Tail-dependence of masting synchrony results in continent-wide seed scarcity

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19 Author Contributions Statement

²⁰ MB and JSz conceived the study, JSz, MB, VJ, JF, AHP designed the study. JSz performed

²¹ the analysis, MB led the writing of the manuscript. All authors contributed critically to the

²² interpretation of the analysis and drafts, and gave final approval for publication.

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24 Declaration of interests

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27 Data accessibility statement

²⁸ The data and R code used in this study have been deposited in the Open Science Framework

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³⁰ full MASTREE+ dataset is available in Hacket-Pain *et al.* (2022).

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38 Abstract

Spatial synchrony may be tail-dependent, meaning it is stronger for peaks rather than troughs, 39 or vice versa. High interannual variation in seed production in perennial plants, called masting, 40 can be synchronized at subcontinental scales, triggering extensive resource pulses or famines. 41 We used data from 99 populations of European beech (Fagus sylvatica) to examine whether 42 masting synchrony differs between mast peaks and years of seed scarcity. Our results revealed 43 that seed scarcity occurs simultaneously across the majority of the species range, extending to 44 populations separated by distances up to 1800 km. Mast peaks were spatially synchronized 45 at distances up to 1000 km and synchrony was geographically concentrated in northeastern 46 Europe. Extensive synchrony in the masting lower tail means that famines caused by beech seed 47 scarcity are amplified by their extensive spatial synchrony, with diverse consequences for food 48 web functioning and climate change biology. 49

keywords: Moran effect | mast seeding | geography of synchrony | pulsed resources | plant
 reproduction | seed production | spatial synchrony | tail-dependent synchrony

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53 Introduction

Spatial synchrony, the tendency of ecological phenomena to manifest correlated fluctuations 54 across diverse locations, impacts regional ecosystem functioning, leading to both large-scale 55 outbreaks and shortages (Ostfeld & Keesing, 2000; Earn et al., 2000; Liebhold et al., 2004; 56 Bjørnstad et al., 2002). One ecological phenomenon exhibiting diverse ecosystem consequences 57 and considerable spatial synchrony is mast seeding, a common reproductive strategy in perennial 58 plants (Koenig & Knops, 2000; Kelly & Sork, 2002; LaMontagne et al., 2020; Journé et al., 59 2023). In masting plants, seed production varies markedly between years, characterized by 60 frequent reproductive failures interspersed with peaks many times greater than the long-term 61 average (Kelly, 1994). This year-to-year variation is synchronized among individuals and can 62 extend to subcontinental scales, thereby triggering extensive disruptions in food webs (Ostfeld 63 & Keesing, 2000; Bogdziewicz et al., 2016; Clark et al., 2019). On the one hand, mast peaks 64 generate a resource pulse resulting in outbreaks of rodents, insects, and other seed consumers 65 that spread up the food web (Schmidt & Ostfeld, 2003), spikes in rodent-borne human diseases 66 (Jones et al., 1998; Bregnard et al., 2021), and elevated concentrations of allergenic pollen 67 (Tseng et al., 2020). Conversely, years devoid of seed production lead to famines, causing 68 rodent population crashes (Zwolak et al., 2018), poor or skipped reproduction in birds and 69 mammals (Ruf et al., 2006; Fidler et al., 2008; Cachelou et al., 2022), the emigration of seed-70

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eating wildlife (Zuckerberg et al., 2020), immigration of birds attracted to habitats where rodents 71 are scarce (Szymkowiak & Thomson, 2019; Maag et al., 2024), or increased human-wildlife 72 conflicts as consumers seek food outside of forests (Kozakai et al., 2011; Bautista et al., 2023). 73 The consequences of pulsed resources from mast peaks and famines resulting from years of seed 74 scarcity are magnified by the spatial scale of their synchrony (Yang et al., 2008; Bogdziewicz 75 et al., 2016; Clark et al., 2019). Novel theory indicates that this spatial synchrony may exhibit 76 tail-dependency, being stronger in either mast peaks or troughs (seed scarcity) (Walter et al., 77 2022), with significant and distinct ecosystem implications, but this has never been explored. 78

A recently introduced framework for investigating tail-dependence in spatial synchrony 79 helps to uncover novel patterns and enhance our understanding of the processes underlying 80 ecological synchrony (Ghosh et al., 2020, 2021; Walter et al., 2022). This framework explores 81 the tendency for the strength of the relationship between two variables to vary across the upper 82 and lower portions of their distribution (high and low tails, see Fig. 1), thereby influencing the 83 synchrony of peaks and troughs (Ghosh et al., 2020, 2021; Walter et al., 2022). In essence, a 84 strong association between an environmental driver and ecological response in either tail should 85 result in higher ecological synchrony within that tail (Ghosh et al., 2020, 2021; Walter et al., 86 2022). For instance, fluctuations in plankton (Ceratium) biomass exhibit greater synchrony 87 when scarce (lower tail) or abundant (upper tail), depending on whether the local relationship 88 with temperature is stronger in the lower or upper tail, respectively (Ghosh et al., 2020). In the 89 case of giant kelp (Macrocystis pyrifera), exposure to intense waves leads to declines in kelp 90 canopy (Walter et al., 2022). The relationship is non-linear, with low impact on kelp biomass 91 when waves are calm, resulting in increased synchrony of kelp population crashes in areas 92 where waves are intense (Walter et al., 2022). The tail-dependent nature of regional synchrony 93 shapes the resilience of regional populations, as regions marked by asynchronous crashes might 94 exhibit greater resilience to disturbance (Walter et al., 2022). Such non-linear relationships 95 between ecological and environmental variables are widespread and were anticipated to be a key 96 mechanism contributing to the emergence of tail-dependence in synchrony (Ghosh et al., 2020, 97 2021; Walter et al., 2022). 98

In masting plants, a major mechanism governing the annual allocation of resources to seed 99 production involves non-linear responses of seed production to weather variations, known as 100 weather cues (Kelly et al., 2013; Pearse et al., 2016). Consequently, the regional synchronization 101 of masting arises from the Moran effect, i.e. spatially correlated fluctuations in environmental 102 drivers of masting (Koenig & Knops, 2013; Ascoli et al., 2017; LaMontagne et al., 2020; Wion 103 et al., 2020; Bogdziewicz et al., 2021a; Reuman et al., 2023). The mechanisms underlying 104 weather cues are species-specific, with a common cue being the impact of summer temperature 105 on the stimulation of flower initiation, which is a primary determinant of subsequent seed 106 production (Satake & Kelly, 2021; Samarth et al., 2021; Journé et al., 2024). Thus, substantial 107 flowering effort, and subsequent large seed production, are triggered when the weather aligns with 108 species-specific criteria (Piovesan & Adams, 2001; Schauber et al., 2002; Fernández-Martínez 109



Figure 1: A graphical representation of the hypothetical association of seed production and weather cue and its consequence for spatial synchrony of mast seeding. In masting plants, seed production commonly increases only weakly at low values of the weather cue. Thus, in the lower tail of the environmental driver, the change in seed production per unit of weather cue is relatively small, promoting high spatial synchrony. In turn, the relationship between masting and weather cues in the upper tail is characterized by a steeper slope. Thus, relatively small spatial variations in weather translate into larger spatial variation in seed production (lower synchrony in the masting upper tail). In addition, in the lower tail, a single factor is sufficient to largely block seeding (i.e. low values of the weather cue), but multiple factors interact to determine the size of the final seed crop once high cue values trigger heavy reproduction (cue alignment). Secondary cues such as weather conditions during spring, which determine pollination success, shape the final seed production and create additional spatial variation.

et al., 2017; Nussbaumer *et al.*, 2018). If individuals and populations collectively respond to

the same cue across different populations, the spatial scale of masting synchrony aligns with the

broad-scale synchronization of weather patterns (Koenig & Knops, 2013; LaMontagne et al.,

¹¹³ 2020; Wion *et al.*, 2020; Bogdziewicz *et al.*, 2023a).

Tail-dependence in masting synchrony may arise from the often exponential or logistic 114 relationships between seed production and weather cues (Fernández-Martínez et al., 2017). For 115 instance, the logistic shape of the relationship between seed production and spring temperature 116 in temperate oaks (Quercus robur and Q. petraea) results in consistently low seed production for 117 temperatures below 12 C°, followed by a sharp increase above it (Schermer et al., 2020). Similar 118 non-linear responses have been observed in New Zealand flora (Kelly et al., 2013), Mediterranean 119 oaks (Koenig et al., 2015), pines (Parmenter et al., 2018), and American hardwoods (Smith et al., 120 2021). Logistic, log-linear, or sigmoidal responses to weather, leading to tail-dependence in the 121 relationship between masting and weather, can consequently drive tail-dependence in masting 122 synchrony. 123

¹²⁴ On one hand, a strong association between masting and weather in the upper tail should

result in stronger synchrony in mast peaks (Walter et al., 2022). Alternatively, synchrony in 125 seed scarcity may outweigh synchrony in masting peaks, owing to the differing influence of 126 weather cues on seed production at low versus high values. (Fig. 1). For low values of the 127 weather cue, a non-linear relationship results in little or no seed production across a relatively 128 broad range of weather cue, promoting spatial synchrony in seed scarcity by buffering against 129 fluctuations within that weather cue range(Fig. 1). In turn, the relationship between masting and 130 weather cues in the upper tail is characterized by a steeper slope (Kelly et al., 2013; Fernández-131 Martínez et al., 2017; Schermer et al., 2020). Due to that hypersensitivity, relatively small 132 spatial variations in weather translate into large spatial variation in seed production, lowering 133 synchrony in the masting upper tail. Another mechanism favoring the dominance of synchrony 134 in seed scarcity over peaks is that, while a single factor, such as the absence of a weather cue, is 135 sufficient to largely inhibit seeding, large mast peaks are contingent upon a sequence of events 136 occurring subsequently, termed cue alignment (Pesendorfer et al., 2016; Ascoli et al., 2021; 137 Yukich-Clendon et al., 2023). In European beech (Fagus sylvatica), for example, once flowers 138 are initiated, spring weather conditions can either enhance (in dry conditions) or impede (in 139 wet conditions) pollination (Ascoli et al., 2017; Journé et al., 2023). Even after successful 140 pollination, summer drought and heatwaves can still reduce the eventual seed crop (Nussbaumer 141 et al., 2020). As a substantial mast peak require the sequential occurrence of all these cues 142 (Ascoli et al., 2021), spatial synchrony diminishes in the upper tail. 143

In this study, we analyzed data from 99 populations of the major European species, European 144 beech (Fagus sylvatica), sampled across the species' range, to investigate the presence of tail-145 dependence in regional masting synchrony. The regional synchronization of beech masting 146 is attributed to the Moran effect (Vacchiano et al., 2017; Bogdziewicz et al., 2021a, 2023a; 147 Journé et al., 2024). Past research investigated mechanisms leading to large-scale synchrony 148 in overall masting pattern (Vacchiano et al., 2017; Bogdziewicz et al., 2023a), but a clear 149 understanding of the relative importance of peaks and troughs synchrony remains unexplored. 150 The forest-forming nature and extensive range of European beech (Leuschner & Ellenberg, 2017) 151 amplify the consequences of potential tail-dependence in beech masting synchrony. Beech's 152 seed production is responsive to consecutive summer temperatures, with sequential cold and 153 hot summers cueing high flowering initiation (Piovesan & Adams, 2001; Journé et al., 2023). 154 Here we used ΔT (difference between summer temperatures in one and two years before seed 155 fall) as the weather cue. First, we estimated the shape of the relationship between weather 156 cue and beech masting, expecting a non-linear response where the relationship is steeper in the 157 upper tail. This implies that seed production will have a weaker association with the weather 158 cue until large cue values are reached. According to the tail-dependence theory (Ghosh et al., 159 2020, 2021; Walter et al., 2022), such upper tail-dependence in the weather-masting association 160 should result in higher synchrony of masting peaks compared to synchrony of seed scarcity. 161 Alternatively, the synchrony of seed scarcity may surpass the synchrony of peaks if non-linear 162 association of masting and weather cue results in little seed production occurs across relatively 163

large temperature ranges in the lower tails. This allows seed scarcity to synchronize as long as
 temperature variations remain within that range (Fig. 1).

Materials and Methods

167 Materials

European beech (Fagus sylvatica L.) is a major forest-forming species in tem-Study system 168 perate Europe. Beech is a model masting species, with seed production characterized by large 169 interannual variation and synchrony of seed production (Nilsson & Wastljung, 1987; Ascoli 170 et al., 2017; Mund et al., 2020). Beech masting allows seeds to escape predation and increases 171 pollination efficiency (Nilsson & Wastljung, 1987; Bogdziewicz et al., 2020). Subsequent cold 172 (two years before seed fall) and hot (one year before seed fall) temperatures in June and July 173 trigger large seed production in European beech (Piovesan & Adams, 2001; Vacchiano et al., 174 2017; Journé et al., 2024). 175

Seed production and environmental data Annual observations of seed production of Eu-176 ropean beech were extracted from MASTREE+, an open-access database of annual records of 177 population-level reproductive effort Hacket-Pain et al. (2022). For our analysis, we restricted 178 the European beech time series to the continuous observations of seed production that covered 179 at least 10 years, observed after 1980. We choose 1980 as a cutoff as data before that date is 180 restricted spatially. We excluded pollen-based and ordinal records. This resulted in 99 time 181 series available for the analysis (median time series length, 29 years; maximum length, 41 years). 182 The number of years per time series is given in Fig. 2. We extracted daily weather data for each 183 site from the corresponding 0.1° grid cell of the E-OBS dataset (Cornes et al., 2018). 184

185 Analysis

Masting and weather cue relationship Our analysis started by estimating the relationship 186 between European beech seed production and the weather cue. Generally, beech seed production 187 is triggered by subsequent cold (two years before seedfall, T2) and hot (one year before seedfall, 188 T1) summers, specifically mean maximum temperature in June and July (Vacchiano et al., 2017). 189 The timing of that cue is conservative across the whole species range, as it is anchored to the 190 summer solstice (Journé et al., 2024). These two parameters (temperature in T1 and T2) can be 191 collapsed into one by taking their difference (ΔT , i.e. the difference between mean maximum 192 June-July temperatures in T1 and T2) (Kelly *et al.*, 2013). Thus, we used ΔT in our analysis as 193 it allows the estimation of masting-cue relationships in the tails with just one parameter. Model 194 fit, as judged with AIC, indicated a similar fit of ΔT model and T1 × T2 model (Table S1). The 195 relationship between masting and ΔT was estimated with a generalized additive mixed model 196



Figure 2: Locations of the 99 time series of annual seed production of European beech used in the study. Point size is scaled to the number of years (N) masting was monitored at the focal location, and the blue shading highlights the European beech natural range. An animation showing the spatiotemporal variation of seed production across the continent is supplemented as Video S1.

(GAMM) in which annual, per-site seed production was included as a response, site ID as a random intercept, and ΔT as a predictor. To accommodate among-site variation in methods of seed production monitoring, we scaled seed production for each site to fall in a set (0, 1) and fitted the GAMM with a beta family error term and logit-link (Journé *et al.*, 2023; Journé *et al.*, 2024). To scale the seed production values, we used the following equation:

$$y_i = \frac{\min - \max}{(\max_y - \min_y) \times (y_i - \min_y)} + \min$$
(1)

where y_i is seed production value of series y in year i, min_y and max_y are, respectively, minimal and maximal seed production values of series y, while min and max are minimal and maximal values to which seed production values of series y are scaled. GAMM model was fitted using R ver. 4.2.3 and mgcv ver. 1.8-42 package (R Core Team, 2023; Wood, 2011).

206 Tail-dependence in regional masting synchrony

Categorization of masting and weather into tails Our framework follows that of Walter *et al.* (2022). First, we divided masting and weather data into tails. Masting lower tail includes annual values of seed production ≤ 0.2 , while upper those > 0.2, for seed production scaled within each site to values between 0 and 1 (see above). That categorization reflects the distribution of annual

values of seed production (Fig. S1), and provides 1347 observations in the lower, and 890 in 211 the upper masting tail. The weather cue was divided into roughly two equal parts, using $\Delta T =$ 212 0 as the threshold. This follows from the nature of the masting-cue association, as high seeding 213 occurs once temperatures in summer T1 are larger than in T2, resulting in positive ΔT values 214 (Kelly et al., 2013; Vacchiano et al., 2017). The thresholds are fairly arbitrary in the sense that 215 neither masting nor weather is a categorical variable. Nonetheless, categorization was tailored 216 to the nature of the data, is biologically justified, and allows the tail-dependence to be analyzed 217 (Ghosh et al., 2021; Walter et al., 2022). Tail separation of masting data is required to estimate 218 the synchrony of mast peaks and seed scarcity. In turn, tail separation of the weather cue is 219 required to estimate the tail-dependent weather effects on masting (Fig. 1). 220

Within-tails masting synchrony We estimated the regional synchrony in masting tails using 221 a partial Spearman correlation, defined as the portion of the standard Spearman rank correlation 222 arising due to the range of values in the two variables being bounded by tails thresholds (Walter 223 et al., 2022). Pairwise correlations were calculated separately for the lower (≤ 0.2) and upper (> 224 0.2) tail of the seed production time series. In cases when the annual value of seed production for 225 the two sites falls into opposite tails, that value was included when calculating partial Spearman 226 correlation in both tails. Thus, if one site experienced a mast peak and the other a year of 227 seed scarcity in the same year, synchrony was reduced in both tails. We calculated pairwise 228 correlations between all pairs of sites, excluding pairs with less than 10 years of overlap. Note 229 that normalization of the masting data does not affect the correlations calculated via Spearman 230 correlation, as these are calculated on ranked data. 231

We calculated distance-decay of within-tail seed production synchrony using non-parametric spatial covariance functions (Bjørnstad & Falck, 2021). We used the matrices of partial Spearman correlations within the lower and upper tails as the response (synchrony variables), explained by the matrices of pairwise geographical distances between sites. To calculate 95% confidence bands for each function, we used standard bootstrapping procedure (Bjørnstad & Falck, 2021).

We used network analysis to visualize the biogeography of tail-dependent masting synchrony. 237 We built spatially-explicit networks of masting synchrony, whereby sites were nodes, while edges 238 were the pairwise synchrony of seed production within the lower and upper tail (scaled between 239 0 and 1). To test for spatial patterns in tail-dependent masting synchrony, we fitted generalized 240 linear mixed models (GLMM) with between-site synchrony (network edges) in either lower 241 or upper tail as a response, while including sites' latitude, longitude, and their interaction as 242 predictors. The models were fitted with beta error structure (logit link) and included site ID as a 243 random intercept. The models were fitted using glmmTMB ver. 1.1.5 R package, while networks 244 were delineated using igraph ver. 1.4.1 R package (Brooks et al., 2017; Csardi & Nepusz, 2006). 245

Drivers of tail-dependent masting synchrony We used the multiple regression quadratic as signment procedure with double-semipartialing (MRQAP) to investigate what drives the spatial

synchrony of mast seeding within the lower and upper tail. The MRQAP is a modeling framework 248 allowing investigation of the relationship between a dependent matrix and independent matrices 249 while considering the non-independence of network data by using permutation techniques to 250 test the significance of effect sizes (Dekker et al., 2007). We built two MRQAP models in which 251 masting synchrony within each tail, as measured by partial Spearman correlations, were the re-252 sponse matrices. In both models, the explanatory matrices were the synchrony (partial Spearman 253 correlations) of the weather cue (ΔT) and between-site similarity in the masting-weather cue 254 relationship in the focal tail. For the latter, we fitted site-level beta regression models with seed 255 production (scaled to fall in a set (0, 1)) as a response and ΔT as a predictor, separately for the 256 lower and upper tail of the weather data series. From these models, we extracted slopes for ΔT , 257 providing a measure of the strength of the masting-weather cue relationship at a given site and in 258 a given tail. We then calculated between-site pairwise similarities of those relationships, by first 259 calculating the Euclidean distances between the slopes for all pairs of sites and then converting 260 them into similarity indices by the formula (1 - [distance/maximum distance] (Bogdziewicz 261 et al., 2021a). Further explanatory matrices were between-site similarities in mean annual tem-262 perature and mean annual precipitation, as well as between-site spatial distance. To facilitate 263 effects comparison, all matrices were linearly scaled between 0 and 1. In both MRQAP models, 264 we tested for statistical significance using t-statistics and 1000 permutations. The models were 265 fitted using asnipe ver. 1.1.16 R package (Farine, 2013). 266

Results

Beech masting and weather. The relationship between beech seed production and summer 268 temperatures was sigmoidal (Fig. 3). We used ΔT (difference in June-July mean maximum 269 temperatures between one and two years before seed fall) as the weather cue because it captures 270 the effects of the past two summers in one parameter (see Methods). Seed production was 271 consistently low for ΔT values below 0 and started to sharply increase above that threshold. 272 Median seed production for ΔT values above 0 was 10-fold larger compared to below (Fig. 3b). 273 Separating weather into lower and upper tails confirmed tail-dependence in the masting-weather 274 cue relationship. Slopes of the relationship between masting and ΔT are higher in the ΔT upper 275 (mean = 3.88, logit scale, n = 99) tail compared to the lower (mean = 2.12, n = 99) tail (Fig. 3c). 276

Tail-dependence in regional masting synchrony. Regional beech masting synchrony was high in both tails. Nonetheless, synchrony was higher in lower tails (little or no seed production) compared to upper tails, despite the upper-tail association of masting and weather (Fig. 4). The mean absolute value of masting synchrony was higher in the lower tail compared to the upper tail at the majority (81 out of 99) of sites (Fig. 4B). Looking at these patterns in space, upper and lower tail masting synchrony were largely similar for distances up to 600 km. Tails started to diverge afterwards, with upper tail-synchrony 95% CI overlapping zero at ~ 1000 km. In



Figure 3: (A) Temporal dynamics of European beech *Fagus sylvatica*) seed production, with grey lines showing population-level data series, and the red line showing the yearly mean across all 99 sites. (B) Relationship between beech seed production and weather cue (ΔT , i.e. the difference between June-July mean maximum temperatures one and two years before seed fall). The solid line shows the model fit and shading associated 95% confidence intervals. The vertical dashed line shows the threshold used to define lower and upper tails in the ΔT . Box plots show the seed production for ΔT values in the lower (pink) and upper (green) tails. C) Distribution of population-level slopes (logit scale) from beta regression models examining masting vs ΔT relationships within the upper and lower tail of ΔT . Seed production is scaled within each site to values between 0 and 1.



Figure 4: Tail dependence in beech masting synchrony. A) distance decay of beech masting synchrony in the upper and lower tail. The lower tail is seed production below 0.2, while the upper is above, for annual values scaled within each site to between 0 and 1. That categorization follows from the shape of the distribution of the annual values (see Methods). Note that the values of synchrony in tails are slightly lower compared to overall regional synchrony (Fig. S2), which follows from categorization into tails and estimation based on partial Spearman correlation. B) Relationship between site-level mean synchrony of seed production in the upper and lower tail, with points scaled according to tail dependence strength (difference between mean synchrony in the upper and lower tail), and color-coded according to whether the mean falls into stronger upper- or lower-tail synchrony. Distribution and spatial synchrony of ΔT is provided in Fig. S3.

²⁸⁴ contrast, synchrony in the lower tail extends 1.8-fold further, with 95% CI overlapping 0 at 1800
²⁸⁵ km (Fig. 4a). Mapping tail-dependence of masting synchrony over Europe revealed that mast
²⁸⁶ peaks synchrony is lower in the southeast of the continent, but that pattern is largely absent in
²⁸⁷ the lower tails (Fig. 5, Table S2).

Drivers of tail-dependent regional masting synchrony. We used the multiple regression 288 quadratic assignment procedure (MRQAP) (Dekker et al., 2007) to investigate the role of weather 289 synchrony and the tail-dependent masting-cue relationship in driving the spatial synchrony of 290 mast seeding (Bogdziewicz et al., 2021a, 2023a). First, in agreement with masting synchrony 291 being driven by the Moran effect, among-site masting synchrony was positively associated with 292 among-site weather synchrony in both tails (Table 1). Interestingly, the effect size of weather 293 synchrony on masting synchrony was almost 3-fold larger in the upper tails ($\beta = 0.42$) than in 294 the lower tails ($\beta = 0.17$), which agrees with the notion that variation in weather synchrony has a 295 larger leverage on masting synchrony in the upper tail (Fig. 1). Masting synchrony in the lower 296 tails was associated with the masting-cue association in the lower tails, but that effect was not 297 significant in the upper tail (Table 1). In other words, the synchrony of seed scarcity between 298 sites was higher if these sites shared a similar response to the weather cue in that tail. Sites 299 that shared more similar mean annual precipitation and mean annual temperature had higher 300



Figure 5: Networks of masting synchrony in A), C) lower and B), D) upper tail. At A) and B) points are sites and lines show their pairwise synchrony in masting. Point sizes are scaled according to site-level mean synchrony of seed production within a given tail. C) and D) show synchrony as estimated with a GLMM model, see Table S2 for the model summary.

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masting synchrony in both tails (Table 1). Moreover, in both tails, the spatial distance between sites negatively correlated with their synchrony. 302

 Table 1: Multiple Regression Quadratic Assignment Procedure (MRQAP) models summaries
 that regressed the matrices of masting synchrony in the lower and upper tails against matrices of similarities in a masting-weather cue relationship in the focal tail, weather synchrony, similarity in climate (MAT: mean annual temperature, MAP: mean annual precipitation), and spatial distance. The weather cue used is ΔT (the difference between June-July mean maximum temperatures in one and two years before seed fall (see Methods). Key definitions of taildependence are provided in Fig. 1.

Predictor	slope	p-value
A) Lower tail synchrony ($R^2 = 0.21$)		
Intercept	0.33	< 0.001
Masting-weather cue similarity	0.10	0.004
Weather synchrony	0.17	< 0.001
Mean annual temperature similarity	0.09	< 0.001
Mean annual precipitation similarity	0.07	< 0.001
Spatial distance	-0.18	< 0.001
B) Upper tail synchrony ($R^2 = 0.38$)		
Intercept	0.21	< 0.001
Masting-weather cue similarity	0.06	0.189
Weather synchrony	0.42	< 0.001
Mean annual temperature similarity	0.14	< 0.001
Mean annual precipitation similarity	0.04	0.037
Spatial distance	-0.18	< 0.001

Discussion 303

Analyzing tail-dependent, regional beech masting synchrony revealed that seed scarcity can 304 occur simultaneously across the vast majority of the species range. Moreover, the synchrony 305 of seed scarcity extends 1.8 times further (approximately 1800 km) than the already extensive 306 synchrony of mast peaks (approximately 1000 km). Contrary to the predictions of the tail-307 dependence theory (Ghosh et al., 2020, 2021; Walter et al., 2022), the upper-tail association 308 between weather cues and masting did not result in stronger synchrony of masting peaks. Upper-309 tail masting synchrony was not higher than lower tail at any distance and started to decline more 310 rapidly. This result appears to stem from the nature of the relationship between beech masting 311 and its weather cue. The relationship translates into a minor increase in seed production across 312 a relatively broad range of cue values, promoting spatial synchrony of seed scarcity by buffering 313 against fluctuations within that cue range. In turn, for large values of the cue, the relatively small 314 among-site variation in temperature results in a large variation in seed production investment 315

across sites. Additionally, while it is sufficient for one factor to veto seed production (Pesendorfer *et al.*, 2016; Abe *et al.*, 2016; Bogdziewicz *et al.*, 2018), various other factors determine the size of a mast peak (Ascoli *et al.*, 2017, 2021; Journé *et al.*, 2023). The identification of taildependency in masting synchrony presented here provides new insights into the consequences of masting association with weather variation, with diverse consequences for food web functioning, masting forecasting, and climate change biology.

Synchrony in seed scarcity is higher compared to synchrony of mast peaks, despite the upper-322 tail association between masting and weather. We attribute that to the specific nature of the 323 relationship between masting and weather. In the case of environmentally-induced synchrony 324 in plankton, aphids, and kelp, a weak correlation with abundance across a specific range of 325 the focal environmental driver results in abundance fluctuations due to other local drivers. In 326 consequence, regional synchrony decreases in the focal tail (Ghosh et al., 2020, 2021; Walter 327 et al., 2022). In the context of masting, including in European beech, when the weather cue 328 values are low, seeding is inhibited or low rather than influenced by other environmental drivers. 329 Second, while the absence of the weather cue largely blocks seeding across distant sites, large 330 mast peaks require the simultaneous occurrence of several events (Ascoli et al., 2017; Journé 331 et al., 2023). These events involve weather conditions throughout the stages of flower and fruit 332 maturation (Pearse et al., 2016). Mast peaks of similar magnitude require more events to align 333 through time and space, potentially diminishing synchrony. This notion is supported by the 334 MRQAP models, which found that similarity in the response of masting to the weather cue in 335 the upper tail was not a strong predictor of between-site masting synchrony. 336

The synchrony of mast peaks exhibits a biogeographical pattern, which is absent in the 337 regional synchrony of seed scarcity. A past study on European beech revealed that the regional 338 synchrony of masting decreases from northwest to southeast Europe (Bogdziewicz et al., 2021a). 339 This biogeographical division mirrors the regional synchrony of temperature, following from 340 the higher complexity of topography in southeastern Europe and the prevalence of a unique 341 mode of climate variability in the northwest (the North Atlantic Oscillation) when compared 342 to the southeast of Europe (East Atlantic pattern, Scandinavian pattern and East Atlantic/West 343 Russia pattern) (Folland et al., 2009; Zuckerberg et al., 2020; Craig & Allan, 2022). Lack of 344 that geographical pattern in the synchrony of seed scarcity supports the notion that synchrony in 345 lower tails is partially buffered against temperature fluctuations compared to the more sensitive 346 synchrony of mast peaks. By segregating masting synchrony into tails, our results demonstrate 347 that mast failures are widespread, and synchronized across the species distribution. In turn, 348 while mast peaks exhibit synchrony across northern Europe, peaks in southeastern Europe are 349 desynchronized from the rest of central and northern Europe. 350

The regional synchrony of seed scarcity surpassing the synchrony of mast peaks holds significant implications for food webs, forest and wildlife management, and seed production forecasting. A key fitness benefit of masting lies in predator satiation (Kelly, 1994; Pearse *et al.*, 2016). The synchronized, substantial year-to-year variation in seed production starves seed

consumers in low-seeding years, facilitating the satiation of these consumer populations in mast 355 years (Zwolak et al., 2022). High regional synchrony of seed scarcity can enhance the starvation 356 of highly mobile seed consumers, such as large mammals or birds (Curran & Leighton, 2000; 357 Bogdziewicz et al., 2022). This finding is especially relevant for southeastern Europe, where 358 synchronous seed scarcity will aid satiation despite desynchronized mast peaks. Forecasting 359 years of high and low seed production emerges as an important tool for aiding management and 360 conservation efforts in ecosystems dominated by masting species (Pearse et al., 2021; Journé 361 et al., 2023). The higher spatial synchrony of seed scarcity means that spatial extrapolation of 362 failure forecasts is more feasible compared to mast peaks. In restoration projects, the extensive 363 spatial scale of seeding failure implies that acquiring material for replanting may be restricted 364 in some years, as poor seeding can extend over the majority of the species range. Finally, in 365 some species, including European beech, tree and leaf growth is reduced in years of high seed 366 production, leading to synchronized fluctuations in carbon sequestration (Hacket-Pain et al., 367 2018; Vergotti et al., 2019; Mund et al., 2020). Dividing masting synchrony into tails indicates 368 less extensive synchrony of peaks, and consequently a more geographically restricted role for 369 masting as a driver of synchrony in forest growth and carbon sequestration variability. 370

Interannual variation in seed production is driven by two major physiological mechanisms: 371 non-linear responses of seed production to weather variation, or weather cues, and endogenous 372 resource dynamics (Satake & Iwasa, 2000; Crone & Rapp, 2014; Pearse et al., 2016). Years 373 with little or no seed production happen either when weather cue was absent, or following high-374 seeding years when resources are depleted after high reproductive effort (Crone et al., 2009; 375 Han & Kabeya, 2017). Thus, high regional synchrony in seed scarcity could also be a direct 376 consequence of the synchrony of mast peaks, to the extent that synchronized resource depletion 377 prevents seeding across extensive scales. Two outcomes of our analysis do not support such a 378 hypothesis. First, at distances above 600 km, synchrony in seed scarcity is higher than synchrony 379 in mast peaks. Second, in southeastern Europe mast peaks showed lower synchronization than 380 elsewhere in Europe, but synchrony of years of seed scarcity showed little difference between 381 southeastern Europe and the rest of the continent. Thus, while resource-related processes appear 382 important as synchronizing factors at local scales (Crone et al., 2009; Abe et al., 2016), they do 383 not drive synchrony of seed scarcity at regional scales. 384

One aspect of regional masting synchrony not addressed in our analysis is its potential 385 variability over time, which may be influenced by climate oscillations. The North Atlantic 386 Oscillation (NAO) is the leading climatic driver in Europe, exerting influence over spatial weather 387 patterns (Ascoli et al., 2021). Consequently, the extent of regional masting synchrony may 388 fluctuate in response to shifts in the NAO phase (Ascoli et al., 2017). Similarly, the positioning 389 of the jet stream over Europe plays an important role in shaping continental weather patterns, 390 thereby impacting forest productivity (Dorado-Liñán et al., 2022). The dynamic weather dipoles 391 created by the jet stream's positioning can thus alter the geographical distribution of masting 392 events. By averaging across temporal conditions, as done in our analysis, these effects are 393

³⁹⁴ overlooked. For instance, the geographical patterns of masting synchrony described in this ³⁹⁵ study are likely to exhibit temporal dynamics, representing an interesting avenue for future ³⁹⁶ research.

In summary, our findings reveal tail-dependence in mast seeding synchrony, stemming from 397 the nature of the relationship between seed production and weather variation. The next step is 398 to investigate whether the described tail dependence in European beech applies to other masting 390 plants. While the logistic or sigmoidal shape of seed production response to weather conditions 400 is common, it is not universal (Fernández-Martínez et al., 2017; Wion et al., 2020). We predict 401 that in cases where the response of masting to weather is linear, tail dependency in synchrony will 402 either not manifest (Walter et al., 2022), or be weaker and follow only from the secondary cue 403 effect (cue alignment; Fig. 3). Additionally, some species exhibit population-specific weather 404 cues (Bogdziewicz et al., 2023a; Fleurot et al., 2023). For example, the main determinants of 405 sessile oaks (Quercus petraea) seed production vary along climatic gradients (Fleurot et al., 406 2023). Exploring how such cue variation affects the regional synchrony of mast peaks versus 407 the synchrony of seed scarcity would provide valuable insights. Furthermore, our results 408 suggest that the synchrony of seed scarcity may absorb larger changes in temperature regimes 409 associated with climate change. In European beech, climate warming disrupts the interannual 410 variation and synchrony of mast seeding, leading to elevated losses to seed predators and lower 411 pollination success (Bogdziewicz et al., 2023b; Foest et al., 2024). Higher temperatures result 412 in an increased frequency of hot summers, which weaken the tree's responsiveness to the cue 413 (Bogdziewicz *et al.*, 2021b). To the extent that this disruption affects the relationship between 414 masting and weather in the upper tail, the synchrony of mast peaks may be more adversely 415 affected than the synchrony of seed scarcity. Applying a tail-dependence framework to masting 416 ecology opens up new research avenues that will contribute to an enhanced understanding of 417 masting biology and how tail dependence manifests across and influences ecological systems. 418

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428 Author Contributions Statement

⁴²⁹ MB and JSz conceived the study, JSz, MB, VJ, JF, AHP designed the study. JSz performed ⁴³⁰ the analysis, MB led the writing of the manuscript. All authors contributed critically to the ⁴³¹ interpretation of the analysis and drafts, and gave final approval for publication.

432

433 Declaration of interests

⁴³⁴ No competing interests to declare.

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436 Data availability statement

⁴³⁷ The data used in this study have been deposited in the Open Science Framework (OSF):
⁴³⁸ https://osf.io/vny4b/?view_only=5e233556ebed48e79fa89ef3ec002544. The full
⁴³⁹ MASTREE+ dataset is available in Hacket-Pain *et al.* (2022).

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Supporting Information

⁶⁶⁶ Szymkowiak et al. Tail-dependence of masting synchrony results in continent-wide seed scarcity

Table S1: AIC-based comparison of generalized linear mixed models testing the relationships between European beech masting and weather cues. We used mean maximum temperatures, T1 is one year while T2 is two years before seedfall. ΔT is the difference between the June-July mean maximum temperature in T1 and T2.

Model	AIC	ΔΑΙΟ	d.f.
June-July temp $(T1) \times$ June-July temp $(T2)$	-3941.0	0	6
June-July temp $(T1)$ + June-July temp $(T2)$	-3937.2	3.8	5
ΔT	np (T1) + June-July temp (T2) -3937.2 -3937.0 -3511.9		4
June-July temp (T2)	-3511.9	429.1	4
June-July temp (T1)	-3475.6	465.4	4

Table S2: Spatial patterns in tail-dependent masting synchrony. Effect sizes were estimated using generalized linear mixed models with beta error structure (logit link) that included between-site synchrony in either lower or upper tail as a response, and sites' geographical coordinates as predictors.

Predictor	β	SE	z-value	p-value
A) Lower tail synchrony				
Intercept	-2.006	0.479	-4.19	< 0.001
Latitude	0.058	0.009	6.08	< 0.001
Longitude	0.107	0.047	2.290	0.022
Latitude*Longitude	-0.002	0.001	-2.51	0.012
B) Upper tail synchrony				
Intercept	-0.142	0.688	-0.21	0.836
Latitude	0.016	0.014	1.19	0.235
Longitude	-0.198	0.066	-2.97	0.003
Latitude*Longitude	0.004	0.001	2.84	0.005



Figure S1: Distribution of annual seed production values used in the analysis, scaled within each site to values between 0 and 1. The vertical dashed line shows the categorization of masting into lower (left) and upper (right) tails.



Figure S2: Distance decay of beech masting synchrony, as measured by pairwise Spearman correlation, if seed production series were not split into tails. The data used is visualized in Fig 2.



Figure S3: (A) Distribution of ΔT values observed during the studied period. Vertical dashed line indicades mean ΔT value of 0.09. (B) Distance decay of spatial synchrony in ΔT , as measured by pairwise Spearman correlation, in the lower and upper tail, as well as if the ΔT time series were not split into tails.