# Phylogeny of *Weinmannia* (Cunoniaceae) reveals the Contribution of the Southern Extratropics to Tropical Andean Biodiversity.

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#### Abstract

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rsegovia@ieb-chile.cl; christine.edwards@mobot.org The Andes are a relatively young mountain range with an impressive concentration of biodiversity. The biogeographic processes contributing to the hyperdiversity of the tropical Andes are still being unraveled. Novel mid- to high-elevation climates may have served as a biological corridor for the immigration of temperate-adapted lineages to lower latitudes, contributing unknown levels of diversity to this region. We tested the hypothesis that the genus Weinmannia corresponds to a lineage of extratropical origin that recently reached and then diversified extensively in tropical Andes. Using a 2bRAD seq approach to generate a time-calibrated phylogenetic tree for the genus, we found that the earliest diverging clades represent southern extratropical species, and a tendency for younger clades to be distributed towards northern latitudes. These results suggest a dispersal route for *Weinmannia* from the southern extratropics to the tropical Andes. As remnants of these lineages of southern origin converge with those that originated in other tropical and extratropical centers of diversification, we provide insights into the multisource origins of hyperdiversity in the modern montane forests of the tropical Andes.

**Keywords**: immigration, diversification, hyperdiversity, tropics, Gondwana

#### Introduction

2 The Andean region of tropical America has one of the world's highest 3 levels of species richness (Balsev 1993), taxonomic endemism (Myers et al. 4 2000) and phylogenetic diversity (Tietje et al. 2023). This hyperdiversity is particularly intriguing given that the modern geomorphology of this area is 5 6 no older than the late Miocene (< 11 Ma) (Gregory-Wodzicki 2000; Siravo et 7 al. 2018). Mountain building is generally thought to have fostered high levels of diversity in a variety of ways, including the speciation of resident lineages 8 (Rahbek et al. 2019) and the immigration of lineages pre-adapted to newly 9 created climatic conditions (Donoghue 2008). Indeed, the Andean orogeny 10 may have increased the rate of lineage diversification (Antonelli and 11 Sanmartin 2011a; Hazzi et al. 2018) and may also have opened a corridor for 12 the immigration of temperate lineages into the lower latitudes of tropical 13 America (Graham 1973; Segovia and Armesto 2015). Comprehensive 14 evolutionary evidence is still being gathered to identify areas of lineage 15 origin and thus unravel the relative influence of these biogeographic 16 processes in shaping the modern pattern of hyperdiversity in the Andes. 17

shows Phylogenetic evidence faster-than-expected rates of 18 diversification for a number of potentially resident plant clades in synchrony 19 with the Andean uplift since the early Miocene (e.g., Luebert and Weigend 20 21 2014; Givnish et al. 2014; Schwery et al. 2015; Spriggs et al. 2015; Perez-22 Escobar et al. 2017, 2022, Dellinger et al. 2024). Moreover, a growing body of phylogenetic evidence shows immigration processes from both the 23 northern and southern extratropics into the tropical Andes. Many lineages 24 25 have immigrated from the northern extratropics, including examples such as Viburnum (Winkworth and Donoghue 2005), Lupinus (Hughes and Eastwood 26 27 2006), and Passiflora section Decaloba (Acha et al. 2021). Conversely, there is less evidence for lineages that likely arrived in the tropical Andes from the 28 southern extratropics, although notable examples include Alstroemeriaceae 29 30 (Chacón et al. 2012), Podocarpus (Quiroga et al. 2016), Gunnera (Bacon et

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al. 2018), and Loranthaceae (Liu et al. 2018). However, most of what we
know about these biogeographic scenarios for the origin of the plant diversity
in the tropical Andes is biased toward clades inhabiting open biomes at highelevations. Evidence regarding the origin and direction of dispersal routes of
the clades that comprise montane forests at intermediate elevations still
remains scarce, and represents a significant gap in our understanding of plant
diversity and evolution in one the most species-rich region on the planet.

38 The idea that lineages from the relatively species-poor extratropics could contribute to the modern hyperdiversity of the tropical Andes is 39 counterintuitive. Traditionally, the highest levels of species richness are 40 thought to be associated with "centers of diversification" (Willis 1922), or 41 areas where a particular lineage originated (Wiens and Donoghue 2004). In 42 addition, there is strong evidence that the American tropics have historically 43 acted as a "species pump" for global plant diversity (Antonelli et al. 2015). 44 However, the immigration from multiple zones, including the extratropics, 45 into tropical Andean forests has also been documented based on taxonomic 46 affinities (Hooghiemstra 1984), fossil records (Graham 1995), and 47 community phylogenetics (González-Caro et al. 2023), rather than just clade 48 reconstructions. These multisource immigration processes, along with rapid 49 lineage diversification, may be key to shaping the modern hyperdiversity of 50 the tropical Andes, increasing not only taxonomic diversity but also 51 evolutionary diversity, according to the "environmental crossroads 52 hypothesis" (Neves et al. 2020). For example, the exceptionally high levels of 53 different measures of phylogenetic diversity in the central and northern 54 Andes (Tietje et al. 2023) may be due to the mixing of deeply isolated biotas 55 with different evolutionary histories (e.g., remnants of paleobiotas from 56 57 Holartic, Austral and Neotropical floristic realms).

Here we investigate the biogeography of *Weinmannia* L. (*sensu* Pillon 58 et al. 2021), formerly *Weinmannia* sect. *Weinmannia* L. (Bradford, 1998), a 59 genus of trees and shrubs that are important components of ecosystem 60

functioning of Andean forests owing to their abundance and diversity. 61 Typically considered a southern hemisphere taxon, Weinmannia consists of 62 two species occurring in the Mascarenes and over 80 species in the Americas 63 (Bradford 1998, 2002; Ulloa-Ulloa et al. 2017; Pillon et al. 2021). The higher 64 diversity of Weinmannia occurs at mid- to high- elevations in the tropical 65 Andes, where it exhibits weak species boundaries and overlapping 66 morphologies due to either recent divergence or hybridization (Bradford 67 1998, 2002). In addition, several species occur in the Pantepui region, and the 68 69 mountain peaks of Central America and the Caribbean Islands. One species is endemic to the subtropical forests of eastern Brazil (i.e., Mata Atlantica), and 70 another species occurs in the temperate forests of southern South America 71 72 (Chile and Argentina) (Gentry 1982, 1995). Previous phylogenies placed the 73 only southern extratropical species (Weinmannia trichosperma Cav.) at the 74 base of the Weinmannia clade (Bradford 1998, 2002), sparking the 75 hypothesis that Weinmannia immigrated recently into the tropical Andes from the southern extratropics (Bradford et al. 2004; Pennington and Dick 76 2004). However, these analyses sampled only a small proportion of the 77 species in the genus and employed only a small number of plastid and/or 78 nuclear regions (Bradford 2002; Pillon et al. 2021), preventing hypothesis 79 testing about the origin and dispersal of Weinmannia across the Andes. 80

To examine the hypothesis of a southern extratropical origin for 81 Weinmannia and a recent immigration into the tropical Andes, we 82 83 constructed a new NGS phylogeny with a comprehensive taxon sampling. First, we tested the prediction that if the genus Weinmannia originated in the 84 southern extratropics, then W. trichosperma from the temperate forests of 85 southern South America should be placed as a sister group to all other 86 Weinmannia species of the Americas. Second, we tested the prediction that if 87 the modern distribution of Weinmannia is a consequence of dispersal from 88 the southern extratropics into the tropical Andes, then the ages of nodes in 89 the phylogeny should show a negative relationship with the reconstructed 90 latitude of the nodes. In other words, the phylogeny should show a pattern in91which younger clades occupy successively more northern latitudes.92

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# **Materials & Methods**

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Sampling and genomic DNA extraction. We collected 896 samples of95Weinmannia from across South America, including the southern Andes96(Chile), the central Andes (Bolivia and Peru), and the northern Andes97(Ecuador and Colombia). We also included four samples from the Mascarene98Islands. For each sample, we preserved leaf tissue in silica and collected an99herbarium voucher specimen. In addition, we included five specimens from100three species of Pterophylla (sensu Pillon et al. 2021) as outgroups.101

To extract DNA, silica-dried tissues were first ground and then 102 cleaned using up to three sorbitol washes following Inglis et al. (2018) to 103 remove mucilage and other secondary compounds. Genomic DNA was 104 extracted using a modified CTAB extraction protocol for plants (based on 105 106 Doyle and Doyle et al. 1987), with additional ethanol washes of precipitated DNA. Finally, following extraction, DNA was purified using KAPA pure 107 Beads (KAPA Biosystems) following manufacturer protocols. DNA 108 concentrations were quantified using a Qubit<sup>™</sup> fluorometer (ThermoFisher). 109

Sequencing. RAD-seq libraries were prepared using a 2b-RAD approach 110 (Wang et al. 2012) and previously published protocols for the method (Linan 111 et al. 2021b; Mashburn et al. 2023). We digested approximately 500 ng of 112 113 purified genomic DNA of each sample, using the type-IIB restriction enzyme BcgI (New England Biolabs), which produced 36 bp DNA fragments from 114 across the nuclear genome of each specimen. To ensure adequate sequence 115 coverage per locus, 5'-NNG-3' selective adapters were used to decrease the 116 number of sequenced loci, as described by Wang et al. (2012). 96 samples 117 were pooled per plate through the use of dual index barcodes, whereby the 118 first barcode applied across columns allow pooling of the 8 rows. Each of 119

these pools was amplified with incorporating the second barcode index (one 120 of 8 unique 6 bp Illumina TruSeg barcodes) and an Illumina primer using 121 high fidelity Phusion PCR mix (New England Biolabs) for 14 cycles of PCR 122 amplification. The resulting amplicons were purified using 2% Agarose gel 123 purification with a MinElute gel purification kit (Qiagen). The purified 124 ligation pools were mixed at equimolar proportions after measuring 125 concentrations in the Qubit fluorometer (Qiagen). The final library obtained 126 was sequenced on an Illumina HiSeq 4000, generating 50 bp single end reads 127 128 at the NUSeq Core facility of Northwestern University.

**RAD locus assembly.** The quality of raw reads was checked using FastQC 129 130 (Andrews, 2010). Sequences were demultiplexed and trimmed to remove barcodes and Illumina adaptors using the trim2bRAD 2barcodes noAdap.pl 131 script available in 2bRAD-Edwards-Lab git repository 132 (https://github.com/Kenizzer/2bRAD-Edwards-Lab). Trimmed reads were 133 assembled *de novo* using the ipyrad v. 0.9.90 pipeline (Eaton 2020). To 134 determine the optimum clustering threshold, we iterated clustering threshold 135 within samples (CTWS) and among samples (CTAS) using every 136 combination of values of 0.86, 0.89, 0.92, and 0.94. The resulting matrices 137 were compared for cluster depth, heterozygosity, the amount of putatively 138 paralogous loci, and the number of SNPs to identify parameters that could 139 lead to assembly errors (Paris et al. 2017). Following this approach, we 140 selected a value of 0.92 for both CTWS and CTAS. We filtered all loci 141 including gaps and retained all loci with no more than five SNPs. 142

Identification of putative hybrids. To identify putative hybrid individuals143that may confound phylogenetic analysis, we assessed admixture using144STRUCTURE v. 2.3.4 (Pritchard et al. 2000) as implemented in ipyrad v.1450.9.90. Due to the large number of samples and putative species, we divided146samples and conducted a series of analyses within two geographically147structured groups (central Andes region and northern Andes region), given148149149

proximal species. For each of these analyses, we conducted an independent 150 assembly with the clustering threshold 0.92 removing all individuals with 151 more than 80% missing sites, retaining loci present in at least 50% of 152 samples, and retaining one SNP per locus. STRUCTURE analyses were run 153 under an admixture model with a burn-in of 300,000 generations and a run 154 length of 700,000 generations. The number of genetic clusters assessed (K) 155 was defined according to the number of species within each sub analysis 156 going from K=2 to 20. A total of 15 repetitions were run for each K value. 157 The optimal K value was selected using the Evanno method (Earl and von 158 Holdt 2011). We defined putative hybrids as individuals with less than 85% 159 assignment to a single genetic cluster, adopting a slightly more conservative 160 criterion compared to the 80% cut-off used by Lindtke et al. (2014), Owusu 161 et al. (2015) and Linan et al (2021). 162

Individual-level tree inference.For phylogenetic inference, we chose 3-5163individuals with no signature of hybridization from each species.This164resulted in 234 accessions, spanning 48 Weinmannia taxa (including 6 that165are not assigned to any described species), plus 3 Pterophylla species as166outgroups (total 51 species).This dataset includes a representative sample of167the 75 (Ulloa-Ulloa et al. 2017) to 90 (Pillon et al. 2021)Weinmannia species168estimated to occur in the Americas.169

170 Maximum likelihood (ML) phylogenetic analysis was conducted using concatenated 36 bp fragments, including invariant sites. First, we 171 performed preliminary analysis to explore the effect of missing data on the 172 173 resulting topologies, varying the percentage of samples at which a locus must 174 be present to be retained from 4% to 48% in increments of 4. We found optimal branching resolution and bootstrap supports when all loci were 175 present in at least 36% (84/234) of samples. The ML phylogeny was inferred 176 in RAxML v. 8.2.12 (Stamatakis 2014) using a rapid hill climbing algorithm 177 and the GTRCAT approximation. Clade supports were calculated using the 178 transfer bootstrap approach with 200 iterations (Lemoine et al. 2018). 179

Species-level Weinmannia phylogeny. To reconstruct a species-level 180 phylogeny, the individual-level phylogeny (Supplementary Fig. 1) was used 181 to select one representative individual from each species (defined as a 182 reciprocally monophyletic group of individuals that had morphological 183 distinctiveness). We choose non-admixed individuals (as indicated by 184 STRUCTURE) with the least amount of missing data. To infer a phylogeny, 185 we employed both a ML analysis of the concatenated loci and SVDQuartets, 186 which is a multi-species coalescent-based approach (Chifman & Kubatko 187 188 2014). For the species-level ML analysis, we followed the same approach as described for the individual-level phylogeny. After preliminary analysis to 189 190 explore the effect of missing data, we prepared a concatenated alignment of all loci present in at least 32% (16/51) of all individuals. This matrix was 191 used to infer a phylogeny in RAxML (Stamatakis 2014). The multi-species 192 193 coalescent-based phylogenetic inference was performed using a randomly 194 selected SNP from each of the 2,879 loci used in the RAxML analysis. We inferred all 249,900 possible quartets for 51 taxa and conducted 500 195 196 bootstrap iterations. The quartet trees were joined into a super tree. Finally, we calculated bootstrap support (BS) for nodes using a transfer bootstrap 197 198 approach (Lemoine et al. 2018). For visualization, the resulting trees were rooted on the branch containing all *Pterophylla* specimens. 199

200 **Time-calibrated phylogeny.** Bayesian inference of divergence times may not always be suitable for RAD-Seq data due to statistical and computational 201 challenges (e.g. Donoghue et al. 2022). We thus inferred divergence times 202 using treePL (Smith and O'Meara 2012), which relies on optimized branch 203 204 length information for estimating divergence times under phylogenetic penalized likelihood. Divergence times were estimated following Maurin 205 206 (2008) using our maximum likelihood (ML) tree. Optimal parameters for 207 treePL were estimated by running the program with the prime option, with the smoothing parameter estimated with a cross-validation analysis. 208 209 Confidence intervals of divergence times were estimated with a bootstrap analysis in RAxML using the option -g and the ML tree to constrain the 210 topology, and the option -k to optimize branch lengths in each of 200211bootstrap iterations performed. These bootstrap trees were time calibrated212using the same treePL parameter as for the ML tree. A consensus tree was213built in TreeAnotator v. 2.5.2 (Drummond and Rambaut 2007) using the214time-calibrated ML tree and the branch-length-bootstrap trees with 0% burn-215in, median heights with the target tree option.216

217 Three calibration points were defined for divergence time estimation. 218 The first was a Weinmannia pollen fossil collected from the northern Andes of Colombia dated to ~3 Ma (Hooghiemstra 1989; Van Der Hammen et al. 219 1973). This age was defined as the minimum age for the most recent 220 221 common ancestor of the clade that encompassed all specimens collected in the Northern Andes. The second point was derived from a fossil pollen 222 223 record of Weinmannia potosina from Potosí, Bolivia (Berry 1917; Graham et al. 2001). Using the age of this record, we established 13.8 Ma as the 224 225 minimum age for the most recent common ancestor (MRCA) of the clade containing all Central and Northern Andean Weinmannia species. The upper 226 limit for this point was set at 33 Ma, aligning with the proposed beginning of 227 the Oligocene epoch (Walker et al. 2012), which is close to previous 228 estimates for the stem node of Weinmannia at 32.3 Ma (Pillon et al. 2021). 229 Additionally, we used the 95% credibility interval from Pillon et al. (2021) to 230 231 estimate the divergence time between Weinmannia and its sister genus Pterophylla, setting this calibration point with a uniform distribution, with a 232 minimum age of 29.99 Ma and a maximum age of 34.4 Ma (Pillon et al. 233 234 2021).

**Testing of our biogeographic hypotheses.** To test our hypothesis, that235*Weinmannia* migrated from south to north, we perform ancestral236reconstruction of latitude using our time-calibrated species-level phylogeny.237We determined the minimum and mean latitude of each species in our238phylogeny based on our geo-referenced occurrence data derived from239herbarium specimens. We conducted ancestral reconstruction of (minimum240

and mean) latitude using a Brownian Motion model on a pruned phylogeny241without outgroups that included 48 Weinmannia species using the 'ace'242function in the phytools v. 2.1 R package (Revell, 2012). We report only the243minimum latitude results here, as the mean latitude estimates were less244reliable due to insufficient records per species.245

Using data from the ancestral reconstruction, we modeled the age of hypothetical ancestors (nodes) as a function of their minimum latitude. For this purpose, we employed two distinct statistical approaches: a Bayesian approach and a frequentist approach based on a Null-Hypothesis Significance Test and non-parametric bootstrap.

Bayesian regression analysis. We developed a hierarchical Bayesian 251 regression that evaluated correlation structures derived from nesting patterns 252 253 between nodes in the phylogeny, allowing us to account for the underlying evolutionary relationship between our observations (i.e., node age and 254 reconstructed latitude). The model was built in Stan v. 2.18.2 (Carpenter et 255 al. 2017), which uses Hamiltonian MCMC to draw posterior samples and it 256 was implemented in R with the rstan package v. 2.26.23 (Stan Development 257 Team 2023). Full stan code can be found in Supplementary Materials 1. First, 258 the model was built by defining a linear predictor function as follows: 259

$$\mu_{n} = \alpha_{int} + \beta * X_n + \theta_n \qquad (1) \qquad 260$$

where  $\mu_n$  is the linear predictor representing the expected value of node age 261 Y<sub>n</sub> for each observation at each node *n*,  $\alpha_{int}$  is the intercept,  $X_n$  is the estimated 262 ancestral latitude of node *n*,  $\beta$  is the estimated slope coefficient representing 263 the change in the expected value of *Y* for a one-unit change in *X*, and  $\theta_n$  is the 264 random effect parameter for each node *n* capturing unexplained variation due 265 to unobserved factors not accounted for by *X*. Random effects were drawn 266 from a multivariate normal distribution according to the following function: 267

$$\theta n \sim multinormal(O_N, \Sigma)$$
 (2) 268

where  $O_N$  is the mean vector, defined as a zero vector of length N 269 (representing the number of nodes), and  $\Sigma$  is a covariance matrix, in this case, 270 given by the phylogenetic covariance matrix. We created this covariance 271 272 matrix using the makeL1 function from the RRphylo package in R (Castiglione et al. 2018). This function constructs an NxN matrix of branch 273 274 lengths for all root-to-node paths given a phylogeny, capturing the hierarchical relationships among each pair of nodes. Consequently, the 275 random effects are modeled to reflect correlations produced by the shared 276 evolutionary history among nodes. This approach allows for a more realistic 277 representation of the data's underlying dependencies. 278

Using the linear predictor function and incorporating random effects 279 (equations 1 and 2), node age  $Y_n$  was modelled using a likelihood function 280 assuming a normal distribution of errors and identity link-function as 281 follows: 282

$$Yn \sim normal(\mu_n, \varepsilon_n) \qquad (3) \qquad 283$$

where  $\varepsilon_n$  is the residual standard deviation capturing the unexplained 284 variation in Y after accounting for the independent variable (X) and random 285 effects ( $\theta$ ). The model was fitted to the data by running 4 independent 286 287 Markov Chain Monte Carlo (MCMC) chains, each for 3 million iterations. For computational efficiency, the chains were thinned every 10 iterations, 288 289 resulting in a total of 300,000 MCMC samples per chain. Of these, we 290 discarded the first 50,000 samples as burn-in. The parameter max\_treedepth was set to a minimum of 10 to handle divergent transitions during MCMC 291 292 sampling.

Model adequacy was evaluated with a posterior predictive check by 293 generating predicted values of node age from the posterior distribution using 294 the calculated  $\mu_n$  and  $\varepsilon_n$  and comparing them with the observed data (Fig. 3B 295 and Supplementary Fig. 5B). The performance and reliability of the MCMC 296 was evaluated with the Gelman-Rubin statistics (Rhat), effective Sample 297 Size, and autocorrelation analysis, among others (Supplementary Materials 2 298 and 3: MCMC diagnostics for ML phylogeny and SVDQ respectively). 299 Results of Bayesian regression were used to test the hypothesis that older 300 ancestors are associated with more southern latitudes leading to a negative 301 302 slope ( $\beta$ <0, given the predominantly negative nature of latitude values), and to assess whether we can reject the null hypothesis ( $\beta \approx 0$ , implying no clear 303 relationship). For this, we obtained the posterior probability distribution of 304 the latitude slope parameter ( $\beta$ ), from which we extracted a 95% credibility 305 306 interval and determined the maximum *a posteriori* estimate. The credibility 307 interval provides a range within which we are reasonably confident the true parameter value lies, and the maximum *a posteriori* estimate represents the 308 309 point in the distribution with the highest posterior probability, serving as a plausible point estimate for the slope parameter. This comprehensive 310 approach allows us to draw informed conclusions regarding the relationship 311 312 between the age of ancestors and their associated latitudes.

Non-parametric bootstrap Null Hypothesis Significance Test (NHST).313As an alternative approach to test our hypothesis under a linear regression314framework, we also used the glm function from the R package *stats* v. 3.6.2315(R Core Team 2023) to model node age as a function of reconstructed316ancestral latitude, following the formula:317

$$Y_{n=} = \alpha_{int} + \beta * X_n + \varepsilon_n \tag{4}$$

319 where  $Y_n$  is the age for each observation at each node *n*,  $\alpha_{int}$  is the intercept,  $X_n$  is the estimated ancestral latitude of node *n*,  $\beta$  is the estimated slope 320 coefficient representing the change in the expected value of *Y* for a one-unit 321 change in *X* and  $\varepsilon_n$  is the error or variation unexplained by other covariates. 322 To find the model with the highest adequacy and fit, we tested combinations 323 of two probability distributions (Gaussian and Gamma) and three link 324 functions (identity, log and inverse) and selected the model with the highest 325 326 fit (lowest AIC value), highest linearity of predicted vs observed values (using qqplots) and better homoscedasticity. When fitting data leveraged 327 from the ML tree, we selected gamma-distributed error and the identity 328 function (Supplementary Materials 4), whereas for data leveraged from the329SVDQ tree, we selected a gaussian distribution with an identity link function330(Supplementary Materials 5).331

We performed a bootstrap analysis within the NHST framework by 332 conducting 10,000 iterations. For each iteration, we randomized the 333 reconstructed latitude values at the nodes, then performed regression analysis 334 to obtain the slope parameter ( $\beta$ ) for each bootstrapped sample. This process 335 generated a null distribution for the slope parameter ( $\beta$ ). Given that the 336 observed slope was negative, we calculated the statistical significance (p-337 value) by finding the proportion of the null distribution that is less than or 338 equal to the observed slope, indicating the probability of getting such a result 339 under the null hypothesis. Following standard statistical procedures, a *p*-340 value smaller than 0.05 indicates significant evidence against the null 341 342 hypothesis (no relationship between node age and ancestral latitude).

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#### Results

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Specimen-level phylogeny. Following the identification and removal of 345 putative hybrids, we obtained a dataset of 234 accessions for the specimen-346 347 level phylogeny (Fig. 1A, Supplementary Fig. 1, Supplementary Table 1). The concatenated alignment was 27,072 bp in length (752 loci) and contained 348 349 31.91% missing data. In the ML phylogeny resulting from this dataset, all 350 accessions of a given species formed monophyletic groups except for Weinmannia reticulata Ruiz & Pav. ex López. Two different subclades, 351 named W. reticulata1 and W. reticulata2, were treated as separate species for 352 353 the purpose of the present analysis (Supplementary Fig. 1). Our analysis also included accessions of undetermined species (sp1 to sp6) that showed 354 morphological and phylogenetic cohesion. Further taxonomic treatment will 355 be required to elucidate these species delimitations, which is beyond the 356 scope of the present study but will be the focus of future work. Overall, the 357

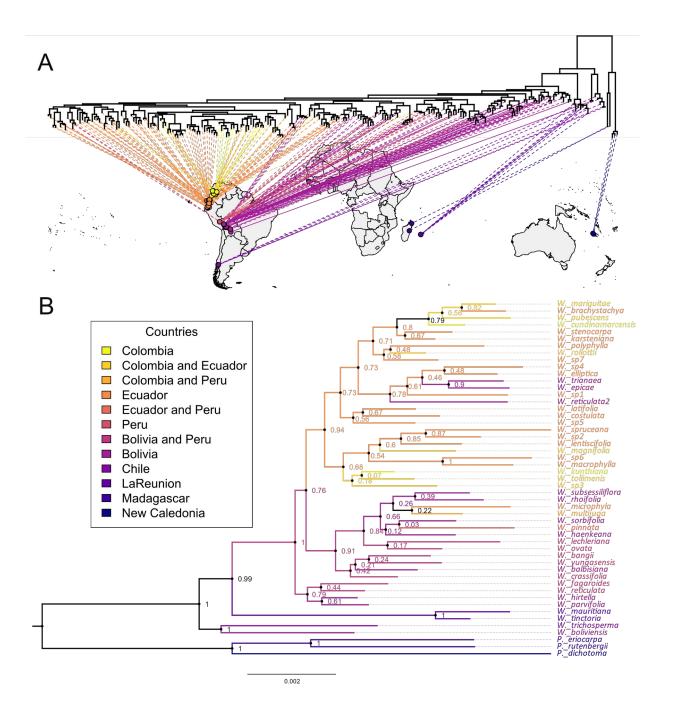


Figure 1. Geographic structure of phylogenetic relationships in Weinmannia. A. Specimen-level phylogeny359with tips projected onto geographic locations. Maximum Likelihood tree inferred from concatenated 2bRAD-seq360data from 234 individuals of Weinmannia plus outgroups. Tips labels and, bootstrap support values can be seen in361Supplementary Figure 1. B. Species-level Phylogeny for Weinmannia. Maximum Likelihood tree inferred from362concatenated 2bRAD-seq data from 48 individuals of Weinmannia plus 3 individuals in the outgroup. Bootstrap363support values are shown as node labels, tip labels and branches are colored by country where species were364collected.365

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specimen-level phylogeny showed a strong geographic structure within South 367
American *Weinmannia*, forming clades containing specimens collected in the 368
same region. Species from the Northern Andes (Ecuador and Colombia) 369
formed a clade nested within the *Weinmannia* crown group that included the 370
species sampled from the Central Andes (Bolivia and Peru), Southern Andes 371
(Chile) and Reunion Island (Fig. 1A, Supplementary Fig. 1). 372

373 Species-level phylogeny. The final concatenated character matrix for the 374 species-tree (51 tips) reconstruction contained 103,676 bp (2,879 loci) and a 48.63% of missing data. The concatenated ML species tree (ML; Fig. 1B;) 375 and the multi-species coalescent model-based species tree (hereafter SVDQ 376 377 tree; Supplementary Fig. 2) both showed strong bootstrap support [Bootstrap Support (BS) = 1 in both cases] for genus Weinmannia, confirming its 378 monophyly. The time-calibrated phylogeny based on the ML topology 379 showed that the MRCA of the genus Weinmannia diverged from the 380 outgroup in the late Eocene around 34.4 Ma and started to diversify ~20.7 381 Ma (Fig. 2), with similar results observed in the SVDQ tree analysis (~21.38 382 Ma; Supplementary Fig. 4). Congruent with our specimen-level phylogeny, 383 the species-level phylogenies also showed a general trend where 384 geographically proximal taxa were found in the same clade (Fig. 2; 385 Supplementary Fig. 4). 386

387 In the ML species tree, Weinmannia trichosperma, the southernmost species located in the temperate, extratropical forests of southern South 388 America, was placed in a clade that was strongly supported as the sister 389 group to the remaining species of *Weinmannia* (BS = 1; Fig. 1B), along with 390 391 Weinmannia boliviensis R. E. Fr. the southernmost species in the central 392 Andes inhabiting the subtropical Tucuman-Bolivian forests. In contrast, our SVDQ tree shows a topology with *W. boliviensis* as a basal branch, followed 393 by a clade consisting of the two Mascarene species, and then W. 394 trichosperma, all as successive sister groups to the remaining Weinmannia 395 species (supplementary Figs. 2 and 3). Despite the topological differences 396 between these two trees, both strongly support the placement of the 397 southernmost lineages *W. trichosperma* and *W. boliviensis*, alongside *W.* 398 *mauritiana* and *W. tinctoria*, as the most basally branching lineages in the 399 phylogeny. 400

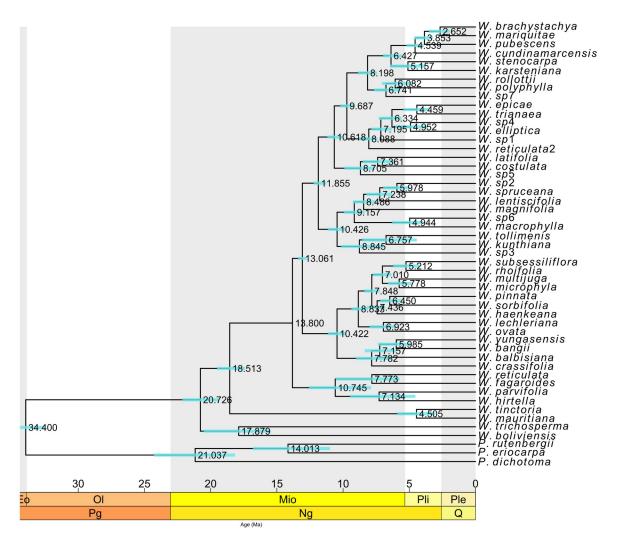


Figure 2. Maximum likelihood phylogeny with estimated divergence times of *Weinmannia* species. Median401divergence age estimates across bootstrap trees with 95% confidence intervals in blue bars.402

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**Tropical Andean clade shows geographic structure along the Central**404**and Northern Andes.**After the divergence of *Weinmannia* in southern405south America and the Mascarenes, the remaining 44 species, which are406found exclusively in the tropical Andes, formed a large, strongly supported407

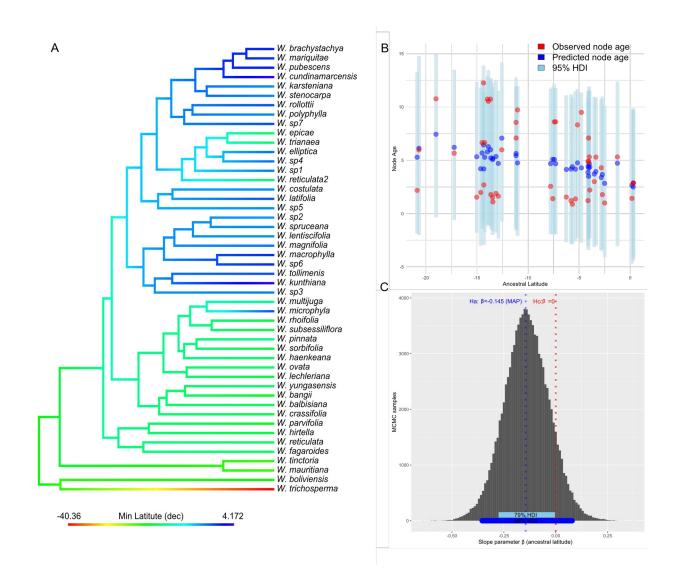
clade (BS=1 in Fig. 1B for ML; and BS= 0.98 in Supplementary Fig. 3 for 408 SVDQ), which started to diversify ~13.8 Mya (Fig. 2 and Supplementary Fig. 409 4). Even accounting for the topological differences among SVDQ and ML, 410 411 the Tropical Andean clade exhibited clear geographic structure. The ML reconstruction shows two clades at its base, the first containing four species 412 413 from Bolivia and Peru (BS= 0.79; Fig. 1B), and its sister clade (BS=0.76; Fig. 1B), which bifurcated into two major clades. The first of these clades 414 was strongly supported (BS=0.91; Fig. 1B) and included 13 species from the 415 416 central Andes, with exceptions such as *Weinmannia multijuga* Killip & A. C. Sm. from Peru and Colombia, Weinmannia pinnata L. from Peru and 417 Ecuador, and Weinmannia microphylla Ruiz & Pav. from Ecuador. The 418 second of these clades was also well supported (BS = 0.84; Fig. 1B) and 419 included most species from the northern Andes (Ecuador and Colombia), 420 421 with the exception of Weinmannia trianea Wedd., Weinmannia epicae A. 422 Fuentes, and what we call Weinmannia reticulata2 from Bolivia (Fig. 1B).

Our SVDQuartets species tree showed a similar pattern for the 423 424 Tropical Andes clade, in that it was divided into two major clades. The first contained 16 species, all found in the central Andes of Bolivia and Peru, with 425 426 the exception of W. multijuga, which is found from Colombia to Peru (BS= 0.83, Supplementary Fig. 2). The second clade was well supported (BS=0.84) 427 and contained the remaining 28 species, all from the northern Andes except 428 W. reticulata2 and W. balbisiana Kunth from Bolivia (Supplementary Fig. 429 2). 430

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Younger clades are distributed towards northern latitudes. Ancestral432state reconstructions for latitudes in internal nodes show that the more basal433branches are more likely to be associated with higher (*i.e.* more southern)434reconstructed latitudes (Fig. 3A and Supplementary Fig. 5A). Likewise, more435recently derived nodes are more likely to be associated with lower436reconstructed latitudes (*i.e.* more northern).437

Our Bayesian model predicting node age as a function of438reconstructed latitude on the ML phylogeny yielded a maximum *a posteriori*439(MAP) slope for latitude (β) of -0.145. The 79% credible interval estimated440



441 Figure 3. Analysis for testing the dispersal from southern latitudes towards the North Andes though the 442 Andes. A. Ancestral character estimate for latitude of hypothetical ancestors (nodes). Ancestral states were 443 reconstructed on the Maximum likelihood timetree using the minimum latitude of each of the 48 putative 444 Weinmannia species considering reviewed accessions. The colors in the figure depict a continuous gradient of 445 latitude, transitioning from southern temperate regions in red to northern tropical regions in blue, with intermediate 446 latitudes in the central Andes represented in green. B. Bayesian linear regression of node age as a function of 447 predicted ancestral latitude: Posterior predictive check. Observed values are represented in red dots. The blue 448 dots represent the maximum a posteriori estimates, skyblue bars represent 95% High density intervals (HDI). C. A 449 posteriori probability distribution for the estimated slope coefficient for latitude as a predictor of node age. 450 Maximum a posteriori (MAP) is equal to  $\beta$ =-0.145 and the 95%HDI in blue segment goes from to -0.354 to 0.078 which includes cero. However, the 79%HDI goes from -0.276 to -0.003. This result shows the slope is different 451 from zero ( $\beta$  =0) rejecting the null hypothesis with a 79% of credibility. 452

for this parameter ranged from -0.276 to -0.003, which does not include zero, 453 providing robust evidence to reject the null hypothesis ( $\beta$ =0; Fig. 3C) of no 454 relationship between node age and latitude. Even at a higher 95% credible 455 456 interval, the value of this parameter ranged from -0.354 to 0.078, showing a strong trend toward negative values (Fig. 3C). Using our SVDQ topology 457 458 with ML-optimized branch lengths, the more basal nodes tended to be associated with more southern latitudes (MAP for slope = -0.089; 459 Supplementary Fig. 5C), with the null hypothesis rejected with a 60% 460 461 credible interval ranging from -0.196 to -0.001 (Supplementary Fig. 5C). The results of the null hypothesis significance test using our nonparametric 462 bootstrap on the slope coefficient also showed that node age tended to be 463 negatively related to ancestral latitude when using our ML tree topology 464 465 (slope = -0.187, p-value = 0.0213). However, when using our SVDQ topology, despite the negative trend of the slope parameter, we did not 466 467 observe a statistically significant relationship between node age and latitude (slope= -0.107; p-value =0.1466). In summary, both Bayesian and NHST 468 approaches for both the ML and SVDQ inferred topologies, showed a 469 negative slope estimate even though the negative nature of the SVDQ 470 topology was not statistically significant. 471

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# Discussion

Our phylogenetic analyses of Weinmannia show that extratropical 474 475 species are preferentially placed at the base of the evolutionary tree. We also find a negative relationship between the age of clades and their latitudinal 476 477 distribution across the phylogeny. This suggests that Weinmannia probably originated from a lineage that was pre-adapted to the climatic conditions of 478 the southern extratropics, and that the diversification of Weinmannia in the 479 American tropics occurred concomitantly with south-to-north dispersal when 480 suitable climates were created as a result of the Andean uplift. 481

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# An extratropical origin for Weinmannia

Our ML phylogeny places Weinmannia trichosperma from the 484 485 temperate forests of southern South America as sister to all the other species 486 of the genus Weinmannia (ML, BS = 1, Fig. 1). This is consistent with previous morphological and genetic reconstructions of the phylogenetic 487 488 relationships of species in the genus (Bradford, 1998, 2002), but is slightly 489 different from results with our SVDQ approach (Supplementary Fig. 2). The 490 paleontological evidence is also consistent with this basal placement of W. trichosperma and suggests that the lineage had a widespread distribution in 491 the southern extratropics during the Paleogene. For example, an early 492 Oligocene (~30 Ma) macrofossil record, Weinmanniaphyllum bernardii 493 494 R.J.Carp. & A.M.Buchan from extratropical Tasmania (Carpenter and Buchanan 1993), where the genus is now extinct, is morphologically similar 495 to W. trichosperma (Bradford, 1998, but see Bradford et al. 2001). Still, the 496 497 extinction of the genus outside of its modern range in the Americas and the Mascarenes may be related to a sharp reduction in forest cover due to the 498 formation of the Antarctic Ice Sheet in the early Oligocene, which has been 499 associated with a massive extinction of the Austral paleoflora across the 500 501 southern hemisphere (Francis, 1996; Truswell and Macphail, 2009). This massive extinction process may also explain why this old, basal clade [stem 502 age 20.7 Ma] is currently relatively poor in species richness (Fig. 1, 2), 503 despite showing evidence of having a wider geographic distribution in the 504 past (Crisp and Cook, 2009; Antonelli and Sanmartín, 2011b). 505

Also consistently placed among the basally diverging lineages is the 506 subtropical species *Weinmannia boliviensis* (Fig. 1a and S2). *W. boliviensis* 507 is distributed in the Tucuman-Bolivian forests (Harling and Fuentes, 2014), 508 which are subtropical montane forests on the eastern slope of the Andes, 509 extending from 23°S to 29°S (Cabrera, 1976, Olson et al. 2001). 510 Interestingly, even though the forests harboring *W. boliviensis* are disjunct 511 from those harboring *W. trichosperma* in southern South America and 512

contiguous with those of the Andean lower latitudes, W. boliviensis is 513 phylogenetically closer to W. trichosperma than to species in the central 514 Andes (Fig. 1). Furthermore, W. trichosperma and W. boliviensis differ 515 morphologically. For example, W. boliviensis has simple leaves whereas W. 516 trichosperma has compound leaves (Harling and Fuentes, 2014). According 517 518 to our ML phylogeny, W. boliviensis and W. trichosperma diverged during the Miocene, 17.9 Ma (Fig. 2). This date precedes the final uplift of the 519 520 central and southern Andes to their present elevations (above 3000 m), which 521 began during the late Miocene (~11 Ma) (Gregory-Wodzicki, 2000; Siravo et al. 2018), and the subsequent expansion of arid areas to the west and east of 522 the mountain range (the southern arid diagonal of South America; Rambo, 523 524 1952, 1953). Therefore, the evolutionary divergence of the base of Weinmannia would have occurred in forests composed of Gondwanan 525 526 lineages that persisted in South America during the Paleogene and Neogene 527 (Romero, 1986). These extratropical paleoforests were isolated from the tropical lowland flora of South America by environmental rather than 528 529 geographic factors (Jaramillo and Cárdenas, 2013) and accumulated their 530 own evolutionary uniqueness (Segovia et al. 2020a). Finally, the species of these forests would have been largely extinguished in the lowlands as a result 531 of the intensification of aridity triggered by the Andean orogeny, leaving 532 remnants in specific locations along the Pacific coast of southern South 533 America, on the Brazilian plateau, and on the slopes of the tropical Andean 534 mountain range, where the genus Weinmannia is now restricted (Villagran 535 and Hinojosa, 1997, 2005). 536

Expanded species sampling, combined with improved phylogenetic 537 resolution, has led to an improved understanding of the origin of the 538 *Weinmannia* clade in the tropical Andes. Bradford (2002) used a maximum 539 of six species from the tropical Andes to reconstruct phylogenetic 540 relationships, using chloroplast and nuclear markers. In contrast, we included 541 45 species from this region (Fig. 1) and a much greater number of variable 542 loci distributed throughout the genome. Moreover, our approach provides 543 much greater resolution than a previous study that employed similar genomic 544 approaches (Angiosperms353) but sampled only one Weinmannia species 545 from tropical America (Pillon et al. 2021). Likewise, formal analyses to test 546 547 the relationship between node age and latitude help resolve the direction of the dispersal of Weinmannia. The Bayesian approach to infer the direction of 548 migration in Weinmannia provides robust results because it avoids binary 549 decisions (e.g. using p-values) and provides a continuum of evidence, which 550 551 reduces the risk of fixed thresholds prone to type I and type II errors.

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# South-to-North dispersal through the Andean Corridor

554 The negative relationship between node age and latitude in our phylogeny reveals a late arrival of the lineage in the tropics and a south-to-555 north dispersal route along the Andes (Fig. 3 and Supplementary Figs. 5 and 556 557 6). We propose that as the Andean uplift created an environmental corridor at mid and high elevations throughout of South America, Weinmannia 558 559 gradually expanded its range from south to north. This scenario is consistent with fossil evidence documenting the oldest pollen records of Weinmannia in 560 the northern Andes are from the late Pliocene and Pleistocene (1.5–3.2 Ma) 561 562 (Van der Hammen et al. 1973, Hooghiemstra 1989). The crown age of the Northern Andes clade recovered in the present study is estimated at 11.9 Ma 563 564 (Fig. 2), predating the earliest fossil evidence for Weinmannia in the region by several million years. This inconsistency may be due to the low 565 preservation potential of the pollen or sampling intensity in the fossil record, 566 which can cause the date of first fossil appearance to be significantly younger 567 than the true arrival date of a taxon in a region (Smith and Peterson, 2002). 568

The dispersal route of *Weinmannia* suggests that the lineage evolved 569 first under the environmental conditions of the southern extratropics and 570 maintained the acquired adaptations during south-to-north dispersal. Plant 571 lineages often exhibit a high degree of phylogenetic niche conservatism 572 (Crisp et al. 2009), and this tendency likely allowed *Weinmannia* and other 573

extratropical lineages to rapidly colonize the similar environments created at 574 mid- and high-elevations following the uplift of the Andes at tropical 575 latitudes (Donoghue, 2008; Segovia and Armesto, 2015). Moreover, modern 576 patterns of phylogenetic turnover across elevational gradients along the 577 tropical Andes reveal a significant influence of ecological sorting of pre-578 adapted clades in shaping tree communities (Ramírez et al. 2019; Griffiths et 579 al. 2020; Linan et al. 2021a). This suggests that the route taken by 580 581 Weinmannia would have been a general corridor of environmentally driven 582 immigration of plant lineages into the tropical Andes. Although the environmental factors currently driving phylogenetic turnover and shaping 583 584 the biogeographic corridor in the past have not been directly tested, the presence of freezing temperatures in a regular year may have some influence. 585 On the one hand, freezing temperatures drive taxonomic turnover and 586 587 differentiate lower and higher elevation montane forests in the northern 588 Andes (Pérez-Escobar et al. 2022). On the other hand, phylogenetic similarities of tree assemblages along the Americas unveil a biological 589 590 corridor that is differentiated by whether freezing temperatures occur 591 regularly (Segovia et al. 2020a). Therefore, the current taxonomic and phylogenetic turnover, and thus the historical corridor shaped by freezing 592 temperatures in the tropical Andes, may exemplify the global pattern of 593 594 conservative evolution (Zanne et al. 2014) and distribution of frost tolerance in plants (Wiens and Donoghue, 2004). 595

596 The tree topologies differed somewhat between RAxML and SVDQ. These differences in phylogenetic tree topologies when using SVDquartet 597 and RAxML (Supplementary Fig. 3) may be due to their different underlying 598 principles and methodologies. SVDquartets uses a quartet-based approach 599 600 that relies on gene coalescence patterns without imposing specific 601 evolutionary models, allowing it to account for complex evolutionary signals such as incomplete lineage sorting or hybridization. In contrast, RAxML is a 602 603 maximum likelihood-based method that operates under defined evolutionary models to estimate relationships and branch lengths, potentially yielding a 604 simpler tree structure. Despite the differences in methods, the overarching 605 606 patterns remain robust across phylogenetic reconstruction approaches, with both SVDquartets and RAxML revealing similar geographic structures. Both 607 608 methods consistently identified a clade containing nearly all species from the northern Andes nested within a broader clade containing all species from the 609 610 tropical Andes (Fig. 1 and Supplementary Fig. 2). Although the trees show slight differences in statistical support, the observed trend of northern 611 lineages appearing more recently suggests that lineage dispersal likely 612 followed a northward progression through the Andes (Fig. 3 and 613 Supplementary Figs. 5 and 6). 614

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#### The intriguing history of the Mascarenes' Weinmannia

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617 Our ML phylogeny shows that a small clade containing two species from the western Indian Ocean, Weinmannia tinctoria Sm. and Weinmannia 618 619 mauritania D.Don, is nested within South American Weinmannia (Fig. 1 and Supplementary Fig. 2). This result confirms previous phylogenetic 620 621 reconstructions (Bradford, 2002), but differs from results obtained with our SVDQ approach (Supplementary Fig. 2). This result is surprising because it 622 implies a long-distance dispersal event that occurred after the last tropical-623 extratropical dispersal event for the lineage within South America, according 624 to our time-calibrated ML phylogeny (Fig. 2). Although cumulative evidence 625 shows high levels of intercontinental dispersal in tropical biomes (Givnish 626 and Renner, 2004, Gagnon et al. 2019), this result is still surprising because 627 the stem age (~18.513 Ma, Fig. 2 and ~ 19.749 Ma, Supplementary Fig. 4) of 628 the Mascarene clade is older than the volcanic origin of the archipelago (less 629 630 than 8 Ma, McDougall and Chamalaun, 1969; McDougall, 1971). Furthermore, this result is unusual because botanical affinities and 631 phylogenetic evidence suggest that Madagascar may have acted as a source 632 of diversity for the Mascarenes (Cadet, 1977; Linan et al. 2019), but no 633 species of Weinmannia are currently found in Madagascar or Africa (Pillon 634 et al. 2021). Thus, any model for the dispersal of *Weinmannia* from South 635 America can only be proposed if Madagascar or Africa are involved as a 636 cryptic stepping-stone to reach the Mascarenes. In any case, further studies 637 are needed to properly address this intriguing disjunction and to clarify 638 possible vicariant or long-distance dispersal events in the origin of the genus 639 *Weinmannia*. 640

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## Conclusion

Weinmannia reflects a pattern in which the area of lineage origin is 643 644 poorer in species richness than the recently colonized area (e.g., 1 species in the extratropics, 2 species in the subtropics, and more than 70 species in the 645 tropical Andes). This pattern, likely due to massive extratropical extinctions 646 and recent diversification in the tropics, has also been proposed for the entire 647 family Cunoniaceae (Segovia et al. 2020b, Pillon et al. 2021), which has 648 traditionally been considered a lineage derived from the "Gondwanan" center 649 650 of plant diversification (Raven and Axelrod, 1974; Mildenhall, 1980; Hill, 1994). Furthermore, the fossil record indicates that the Cunoniaceae family 651 was present in Antarctica (i.e., western Gondwana) during the Late 652 Cretaceous (~70 Ma) (Poole et al. 2000), along with a highly diverse 653 vegetation similar in taxonomic composition to the temperate forests of 654 southern South America today (Poole et al. 2003). This suggests that the 655 biogeographic history of Weinmannia and Cunoniaceae may have been 656 shared with other lineages from the so-called Austral Floristic Realm 657 (Morrone, 2002; Moreira-Muñoz, 2007). Thus, our results provide insight 658 into an overlooked and seemingly counterintuitive phenomenon: the 659 taxonomic and phylogenetic enrichment of the hyperdiverse tropical Andean 660 floras by the immigration of remnant lineages from the ancient southern 661 extratropical floras of the Southern Hemisphere. 662

# Acknowledgements

The study was supported by FONDECYT 11200967 and the National 665 Science Foundation (DEB 1836353). R.A.S is supported by Institute of 666 Ecology and Biodiversity (IEB) ANID grant FB210006, F.F.G. was 667 supported by the Shirley A. Graham Fellowships in Systematic Botany and 668 Biogeography of the Missouri Botanical Garden. C.E.E. and the 669 Conservation Genetics Program at the Missouri Botanical Garden is 670 supported by donations from Stephen and Camilla Brauer, Philip and Sima 671 Needleman, and the Bellwether Foundation. Collections in Ecuador were 672 partially funded by Universidad Tecnológica Indoamérica, NHO. We thank 673 674 the Ministerio del Ambiente, Agua y Transición Ecológica in Ecuador for collecting permits (contrato marco MAE-DNB-CM-2019-011). 675

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Supplementary Figures for:

# Phylogeny of *Weinmannia* (Cunoniaceae) Reveals the Contribution of the Southern Extratropics to Tropical Andean Biodiversity.

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# 2024-08-13

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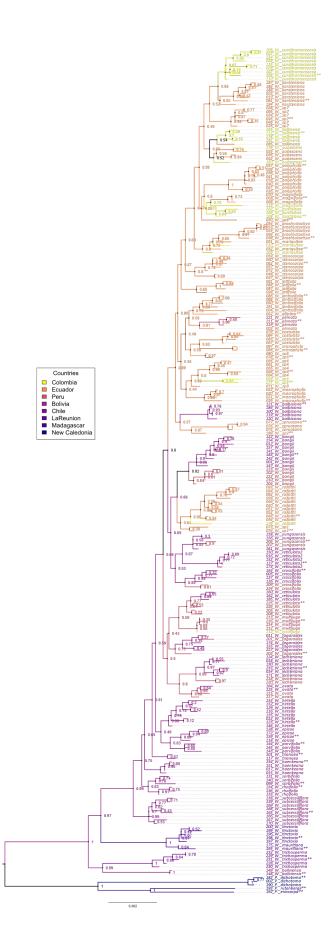
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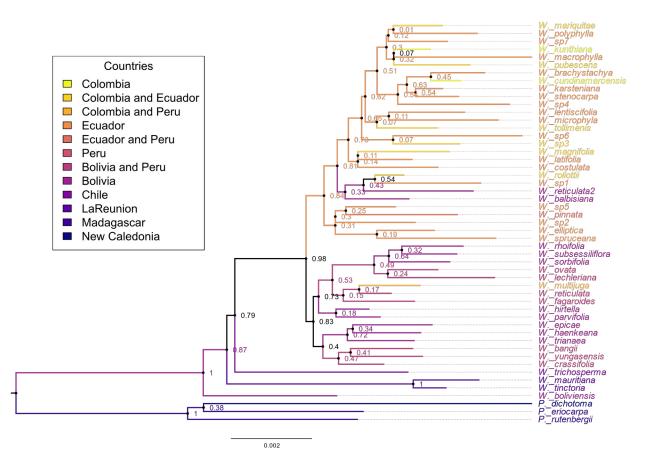
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Keywords: immigration, diversification, hyperdiversity, tropics, Gondwana



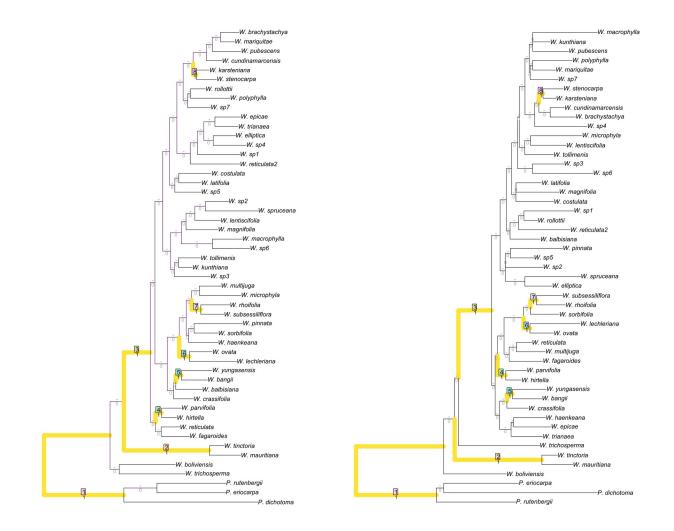
# Supplementary Figure 1. Individual-level 2bRAD-seq tree for Weinmannia.

Maximum Likelihood tree inferred from concatenated 2bRAD-seq data from 234 individuals of *Weinmannia* plus outgroups. Tip labels contain the Specimen\_ID\_code (see supplementary table 1) and Species Name. Tips labels are colored by country of origin (see legend). Bootstrap support values are shown as branch labels next to nodes. Specimens marked with \*\* were included in the species-level phylogenetic analyses.



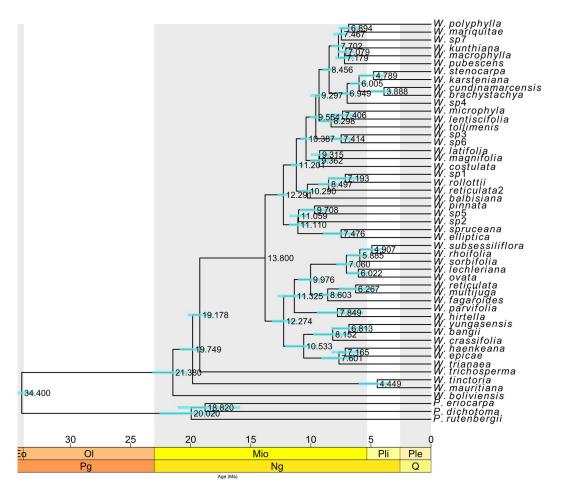
# Supplementary Figure 2. Species-level Phylogeny for Weinmannia.

SVDQuartets tree inferred from concatenated 2bRAD-seq data from 48 individuals of *Weinmannia* plus 3 individuals in the outgroup. Bootstrap support values are show as node labels, tip labels and branches are colored by country where species were collected. Branch labels were estimated using RAxML using this topology, (see methods).

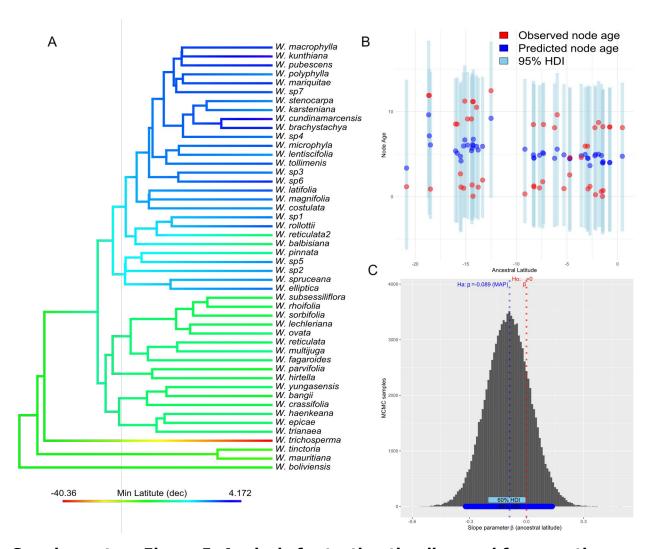


# Supplementary Figure 3. Comparison of RAxML (left) and SVDQuartets (right) species-level trees. *Weinmannia* phylogeny for both methods are compared by highlighting in yellow common splitting patterns. The number of ceros

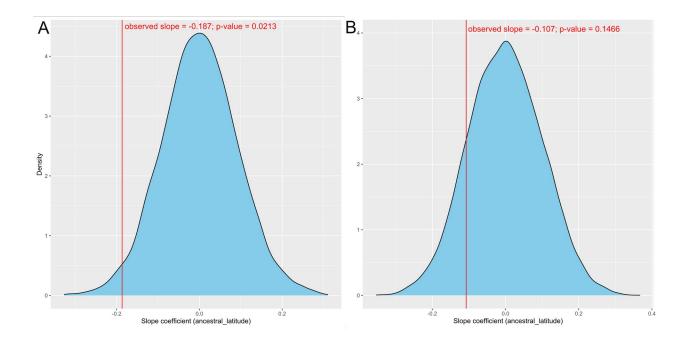
depicted at the trees' nodes represent Robinson-Fould distance among both trees. (Figure generated with TreeDist Package in R)



**Supplementary Figure 4.** SVDQuartets phylogeny with estimated divergence times of *Weinmannia* species. Median divergence age estimates across bootstrap trees with 95% confidence intervals in blue bars.



Supplementary Figure 5. Analysis for testing the dispersal from southern latitudes towards the North Andes though the Andes using topology inferred with SVDQuartets. A. Ancestral character estimate for latitude of hypothetical ancestors (nodes). Ancestral states were reconstructed on the SVDQuartets timetree using the minimum latitude of each of the 48 putative Weinmannia species considering reviewed accessions. The colors in the figure depict a continuous gradient of latitude, transitioning from southern temperate regions in red to northern tropical regions in blue, with intermediate latitudes in the central Andes represented in green. B. Bayesian linear regression of node age as a function of predicted ancestral latitude: Posterior predictive check. Observed values are represented in red dots. The blue dots represent the maximum a posteriori estimates, skyblue bars represent 95% High density intervals (HDI). C. A posteriori probability distribution for the estimated slope coefficient for latitude as a predictor of node age. Maximum a posteriori (MAP) is equal to  $\beta$ =-0.089 and the 95%HDI in blue segment goes from to -0.320 to 0.135 which includes cero. However, the 60%HDI goes from -0.196 to -0.001. This result shows the slope is different from zero ( $\beta = 0$ ) rejecting the null hypothesis with a 60% of credibility.



Supplementary Figure 6. Null-Hypothesis test for the slope coefficient when modelling Node Age as a function of ancestral latitude based on non-parametric bootstrap. A. Test performed using the Maximum likelihood Species-tree. Slope coefficient estimate of -0.187 was significantly different than cero p-val < 0.05 suggestin older lineages tend to be southwards. B. Test performed using the SVDQuartet Species-tree. In contrast with the previous estimate Slope coefficient estimate of -0.107 was not significantly different than cero p-val > 0.05 using this non-parametric bootstrap approach.