional

512 quantitative genetic approach based on regression where we distinguish traits as
513 independent variables (predictors) and fitness components as dependent variables.
514 Oftentimes, the joint distribution for traits can be treated with a multivariate normal
515 distribution. However, fitness components are usually non-normal (e.g. binary for
516 survivorship, negative binomial for fecundity, etc). It may be easier to accommodate the
517 differing distributions for transcript variation and fitness components in a regression
518 framework. A second reason to separate fitness from characterization of transcriptome
519 variation is that we often expect the relationship between trait values and fitness to be
520 non-linear due to stabilizing, disruptive, or correlational selection.

521

522 **Conclusions**

523

524 Several common themes emerged from our overview of techniques for characterizing
525 selection on the transcriptome even though many techniques are still in areas of active
526 development. First, at their heart, most of the approaches we have discussed approach
527 the N - p problem through some form of compression and reduction in the number of
528 parameters that have to be estimated. As long as the sample sizes for the number of
529 genes for which expression is measured with sequencing technologies exceeds the
530 number of individuals in experiments, some form of data reduction or compression will
531 remain a requirement.

532

533 Second, we perceive distinct analysis paths which investigators can take, based on the
534 data in hand and the tractability of the system. For species in which it is possible to
535 perform breeding designs, create known and replicated genotypes, and/or generate
536 inbred lines, analyses based on breeding values should be pursued. In these systems,
537 expression can be measured in the greenhouse or growth chamber and fitness
538 estimates obtained from the same genotypes (or relatives with predictable breeding
539 values). In the case of inbred lines, successive estimates of transcriptomes,
540 performance, and fitness could be obtained from immortalized genotypes that are
541 exposed to a variety of growth conditions. In contrast, for species or systems where it is
542 difficult to obtain immortalized genotypes— or where cost constraints preclude
543 characterizing the transcriptomes of many genotypes— estimates of selection on the
544 transcriptome are more akin to the field studies of selection on macro traits that followed
545 the Lande and Arnold (1983) paper. The rich picture of how natural selection acts on
546 morphological, behavioral, and life-history phenotypes is from a set of studies similar in
547 design to a single-instance measurement of selection on the transcriptome (Henry and
548 Stinchcombe 2025). We have drastically fewer estimates of selection on transcriptomes
549 to characterize its strength, mode, and spatial or temporal consistency, perhaps
550 because the approach and technology are in early development. More than 40 years
551 ago, Arnold (1983) coined the expression “morphology, performance, fitness” in a

552 landmark paper describing how to understand variation in, and selection on,
 553 morphology. We suggest that an important area of research in the next 40 years of
 554 evolutionary biology will be to explore the mapping from gene expression to phenotype
 555 to fitness.

556

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558

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566

567

568 **References**

569

570 **Arnold SJ. 1983.** Morphology, performance and fitness. *American zoologist* **23**: 347–
 571 361.

572 **Blows MW, Allen SL, Collet JM, Chenoweth SF, McGuigan K. 2015.** The phenome-
 573 wide distribution of genetic variance. *The American naturalist* **186**: 15–30.

574 **Brown KE, Kelly JK. 2022.** Genome-wide association mapping of transcriptome
 575 variation in *Mimulus guttatus* indicates differing patterns of selection on *cis*-
 576 versus *trans*-acting mutations. *Genetics* **220**: :iyab189.

577 **Campbell-Staton SC, Cheviron ZA, Rochette N, Catchen J, Losos JB, Edwards SV.**
 578 **2017.** Winter storms drive rapid phenotypic, regulatory, and genomic shifts in the
 579 green anole lizard. *Science* **357**(6350): 495-498.

580 **Chong VK, Fung HF, Stinchcombe JR. 2018.** A note on measuring natural selection
 581 on principal component scores. *Evolution letters* **2**: 272–280.

582 **Frichot E, Schoville SD, Bouchard G, François O. 2013.** Testing for associations
 583 between loci and environmental gradients using latent factor mixed models.
 584 *Molecular biology and evolution* **30**: 1687–1699.

585 **Ghalambor CK, Hoke KL, Ruell EW, Fischer EK, Reznick DN, Hughes KA. 2015.**
 586 Non-adaptive plasticity potentiates rapid adaptive evolution of gene expression in
 587 nature. *Nature* **525**(7569): 372-375.

588 **Ghalambor CK, Hoke KL, Ruell EW, Fischer EK, Reznick DN, Hughes KA. 2018.**
 589 Erratum: Non-adaptive plasticity potentiates rapid adaptive evolution of gene
 590 expression in nature. *Nature* **555**: 688.

591 **Gilad Y, Oshlack A, Rifkin SA. 2006.** Natural selection on gene expression. *Trends in*
 592 *Genetics* **22**: 456–461.

- 593 **Groen SC, Čalić I, Joly-Lopez Z, Platts AE, Choi JY, Natividad M, Dorph K, Mauck**
 594 **WM 3rd, Bracken B, Cabral CLU, et al. 2020.** The strength and pattern of
 595 natural selection on gene expression in rice. *Nature* **578**: 572–576.
- 596 **Hadfield, J. D. 2008.** Estimating evolutionary parameters when viability selection is
 597 operating. *Proceedings of the Royal Society B: Biological Sciences* **275**:723–734.
- 598 **Hajduk GK, Walling CA, Cockburn A, Kruuk LEB. 2020.** The ‘algebra of evolution’:
 599 the Robertson-Price identity and viability selection for body mass in a wild bird
 600 population. *Philosophical transactions of the Royal Society of London. Series B,*
 601 *Biological sciences* **375**: 20190359.
- 602 **Hamann E, Pauli CS, Joly-Lopez Z, Groen SC, Rest JS, Kane NC, Purugganan MD,**
 603 **Franks SJ. 2021.** Rapid evolutionary changes in gene expression in response to
 604 climate fluctuations. *Molecular Ecology* **30**(1): 193-206.
- 605 **Henry GA, Stinchcombe JR. 2023.** Strong selection is poorly aligned with genetic
 606 variation in *Ipomoea hederacea*: implications for divergence and constraint.
 607 *Evolution* **77**: 1712–1719.
- 608 **Henry GA, Stinchcombe JR. 2025.** Predicting fitness-related traits using gene
 609 expression and machine learning. *Genome biology and evolution* **17**: evae275.
- 610 **Hill WG. 2010.** Understanding and using quantitative genetic variation. *Philosophical*
 611 *transactions of the Royal Society of London. Series B, Biological sciences* **365**:
 612 73–85.
- 613 **Jolliffe IT. 2002.** *Principal Component Analysis*. NY, NY, USA: Springer.
- 614 **Josephs EB, Lee YW, Stinchcombe JR, Wright SI. 2015.** Association mapping
 615 reveals the role of purifying selection in the maintenance of genomic variation in
 616 gene expression. *Proceedings of the National Academy of Sciences of the*
 617 *United States of America* **112**: 15390–15395.
- 618 **Josephs EB, Lee YW, Wood CW, Schoen DJ, Wright SI, Stinchcombe JR. 2020.**
 619 The evolutionary forces shaping *cis*- and *trans*-regulation of gene expression
 620 within a population of outcrossing plants. *Molecular biology and evolution* **37**:
 621 2386–2393.
- 622 **King M-C, Wilson AC. 1975.** Evolution at two levels in humans and chimpanzees.
 623 *Science* **188**: 107–116.
- 624 **Kingsolver JG, Hoekstra HE, Hoekstra JM, Berrigan D, Vignieri SN, Hill CE, Hoang**
 625 **A, Gibert P, Beerli P. 2001.** The strength of phenotypic selection in natural
 626 populations. *The American naturalist* **157**: 245–261.
- 627 **Kingsolver JG, Diamond SE, Siepielski AM, Carlson SM. 2012.** Synthetic analyses
 628 of phenotypic selection in natural populations: lessons, limitations and future
 629 directions. *Evolutionary ecology* **26**: 1101–1118.
- 630 **Kruuk LEB, Slate J, Pemberton JM, Brotherstone S, Guinness F, Clutton-Brock T.**
 631 **2002.** Antler size in red deer: Heritability and selection but no evolution. *Evolution*
 632 **56**: 1683–1695.
- 633 **Lande R, Arnold S. 1983.** The measurement of selection on correlated characters.
 634 *Evolution* **37**: 1210-1226.
- 635 **Langfelder P, Horvath S. 2008.** WGCNA: an R package for weighted correlation
 636 network analysis. *BMC bioinformatics* **9**: 559.

- 637 **Liu, Xuanyao, Li, Yang I. and Pritchard, Jonathan K. 2019.** Trans effects on gene
638 expression can drive omnigenic inheritance. *Cell* 177 (4): 022-1034.e6
639 <https://doi.org/10.1016/j.cell.2019.04.014>
- 640 **Lynch M, Walsh B. 1998.** *Genetics and the analysis of quantitative traits*. Sunderland,
641 MA, USA: Sinauer Associates.
- 642 **McGuigan K, Blows MW. 2010.** Evolvability of individual traits in a multivariate context:
643 partitioning the additive genetic variance into common and specific components.
644 *Evolution* 64: 1899–1911.
- 645 **Mitchell-Olds T, Shaw RG. 1987.** Regression analysis of natural selection: statistical
646 inference and biological interpretation. *Evolution* 41: 1149–1161.
- 647 **Mojica JP, Kelly JK. 2010.** Viability selection prior to trait expression is an essential
648 component of natural selection. *Proceedings of the Royal Society B: Biological*
649 *Sciences* 277: 2945–2950.
- 650 **Morrissey MB, Kruuk LEB, Wilson AJ. 2010.** The danger of applying the breeder's
651 equation in observational studies of natural populations. *Journal of evolutionary*
652 *biology* 23: 2277–2288.
- 653 **Morrissey MB, Parker DJ, Korsten P, Pemberton JM, Kruuk LEB, Wilson AJ. 2012.**
654 The prediction of adaptive evolution: Empirical application of the secondary
655 theorem of selection and comparison to the breeder's equation. *Evolution* 66:
656 2399–2410.
- 657 **Morrissey MB. 2014.** In search of the best methods for multivariate selection analysis.
658 *Methods in Ecology and Evolution* 5: 1095–1109.
- 659 **Oleksiak MF, Churchill GA, Crawford DL. 2002.** Variation in gene expression within
660 and among natural populations. *Nature genetics* 32: 261–266.
- 661 **Palakurty SX, Stinchcombe JR, Afkhami ME. 2018.** Cooperation and coexpression:
662 How coexpression networks shift in response to multiple mutualists. *Molecular*
663 *ecology* 27: 1860–1873.
- 664 **Price GR. 1970.** Selection and covariance. *Nature* 227: 520–521.
- 665 **Price GR. 1972.** Extension of covariance selection mathematics. *Annals of human*
666 *genetics* 35: 485–490.
- 667 **Price T, Kirkpatrick M, Arnold SJ. 1988.** Directional selection and the evolution of
668 breeding date in birds. *Science* 240: 798–799.
- 669 **Rausher MD. 1992.** The measurement of selection on quantitative traits: biases due to
670 the environmental covariances between traits and fitness. *Evolution* 46: 616-626.
- 671 **Rifkin SA, Houle D, Kim J, White KP. 2005.** A mutation accumulation assay reveals a
672 broad capacity for rapid evolution of gene expression. *Nature* 438: 220–223.
- 673 **Robertson A. 1966.** A mathematical model of the culling process in dairy cattle. *Animal*
674 *production* 8: 95–108.
- 675 **Runcie DE, Mukherjee S. 2013.** Dissecting high-dimensional phenotypes with
676 Bayesian Sparse Factor Analysis of genetic covariance matrices. *Genetics* 194:
677 753–767.
- 678 **Runcie DE, Qu J, Cheng H, Crawford L. 2021.** MegaLMM: Mega-scale linear mixed
679 models for genomic predictions with thousands of traits. *Genome biology* 22:
680 213.

- 681 **Schrider DR, Kern AD. 2018.** Supervised machine learning for population genetics: A
682 new paradigm. *Trends in genetics* 34: 301–312.
- 683 **Siepielski AM, Gotanda KM, Morrissey MB, Diamond SE, DiBattista JD, Carlson**
684 **SM. 2013.** The spatial patterns of directional phenotypic selection. *Ecology*
685 *letters* 16: 1382–1392.
- 686 **Slotte T, Bataillon T, Hansen TT, St Onge K, Wright SI, Schierup MH. 2011.**
687 Genomic determinants of protein evolution and polymorphism in *Arabidopsis*.
688 *Genome biology and evolution* 3: 1210–1219.
- 689 **Stinchcombe JR, Rutter MT, Burdick DS, Tiffin P, Rausher MD, Mauricio R.**
690 **2002.** Testing for environmentally induced bias in phenotypic estimates of natural
691 selection: theory and practice. *The American naturalist* 160: 511–523.
- 692 **Svensson EI. 2023.** Phenotypic selection in natural populations: what have we learned
693 in 40 years? *Evolution* 77: 1493–1504.
- 694 **Sztepanacz JL, Houle D. 2024.** Regularized regression can improve estimates of
695 multivariate selection in the face of multicollinearity and limited data. *Evolution*
696 *letters* 8: 361–373.
- 697 **Troth A, Puzey JR, Kim RS, Willis JH, Kelly JK. 2018.** Selective trade-offs maintain
698 alleles underpinning complex trait variation in plants. *Science* 361(6401): 475-
699 478.
- 700 **Wright SI, Yau CBK, Looseley M, Meyers BC. 2004.** Effects of gene expression on
701 molecular evolution in *Arabidopsis thaliana* and *Arabidopsis lyrata*. *Molecular*
702 *biology and evolution* 21: 1719–1726.