

# Quantifying soil microbial effects on plant species coexistence: a conceptual synthesis

Gaurav S. Kandlikar (contact: [gkandlikar@lsu.edu](mailto:gkandlikar@lsu.edu))

Department of Biological Sciences, Louisiana State University

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## Abstract

Soil microorganisms play a critical role in shaping the biodiversity dynamics of plant communities. These microbial effects can arise through direct mediation of plant fitness by pathogens and mutualists, and over the past two decades, numerous studies have shined a spotlight on the role of dynamic feedbacks between plants and soil microorganisms as key determinants of plant species coexistence. Such feedbacks occur when plants modify the composition of the soil community, which in turn affects plant performance. Stimulated by a theoretical model developed in the 1990s, a bulk of the empirical evidence for microbial controls over plant coexistence comes from experiments that quantify plant growth in soil communities that were previously conditioned by conspecific or heterospecific plants. These studies have revealed that soil microbes can generate strong negative to positive frequency-dependent dynamics among plants.

Even as soil microbes have become recognized as a key player in determining plant coexistence outcomes, the past five years have seen a renewed interest in expanding the conceptual foundations of this field. New results include re-interpretations of key metrics from classic two-species models, extensions of plant-soil feedback theory to multi-species communities, and frameworks to integrate plant-soil feedbacks with processes like intra- and inter-specific competition. Here, I review the implications of theoretical developments for interpreting existing empirical results, and highlight proposed analyses and designs for future experiments that can enable a more complete understanding of microbial regulation of plant community dynamics.

## Keywords

multispecies coexistence, mutualists, mycorrhiza, pathogens, plant-soil feedback, theory-data integration, soil microbiome

The environment is not a structure imposed on living beings from the outside but is in fact a creation of those beings. Just as there is no organism without an environment, there is no environment without an organism.

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Richard Lewontin

*The Organism as the Subject and Object of Evolution*

## 1 Introduction

2 Like all organisms, plants simultaneously respond to and shape their environment. One  
3 aspect of the environment that is especially dynamic is the microbial community in the  
4 soil. Plants can actively alter the structure of the soil community, for example by secret-  
5 ing root exudates that promote the growth of some microbes over others. Plants can also  
6 affect the soil community more passively, for example by creating leaf litter that favors  
7 certain decomposing microbes over others. The soil community is itself a heterogeneous  
8 entity, comprising a diversity of microbes that can interact with plants directly as mutual-  
9 ists or pathogens, or indirectly by regulating nutrient dynamics and other soil properties  
10 in their role as decomposers. Through these complex networks of interactions, soil mi-  
11 crobes likely play an important role in structuring biodiversity and community dynamics  
12 in all terrestrial ecosystems (Van Der Heijden et al., 2008).

13 One plant community outcome for which there is growing interest and evidence  
14 of microbial regulation is that of plant species coexistence. A hallmark of this research  
15 has been a tight integration of theory and experiment (e.g. Bever et al., 1997; Kulmatiski  
16 et al., 2011; Stein and Mangan, 2020). Theory suggested a streamlined experimental de-  
17 sign for quantifying microbial effects on plant coexistence (Bever et al., 1997), and through  
18 meta-analysis of numerous such experiments, we now know that microbes can affect plant  
19 coexistence outcomes in a wide range of ecosystems (Crawford et al., 2019). Coexistence-  
20 promoting negative feedbacks most strongly arise among plant pairs that are distantly  
21 related, associate with similar mycorrhiza, and interact in soils to which they are native  
22 (Crawford et al., 2019), but this negative feedback is seldom strong enough to overcome the  
23 fitness imbalances between plants that microbes simultaneously generate (Yan et al., 2022).  
24 As a result, soil communities by themselves are unlikely to explain observed coexistence  
25 in plant communities, and building on simple pairwise pot experiments to understand  
26 how these effects play out in nature remains a challenge. To help foster continued inter-  
27 play between theoretical and empirical research as we address this challenge, I use this  
28 Synthesis as an opportunity to review recent theoretical advances and their implications  
29 for empirical work.

## 30 **Pairwise plant coexistence under soil microbial feedbacks**

31 Experimental research on soil microbial regulation of plant species coexistence was cat-  
32 alyzed by the theoretical framework of Bever et al. (1997), which evaluates microbial ef-  
33 fects on the dynamics of two plant species. In this model, each plant population grows  
34 exponentially at a rate determined by the composition of the soil microbial community.  
35 The composition of the soil community, in turn, is determined by the composition of the  
36 plant community, along with the strength of each species' conditioning effect. These inter-  
37 actions gives rise to feedbacks in the plant-soil system, such that the growth rate (fitness)  
38 of each plant species depends on its own frequency in the system. A formal model descrip-  
39 tion is available in the original publication (Bever et al., 1997) and in Appendix S1. Briefly,  
40 the model follows the the dynamics of two plant species 1 and 2, and the distinct soil mi-  
41 crobial communities  $A$  and  $B$  that each species cultivates (Fig. 1A). The rate at which plant  
42 1 conditions the soil towards community  $A$  is set to 1, and the relative rate at which plant  
43 2 conditions the soil towards  $B$  is denoted  $v$ . The effects of microbial community  $A$  on the  
44 growth rate of plants 1 and 2 are denoted  $m_{1A}$  and  $m_{2A}$ , respectively, and  $m_{1B}$  and  $m_{2B}$   
45 capture the effect of microbial community  $B$  on plants 1 and 2. Positive values of  $m_{iX}$   
46 indicate that plant species  $i$  perform better in soils with microbial community  $X$  than in  
47 soils without this microbial community; negative values indicate that plant  $i$  is suppressed  
48 by microbial community  $X$  (Fig. 1A-B).

49 Bever et al. (1997) presented two key insights about this model that set the stage  
50 for the design and analysis of subsequent empirical studies of microbially mediated plant  
51 coexistence. First, the authors showed that whether microbes drive positive or negative  
52 feedback in plant population dynamics is captured by the sign of a metric termed  $I_S$ :

$$I_S = (m_{1A} + m_{2B}) - (m_{1B} + m_{2A}) \quad (\text{Eqn. 1})$$

53 Positive feedback arises when microbial communities benefit their conditioning plant  
54 species more than they benefit the other species, or when microbes hurt the condi-  
55 tioning plant less than they hurt the other plant. Mathematically, this requires that  
56  $m_{1A} + m_{2B} > m_{1B} + m_{2A}$ . In nature, positive feedback is more likely among pairs  
57 of ectomycorrhizal plant species than among plants that associate with arbuscular  
58 mycorrhiza (Bennett et al., 2017; Teste et al., 2017; Van Nuland et al., 2023). On the  
59 other hand, negative feedback arises when conditioned soil communities benefit the  
60 conditioning species less than the other plant (or hurt the conditioning species more  
61 than the other plant). In nature, negative feedback is likely driven by the accumulation

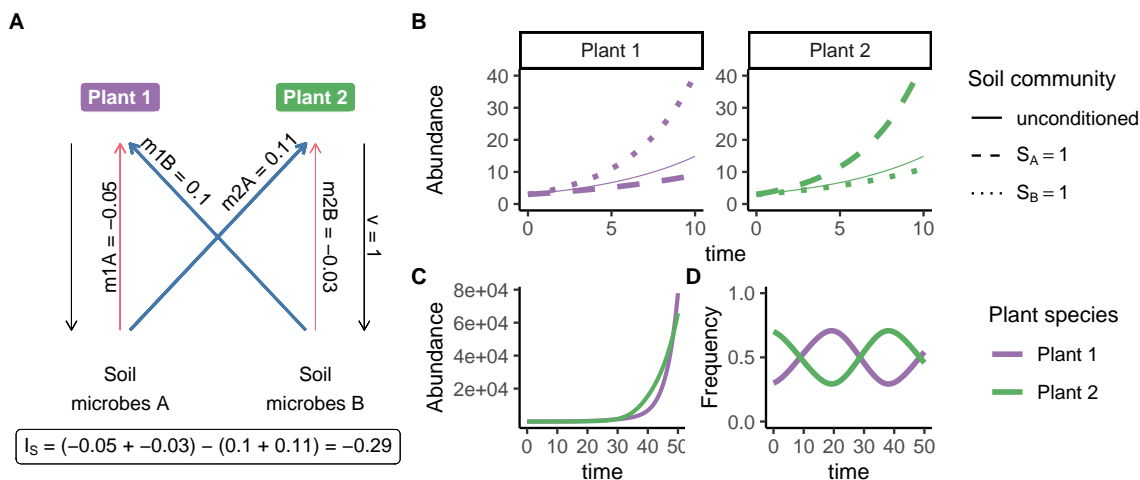
62 of host-specific soil pathogens, especially including fungi and oomycetes (Crawford et  
 63 al., 2019; Domínguez-Begines et al., 2021), but it can also arise when conditioned soil  
 64 microbes increase plant growth, provided that soil microbes benefit the conditioning  
 65 plant less than they benefit the other plant (Bever, 2002). Positive feedback hinders plant  
 66 diversity, because microbes provide a relative advantage to whichever species is more  
 67 frequent in the community. Negative feedback promotes diversity, because microbes  
 68 provide an advantage to whichever species is rare, allowing it to rise in frequency and  
 69 avoid extinction (Fig. 1C-D). Subsequent descriptions of this model further extended the  
 70 implications of  $I_S$  for species coexistence, as in Bever (2003):

71 “When the interaction coefficient is positive ( $I_S > 0$ ), the soil community  
 72 dynamics generate net positive feedback on plant growth and the compet-  
 73 ing plant species do not coexist. When the interaction coefficient is negative  
 74 ( $I_S < 0$ ), the soil community dynamics generate net negative feedback on  
 75 plant growth, and, as a result the competing plant species do coexist.”

76 The second key contribution of Bever et al. (1997) was a clear explanation of the  
 77 steps necessary for quantifying  $I_S$  empirically. This experimental design builds on im-  
 78 portant features of the parameters  $m_{iX}$ , and of the interaction coefficient  $I_S$ . Recall that  
 79 in this model, microbes only affect the rate of exponential population growth for the two  
 80 plant species (Fig. 1B-C). Assuming that biomass accumulation dynamics of individual  
 81 plants mirror the population growth process, and provided that all plants were exposed  
 82 to the soil community for an equal amount of time, one can estimate the  $m$  parameters  
 83 with the log-transformed biomass of plants grown in different soil microbial contexts:  
 84  $m_{iX} = \log(B_{iX}) - \log(B_{i0})$ . Here,  $B_{iX}$  is the biomass of plant  $i$  in soil community  $X$ ,  
 85 and  $B_{i0}$  is plant  $i$ 's biomass in reference (unconditioned) soil. In fact, Bever et al. (1997)  
 86 showed that the data requirements for quantifying  $I_S$  simplify even further. Due to the  
 87 arrangement of the  $m_{iX}$  terms, empirical quantification of  $I_S$  only requires biomass data  
 88 of plants grown with a conspecific- or heterospecific-conditioned soil community; growth  
 89 in unconditioned soils cancels out altogether:

$$I_S = \left[ \overbrace{(\log(B_{1A}) - \log(B_{10}))}^{m_{1A}} + \overbrace{(\log(B_{2B}) - \log(B_{20}))}^{m_{2B}} \right] - \left[ \overbrace{(\log(B_{1B}) - \log(B_{10}))}^{m_{1B}} + \overbrace{(\log(B_{2A}) - \log(B_{20}))}^{m_{2A}} \right]$$

90 Building on this insight, Bever et al. (1997) proposed a two-phased empirical design that  
 91 yields all the necessary  $B_{iX}$  terms for quantifying  $I_S$ . This design has been described in  
 92 detail elsewhere (e.g. Bever et al., 2012), and is summarized in Fig. S.1.



**Fig 1: Schematic and simulated model dynamics from Bever et al. (1997)'s canonical framework for plant-soil feedback.**

**A.** The model simulates the dynamics of two plant species (1 and 2) that cultivate distinct soil communities (*A* and *B*). Both plant species have some growth rate in unconditioned soils (set to 0.16 for the simulations in panels B-D), which is increased or decreased depending on the state of the microbial community, as described by the arrows (e.g. when the soil only reflects microbial community *A*, plant 1's growth rate decreases by 0.05, and plant 2's growth rate increases by 0.11). Following Eqn. 1, microbes generate  $I_S < 0$  (negative feedback) for this set of parameters. **B.** Plant population dynamics when each species is growing separately in soils that are unconditioned (thin solid line), wholly conditioned by Plant 1 ( $S_A = 1$ , dashed line), or wholly conditioned by Plant 2 ( $S_B = 1$ , dotted line). Note that these scenarios are only illustrative and not biologically plausible dynamics - for example, as plant 1 grows, it should become impossible for the soil state  $S_B = 1$  to persist, as plant 1's conditioning effects become evident. **C.** When both plants grow together, the soil community dynamically changes between  $S_A = 1$  and  $S_B = 1$ , as determined by the plant composition. Both plants experience exponential growth, at a rate determined by the composition of the soil. In this simulation each species' abundance periodically rises above the other's. **D.** The relative abundance (frequency) of each plant species. Microbes promote coexistence in this system by generating neutral oscillations. Model simulations were conducted with `deSolve::ode()` (Soetart et al. 2010) and visualized with Tidyverse packages (Wickham et al. 2019) in R v. 4.3.2 (R Core team 2023)

93 **Limits to inferring coexistence from  $I_S < 0$**

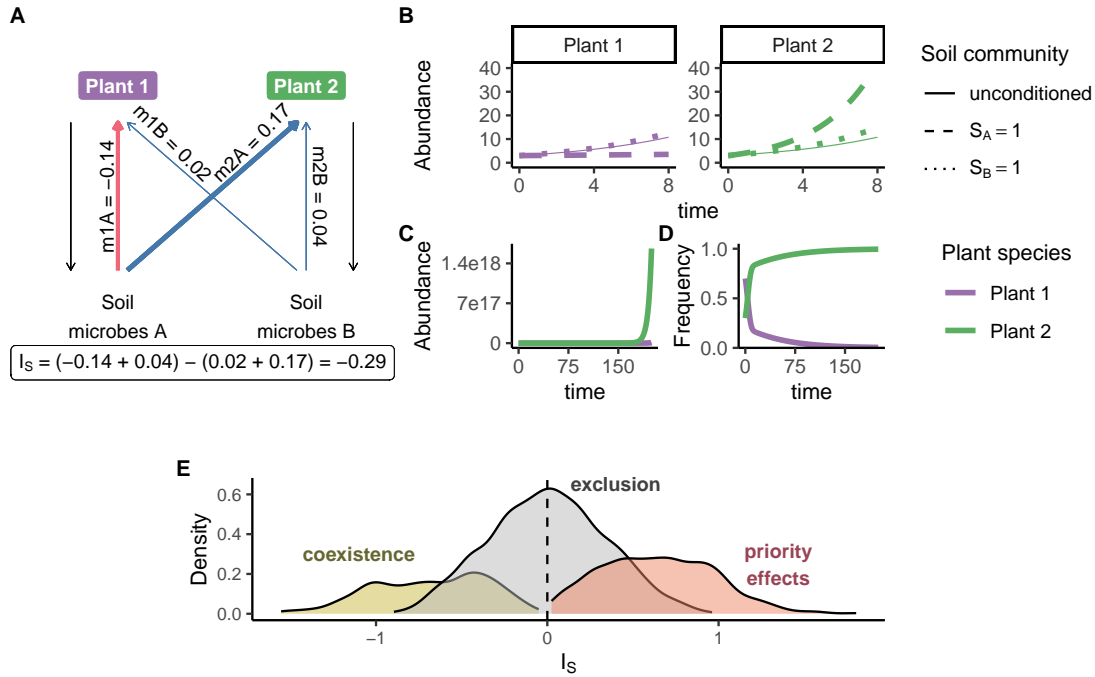
94 While the insights from Bever et al. (1997) have enabled a vast body of empirical work  
 95 (synthesized most recently in Crawford et al., 2019; see also Kulmatiski et al., 2008; Bever  
 96 et al., 2012), several recent studies have highlighted limitations to inferring microbially  
 97 mediated plant coexistence on the basis of negative feedback alone (Ke and Miki, 2015;  
 98 Kandlikar et al., 2019; Broekman et al., 2019; Beckman et al., 2023). The main takeaway

99 from this work is that while  $I_S < 0$  is a necessary condition for coexistence in the Bever  
100 et al. (1997) model, stabilizing effects of microbes do not *guarantee* long-term plant coexis-  
101 tence (Fig. 2). Part of the issue is that additional information that is not captured in  $I_S$  is  
102 required for accurate inferences of coexistence. This is not a new result per se: the original  
103 analysis and interpretation of  $I_S$  operates within the assumption that the soil microbes do  
104 not harm (or benefit) one species substantially more than the other (see pp. 563 of Bever et  
105 al. (1997)). However, in practice, this assumption is rarely tested, and the renewed clarity  
106 that one species can exclude the other despite  $I_S < 0$  represents a departure from the  
107 longstanding interpretation that the sign of this metric reflects whether or not microbes  
108 drive species coexistence. I discuss theoretical metrics and experimental designs that help  
109 overcome this assumption in the [following section](#).

110 Bever et al. (1997)'s analysis also builds on the assumption that both species condi-  
111 tion the soil community with roughly equal strengths ( $v \approx 1$ ). While explicit tests of this  
112 assumption are scarce, recent results raise questions about its generality. For example,  
113 low-abundance non-native species can have outsized effects on soil microbial communi-  
114 ties (Peltzer et al., 2009), which points to substantial interspecific variation in soil condi-  
115 tioning strength. Moreover, Chen et al. (2019) found that variation in the rates at which  
116 tree species accumulate pathogenic vs. mutualistic (ectomycorrhizal) soil fungi explains  
117 variation in the observed strength of conspecific negative density-dependence. While the  
118 magnitude of  $v$  does not change the coexistence criteria in Bever et al. (1997) model (see  
119 Appendix S1), strong asymmetries in conditioning strengths have important implications  
120 for the system's temporal dynamics. For example, for a given set of  $m_{iX}$  parameters that  
121 should result in coexistence,  $v \gg 1$  or  $v \ll 1$  result in extended periods of dominance by  
122 one species (Fig. S.2). This increases the risk of stochastic extinction of the rare species.  
123 Very few studies have systematically evaluated the consequences of varying conditioning  
124 strengths on the feedback process (but see Ke and Levine (2021)), and further theoretical  
125 and empirical evaluation of microbial conditioning dynamics should yield fruitful insights  
126 (see also Chung (2023)).

### 127 ***How to more thoroughly evaluate plant coexistence with soil feedbacks?***

128 Given that  $I_S < 0$  does not guarantee plant coexistence in the Bever et al. (1997), what  
129 other information can help generate more reliable inferences? At least two analytical ap-  
130 proaches address this question, yielding complementary insights. Both approaches are  
131 detailed in Appendix S1 and summarized here. The first approach was outlined in the



**Fig 2: Soil microbes can drive plant species exclusion even when they generate negative feedback.**

textbf A. In this simulation, plant 1's performance is suppressed by its own microbial community *A*, but boosted by microbial community *B*. On the other hand, plant 2's performance is increased by either conditioned community. The *m* terms yield the same negative  $I_S$  as in Fig. 1. **B.** Population growth dynamics for the two plants when soil is entirely conditioned by one plant or the other. **C.** When the two plants grow together, plant 2 has a higher rate of exponential growth than plant 1 at all times. **D.** Due to differences in the rates of exponential growth, the gulf in the two species' relative frequencies grows until the system is effectively entirely dominated by plant 2, and plant 1's frequency is nearly zero. **E.** The result shown in panels A-D is not exceptional: when microbes drive negative feedback ( $I_S < 0$ ), the two plant species coexist in only about half of the simulation runs; in the other half, only one species persists. While coexistence is never possible under positive feedback ( $I_S > 0$ ), inferring plant dynamics on the basis of  $I_S$  alone obscures the fact that in some cases, microbes give rise to frequency-dependent priority effects (species that is initially more abundant excludes the other), while in other cases, the same plant wins regardless of its initial frequency. Values of  $m_{iX}$  were drawn from a uniform distribution (minimum value: -0.5, maximum value: 0.5). The density graph summarizes outcomes from 2000 simulation runs, which were generated with `rootSolve::runsteady()` (Soetart and Herman 2009; Soetart 2009).

132 original model analysis, but has received little empirical attention. This approach pro-  
 133 ceeds by identifying parameter combinations that allow for equilibrium conditions that  
 134 are both feasible (meaning that all players are present with frequency  $> 0$ ) and neutrally  
 135 stable (meaning that perturbations to the equilibrium do not cause the system to collapse to  
 136 monodominance). A second approach for identifying coexistence outcomes in the Bever et  
 137 al. (1997) model was implemented in Kandlikar et al. (2019), and builds on the mutual in-  
 138 vasibility requirement for pairwise species coexistence (Turelli, 1978; Chesson and Ellner,  
 139 1989; Grainger et al., 2019). Applying the invasion criterion to the Bever et al. (1997) model  
 140 means that the plants can coexist if each species can successfully establish a foothold into  
 141 an equilibrium monoculture of the other plant (and its corresponding soil community).  
 142 Each species' population growth rate as it begins (or fails) to establish in its competitor's  
 143 monoculture is its "low-density growth rate", or LDGR. (This term is more commonly  
 144 called the "invasion growth rate" in the coexistence literature, but given the potential con-  
 145 fusion between this abstract property and the separate process of ecological invasions by  
 146 non-native plants, where soil microbes can also play an important role, I follow Lavorel  
 147 and Chesson (1995) and Hallett et al. (2023) in using the term "low-density growth rate").  
 148 For the Bever et al. (1997) model, the LDGR for each species is given by the following:

$$\text{LDGR}_{1 \rightarrow 2} = m_{1B} - m_{2B} \quad \text{and} \quad \text{LDGR}_{2 \rightarrow 1} = m_{2A} - m_{1A}$$

149 Here,  $\text{LDGR}_{1 \rightarrow 2}$  is the growth rate of plant 1 in a monoculture of plant 2, and vice-versa  
 150 for  $\text{LDGR}_{2 \rightarrow 1}$ . Coexistence requires that each species have a positive LDGR, meaning that  
 151 the following inequalities should be true:

$$m_{2B} < m_{1B} \quad \text{and} \quad m_{1A} < m_{2A} \quad (\text{Eqn. 2})$$

152 As shown in Appendix S1, this is identical to the coexistence requirements identified  
 153 through the feasibility analysis in Bever et al. (1997), and it mirrors the well-established  
 154 criteria for two-species coexistence in a Lotka-Volterra competition model.

155 Evaluating Eqn. 2 is enough for evaluating whether or not species can coexist  
 156 in the Bever et al. (1997) model, but further decomposing the LDGRs can yield useful  
 157 insights into the biological basis for coexistence outcomes. Specifically, following the ap-  
 158 proach described in Chesson (2000) and Chesson (2018), one can further decompose LD-  
 159 GRs into two terms. One term captures the degree to which the soil communities increase  
 160 (or decrease) the LDGR of both species, thereby favoring (or disfavoring) coexistence. The



161 second term captures the degree to which the microbial communities disproportionately  
 162 favor one plant species over the other, thereby increasing the LDGR of one species and  
 163 decreasing the LDGR for the other. Kandlikar et al. (2019) derived these terms for Bever  
 164 et al. (1997)'s model which, following convention (Chesson, 2000, 2018), are termed as  
 165 the microbially mediated “stabilization” and “fitness difference”, respectively. Whether  
 166 or not species can coexist is determined by the balance of these two effects. Specifically,  
 167 coexistence requires the following to be true:

$$\overbrace{-\frac{1}{2}((m_{1A} + m_{2B}) - (m_{2A} + m_{1B}))}^{\text{stabilization}} > \text{abs}\left(\overbrace{\left(\frac{1}{2}(m_{1A} + m_{1B}) - \frac{1}{2}(m_{2A} + m_{2B})\right)}^{\text{fitness difference}_{1,2}}\right)$$

(Eqn. 3)

168 Algebraically, the expression above is equivalent to Eqn. 2 (see Box S1.2 in Ap-  
 169 pendix S1). When this inequality is met, both species have positive LDGRs. Alter-  
 170 nately, when microbes primarily act to destabilize plant interactions (stabilization <  
 171 0 and abs(stabilization) > abs(fitness difference)), both species have negative LDGRs,  
 172 and microbes give rise to frequency-dependent priority effects (either species can form a  
 173 monoculture, but the two species cannot coexist (Yan et al., 2022; Zou and Rudolf, 2023)).  
 174 When fitness differences overwhelm the strength of (de)stabilization, one species has  
 175 negative LDGR, and the other has a positive LDGR. In this case, microbes drive exclusion  
 176 of the species with negative LDGR.

177 Evaluating microbial effects on the basis of the (de)stabilization and fitness differ-  
 178 ences provides valuable insight into how their net effects arise. For example, the accumu-  
 179 lation of species specific pathogens favors stabilization, but host-specific pathogens can  
 180 nevertheless drive exclusion if one plant suffers more from its pathogens than the other  
 181 (strong fitness differences). On the other hand, when plants are equally susceptible to  
 182 pathogens, even a small amount of host specificity can promote stable plant coexistence.  
 183 Moreover, framing soil microbial effects in terms of the degrees to which they generate  
 184 stabilization and fitness differences unlocks the potential to integrate soil microbes into a  
 185 broader theoretical framework that is actively being applied for studying how plant coex-  
 186 istence is mediated by pollinators (Lanuza et al., 2018; Johnson et al., 2022), seed consumers  
 187 (Petry et al., 2018), foliar pathogens (Uricchio et al., 2019), facilitation (Bimler et al., 2018),  
 188 and a host of other abiotic and biotic processes.

189 **Implications for empirical studies**

190 As with  $I_S$ , the complete coexistence criterion in Eqn. 3 is simply a linear combination of  
 191 the four  $m_{iX}$  terms that capture microbial effects on plant performance. In principle, this  
 192 might suggest that evaluating coexistence requires the same data as is required for quan-  
 193 tifying  $I_S$ . However, in practice, evaluating coexistence requires more information. This  
 194 distinction has to do with the role that plant performance in reference (uncultivated) soils  
 195 plays in determining  $m_{iX}$ . As shown above, plant biomass in reference soil cancels out of  
 196 the equation for  $I_S$ . This is also true for calculating stabilization; indeed, stabilization is  
 197 simply equal to  $-\frac{1}{2}I_S$ . However, plant growth in reference soil does not cancel out of the  
 198 fitness difference expansion:

$$\text{fitness difference}_{1,2} = \frac{1}{2} \left[ \overbrace{\log(B_{1A}) - \log(B_{10})}^{m_{1A}} + \overbrace{\log(B_{1B}) - \log(B_{10})}^{m_{1B}} \right] - \frac{1}{2} \left[ \overbrace{\log(B_{2A}) - \log(B_{20})}^{m_{2A}} + \overbrace{\log(B_{2B}) - \log(B_{20})}^{m_{2B}} \right]$$

199 The trivial implication of this result is that experiments aiming to infer plant coexistence  
 200 in the Bever et al. (1997) model should include an additional response phase treatment in  
 201 which plants are grown with a reference soil community (Kandlikar et al., 2019; Beckman  
 202 et al., 2023). However, theory alone does not provide an unambiguous guide for defining  
 203 the “correct” reference soil to use in an experiment. The original parameter descriptions  
 204 only define the reference soil by negation, as soil *without* a conditioning history of either  
 205 focal plant (Bever et al., 1997). In principle, this definition could apply equally well to any  
 206 soils where the focal species have not grown. Kandlikar et al. (2021) suggest that the ideal  
 207 reference soil for experiments reflects the microbial community that would exist in the rel-  
 208 evant field system when the focal plant species are absent. Alternatively, Beckman et al.  
 209 (2023) suggest soils conditioned by plants that associate with different mycorrhizal types  
 210 or that have different geographic origins than the focal species as potential references.  
 211 However, such soils are unlikely to include even low abundances of specialist pathogens  
 212 or mutualists that the focal species might encounter in nature, which could affect the esti-  
 213 mation of fitness differences and stabilization. When studies replace a specific condition-  
 214 ing phase and instead inoculate response phase pots with soils from adults in the field,  
 215 soil from bare patches devoid of vegetation may be an appropriate reference. Many past  
 216 studies included controls of plants growing in sterilized soils, but such soil is not an ap-  
 217 propriate reference for studies aiming to isolate the *effects of the conditioning/feedback process*

218 *itself* (Abbott et al., 2021). This is because the consequences of microbial feedbacks arise  
219 from differences in plant performance with a species-specific soil community vs. perfor-  
220 mance in soil that does not reflect the conditioning effect of the focal plant species. Thus,  
221 comparisons of growth in conditioned vs. sterile soils conflate the presence of a microbial  
222 community with the process of plant conditioning. Nevertheless, as I discuss later in this  
223 manuscript (see Fig. 3), plant growth in sterile soils provides an important baseline for  
224 studies aiming to quantify the coexistence consequences of soil microbes generally, rather  
225 than the conditioning/response process specifically (see also Yan et al., 2022; Ke and Wan,  
226 2023).

227 It is worth noting although the importance of reference soil growth is underscored  
228 by its prominence in the fitness difference calculation, the choice of reference soil also af-  
229 fects empirical estimates of  $I_G$ . For example, conditioning soils from a reference that con-  
230 tains low densities of the focal species' specialist pathogens can drive stronger stabilization  
231 (if the specialist pathogens proliferate during the conditioning phase) than conditioning  
232 from a reference that is completely lacking in specialist pathogens. In other words, all  
233 two-phase studies are built on implicit choices of a reference soil state. When the goal  
234 is to evaluate coexistence, plant growth in this same baseline soil community should be  
235 used to estimate  $B_{i0}$ . Preserving the reference soil community during the conditioning  
236 phase presents methodological challenges, as microbial communities are dynamic entities  
237 whose members grow and die (Abbott et al., 2021). Thus, future studies that couple refer-  
238 ence soil treatments with assays of microbial activity/composition (especially approaches  
239 that also quantify microbial abundances (Tkacz et al., 2018)), and/or include carefully de-  
240 signed controls to evaluate the effects of such microbial dynamics will help paint a more  
241 complete picture of how soil communities shape plant coexistence.

242 The specification of Bever et al. (1997)'s model also has implications for the analy-  
243 sis of data from two-phased plant–soil feedback studies. As explained above, the canonical  
244 approach for evaluating coexistence consequences of microbes is through comparisons of  
245 plant growth in different soil contexts (e.g.  $m_{1A} = \log(\text{biomass}_{1A}) - \log(\text{biomass}_{10})$ ).  
246 There is some ambiguity in the literature about the importance of log-transforming  
247 biomass measurements in such analyses.<sup>1</sup> Some authors omit this step entirely (or  
248 omit it from the reported methods, e.g. Bauer et al. (2017)), while some employ other  
249 transformations (e.g. square root transformation, Smith and Reynolds (2015)). When

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<sup>1</sup>Note: A previous version of this manuscript mistakenly suggested that Dudenhöffer et al. (2022)'s analysis featured a double-log transformation. These authors avoided the double-log through a back-transformation step that I had overlooked.

250 log-transformation *is* reported, it is often justified on the basis of the statistical properties  
251 (skewness) of the data (e.g. Duell et al., 2023). While ensuring that the appropriate data  
252 transformations are applied prior to model fitting is of course essential, log-transforming  
253 biomass data from plant-soil feedback experiments before calculating  $I_S$  serves more than  
254 just a statistical purpose. A key but implicit assumption to parameterizing the Bever et al.  
255 (1997) model with individual plant biomass data is that plant growth during the response  
256 phase of experiments is an exponential process, mirroring the exponential population  
257 growth of the underlying population dynamics model. Assuming that the final biomass  
258 value is the result of an exponential growth process, log-transforming the final biomass  
259 enables comparisons of the *rates* of biomass accumulation (Blackman, 1919) - i.e., the  
260 biomass analogs of the exponential population growth rate parameters  $m_{iX}$ . Moreover,  
261 whether or not microbial impacts on individual plant growth translate into corresponding  
262 effects on long-term population dynamics remains unclear. Such mismatch between  
263 microbial effects on biomass vs. population dynamics can arise for at least two reasons.  
264 First, while it is true that large plants generally achieve higher demographic fitness than  
265 do smaller conspecifics (Younginger et al., 2017), examples abound of species for which  
266 differences in biomass are poor predictors of fitness (reviewed in Fridley, 2017). Recent  
267 empirical results also suggest that microbial effects vital rates related to growth, survival  
268 and reproduction can be largely uncorrelated (Dudenhöffer et al., 2018; Chung et al.,  
269 2023b), which further complicates inferences of long-term dynamics from measurements  
270 of plant biomass. Thus, theoretical and empirical efforts to more thoroughly incorporate  
271 microbial effects on key life history processes are likely to yield important insights.

## 272 **Soil microbial feedbacks in more diverse plant communities**

273 Studies of plant coexistence are often motivated by diverse communities, but microbial  
274 mediation of plant coexistence is usually evaluated among species pairs. While pairwise  
275 analyses provide important insights, extending these results to interpret microbial effects  
276 on diverse plant systems can be challenging (Barabás et al., 2016; Levine et al., 2017). Sev-  
277 eral studies have addressed this gap through extensions to the classic two-species plant-  
278 soil feedback model. An early advance was that of Kulmatiski et al. (2011), who developed  
279 a model of three plant species and showed that the additional complexity of such a system  
280 can yield routes to coexistence that are not identified from pairwise analyses. For exam-  
281 ple, cyclic plant dynamics can arise even when each species performs better in its own soil  
282 community than in other species' soil (i.e.  $m_{1A} > \{m_{1B}, m_{1C}\}$ ) – an outcome that seem-

283 ingly contradicts the two-species coexistence criteria (Eqn. 2). More recently, Miller et al.  
284 (2022) extended the classic plant-soil feedback model to an arbitrary number of species  
285 and found that without any additional assumptions beyond those in Bever et al. (1997),  
286 robust coexistence of more than two species is virtually impossible. While it is possible to  
287 identify precise parameter sets yield oscillatory coexistence in this  $n$ -species model, this  
288 coexistence is fragile: minuscule perturbations to plant frequencies or to parameters cause  
289 the system to collapse to low-diversity (1 or 2 species). They conclude that stable multi-  
290 species coexistence is unlikely without accounting for other processes that regulate the dy-  
291 namics of plants or of soil microbes. One such source of regulation is to more thoroughly  
292 integrate plant-microbe interactions and plant competition into a unified framework, a  
293 topic I return to in a [following section](#).

294 Another potential source of regulation is through incorporating density-  
295 dependence in the microbial dynamics. This approach was implemented for two-species  
296 systems in Eppinga et al. (2006) and Aguilera (2011), and further extended to a multi-  
297 species plant system by Mack et al. (2019). This analysis identified a range of pathways  
298 through which microbes can enable multispecies plant coexistence, ranging from strict  
299 negative feedback to strict intransitivity in the system. Building on this model, Eppinga et  
300 al. (2018) analytically derived an  $n$ -species analogue of the pairwise stabilization metric  
301 termed  $I_C$ . As with  $I_S$ , negative values of  $I_C$  predict negative community-wide feedback,  
302 which is necessary for all  $n$  species persist in the system (see Appendix S2). Similar  
303 caveats also apply: while coexistence of all species is promoted by  $I_C < 0$ , negative com-  
304 munity feedback does not guarantee coexistence. Importantly for empirical application,  
305 quantifying  $I_C$  only requires a complete performance matrix (i.e. all combinations of  
306  $B_{iX}$ ), the likes of which are generated from fully-factorial pairwise plant-soil feedback  
307 studies of  $>2$  species.

### 308 ***Implications for empirical studies***

309 To date, the vast majority of experiments interested in evaluating microbial effects in di-  
310 verse communities have done so by inferring system-wide feedback from contrasts of pair-  
311 wise  $I_S$  at the species level (statistical summary of all  $I_S$  values involving each focal species  
312 (Mangan et al., 2010; Bauer et al., 2015)), or whole-community level (Pizano et al., 2019;  
313 Stein and Mangan, 2020; Dudenhöffer et al., 2022). While such statistical averaging of  
314 pairwise coefficients seems promising, theory suggests that inferring multi-species effects  
315 from such calculations comes with pitfalls (Barabás et al., 2016; Spaak and Schreiber, 2023)

316 that have not yet been formally evaluated in the context of plant-soil feedback. The theo-  
317 retical advances in Eppinga et al. (2018) suggest a robust alternative that is also frictionless,  
318 in that it does not require changing the two-phase design (Fig. S.1). Specifically, in sys-  
319 tems where the model’s assumptions regarding self-regulation of microbial dynamics are  
320 expected to apply, quantifying community-wide feedback through  $I_C$  provides a theoret-  
321 ically justified measure of microbial feedbacks on multispecies plant community structure.  
322 Moreover, doing so only requires the same information necessary to quantify species- or  
323 community-average  $I_S$ , and can yield surprising results. For example, Dudenhöffer et  
324 al. (2022) find that soil microbes most strongly *stabilize* pairwise plant coexistence under  
325 drought, but quantifying  $I_C$  for species triplets suggest that microbes most strongly *desta-*  
326 *bilize* multispecies systems under drought (Fig. S.3 and Appendix S2).  
327 Such analyses point to the value of future studies linking data with theoretically rigorous  
328 metrics of multispecies coexistence dynamics for advancing our understanding microbial  
329 regulation of plant dynamics in diverse systems.

### 330 **Contextualizing plant-microbe interactions relative to plant-plant interactions**

331 Plant-microbe interactions are one of many processes that simultaneously structure  
332 plant communities. Thus, while models and experiments that isolate the soil condition-  
333 ing/response process help establish the *potential* role of soil microbes in regulating plant  
334 communities, quantifying their contributions to plant coexistence in nature requires  
335 contextualizing their effects relative to those of other processes. Apart from their effects  
336 on plants’ intrinsic fitness – the primary focus of plant-soil feedback studies – soil  
337 microbes can also affect plant coexistence and community dynamics by altering the  
338 nature of plant interactions with herbivores (Koricheva et al., 2009), pollinators (Barber  
339 and Soper Gorden, 2015), and con- and hetero-specific neighbors (Lekberg et al., 2018).  
340 Projecting the long-term consequences of such interactive effects on plant coexistence  
341 requires careful coupling of experimental data with population dynamics models. For  
342 example, consider Bever (2003)’s integration of the classic plant-soil feedback process  
343 (Bever et al., 1997) into a system of plants that also interact via Lotka-Volterra competition.  
344 As detailed in Appendix S3, this relatively simple extension implies subtle changes in  
345 the assumptions about the effect that microbes have on plant intrinsic growth rates –  
346 assumptions that also change the empirical estimation and interpretation of the  $m_{iX}$   
347 parameters. Specifically, while the  $m_{iX}$  parameters in the 1997 model represent direct  
348 (additive) effects of microbes on plants’ baseline growth rates in unconditioned soils, the

349 specification of the 2003 model means that the corresponding parameters represent the  
350 *proportional* (multiplicative) effects of microbes on baseline growth. Thus,  $m_{iX}$  terms in  
351 this model have different units than in the 1997 model, and estimating these parameters  
352 from empirical data requires scaling microbial effects relative to the baseline growth  
353 ( $m_{iX} = \frac{B_{iX} - B_{i0}}{B_{i0}}$ , see Appendix S3 for details).

354 More fundamentally, contextualizing microbial effects relative to those of other  
355 processes requires us to use modelling frameworks that reflect the diversity of pathways  
356 through which microbes impact plant communities in nature. One such possibility is that  
357 the effect of the microbial community scales with changes in the abundance (or density)  
358 of microbes, rather than changes in relative frequency. Empirical studies suggest that  
359 such density-dependent microbial effects may be common among fungal or oomycete  
360 pathogens (Liang et al., 2016; Lamichhane et al., 2017). In these communities, if both fo-  
361 cal plant species are present at low abundances, then their cultivated soil microbes would  
362 also be expected at low abundances and thus exert only weak effects on plant growth. This  
363 is difficult to envisage in the feedback framework due its constraint that  $S_A + S_B = 1$ ,  
364 meaning that if plant 1's microbes are rare, then plant 2's microbes are automatically dom-  
365 inant. Modeling how the absolute densities of microbes (rather than relative frequencies)  
366 affect plant population dynamics can help overcome some of these limitations (Kandlikar  
367 et al., 2019; Ke and Wan, 2020). These abundance-based frameworks also help more di-  
368 rectly capture scenarios in which microbial effects on plant dynamics arise primarily by  
369 changing the nature of density-dependence, as has been shown to be the case for some ec-  
370 tomycorrhizal fungi, which benefit some plant hosts by reducing the competitive effects  
371 from neighbors rather than by increasing plants' intrinsic fitness per se (e.g. Liang et al.,  
372 2021). Importantly, although these changes to the feedback framework come with some  
373 cost of empirical tractability, the most relevant parameters can nevertheless be quantified  
374 from pot experiments tracking plant growth without explicit measurements of microbe  
375 dynamics (see Fig. 3, Ke and Wan (2020), and Ke and Wan (2023) for details).

376 Further departures from the feedback framework, which implies a strict corre-  
377 spondence between the number of plant species and microbial communities, also yield  
378 important insights. For example, tracking plant species' interactions with different types  
379 of mutualistic vs. harmful microbes can yield a more predictive understanding of condi-  
380 tions under which soil microbes contribute to coexistence vs. species replacement (Jiang  
381 et al., 2020; Schroeder et al., 2020). Integrating the role of soil microbes as mutualists,  
382 pathogens, and decomposers into a mechanistic framework for resource competition is

383 also a compelling path towards understand the relative importance of microbes on biodi-  
384 versity maintenance in nature (Chung et al., 2023a). Parameterizing such models requires  
385 experiments that carefully manipulate the abundance of different types of microbes and/or  
386 manipulate relevant abiotic factors like nutrient availability. These experiments are likely  
387 impractical in most field systems – especially in systems where the microbial communities  
388 themselves remain poorly understood. Nevertheless, insights from a few such studies in  
389 relatively tractable systems may enable a more mechanistic and generalizable understand-  
390 ing of how microbes shape plant communities across natural contexts.

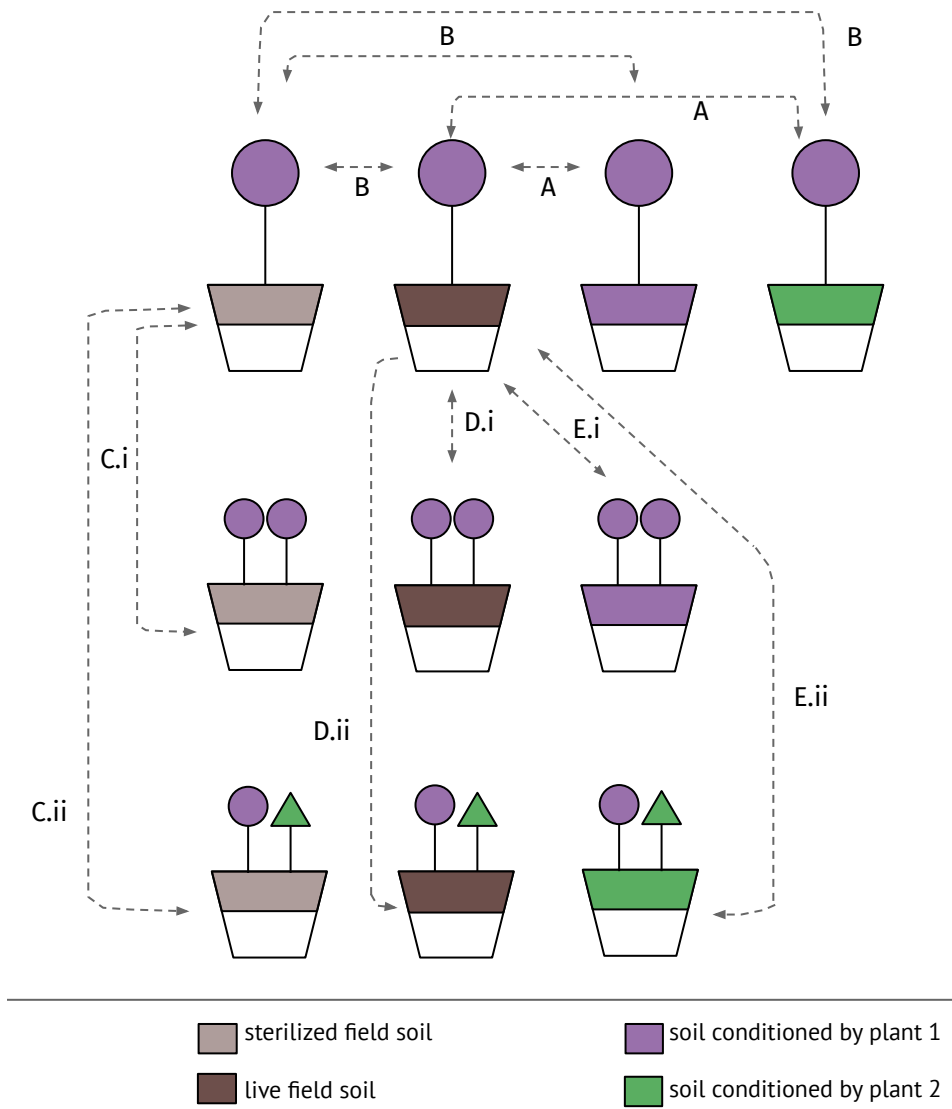
### 391 *Implications for empirical tests*

392 Evidence that soil microbial effects scale up to structure whole plant communities largely  
393 comes from studies that correlate outcomes from feedback experiments to properties like  
394 species' relative abundance (meta-analyzed by Reinhart et al., 2021), community stabil-  
395 ity (Chung et al., 2019), or productivity (Forero et al., 2022). While such work provides  
396 compelling evidence for the importance of soil microbes, the lack of an integrative frame-  
397 work for studying their effects stymies our ability to make sense of seemingly contradic-  
398 tory results. For example, species whose performance is more strongly suppressed by  
399 conspecific-conditioned soil communities tend to be less abundant on the landscape in  
400 some systems (Klironomos, 2002; Mangan et al., 2010), but the opposite pattern arises else-  
401 where (Corrales et al., 2016; Maron et al., 2016). In yet other systems, feedback strength  
402 and abundance are unrelated (Reinhart, 2012). A unified framework that integrates micro-  
403 bial effects with other processes structuring plant communities can offer useful insights  
404 for making sense of the diversity of patterns observed in nature. For example, explicitly in-  
405 tegrating plant-soil feedbacks and resource competition suggests that soil microbes drive  
406 plant dynamics most strongly when nutrients are less limiting; microbial effects are un-  
407 likely to affect plant competition when resource dynamics are slow (Kandlikar et al., 2019).  
408 Qualitatively, such a result is consistently with Corrales et al. (2016)'s conclusion that ef-  
409 fects of slow soil nitrogen cycling override any negative plant-soil feedbacks in driving  
410 the monodominance of an ectomycorrhizal tree in a tropical montane forest. Such con-  
411 gruence between theoretical and empirical conclusions points to the potential value of  
412 studies that directly integrate experimental data and demographic models for projecting  
413 how microbes shape the dynamics of natural plant communities.

414 As mentioned in the previous section, empirically quantifying the interactive role  
415 of plant–microbe interactions and other processes will require experimental treatments



416 beyond those in the classic two-phase design (Fig. 3). The particular design and analysis  
417 of these experiments will vary depending on the scope of which microbes' effects are being  
418 evaluated (e.g. is the goal to quantify the effects of soil microbes as a whole, or of the con-  
419 ditioning process specifically?), which demographic processes are affected by microbes  
420 (e.g. do microbes primarily affect intrinsic growth rates, density dependence, or both pro-  
421 cesses?), and whether the goal of the study is to evaluate long-term pairwise coexistence  
422 or a different outcome (e.g. do microbes enable coexistence, vs. do microbes help stabilize  
423 plant dynamics?). For example, studies that are focused on experimentally evaluating the  
424 role of soil of soil microbes in shaping conspecific density dependence specifically rather  
425 than coexistence more broadly can focus their efforts on the comparisons labelled C.i, D.i,  
426 and E.i in Fig. 3. If such a study were to find strong positive effects of neighbors in com-  
427 parison D.i but negative effects of neighbors in C.i, this would indicate that environmental  
428 microbes contribute to intraspecific facilitation. On the other hand, studies aiming to eval-  
429 uate the contribution of soil microbes to pairwise coexistence among species that also com-  
430 pete with one another would require the full complement of comparisons in Fig. 3. In this  
431 case, comparisons C.i and C.ii would measure intra- and inter-specific competition with-  
432 out microbes, while D.i/D.ii and E.i/E.ii would measure intra- and inter-specific competi-  
433 tion environmental and conditioned microbes, respectively. Whether one needs to include  
434 individual plant growth in sterile, live, and conditioned soils, or if only one of these is re-  
435 quired, depends on whether soil microbes are expected to only affect density-dependence,  
436 or also affect intrinsic growth. Data from such studies can be used to parameterize popu-  
437 lation dynamics models that enable long-term predictions of pairwise coexistence under  
438 different microbial contexts (e.g. Chung and Rudgers, 2016; Van Nuland et al., 2023). The  
439 rapidly developing toolkit for evaluating multispecies coexistence dynamics (e.g. Song et  
440 al., 2018) also provides a promising avenue for building towards a more comprehensive  
441 understanding of how microbes shape the dynamics of diverse plant communities. As  
442 studies scale up to include more experimental treatments necessary for parameterizing  
443 such models, ensuring that the observed patterns are consistent with model assumptions  
444 and that uncertainty in parameter estimates is properly propagated to long-term conclu-  
445 sions should remain a priority (Terry and Armitage, 2023).



**Fig 3: Potential design of a pot experiment that yields a more complete understanding of how microbes shape plant interactions.**

Arrows labelled **A** fall under the purview of the classic pairwise feedback framework; these comparisons help predict coexistence when plants only interact with one another through the soil community. Arrows **B** provide insight on how the whole microbial community - and not just the conditioning process - shapes coexistence. Arrows **C**, **D**, and **E** quantify plant-plant interactions (both intra- and inter-specific) in the absence of microbes, in the absence of the conditioning process, and when microbes are present and conditioned, respectively. Differences in arrows C-E can be used to infer how direct plant interactions and soil microbes jointly shape coexistence outcomes. For simplicity this figure only illustrates the soil treatments for one plant species; similar soil treatments are also required with plant 2 as the focal species for evaluating coexistence. Note that this design differs from the ‘minimal design’ of Ke and Wan (2020) by including individual plant growth in different soil backgrounds; these treatments can be omitted if microbes are thought to only affect the nature of density dependence rather than plants’ intrinsic growth. As highlighted in Ke and Wan (2023), additional density treatments may be required to evaluate the nature of density dependence in some systems.

## 446 **Conclusion**

447 Soil microbes play a key role in the dynamics of all terrestrial ecosystems. A tight integra-  
448 tion of theory and experiments over the past few decades has enabled rapid and sustained  
449 progress in our understanding of how soil microbes shape plant species coexistence. The  
450 theoretical advances reviewed here point to three areas of empirical research that should  
451 yield important insights:

- 452 1. While we now know that soil microbes can drive positive or negative feedback in  
453 a wide range of ecosystems, existing evidence also suggests that any such negative  
454 feedback rarely results in long-term coexistence (Yan et al., 2022). Evaluating the  
455 conditions under which soil microbes themselves give rise to pairwise *coexistence*  
456 (versus exclusion or priority effects) remains an open question.
- 457 2. While statistical averaging of pairwise metrics can provide useful insights into mi-  
458 crobial effects in diverse communities, theory shows that such analyses come with  
459 some pitfalls. Eppinga et al. (2018)'s analytically-derived community-wide stabi-  
460 lization metric can be parameterised with data from fully factorial feedback studies,  
461 and doing so has the potential to yield insights into microbial effects on multispecies  
462 systems that are masked in pairwise analyses.
- 463 3. Designing pot experiments with treatments informed by theoretical models that in-  
464 tegrate soil microbial effects with those of other processes like resource competition  
465 (e.g. Ke and Wan, 2020, 2023) will enable a more complete understanding of the  
466 conditions under which soil microbial effects scale up to affect plant community  
467 structure.

468 Continuing the interplay between theory and data is critical not only to improve our fun-  
469 damental understanding of how soil microbes shape plant coexistence, but also promises  
470 to generate actionable insights into the role of soil microbes in pressing environmental  
471 challenges like invasive species management habitat restoration.

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### 483 **Author contributions**

484 G.S.K: Conceptualization; Writing - original draft; Writing - review & editing.

### 485 **Data availability**

486 No data were used in this manuscript. Code for rendering all figures and manuscript  
487 documents is available at <https://gitlab.com/gklab/ajb-synthesis-public>.

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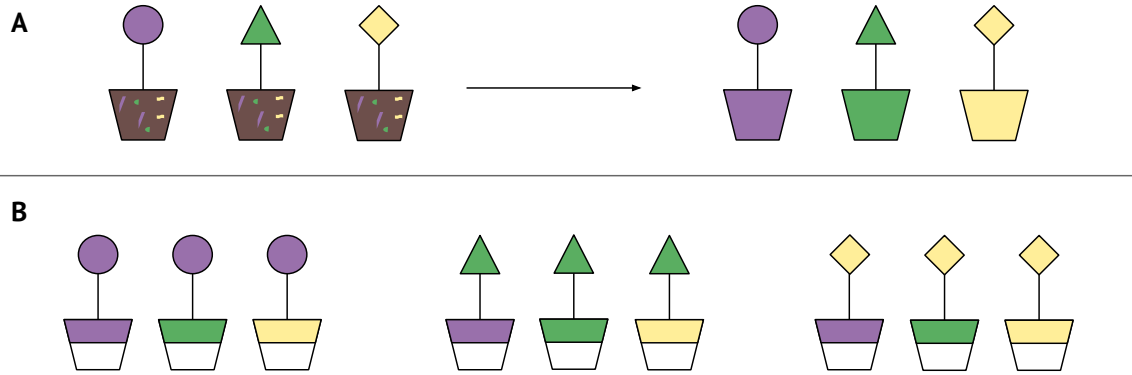
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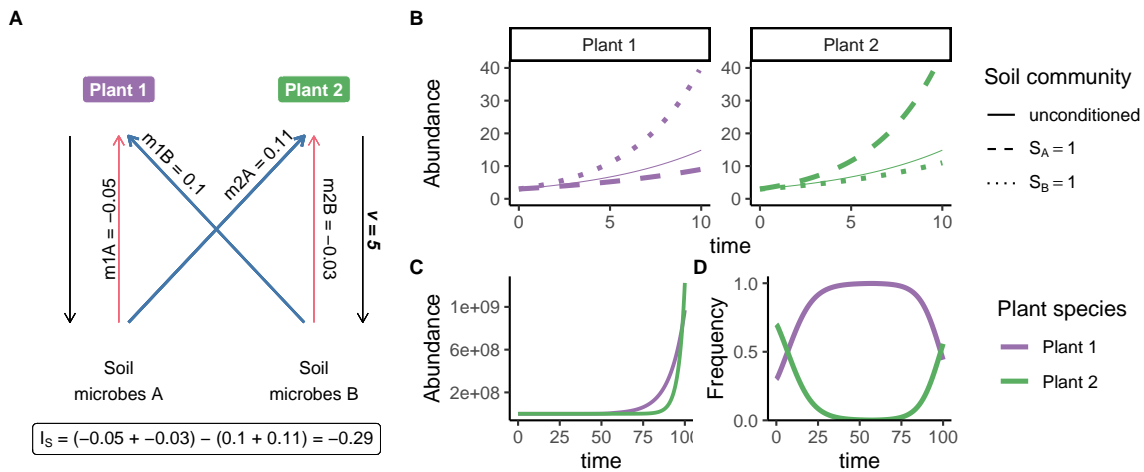
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## Supplemental Figures

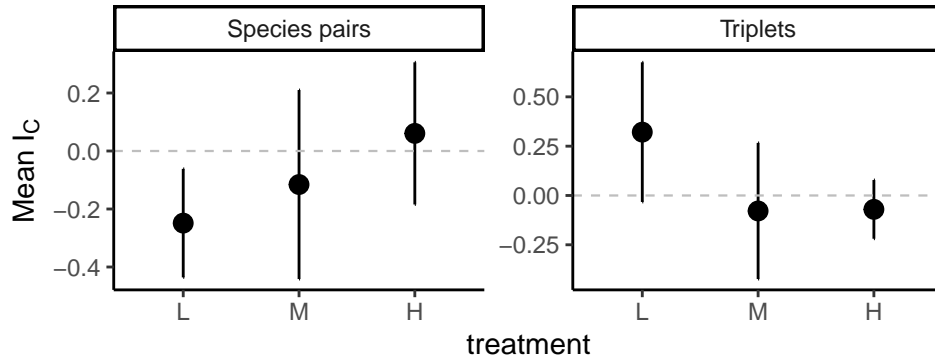


**Fig S.1: Schematic of the two-phase feedback experimental design.** **A.** In the first phase of the experiment, individuals (or monocultures) of each species are grown in soils that are identical at the beginning of the experiment. Over time, the plants grow, and the soil microbial community changes to reflect each species' unique conditioning effect (represented by distinct soil colors). **B.** In the second phase of the experiment, individuals of each species are grown, this time soils conditioned by conspecifics or by heterospecifics in the previous phase. A small volume of the conditioned inoculum is added to pots that primarily contain a common sterilized background soil (often  $\leq 10\%$  of the total soil volume in the pot is live conditioned inoculum, and the rest is bulk sterilized soil). Thus, soils should only differ in terms of their microbial community, and any nutritional differences that arise during the conditioning phase should not have a strong effect on plant growth in the response phase.



**Fig S.2: Variation in species' conditioning strengths affects the temporal dynamics of species coexistence**

**A.** This simulation uses identical  $miX$  parameters as in Fig. 1 of the main text, but now,  $v = 5$ , which means that plant 2 conditions the soil towards  $S_B$  more strongly than does plant 1 towards  $S_A$ . **B.** Population growth dynamics for the two plants when soil is entirely conditioned by one plant or the other; this is identical to Fig. 1B. **C.** When the two plants grow together, both plants have periods when they overtake the other in abundance, but there is an extended period of time when plant 1 is substantially more abundant than plant 2 - punctuated by brief periods during which plant 2 overtakes plant 1 in abundance. **D.** Due to differences in the rates of exponential growth of the species over extended periods of time, the gulf in the two species' relative frequencies grows until the system appears to be effectively entirely dominated by plant 1. Only over long periods of time does it become evident that plant 2 can rebound in abundance.



**Fig S.3: Microbial stabilization of species pairs and triplets under low, medium, and high watering regimes.** This figure shows results from an analysis estimating Eppinga et al. (2018)'s  $I_C$ , the multi-species analog of the pairwise interaction metric  $I_S$ , using data from Dudenhöffer et al. (2022). As in the original publication, we find that among species pairs, microbes exert stronger stabilization under drought ("low watering") than under high-watering regimes. However, among species triplets, the trend is reversed, with microbes generating slightly positive (destabilizing) feedbacks under drought, and slight negative (stabilizing) feedback under high-watering. Analysis details are available in Supplement S3, as are similar figures for 4 to 8 species communities that can be assembled from Dudenhöffer et al. (2022)'s study; these show that the result shown for triplets here generally extends to more diverse communities as well.

## Appendix S1: Conditions for coexistence in the classic plant-soil feedback model

Contact: Gaurav S. Kandlikar, [gkandlikar@lsu.edu](mailto:gkandlikar@lsu.edu)

This appendix begins with an overview of the dynamics model from Bever et al. (1997), including detailed steps to convert the underlying exponential growth equations for plants and microbes into equations for tracking changes in plant and microbe frequencies. After describing the model, I outline two approaches for identifying the conditions that allow long-term persistence of both plant species in this model. Note that throughout this appendix, I use  $N$  to denote state variables that reflect abundances, and  $F$  to denote frequency. The subscripts 1 and 2 refer to the plant species, and the subscripts  $A$  and  $B$  refer to their associated soil communities.

### Model description

The Bever et al. (1997) framework begins by considering a system comprising two plant species whose populations grow exponentially at a rate determined by the composition of the soil microbial community:

$$\frac{dN_1}{dt} = W_1 N_1 \quad \text{and} \quad \frac{dN_2}{dt} = W_2 N_2 \quad (\text{S1.1})$$

$W_i$ , the per-capita population growth rate of species  $i$ , is determined by the relative frequency of each microbial community ( $F_A$  and  $F_B$ ), and by the effect of each microbial community on plant  $i$  ( $m_{iA}$  and  $m_{iB}$ ):

$$W_i = m_{iA} F_A + m_{iB} F_B \quad (\text{S1.2})$$

Here, the two  $m$  terms have the units of  $\frac{1}{\text{microbe frequency} \cdot \text{time}}$ .  $F_A$  and  $F_B$  represent the relative frequency of each microbial community, rather than their absolute abundance. Thus,  $F_A + F_B = 1$ , and Eqn. S1.2 can also be written as  $W_i = m_{iA} F_A + m_{iB} (1 - F_A)$ , and  $W_i$  has units of  $\frac{1}{\text{time}}$ . Substituting this into the plant dynamics equation (S1.1) gives the full equations for plant population dynamics:

$$\frac{dN_1}{dt} = N_1 (m_{1A} F_A + m_{1B} (1 - F_A)) \quad \text{and} \quad \frac{dN_2}{dt} = N_2 (m_{2A} F_A + m_{2B} (1 - F_A)) \quad (\text{S1.3})$$

The abundance of soil microbial communities  $N_A$  and  $N_B$  also experiences exponential growth, with the rate of growth determined by the relative frequency of each plant<sup>2</sup>:

$$\frac{dN_A}{dt} = N_A \frac{N_1}{N_1 + N_2} \quad \text{and} \quad \frac{dN_B}{dt} = v N_B \frac{N_2}{N_1 + N_2} \quad (\text{S1.4})$$

The parameter  $v$  defines how strongly soil microbial community B accumulates with plant 2, relative to how strongly soil community A accumulates with plant 1.

Recognizing that plant population growth rates depend on the composition of the microbial community, which in turn depend on the relative frequency of each plant, we can express the system in terms of plant frequencies. This lets us simplify from the two equations in S1.3, to one equation for the frequency of plant 1 ( $F_1 = \frac{N_1}{N_1 + N_2}$ ):

$$\frac{dF_1}{dt} = F_1(1 - F_1)[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)] \quad (\text{S1.5})$$

By definition,  $F_2 = 1 - F_1$ , and  $\frac{dF_2}{dt} = -\frac{dF_1}{dt}$ .

Similarly, from the equations for tracking change in soil community abundance (Eqns. Equation S1.4), we can derive equations for the change in the frequency of microbial community ( $F_A = \frac{N_A}{N_A + N_B}$ ):

$$\frac{dF_A}{dt} = F_A(1 - F_A)(F_1 - v(1 - F_1)) \quad (\text{S1.6})$$

By definition,  $F_B = 1 - F_A$ , and  $\frac{dF_B}{dt} = -\frac{dF_A}{dt}$ .

Deriving Eqn. S1.5 from Eqn. S1.3, and deriving Eqn. S1.6 from Eqn. S1.4 requires application of the quotient rule. To make this derivation more accessible, I provide detailed steps in Box S1. After Box S1, I outline two complementary ways to evaluate the conditions for coexistence in this model (via [evaluating feasibility and stability of equilibria](#), or via [evaluating the low-density growth rates](#)).

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<sup>2</sup>Note that on p. 563 of Bever et al. (1997), the authors write (using slightly different notation) that  $dN_A/dt = N_A N_1$ , implying that the growth rate of microbial community  $A$  depends on the *abundance* rather than *frequency* of plant 1. I believe this was a typo.

**Box S1: Deriving the equation for plant frequency dynamics from exponential growth equations**

This box details the steps for expressing plant and soil microbial frequency dynamics (Eqns S1.5 and S1.6) from the exponential growth models (Eqns S1.3 and S1.4).

**Plant frequency dynamics** To derive the plant frequency dynamics equation, we first define  $F_1$  as the relative abundance of plant 1:  $F_1 = \frac{N_1}{N_1 + N_2}$ . Our goal now is to derive the equation for change in  $F_1$  over time:  $\frac{dF_1}{dt}$ .

We proceed by applying the quotient rule (for  $h(x) = \frac{f(x)}{g(x)}$ ,  $h'(x) = \frac{f'(x)g(x) - g'(x)f(x)}{g(x)^2}$ ) to get

$$\frac{dF_1}{dt} = \frac{d\frac{N_1}{N_1 + N_2}}{dt} = \frac{\frac{dN_1}{dt}(N_1 + N_2) - N_1(\frac{dN_1}{dt} + \frac{dN_2}{dt})}{(N_1 + N_2)^2}$$

Recalling that  $\frac{dN_1}{dt} = N_1(m_{1A}F_A + m_{1B}F_B)$  and  $\frac{dN_2}{dt} = N_2(m_{2A}F_A + m_{2B}F_B)$ , we can rewrite the equation as follows:

$$\frac{dF_1}{dt} = \frac{N_1(m_{1A}F_A + m_{1B}F_B)}{N_1 + N_2} - \frac{N_1(N_1(m_{1A}F_A + m_{1B}F_B) + N_2(m_{2A}F_A + m_{2B}F_B))}{(N_1 + N_2)^2}$$

Recalling that by definition,  $F_1 = \frac{N_1}{N_1 + N_2}$  and  $F_2 = \frac{N_2}{N_1 + N_2}$ , this equation simplifies as follows:

$$\frac{dF_1}{dt} = F_1[(m_{1A}F_A + m_{1B}F_B) - F_1(m_{1A}F_A + m_{1B}F_B) - F_2(m_{2A}F_A + m_{2B}F_B)]$$

Combining the first two terms in the square brackets gives:

$$\frac{dF_1}{dt} = F_1[(1 - F_1)(m_{1A}F_A + m_{1B}F_B) - F_2(m_{2A}F_A + m_{2B}F_B)]$$

Now, recognizing that  $(1 - F_1) = F_2$ , we can simplify this to:

$$\frac{dF_1}{dt} = F_1[F_2[(m_{1A}F_A + m_{1B}F_B) - (m_{2A}F_A + m_{2B}F_B)]]$$

Moving  $F_2$  outside the brackets, recognizing that  $F_2 = 1 - F_1$ , and recognizing that  $F_B = (1 - F_A)$  gives the frequency dynamics equation as stated in Eqn. 2 of Bever et al. (1997) (see also Eqn. S1.5 above):

$$\frac{dF_1}{dt} = F_1(1 - F_1)[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)]$$

*continued on next page*



### Soil frequency dynamics

Next, we derive the microbial frequency dynamics (Eqn S1.6) from the equations for change in microbial abundance (Eqn S1.4). As above, we first define  $F_A$  as the relative abundance of soil community A:  $F_1 = \frac{N_A}{N_A + N_B}$ . Our goal now is to derive the equation for change in  $F_A$  over time:  $\frac{dF_A}{dt}$ .

As above, applying the quotient rule yields:

$$\frac{dF_A}{dt} = \frac{d\frac{N_A}{N_A + N_B}}{dt} = \frac{\frac{dN_A}{dt}(N_A + N_B) - N_A(\frac{dN_A}{dt} + \frac{dN_B}{dt})}{(N_A + N_B)^2}$$

Recalling from above that  $\frac{dN_A}{dt} = N_A F_1$  and canceling terms gives:

$$\frac{dF_A}{dt} = \frac{N_A F_1}{N_A + N_B} - \frac{N_A(N_A F_1 + v N_B F_2)}{(N_A + N_B)^2}$$

Recognizing that  $F_A = \frac{N_A}{N_A + N_B}$ , and expanding out the second term, we can rewrite the equation as follows:

$$\frac{dF_A}{dt} = F_A F_1 - \frac{F_A(N_A F_1)}{N_A + N_B} - \frac{F_A(v N_B F_2)}{N_A + N_B}$$

Once again recognizing that  $F_A = \frac{N_A}{N_A + N_B}$ , we can further simplify the equation:

$$\frac{dF_A}{dt} = F_A F_1 - F_A^2(F_1) - v F_A F_B(F_2)$$

Factoring out  $F_A$  gives

$$\frac{dF_A}{dt} = F_A(F_1 - F_A F_1 - v F_B F_2)$$

We can further factor out  $F_1$  in the parenthetical term to rewrite the equation:

$$\frac{dF_A}{dt} = F_A(F_1(1 - F_A) - v F_B F_2)$$

Recognizing that  $1 - F_A = F_B$ , we can write:

$$\frac{dF_A}{dt} = F_A(F_1 F_B - v F_B F_2) = \boxed{F_A(1 - F_A)[F_1 - v(1 - F_1)]}$$

This is the same as Eqn. 3 in Bever et al. (1997), and Eqn. S1.6 above.

## Evaluating coexistence by analysing the feasibility and stability of equilibrium points

The first approach to deriving the conditions necessary for coexistence of the two plant species involves identifying the conditions under which the model's equilibria are both *feasible* and *stable*. Feasible equilibrium points mean that all the components of the system (in this case, the two plant species and their associated microbial communities) are present in the system at equilibrium; stability means that slight perturbations from the equilibrium do not push the system towards exclusion of one plant or the other. This approach builds on the insight that for two-species Lotka-Volterra models, feasible equilibria that are locally stable guarantee coexistence (Goh, 1976).

### Identifying the equilibrium conditions

The first step in this analysis is to find the equilibrium points of the model. To do so, we set Eqns. S1.5 and S1.6 equal to zero. We can start by evaluating the plant dynamics equation:

$$\frac{dF_1}{dt} = F_1(1 - F_1)[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)] = 0$$

This condition is satisfied when  $F_1 = 0$  or when  $F_1 = 1$ , which corresponds to cases in which the plant community is a monoculture of species 2 or 1 respectively. However, equilibrium can also arise when the third term (i.e. the term in square brackets) is equal to zero:

$$[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)] = 0 \quad (\text{S1.7})$$

Solving this for  $F_A$  shows that equilibrium is achieved when the following is true:

$$F_A^* = \frac{m_{2B} - m_{1B}}{m_{1A} - m_{2A} - m_{1B} + m_{2B}} = \frac{m_{2B} - m_{1B}}{I_S} \quad (\text{S1.8})$$

Given that  $dF_2/dt = -dF_1/dt$ , Equation S1.8 also implies that  $dF_2/dt = 0$ .

For the whole system to be at equilibrium, the microbial communities also need to be at equilibrium:

$$\frac{dF_A}{dt} = F_A(1 - F_A)[(F_1 - v(1 - F_1))] = 0 \quad (\text{S1.9})$$

As above, the microbial community can equilibrate when it comprises entirely of microbial community *A* or *B*, corresponding to  $F_A = 1$  or  $F_A = 0$ , respectively. The community is also at equilibrium when the third term (in square brackets) is equal to zero:

$$F_1 - v(1 - F_1) = 0$$

Solving this for  $F_1$  shows that equilibrium requires the following to be true:

$$F_1^* = \frac{v}{1 + v} \quad (\text{S1.10})$$

### **Identifying feasible equilibrium points**

Having identified the equilibrium conditions (Eqns. S1.8 and S1.10) can now evaluate the conditions under which this equilibrium is *feasible*, i.e. what is required for the equilibrium frequency of both plants and microbes to be between 0 and 1 ( $0 < F_A^* < 1$  and  $0 < F_1^* < 1$ ).

For simplicity, we begin with  $F_1^*$ . The value of Eqn. S1.10 will be between 0 and 1 for any  $v > 0$ . In other words, so long as both plant species condition the soil community, this condition is satisfied.

Next we move to Eqns. S1.8. Two sets of conditions can allow for  $0 < F_A^* < 1$ :

**Condition 1:** The numerator and denominator of Eqn. S1.8 are both positive ( $m_{2B} - m_{1B} > 0$  and  $m_{1A} - m_{2A} - m_{1B} + m_{2B} > 0$ ), and the magnitude of the numerator is smaller than that of the denominator ( $m_{2B} - m_{1B} < m_{1A} - m_{2A} - m_{1B} + m_{2B}$ ).<sup>3</sup>

**Condition 2:** The numerator and denominator of Eqn. S1.8 are both negative ( $m_{2B} - m_{1B} < 0$  and  $m_{1A} - m_{2A} - m_{1B} + m_{2B} < 0$ ), and the magnitude of the numerator is smaller than that of the denominator ( $\text{abs}(m_{2B} - m_{1B}) < \text{abs}(m_{1A} - m_{2A} - m_{1B} + m_{2B})$ )

If either condition is met (along with the condition that  $v > 0$ ), the system has a feasible equilibrium point at which all players (both plants and both microbes) are present in the system at a frequency between 0 and 1. If neither of these conditions is met (e.g. if  $I_S < 0$  but  $m_{2B} - m_{1B} > 0$ ), the system does not have an internal equilibrium; in other words, the system only has a boundary equilibrium corresponding to only one species being present in the system.

The next step for understanding the coexistence conditions in this model is to evaluate the dynamic stability of these equilibrium points.

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<sup>3</sup>Note that due to algebra, if  $m_{2B} > m_{1B}$  and  $I_S > 0$ ,  $m_{1A} > m_{2A}$  is implied; likewise, if  $m_{2B} < m_{1B}$  and  $I_S < 0$  (Condition 2) is satisfied,  $m_{1A} < m_{2A}$  is implied.

### Evaluating the dynamic stability of equilibrium points

While the above expressions (conditions 1 and 2, along with  $v > 0$ ) capture the conditions necessary for the existence of feasible equilibrium points, long-term coexistence also requires that these points are dynamically stable (i.e. that the system recovers equilibrium from slight perturbations away from the equilibrium state, Goh (1976)).

We can evaluate the local stability of the equilibria by creating the Jacobian Matrix of the system, which is denoted  $J$ . The Jacobian matrix helps us evaluate whether or not a system that is at equilibrium returns to the equilibrium when it is perturbed slightly. Each element in  $J$  is the partial derivative of one of the dynamics equations (Eqns S1.5 and S1.6) with respect to one of the components:

$$J = \begin{bmatrix} \frac{\partial \dot{F}_1}{\partial F_1} & \frac{\partial \dot{F}_1}{\partial F_A} \\ \frac{\partial \dot{F}_A}{\partial F_1} & \frac{\partial \dot{F}_A}{\partial F_A} \end{bmatrix}$$

Note that above,  $\dot{F}_1 = \frac{dF_1}{dt}$ , and  $\dot{F}_A = \frac{dF_A}{dt}$ . Taking the respective partial derivatives gives us the following expressions for the four elements of the matrix:

$$\begin{aligned} \frac{\partial \dot{F}_1}{\partial F_1} &= (1 - 2F_1)[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)] \\ \frac{\partial \dot{F}_1}{\partial F_A} &= F_1(1 - F_1)(m_{1A} - m_{2A} - m_{1B} + m_{2B}) = F_1(1 - F_1)I_S \\ \frac{\partial \dot{F}_A}{\partial F_1} &= F_A(1 - F_A)(1 + v) \\ \frac{\partial \dot{F}_A}{\partial F_A} &= [F_1 - v(1 - F_1)](1 - 2F_A) \end{aligned}$$

These four terms define the entries of the Jacobian matrix, which we can now evaluate at the system's equilibrium points to determine their local stability.

Recall from our analysis of Eqn. S1.7 that at equilibrium,  $[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)] = 0$ ; thus,  $\frac{\partial \dot{F}_1}{\partial F_1}$  also equals 0 at equilibrium.

Similarly, recall from the analysis of Eqn. S1.9 that  $[F_1 - v(1 - F_1)] = 0$  at equilibrium; thus,  $\frac{\partial \dot{F}_A}{\partial F_A}$  also equals zero at equilibrium.

The system's Jacobian evaluated at its equilibrium  $(F_1^*, F_A^*)$  thus simplifies as follows:

$$J|_{F_1^*, F_A^*} = \begin{bmatrix} 0 & F_1(1 - F_1)I_S \\ F_A(1 - F_A)(1 + v) & 0 \end{bmatrix}$$

We can evaluate the local stability of the equilibrium points on the basis of the trace and determinant of the matrix  $J$  (Panvilov et al., 2021). The trace (tr) for a square matrix is the

sum of its diagonal entries, so  $\text{tr}(J|_{F_1^*, F_A^*}) = 0$ .

Given that the trace of the matrix is zero, the equilibrium can have one of two properties:

1. The equilibrium is a “center equilibrium” if the determinant is positive (Panvilov et al., 2021). A center equilibrium implies that the system is neutrally stable, meaning that the system never returns to the equilibrium point itself after perturbation; it remains in a perpetual cycle. For our purposes, we interpret this as a coexistence equilibrium, because it implies that both species have cyclical dynamics of their frequency in the system.
2. The equilibrium is a saddle node if the determinant is negative (Panvilov et al., 2021). This means that once perturbed from equilibrium, the system continues moving away from the equilibrium (perturbations in favor of species 1 eventually lead to monodominance by species 1, and vice-versa for perturbations in favor of species 2).

Thus, whether or not any feasible equilibrium point corresponds to stable coexistence is determined by the sign of the determinant.

Recalling that the determinant of a generic two-by-two matrix  $\begin{pmatrix} a & b \\ c & d \end{pmatrix}$  is equal to  $(ad) - (bc)$ , the determinant of  $J$  is as follows:

$$\det(J|_{F_1^*, F_A^*}) = 0 - \left[ \overbrace{(F_A(1 - F_A)(1 + v))}^{\text{term 1}} * \overbrace{(F_1(1 - F_1)I_S)}^{\text{term 2}} \right]$$

Given that we are evaluating feasible equilibrium points where  $0 < F_A, F_B < 1$ , and  $v > 0$ , term 1 is always positive. Additionally, given that by definition at the feasible equilibrium  $0 < F_1, F_2 < 1$ , the sign of term 2 - and thus, the sign of the determinant as a whole - is determined by the sign of  $I_S$ . Specifically, negative values of  $I_S$  correspond to a positive determinant, while positive values of  $I_S$  correspond to a negative determinant.

Building on the two potential properties listed above, this means that the equilibrium is neutrally stable if  $I_S < 0$ , or is a saddle node if  $I_S > 0$ .

#### ***Combining the criteria for feasibility and stability***

From the above analysis, we see that only the equilibrium points in which both species can coexist with neutral stability satisfy [Condition 2 for feasible equilibria](#):

$$m_{2B} - m_{1B} < 0 \text{ and } m_{1A} - m_{2A} - m_{1B} + m_{2B} < 0.$$

Note that the above inequality implies that  $m_{1A} < m_{2A}$ . Thus, we can express the coexistence conditions simply as:

$$m_{2B} < m_{1B} \text{ and } m_{1A} < m_{2A} \tag{S1.11}$$

## Evaluating coexistence by analysing the requirements for mutual invasion

While the above approach derives the coexistence criteria by evaluating the conditions for local stability around feasible equilibria, one can also approach coexistence criteria by evaluating the conditions that allow mutual invasibility (Turelli, 1978; Chesson and Ellner, 1989; Grainger et al., 2019). As explained in the main text of the manuscript, this approach builds on the insight that coexistence requires that each species can gain a foothold (i.e. achieve a positive low-density growth rate, or LDGR) as it grows into an equilibrium monoculture of the other. Following Chesson (2000) and Chesson (2018), one can further decompose the LDGRs into two terms - one that captures the microbially mediated stabilization (which promotes both species' invasion growth rates, and thus favors coexistence), and a second term that captures the microbially mediated fitness difference (which benefits one plant's invasion growth rate but suppresses the other, and thus favors exclusion). The details of this analysis are provided in the appendix of Kandlikar et al. (2019), and summarized below.

We begin the analysis with Eqn. S1.5, which defines the dynamics of each plant's frequency in the system:

$$\frac{dF_1}{dt} = F_1(1 - F_1)[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)]$$

We first evaluate the case where the system is an equilibrium monoculture of plant 2 (and its corresponding soil community). Plant 1 and its soil community are absent, meaning that  $F_1 = F_A = 0$ . We can now quantify plant 1's per-frequency growth rate ( $\frac{1}{F_1} \frac{dF_1}{dt}$ ) as follows:

$$\text{LDGR}_{1 \rightarrow 2} = \frac{1}{F_1} \frac{dF_1}{dt} = (1 - F_1)[(m_{1A} - m_{2A})F_A + (m_{1B} - m_{2B})(1 - F_A)] \quad (\text{S1.12})$$

Given that  $F_1 = F_A = 0$ , Eqn. S1.12 simplifies as follows:

$$\text{LDGR}_{1 \rightarrow 2} = m_{1B} - m_{2B} \quad (\text{S1.13})$$

Through a similar analysis of plant 2's growth into a monoculture of plant 1, we get the invasion growth rate of plant 2:

$$\text{LDGR}_{2 \rightarrow 1} = m_{2A} - m_{1A} \quad (\text{S1.14})$$

If both of these conditions are satisfied, both species have positive low-density growth rates and can coexist provided that the following is true. Thus, this analysis yields the coexistence criteria:

$$m_{1B} > m_{2B} \quad \text{and} \quad m_{2A} > m_{1A} \quad (\text{S1.15})$$

The inequalities in Eqn. S1.15 are identical to those that we derived through the feasibility

analysis above Eqn. S1.11, showing the inherent complementarity of these two approaches. If our goal were to simply evaluate coexistence in the Bever et al. (1997) model, simply evaluating LDGRs is a perfectly valid ending: if both LDGRs are positive, the two species can coexist; if the LDGRs are of opposite signs, the species with a positive LDGR outcompetes the other; and if both LDGRs are negative, the species experience frequency-dependent priority effects such that either species can establish a monoculture, but both cannot coexist.

However, we can extend our analysis further to generate additional insights. Specifically, decomposing the LDGRs into microbially mediated stabilization and fitness differences allows us to integrate plant-microbe interactions into a wider body of work that seeks to understand how plant coexistence is structured by competition, pollinators, herbivores, etc. (see main text for citations to specific examples).

As explained in Chesson (2018) and in Appendix S1 of Kandlikar et al. (2019), the first step in this decomposition is to define the species-level average fitness. In the case of the Bever et al. (1997) model, we can define the average fitness of species 1 as its average growth rate at all possible soil states (from  $F_A = 0$  to  $F_A = 1$ ):

$$\text{fitness}_1 = \frac{\int_0^1 m_{1B} + (m_{1A} - m_{1B})F_A dF_A}{\int_0^1 dF_A} = m_{1B}F_A + \frac{m_{1A} - m_{1B}}{2}F_A^2 \Big|_0^1 = \frac{m_{1A} + m_{1B}}{2}$$

Similarly,  $\text{fitness}_2 = \frac{m_{2A} + m_{2B}}{2}$ . With these definitions of species 1 and 2's average fitness, we can express each species' invasion growth rate as the sum of the fitness difference and stabilization:

$$\text{LDGR}_1 = \text{fitness difference}_{1,2} + \text{stabilization} \quad (\text{S1.16})$$

$$\text{LDGR}_2 = \text{fitness difference}_{2,1} + \text{stabilization} \quad (\text{S1.17})$$

Note that  $\text{fitness difference}_{1,2}$  is simply the difference between species 1 and 2's average fitness as defined above:

$$\text{fitness difference}_{1,2} = \left( \overbrace{\frac{m_{1A} + m_{1B}}{2}}^{\text{plant 1 fitness}} \right) - \left( \overbrace{\frac{m_{2A} + m_{2B}}{2}}^{\text{plant 2 fitness}} \right)$$

The order of the two terms is flipped for calculating  $\text{fitness difference}_{2,1}$ . Thus, in the absence of stabilization, only one species can have a positive invasion growth rate, and coexistence is not possible.

Above, we saw that  $\text{LDGR}_{1 \rightarrow 2} = m_{1B} - m_{2B}$  (Eqn S1.13). Substituting this into Eqn. S1.16, we get:

$$m_{1B} - m_{2B} = \left(\frac{m_{1A} + m_{1B}}{2}\right) - \left(\frac{m_{2A} + m_{2B}}{2}\right) + \text{stabilization}$$

Algebra (detailed in [Box S2](#)) yields the expression for stabilization:

$$\text{stabilization} = -\frac{1}{2}(m_{1A} - m_{1B} - m_{2A} + m_{2B}) = -\frac{1}{2}I_S$$

For both species to have a positive LDGR, the strength of stabilization should exceed the absolute value of the fitness difference<sup>4</sup>:

$$\text{stabilization} > \text{abs}(\text{fitness difference})$$

When fitness differences exceed stabilization, only the species with the higher fitness can invade into a monoculture of the other; this corresponds to species exclusion.

Negative stabilization (destabilization) suppresses each species' LDGR. If it does so to the point that neither species has a positive LDGR, the system experiences priority effects: whichever species is present at a higher frequency will dominate, and the species with initially low frequencies eventually gets excluded.

### Box S2: Deriving the stabilization term

Above, we saw that the  $\text{LDGR}_{1 \rightarrow 2}$  can be expressed as follows:

$$\text{LDGR}_{1 \rightarrow 2} = m_{1B} - m_{2B} = \left(\frac{m_{1A} + m_{1B}}{2}\right) - \left(\frac{m_{2A} + m_{2B}}{2}\right) + \text{stabilization}$$

We can rewrite this as follows:

$$m_{1B} - m_{2B} = \frac{1}{2}m_{1A} + \frac{1}{2}m_{1B} - \frac{1}{2}m_{2A} - \frac{1}{2}m_{2B} + \text{stabilization}$$

Moving the terms to the left of the equal sign to the right, and moving stabilization to the left gives

$$-\text{stabilization} = \frac{1}{2}m_{1A} - \frac{1}{2}m_{1B} - \frac{1}{2}m_{2A} + \frac{1}{2}m_{2B}$$

This equation simplifies to the expression for stabilization:

$$\text{stabilization} = -\frac{1}{2}(m_{1A} - m_{1B} - m_{2A} + m_{2B})$$

<sup>4</sup>the absolute value of  $(\text{fitness difference})_{1,2}$  equals that of  $(\text{fitness difference})_{2,1}$ , so subscripts are not required



**The decomposition also applies to LDGR<sub>2→1</sub>**

While we derived stabilization from plant 1's LDGR, we can show that this applies equally well to plant 2's low density growth:

$$\text{LDGR}_{2 \rightarrow 1} = \text{fitness difference}_{2,1} + \text{stabilization}$$

Substituting the expressions for fitness difference<sub>2,1</sub> and stabilization gives us:

$$\text{LDGR}_{2 \rightarrow 1} = \frac{m_{2A} + m_{2B}}{2} - \frac{m_{1A} + m_{1B}}{2} - \frac{1}{2}(m_{1A} - m_{1B} - m_{2A} + m_{2B})$$

Through algebra, we recover Eqn. S1.17 as above:

$$\text{LDGR}_{2 \rightarrow 1} = m_{2A} - m_{1A}$$

**The coexistence criteria in terms of stabilization/fitness difference is equivalent to that from the LDGR analysis**

Finally, we can show that the coexistence criteria expressed as “stabilization > abs(fitness difference)” is equivalent to the criteria in Eqn. S1.15.

Recall the coexistence criteria in terms of stabilization and fitness difference:

$$\overbrace{-\frac{1}{2}((m_{1A} + m_{2B}) - (m_{2A} + m_{1B}))}^{\text{stabilization}} > \text{abs} \left( \overbrace{\left( \frac{1}{2}(m_{1A} + m_{1B}) - \frac{1}{2}(m_{2A} + m_{2B}) \right)}^{\text{fitness difference}_{1,2}} \right) \quad (\text{S1.18})$$

By dividing through by  $-\frac{1}{2}$ , this can be reexpressed as follows:

$$m_{1A} + m_{2B} - m_{2A} - m_{1B} < \text{abs}(m_{1A} + m_{1B} - m_{2A} - m_{2B})$$

To accounting for the absolute value function on the right, this inequality can be written as two separate inequalities:

$$m_{1A} + m_{2B} - m_{2A} - m_{1B} < m_{1A} + m_{1B} - m_{2A} - m_{2B} \quad (\text{S1.19})$$

$$m_{1A} + m_{2B} - m_{2A} - m_{1B} > -m_{1A} - m_{1B} + m_{2A} + m_{2B} \quad (\text{S1.20})$$

Cancelling like terms in Eqn. S1.19 gives  $m_{1B} > m_{2B}$ , and doing the same in Eqn. S1.20  $m_{1A} > m_{2A}$ . Together, these are identical to Eqn. S1.15.

## Appendix S2: Quantifying the community-wide stabilization metric $I_C$ with empirical data

**Contact:** Gaurav S. Kandlikar, [gkandlikar@lsu.edu](mailto:gkandlikar@lsu.edu)

With some modifications to the nature of soil microbial dynamics, Mack et al. (2019) and Eppinga et al. (2018) extended the two-species feedback model of Bever et al. (1997) to evaluate microbial effects on coexistence in multi-species plant communities. Details of this multispecies model derivation and analysis are provided in the original publications. In this appendix, I demonstrate a practical application of quantifying the feedback metric from Eppinga et al. (2018) using empirical data from Dudenhöffer et al. (2022).

By analyzing an  $n$ -species plant-soil feedback model, Eppinga et al. (2018) showed that whether microbes generate positive or negative feedback is determined by the sign of the metric  $I_C$ , which serves as a community-wide analog of the two-species term  $I_S$ . Extending from the notation of the two-species model used in the main text, plant species are denoted  $1, 2, \dots, n$ , and the corresponding microbial communities are denoted  $A, B, \dots, X$ . The effect of a given microbial community  $x$  on plant  $i$  is denoted  $m_{ix}$ . One can arrange the  $m$  terms into an interaction matrix  $\mathbf{A}$ :

$$\mathbf{A} = \begin{bmatrix} m_{1A} & m_{1B} & \dots & m_{1X} \\ m_{2A} & \dots & \dots & m_{2X} \\ \dots & \dots & \dots & \dots \\ m_{nA} & \dots & \dots & m_{nX} \end{bmatrix}$$

One can use this interaction  $\mathbf{A}$  to calculate the community-wide stabilization  $I_C$  as follows:

$$I_C = (-1)^n \sum_{j=1}^n \det \mathbf{A}_j \quad (\text{S2.21})$$

Here,  $\det$  represents the matrix determinant, and  $\mathbf{A}_j$  denotes the interaction matrix  $\mathbf{A}$  with the  $j$ 'th column replaced with a vector of 1s. Note that in two-species systems ( $n = 2$ ),  $I_C = I_S$ , as detailed in Box S2.1.

The main text presents the caveats of using  $I_S$  to predicting pairwise species coexistence. Such caveats also exist for  $I_C < 0$ . While negative values of  $I_C$  indicate negative feedback (stabilizing effects on community dynamics), they do not guarantee that all species can coexist. In addition to  $I_C < 0$ , community-wide coexistence also requires that each species can persist at non-zero frequency at equilibrium:

$$0 < \hat{P}_i = \frac{\det \mathbf{A}_i}{(-1)^n \sum_{j=1}^n \det \mathbf{A}_j} < 1 \quad (\text{S2.22})$$

For the same reason that growth in reference soil is essential for calculating pairwise fit-

ness differences (see Main Text), such data is also required for quantifying the equilibrium frequency.

**Box 1: Correspondence between  $I_C$  and  $I_S$  when  $n = 2$**

The interaction matrix for two species is as follows:

$$\mathbf{A} = \begin{bmatrix} m_{1A} & m_{1B} \\ m_{2A} & m_{2B} \end{bmatrix}$$

Following Eqn. S2.21 above,  $I_C$  for this 2-species system ( $I_C$ ) is calculated as follows:

$$I_C = (-1)^2 \sum_{j=1}^2 \det A_j = (-1)^2 \left( \det \begin{pmatrix} 1 & m_{1B} \\ 1 & m_{2B} \end{pmatrix} + \det \begin{pmatrix} m_{1A} & 1 \\ m_{2A} & 1 \end{pmatrix} \right) \quad (\text{S2.23})$$

Given that  $\det \begin{pmatrix} a & b \\ c & d \end{pmatrix} = ad - bc$ , Eqn. S2.23 simplifies as:

$$I_C = (-1)^2 ((1 * m_{2B} - m_{1B} * 1) + (m_{1A} * 1 - 1 * m_{2A}))$$

Through algebra, this simplifies to  $I_C = m_{2B} - m_{1B} + m_{1A} - m_{2A}$ , which is equivalent to the pairwise  $I_S$ .

**Quantifying  $I_C$  with empirical data**

This subsection provides R code for calculating  $I_C$  from the data collected for Dudenhöffer et al. (2022)'s study, which evaluated how drought affects plant-soil feedback outcomes.

**Note:** The goal of this code is not to be universally applicable in its current form to all datasets; rather, this code can merely serve as a starting point for future studies aiming to evaluate community-wide stability with  $I_C$ . The code below makes a number of simplifying assumptions (Box S2) which may not be appropriate in other contexts.

**Assumptions embedded in the code**

- Microbial effects on plant performance arise primarily through modification of plant biomass (survival not impacted; note that this diverges from Dudenhöffer et al. (2022)'s original analysis)
- In cases where an estimate of  $B_{ix}$  was unable, I use the average value of  $B_{ix}$  from all other pots in the same environmental (watering) treatment.

```

library(tidyverse)
library(readxl)
library(osfr) # for downloading dataset

# Download dataset if it is not available
if(!("data_PSF_response_phase.xlsx" %in% list.files())) {
  osf_retrieve_file("https://osf.io/nx2e6") %>%
  osf_download()
}

psf_data <- read_xlsx("data_PSF_response_phase.xlsx")

# Structure of the dataset
colnames(psf_data)

[1] "block"      "soil"      "treatment" "species"   "part"      "bm"
[7] "dead"

unique(psf_data$block) # There are sterile soils in this; we can filter them

[1] "A" "B" "C" "D" "E" "F" "G" "H" "I" "S1" "S2" "S3"

table(psf_data$dead)/2

  0    1
2081 223

# There are >200 dead plants; for now we can set aside
# microbial effects on mortality and focus instead on growth
# This differs from the authors of this study, but is consistent
# with lots of other work on PSF.
# For now, we can just assign dead plants to have the mean biomass
# across other replicates of the same species/soil/treatment combo.

# Data reformatting
interaction_matrices <-
  psf_data %>%
  # Change the species names to be in sentence case i.e. "AT" becomes "At"
  mutate(soil = str_to_sentence(soil),
         species = str_to_sentence(species)) %>%
  # filter out sterile soils treatment - not relevant for I_C
  filter(!str_detect(block, "S")) %>%
  # group by the relevant categories
  group_by(block, soil, treatment, species) %>%
  # Right now, biomass is separated agb/bgb;

```

```

# This chunk combines the two into whole-plant biomass ("combined_bm")
summarize(combined_bm = sum(bm),
          dead = max(dead)) %>%
ungroup() %>%
# Next, we replace NAs (biomass of dead plant) with mean values of that group
# NOTE that this is not a unvierversally good choice; but needed to do something
# like this here because if any one entry of the matrix A is missing,
# that matrix cannot be used for calculating I_C.
group_by(soil, treatment, species) %>%
mutate(combined_bm = ifelse(dead ==1, NA, combined_bm),
       combined_bm = ifelse(dead == 1,
                             mean(combined_bm, na.rm = T),
                             combined_bm)) %>%

# filter out dead individuals
select(-dead) %>%
# calculate log biomass
mutate(combined_bm = log(combined_bm)) %>%
# make treatment into a factor vector, with levels L/M/H
mutate(treatment = as_factor(treatment),
       treatment = fct_relevel(treatment, c("L", "M", "H"))) %>%
ungroup() %>%
# Now, we can work within each treatment & block to make interaction matrices.
group_by(treatment, block) %>%
arrange(treatment) %>%
nest() %>%
# The next chunk uses the biomass values B_{ij} and makes an
# interaction matrix (A) for each replicate block/treatment combo
mutate(interaction_matrix =
       map(data,
           ~pivot_wider(.x,
                        # row is a species and each column is a soil type:
                        names_from = soil,
                        values_from = combined_bm) %>%
           column_to_rownames('species') %>%
           as.matrix()))

# We can look at this new object:
interaction_matrices

# A tibble: 27 x 4
# Groups:   treatment, block [27]
  block treatment data          interaction_matrix
  <chr> <fct>      <list>          <list>
1 A     L           <tibble [64 x 3]> <dbl [8 x 8]>
2 B     L           <tibble [64 x 3]> <dbl [8 x 8]>
3 C     L           <tibble [64 x 3]> <dbl [8 x 8]>
4 D     L           <tibble [64 x 3]> <dbl [8 x 8]>

```

```

5 E    L    <tibble [64 x 3]> <dbl [8 x 8]>
6 F    L    <tibble [64 x 3]> <dbl [8 x 8]>
7 G    L    <tibble [64 x 3]> <dbl [8 x 8]>
8 H    L    <tibble [64 x 3]> <dbl [8 x 8]>
9 I    L    <tibble [64 x 3]> <dbl [8 x 8]>
10 A   M    <tibble [64 x 3]> <dbl [8 x 8]>
# i 17 more rows

```

```
# We can look at what an interaction matrix looks like:
```

```
# This is the interaction matrix for Block A, low water treatment:
```

```

interaction_matrices %>%
  filter(block == "A", treatment == "L") %>%
  pull(interaction_matrix) %>% pluck(1) %>% round(., 2)

```

	At	Bi	Rc	Rh	Sh	Sn	Ss	Vb
At	-1.33	-1.51	-0.89	-0.61	-1.69	-0.90	-0.67	-1.97
Bi	-1.45	-3.86	-2.70	-0.31	-4.34	-4.42	-3.00	-2.60
Rc	-1.31	-1.26	-0.87	-1.90	-4.02	-1.66	-1.14	-1.01
Rh	-0.42	-0.85	-1.46	-0.94	-4.96	-1.12	-0.59	-1.43
Sh	-2.90	-2.54	-2.02	-2.47	-3.73	-2.26	-0.49	-3.69
Sn	-1.04	-3.44	-1.43	-1.96	-1.26	-1.39	-1.93	-0.80
Ss	-1.41	-1.60	-1.85	-1.38	-3.32	-2.60	-1.16	-1.82
Vb	-0.47	-0.77	0.06	-0.97	-3.61	-0.84	-1.01	-1.42

```

# Conceptually it is similar to Fig. 2A from the paper, but note that this matrix
# is for Block A only; that one averages from across blocks (and also accounts
# for microbial effects on mortality).

```

Now that we have made the interaction matrices (1 matrix per block, per treatment), we can use this matrix to calculate the  $I_C$  for every possible 2, 3, 4, 5, 6, 7, and 8 species combination. That will represent the  $I_C$  for a given community, in a given treatment, in a given block. We can then summarize over different blocks to get a mean  $I_C$  for each community in each treatment.

```

# To make calculations easier, we can write a function that calculates
# I_C for all possible subcommunities, given a complete interaction matrix:

```

```

Ic_for_all_subs <- function(intmat) {

  # 1. Helper function: Given a species combination, make a submatrix
  make_submatrix <- function(intmat, indices) {
    to_return <- apply(indices, 2, function(x) intmat[x,x], simplify = F)
    names(to_return) <- apply(indices, 2, function(x)
      paste(rownames(intmat)[x], collapse=""))
    to_return
  }
}

```

```

# 2. Helper function: Given an interaction matrix, calculate I_C
Ic <- function(intmat) {
  # 2.1. make a variable that holds species number
  nsp = ncol(intmat)

  # 2.2. Define a sub-helper function for making matrices A_j
  # This function replaces the j'th column
  # in the interaction matrix (intmat) with a column of 1s
  make_Aj_mat <- function(intmat, j) {
    temp_mat <- intmat # Define a temporary holder matrix
    temp_mat[,j] <- 1 # Return the j'th column with 1
    return(temp_mat) # return the holder matrix
  }

  # 2.3. Make Aj vectors using the sub-helper function above
  Ajs <- map(1:nsp, ~make_Aj_mat(intmat, .x))
  # 2.4. Calculate determinants of all Aj matrices
  dets <- map_dbl(Ajs, det)
  # 2.5. Calculate IC
  ((-1)^nsp)*(sum(dets))
}

# 3. Helper function: Given a list of submatrices, calculate their IC
# This function returns a vector; each element in the vector is I_C
make_Ic_vec <- function(submats) {
  map_dbl(submats, Ic)
}

# 4. Define the number of species (total), all possible 2:n species combinations,
# and make all possible sub-matrices of 2:n species using Helper Fn 1 above.
nsp <- nrow(intmat)
possible_combns <- map(2:nsp, ~combn(nsp, .x))
submats <- map(possible_combns, ~make_submatrix(intmat, .x))

# 5. calculate Ic for all submatrices
all_Ics <- map(submats, make_Ic_vec)
all_Ics
}

# We can now use this function to calculate all ICs:
# Here, we use the interaction_matrices object and make a new column
# called all_Ics; each element in this column will have all
# possible I_C values for a given matrix.

```

```

interaction_matrices_with_ICs <-
  interaction_matrices %>%
  mutate(all_Ics = map(interaction_matrix, Ic_for_all_subs))

head(interaction_matrices_with_ICs)

# A tibble: 6 x 5
# Groups:   treatment, block [6]
  block treatment data          interaction_matrix all_Ics
  <chr> <fct>    <list>          <list>          <list>
1 A     L        <tibble [64 x 3]> <dbl [8 x 8]>    <list [7]>
2 B     L        <tibble [64 x 3]> <dbl [8 x 8]>    <list [7]>
3 C     L        <tibble [64 x 3]> <dbl [8 x 8]>    <list [7]>
4 D     L        <tibble [64 x 3]> <dbl [8 x 8]>    <list [7]>
5 E     L        <tibble [64 x 3]> <dbl [8 x 8]>    <list [7]>
6 F     L        <tibble [64 x 3]> <dbl [8 x 8]>    <list [7]>

# The first entry has all I_Cs for Block A/treatment L:
# (This will be a list; the first element in the list is a vector
# of the two-species I_Cs; the second element is a vector of the 3-species
# I_Cs, and so on)

# Two species I_C (AKA I_S), only printing first 10
interaction_matrices_with_ICs$all_Ics[[1]][[1]][1:10]

      AtBi      AtRc      AtRh      AtSh      AtSn      AtSs
-2.23727667  0.00289922 -1.24613067 -0.46826601 -0.78187924 -0.41424075
      AtVb      BiRc      BiRh      BiSh
-0.30827442 -0.76536237 -3.64484659 -0.71182094

# I_S for triplets, only printing first 10
interaction_matrices_with_ICs$all_Ics[[1]][[2]][1:10]

      AtBiRc      AtBiRh      AtBiSh      AtBiSn      AtBiSs      AtBiVb      AtRcRh
-0.3814381 -2.8975229 -2.4155463  5.8255029 -0.8931710 -0.6285068  1.9405044
      AtRcSh      AtRcSn      AtRcSs
-1.0305773  0.6521919  0.4319539

We now have a value of  $I_C$  for every  $n = 2, 3, \dots, 8$  species combination in each
block/treatment combination. There are various ways one can summarize this informa-
tion; for simplicity, I will just summarize the mean value of pairwise  $I_C$ , triplet  $I_C$ , ...
8-species  $I_C$  in each treatment.

interaction_matrices_with_ICs %>%
  # This next mutate call takes the big list of Ics and splits the
  # information into columns for 2, 3, ... 8 species communities

```



```

mutate(coms2 = map(all_Ics, ~pluck(.x,1)),
       coms3 = map(all_Ics, ~pluck(.x,2)),
       coms4 = map(all_Ics, ~pluck(.x,3)),
       coms5 = map(all_Ics, ~pluck(.x,4)),
       coms6 = map(all_Ics, ~pluck(.x,5)),
       coms7 = map(all_Ics, ~pluck(.x,6)),
       coms8 = map(all_Ics, ~pluck(.x,7))) %>%
# data managing: we can get rid of a few things and only focus on the IC columns
select(-data, -interaction_matrix, -all_Ics) %>%
unnest(c(treatment, block)) %>%
# At this step, we get the mean value of $I_C$ for each n-species community
mutate(across(coms2:coms8, ~map_dbl(.x, mean))) %>%
# Clean out some columns that we don't need
select(block, treatment, coms2:coms8) %>%
# Calculate mean and SD of IC for each community size in each treatment
group_by(treatment) %>%
# Get the mean, standard deviation, and replicates per calculation
summarise(across(coms2:coms8, mean, .names = "{.col}_mean"),
          across(coms2:coms8, sd, .names = "{.col}_sd"),
          across(coms2:coms8, length, .names = "{.col}_nreps")) %>%
# Clean out some unused columns
select(-(coms2_nreps:coms7_nreps), nreps = coms8_nreps) %>%
# Reshape the data in two steps:
# First, pivot it longer so that each sd/mean ends up on its own row
pivot_longer(coms2_mean:coms8_sd) %>%
# Then, clean up the names and pivot it wider so that mean/sd are in different columns
separate(name, into = c("which_comm", "which_value"), sep = "_") %>%
pivot_wider(names_from = which_value, values_from = value) %>%
# calculate SEM as sd/sqrt(n)
mutate(sem = sd/sqrt(nreps)) %>%

# NOTE: uncomment the following lines to make the Main Text Fig. S3;
# which focuses only on two- and three-species communities
# filter(which_comm %in% c("coms2", "coms3")) %>%
# mutate(which_comm = ifelse(which_comm == "coms2", "Species pairs", "Triplets")) %>%

mutate(which_comm = case_when(which_comm == "coms2" ~ "2 species communities",
                             which_comm == "coms3" ~ "3 species communities",
                             which_comm == "coms4" ~ "4 species communities",
                             which_comm == "coms5" ~ "5 species communities",
                             which_comm == "coms6" ~ "6 species communities",
                             which_comm == "coms7" ~ "7 species communities",
                             which_comm == "coms8" ~ "8 species communities")) %>%

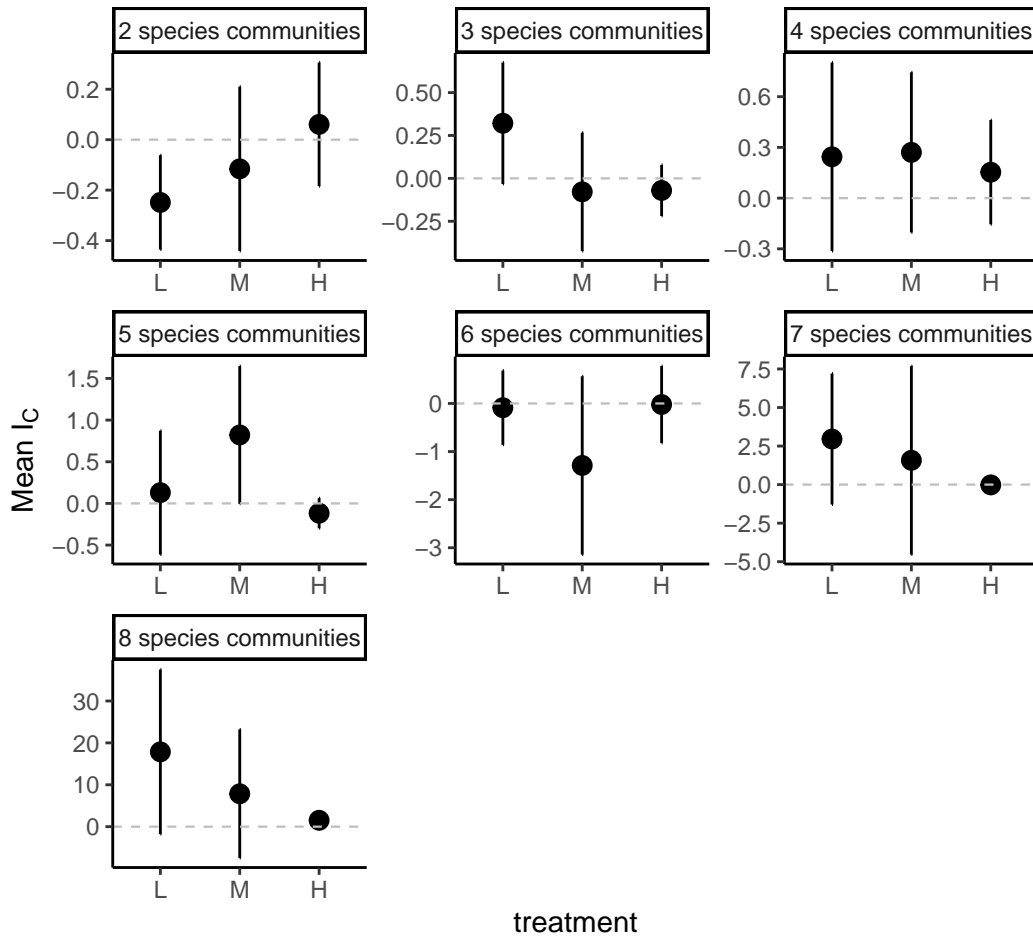
ggplot(aes(x = treatment, y = mean, ymin = mean-sem*2, ymax = mean+sem*2)) +
geom_point(size = 3) +

```

```

geom_errorbar(width = 0) +
facet_wrap(~which_comm, scales = "free") +
ylab(latex2exp::TeX("Mean  $I_C$ ")) +
geom_hline(yintercept = 0, linewidth = 0.4, linetype = "dashed", color = "grey") +
theme_classic()

```



```

# ggsave(filename = "figures/figS4.pdf", width = 5, height = 2)

```

## Appendix S3: Derivation and implications of Bever (2003)'s model for plant–soil feedback among competing plants

**Contact:** Gaurav S. Kandlikar, [gkandlikar@lsu.edu](mailto:gkandlikar@lsu.edu)

Bever (2003) presented the first modeling framework for integrating plant–soil feedback with the effects of direct competition among plants. One of the key results from this analysis was that sufficiently strong negative plant–soil feedback can enable coexistence of species pairs where the competitive dynamics by themselves result in exclusion.

In this Appendix, I briefly explain the derivation of this model, and focus on some implications of its specification for empirical applications. A more complete analysis of this model is presented in the original publication and in Revilla et al. (2013).

In the Bever (2003) model, the effects of plant–soil feedback and plant competition are integrated into a population dynamics model as follows:

$$\frac{dN_1}{dt} = r_1 N_1 \left( 1 + m_{1A} S_A + m_{1B} S_B - \frac{N_1 + c_{12} N_2}{K_1} \right) \quad (\text{S3.1})$$

Following the original parameter definitions,  $N_i$  is the abundance of plant  $i$ ,  $r_i$  is the intrinsic population growth rate of plant  $i$  in the absence of soil conditioning (units:  $\frac{1}{\text{time}}$ ),  $m_{iX}$  is the effect of microbial community  $X$  on plant species  $i$  (units:  $\frac{1}{\text{microbe frequency}}$ ),  $K_i$  is plant  $i$ 's carrying capacity in the absence of conditioning (units: abundance), and  $c_{ji}$  is the per-capita competitive effect of plant  $i$  on plant  $j$ , relative to the strength of plant  $j$ 's intraspecific competitive effect (units: unitless).

While this model is in principle a simple extension of Bever et al. (1997)'s exponential growth model to incorporate intra- and interspecific competition, the specific formulation of the model implies subtle but important changes in our assumptions of how microbes alter plant population dynamics. A practical example helps illustrate the distinction.

Recall from Appendix 1 that in the Bever et al. (1997) model, the composition of the soil community directly (additively) increases or decreases the intrinsic growth rate of the plants. Thus, for example, the realized per-capita growth rate of plant 1 when growing alone (i.e. in soil entirely conditioned by plant 1, such that  $S_A = 1$ ) is:

$$\text{Realized growth rate of plant 1 in soil A (1997 model)} = r_{1, \text{unconditioned soil}} + m_{1A} \quad (\text{S3.2})$$

where  $r_1$  is plant 1's growth rate in unconditioned soil.

In contrast, we can evaluate how soil microbes affect plant growth in the Bever (2003) model. To recreate a similar scenario of plant 1 growing alone in conspecific-conditioned soil, we set  $S_A = 1$ , and both plant densities to zero (i.e.  $N_1 = N_2 = 0$ ). With these values, the per-capita growth rate of plant 1 is:

$$\begin{aligned} \text{Realized growth rate of plant 1 in soil A (2003 model)} &= r_{1,\text{unconditioned soil}}(1 + m_{1A}) \\ &= r_{1,\text{unconditioned soil}} + m_{1A} \times r_{1,\text{unconditioned soil}} \quad (\text{S3.3}) \end{aligned}$$

In other words, the specification of microbial effects in Eqn. S3.1 implies that the composition of the soil community causes a *proportional* increase or decrease in the intrinsic growth rate of the plant species, rather than an additive increase as in the exponential growth model. Thus, for a plant that grows in unconditioned soil at a rate of  $r_{1,\text{unconditioned soil}} = 0.5$ , a microbial effect of  $m_{1A} = 0.2$  implies that it grows at a rate of  $r_{1,\text{soil A}} = 0.7$  in the 1997 model, but  $r_{1,\text{soil A}} = 0.6$  in the 2003 model.

The difference in model specification also has implications for how one analyzes experimental data, based on whether the goal is to parameterize the exponential growth model of Bever et al. (1997) or the competition model of Bever (2003). As explained in the main text and in Appendix 1,  $m_{iX}$  terms can be simply calculated as the difference in log-biomass of plants in conditioned vs. unconditioned soils, without any additional scaling (i.e.  $m_{iX} = \log(B_{iX}) - \log(B_{i0})$ ). However, Eqn. S3.3 implies that when using these same data to parameterize the 2003 model, scaling by the growth rate in unconditioned soil becomes essential:

$$m_{iX} = \frac{\log(B_{iX}) - \log(B_{i0})}{\log(B_{i0})}$$