

1 **Measuring biological generality in meta-analysis: a pluralistic approach to**
2 **heterogeneity quantification and stratification**

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4 Yefeng Yang¹, Daniel W. A. Noble², Rebecca Spake³, Alistair M. Senior⁴, Malgorzata
5 Lagisz^{1,#}, Shinichi Nakagawa^{1,5,#}

6

7 ¹Evolution & Ecology Research Centre and School of Biological, Earth and Environmental
8 Sciences, University of New South Wales, Sydney, NSW 2052, Australia

9 ²Division of Ecology and Evolution, Research School of Biology, The Australian National
10 University, Canberra, ACT 2600, Australia

11 ³Ecology and Evolutionary Biology Research Division, School of Biological Sciences,
12 University of Reading, RG6 6EX, Reading, UK

13 ⁴Charles Perkins Centre, Sydney Precision Data Science Centre, School of Life and
14 Environmental Sciences and School of Mathematics and Statistics, The University of Sydney,
15 Sydney, NSW 2006, Australia

16 ⁵Theoretical Sciences Visiting Program, Okinawa Institute of Science and Technology
17 Graduate University, Onna, 904-0495, Japan

18

19 Correspondence

20 E-mail: s.nakagawa@unsw.edu.au (SN)

21 # Equal senior author

22

23 **ORCID**

24 Yefeng Yang 0000-0002-8610-4016

25 Daniel W. A. Noble 0000-0001-9460-8743

- 26 Rebecca Spake 0000-0003-4671-2225
- 27 Alistair M. Senior 0000-0001-9805-7280
- 28 Malgorzata Lagisz 0000-0002-3993-6127
- 29 Shinichi Nakagawa 0000-0002-7765-5182
- 30

31 **Abstract**

32 Uncovering general rules enhances the predictive capabilities in ecology and evolution. Meta-
33 analytic approaches play a critical role in this endeavour, examining the extent to which
34 phenomena can be replicated, generalized, and transferred. However, ecologists and
35 evolutionary biologists have largely overlooked the role of meta-analytic heterogeneity in
36 informing generality. To reform this situation, we introduce a pluralistic approach aimed at
37 quantifying and stratifying various heterogeneity metrics, such as I^2 , CV , M , and predictive
38 distribution. These metrics offer complementary information, revealing the source,
39 magnitude, and visual representation of heterogeneity. Our analysis of 512 meta-analyses
40 demonstrates that heterogeneity is, on average, ten times larger than statistical noise,
41 contributing to 91% of the observed variance (median $I^2 = 91\%$). This amount of
42 heterogeneity is nearly twice the size of the meta-analytic mean effect (median $CV = 1.8$, $M =$
43 0.6), indicating substantial total heterogeneity in ecology and evolution. Surprisingly, in half
44 of the cases, focal effects could generalize across studies even with high total heterogeneity
45 by controlling for within-study variation. Our synthesis also visualises empirical distributions
46 of various heterogeneity metrics, potentially serving as new benchmarks for informed
47 interpretation. Our proposed pluralistic approach will accelerate the future quest for general
48 rules via meta-analyses.

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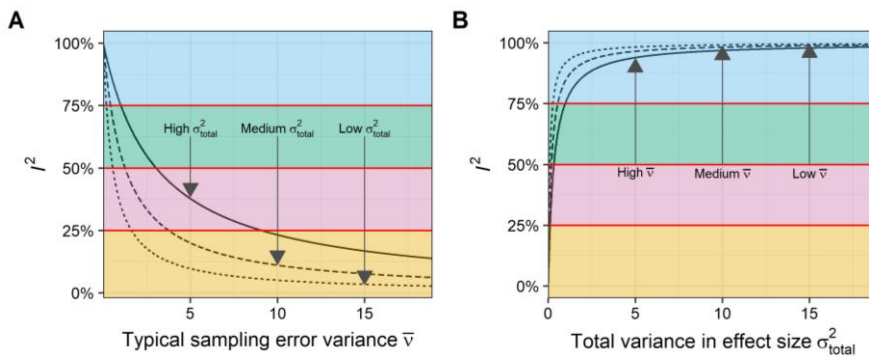
50 **Main**

51 Uncovering general patterns holds immense significance in ecology and evolution ¹. This
52 enables scientists, practitioners, and policymakers to transfer findings across diverse systems,
53 taxonomic groups, and spatiotemporal contexts. This pursuit enhances predictive capabilities
54 and facilitates more precise management, intervention, and conservation practices. Ecologists
55 and evolutionary biologists strive to unveil general processes and patterns using a range of
56 approaches ². Notably, meta-analytic modelling has emerged as a natural route to assess the
57 generality or context dependence of an effect of interest. By synthesizing a collection of
58 conceptual replications ³, meta-analyses can scrutinize the extent to which inferences drawn
59 from a specific context can be replicated (replication), extended beyond the reference context
60 to a new context of interest (transferred), and extrapolated to the broader target population
61 (generalized) as requested by stakeholders ^{2,4}.

62
63 Meta-analyses play a crucial role in evaluating the generality of patterns ³. Firstly, they
64 quantitatively estimate the population mean effect across studies ⁵⁻⁷, characterising the central
65 tendency of a focal effect. Secondly, they can identify effect modifiers or moderators
66 contributing to context dependence ⁵ and provide tailored estimates for target contexts ⁴.
67 Third, meta-analyses can quantify variability in study outcome, the “heterogeneity” among
68 effect sizes. Without quantifying heterogeneity, it is difficult to interpret both the overall
69 trends and context-specific effects ⁸. Heterogeneity can help to indicate the degree of
70 inconsistency, or context dependence, of study findings, with high heterogeneity signalling a
71 need to investigate the drivers of the variation. Lower heterogeneity can indicate high
72 generality. Specifically, the mean effect size is highly transferable across the contexts
73 characterised by the study pool without the need to consider effect modifiers ². Until now, the

74 significance of heterogeneity in informing generality has been largely overlooked. Indeed,
75 surveys have revealed that heterogeneity statistics are not routinely reported⁷⁻⁹.

76



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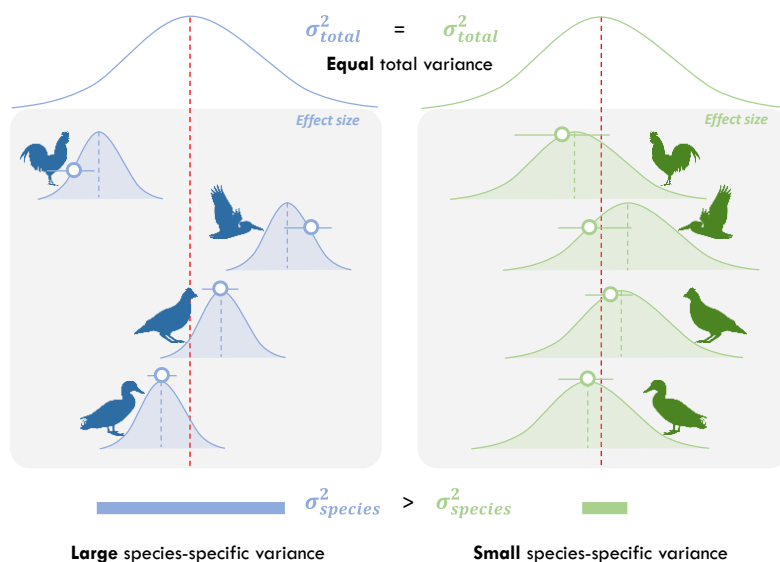
78 **Fig. 1:**

79 The interpretation of total I^2 can be ambiguous and can lead to incorrect conclusions about the
80 magnitude of heterogeneity. (A) A large estimated total I^2 value could be due to small sampling error
81 variances \bar{v} (i.e., low statistical noise). (B) On the other hand, a large total I^2 value could also result
82 from a large true heterogeneity. Values of σ_{total}^2 and \bar{v} were derived from their empirical distributions
83 based on 512 meta-analyses (see Figs. S1 and S2). Total I^2 values were calculated using Equations 2
84 and 3. High, medium, and low σ_{total}^2 (and \bar{v}) denote the 25%, 50%, and 75% percentiles of their
85 empirical distributions (Table 1). Three horizontal lines denote the conventional thresholds for the use
86 of I^2 to interpret the magnitude of heterogeneity¹⁰.

87

88 Currently, measuring and interpreting meta-analytic heterogeneity faces two major
89 limitations. First, no single heterogeneity metric provides a holistic interpretation of
90 generality¹¹. Currently, the I^2 statistic is a popular metric that quantifies the proportion of
91 variance due to differences between effect sizes rather than by statistical noise (i.e., sampling
92 variance)^{12,13}. The biological interpretation of I^2 , however, is ambiguous¹⁴ because a small
93 absolute heterogeneity can lead to a high I^2 due to small statistical noise (see Fig. 1)^{12,14,15}. In

94 addition, I^2 is a point estimate and cannot reflect the whole distribution of context-specific
 95 effects ¹⁶. Second, meta-analyses typically focus on estimating total heterogeneity only ⁵,
 96 despite the hierarchical nature of real biological data structures ^{6,9}. Explicitly decomposing
 97 effect size heterogeneity across hierarchical levels (i.e., stratification) enables a more nuanced
 98 assessment of generality, and helps in identifying contextual factors ⁵ that drive context
 99 dependence. For example, in a multi-taxon meta-analysis, if stratification of studies by
 100 species yields low heterogeneity at the taxon level, the focal effect still can be generalizable
 101 across taxon (in terms of accounting for within-taxon variation; Fig. 2). This is so, even if the
 102 total heterogeneity remains high ⁸.



103

104 **Fig. 2:**

105 A cross-taxa meta-analysis with a high total variance can have a small amount of species-specific
 106 heterogeneity. The focal effect is still possible to be generalizable at the species level. The circles
 107 represent the replication of species-specific effect. The red dashed lines denote the meta-analytic
 108 mean effects. See a real example in **Modelling additional source heterogeneity**.

109

110 Here, we present solutions to the aforementioned limitations, offering pluralistic pathways to
111 biological generality and transferability. We begin by reformulating the concept of
112 heterogeneity within the multilevel meta-analytic model and evaluating commonly used
113 heterogeneity measures. Building on this foundation, we take currently underused
114 heterogeneity metrics and propose new, stratified versions. After introducing the theoretical
115 background, we leverage a big dataset spanning 512 meta-analyses from the fields of ecology
116 and evolutionary biology (cf. ^{17,18}) to unveil empirical patterns of heterogeneity using these
117 measures and establish meta-scientific evidence on their (in)congruence. Next, we show ways
118 to visualise measures of heterogeneity using predictive distributions. Finally, we provide
119 practical recommendations and a tutorial with R functions for researchers to navigate the
120 complex landscape of heterogeneity (https://yefeng0920.github.io/heterogeneity_guide/). Our
121 synthesis highlights the significance of adopting a pluralistic framework for a comprehensive
122 understanding of meta-analytic findings in ecology and evolutionary biology.

123 **Discerning biological generality**

124 **Heterogeneity in multilevel meta-analytic modelling framework**

125 Data used in meta-analyses often exhibit a complex hierarchical structure^{5,19}, with study
126 identity serving as a typical clustering variable, forming two strata (or more). Ecological and
127 evolutionary meta-analyses typically report around eight effect sizes per study²⁰. However,
128 Traditional random-effects meta-analytic approaches do not account for heterogeneity driven
129 by such data stratification^{6,7,9}, and multi-level meta-analysis is required to model
130 heterogeneity at different strata or multi-levels in a meta-analysis (see **Methods**).

131

132 In the simplest multilevel model, the effect size estimate $ES_{[i]}$ is modelled as a combination
133 of the population mean effect or meta-analytic mean effect size μ , random effects at two
134 strata (i.e., between- and within-study levels), and statistical noise:

$$135 \quad ES_{[i]} = \mu + u_{between[j]} + u_{within[i]} + e_{[i]}, (1)$$

136 The typical assumptions for Equation 1 is as follows: (i) between-study-level random effect
137 $u_{b[j]}$ follows a normal distribution with mean zero and variance $\sigma_{between}^2$: $u_{between[j]} \sim$
138 $\mathcal{N}(0, \sigma_{between}^2)$, (ii) within-study-level random effect $u_{within[i]}$ follows a normal distribution
139 with mean zero and variance σ_{within}^2 : $u_{within[i]} \sim \mathcal{N}(0, \sigma_{within}^2)$, and (iii) sampling error $e_{[i]}$
140 follows a normal distribution with mean zero and variance in effects defined by the sampling
141 variance ($v_{[i]}$) associated with each effect size, i , such that $e_{[i]} \sim \mathcal{N}(0, v_{[i]})$. The assumption
142 of homogeneous variances for the random effects can be relaxed to allow for
143 heteroscedasticity²¹. Similarly, the assumption of independent sampling errors ($e_{[i]}$) can be
144 relaxed to allow for sampling error covariance $v_{[i]}$ ⁷. In the following sections, we will
145 elaborate on how to stratify heterogeneity information using Equation 1.

146

147 **Unstandardised heterogeneity metrics**

148 Cochran's Q is a widely used metric for assessing heterogeneity in meta-analyses²². It serves
149 as a test statistic to determine whether the true effects are homogeneous or not, informing a
150 binary decision as to whether the effect sizes come from a common underlying population, or
151 not (i.e., is heterogeneity 'non-zero'?). In contrast, the variance of true effects ($\sigma_{total}^2 =$
152 $\sigma_{between}^2 + \sigma_{within}^2$) provides a direct measure of absolute heterogeneity. Equation 1 offers a
153 general way to partition the variance of the observed effects into sampling error variance, and
154 that of true effects at different strata, such as between-study ($\sigma_{between}^2$) and within-study
155 strata (σ_{within}^2). By considering additional strata, such as variation in effects among species or
156 geographical locations, the total variance in true effects (σ_{total}^2) can be further decomposed to
157 assess generality at these specific strata (See **Model additional source heterogeneity**). For
158 example, high variation among studies implies lack of generality from one study to another
159 while low variation among species implies effects are similar, on average, across species.
160 Nonetheless, relying solely on such absolute variance may not provide practical intuition
161 regarding the magnitude of heterogeneity. For example, in a meta-analysis with $\sigma_{total}^2 = 1$, it
162 is unclear whether this amount of variance is large and meaningful because absolute variance
163 is not unit-free and not comparable across effect size measure used.

164

165 **Variance-standardised heterogeneity metrics**

166 The heterogeneity index, I^2 has emerged as the most popular metric as it provides a
167 standardized measure of heterogeneity that accounts for the scale dependence (i.e., unit-free)
168 ¹⁰. I^2 is a variance-scaled heterogeneity metric that measures the proportion of total variance
169 beyond statistical noise¹³. The total I^2 can be computed by dividing the variance in the true
170 effects (σ_{total}^2) by the variance in the observed effects ($\text{Var}[ES_{[i]}]$), as follows:

171
$$I_{total}^2 = \frac{\sigma_{total}^2}{\text{Var}[ES_{[i]}]} = \frac{\sigma_{total}^2}{\sigma_{total}^2 + \bar{v}}, (2)$$

172 where \bar{v} represents the “typical” sampling error variance. \bar{v} can be computed using different
 173 estimators^{23,24}, with the common one being¹³:

174
$$\bar{v} = \frac{(k-1) \sum_{i=1}^k 1/v_{[i]}}{(\sum_{i=1}^k 1/v_{[i]})^2 - \sum_{i=1}^k 1/v_{[i]}^2}, (3)$$

175 Within the multilevel modelling framework, the total I^2 can be stratified at different strata
 176 ^{5,24}, for example, by estimating I^2 at between-study ($I_{between}^2$) and within-study (I_{within}^2)
 177 levels:

178
$$I_{between}^2 = \frac{\sigma_{between}^2}{\text{Var}[ES_{[i]}]} = \frac{\sigma_{between}^2}{\sigma_{total}^2 + \bar{v}}, (4)$$

179
$$I_{within}^2 = \frac{\sigma_{within}^2}{\text{Var}[ES_{[i]}]} = \frac{\sigma_{within}^2}{\sigma_{total}^2 + \bar{v}}, (5)$$

180 However, as mentioned earlier, large I^2 values do not necessarily imply a practically relevant
 181 amount of heterogeneity (see Fig. 1; also see a case study in **Model additional source of**
 182 **heterogeneity**). Statistical noise can sometimes inflate I^2 values, which is a common
 183 occurrence in ecology and evolutionary meta-analyses (see **Empirical patterns of**
 184 **heterogeneity in ecology and evolution**). Stratified I^2 metrics range from 0 to 100% (but
 185 together sum to 100%), providing a clearer intuition of the source of heterogeneity and aiding
 186 in assessing the drivers of context dependence at different strata. For example, a I_{within}^2 of
 187 90% means within-study variation can account for 90% of heterogeneity, therefore, indicating
 188 that within-study level predictors are more likely to drive context dependence. I^2 and its
 189 stratified variants can also be transformed into the ratio of the variance of true effect to
 190 typical sampling error variance ($\frac{\sigma^2}{\bar{v}} = \frac{I^2}{(1-I^2)}$ or $\log\left(\frac{\sigma^2}{\bar{v}}\right) = \text{logit}(I^2)$), which represents
 191 heterogeneity as a proportion of the statistical noise (sampling error variance).

192

193 **Mean-standardised heterogeneity metrics**

194 Evolutionary biologists and behavioural ecologists are familiar with the variance-scaled
195 metrics such as heritability (h^2) and repeatability (R), which are statistically comparable to
196 the heterogeneity index, I^2 . Although less commonly used, there also exists the mean-scaled
197 counterparts, such as evolvability or the coefficient of variation (CV) for additive genetic
198 variance (CV_A) and CV for between-individual variance (CV_B)²⁵. In a similar manner, there
199 exists a mean-scaled heterogeneity metric that can provide a standardized measure of
200 heterogeneity, denoted as CV_{total} , that compares the standard deviation σ_t to the magnitude
201 of its population mean (μ)²³:

202
$$CV_{total} = \frac{\sigma_{total}}{|\mu|}, (6)$$

203 CV_t expresses the total heterogeneity as a proportion of the meta-analytic mean effect (or as a
204 percentage of change in the meta-analytic mean effect when multiplied by 100). To provide a
205 more precise quantification of heterogeneity at different strata, we propose stratified versions
206 of CV_t . Under the simplest multilevel model framework (Equation 1), we propose estimating
207 between-study, CV_b , and within-study, CV_w , as follows:

208
$$CV_{between} = \frac{\sigma_{between}}{|\mu|}, (7)$$

209
$$CV_{within} = \frac{\sigma_{within}}{|\mu|}, (8)$$

210 Notably, these mean-scaled variance metrics have the limitation of becoming arbitrarily large
211 as the magnitude of meta-analytic mean effect $|\mu|$ approaches zero²⁶. It is this limitation that
212 has probably prevented the widespread adoption of the mean-scaled variance in the field of
213 evolutionary quantitative genetic and animal personality research^{25,27}.

214

215 **Variance-mean-standardised heterogeneity metrics**

216 To remedy the problems of I^2 and CV_{total} as illustrated above, there is a more robust measure
217 of heterogeneity M_{total} that combines the strengths of mean-scaled and variance-scaled
218 metrics ¹¹:

$$219 \quad M_{total} = \frac{\sigma_{between} + \sigma_{within}}{\sigma_{between} + \sigma_{within} + |\mu|}, (9)$$

220 Here we propose between-study ($M_{between}$) and within-study (M_{within}) versions by
221 stratifying M_t , which allows for a more precise quantification of heterogeneity at specific
222 strata:

$$223 \quad M_{between} = \frac{\sigma_{between}}{\sigma_{between} + \sigma_{within} + |\mu|}, (10)$$

$$224 \quad M_{within} = \frac{\sigma_{within}}{\sigma_{between} + \sigma_{within} + |\mu|}, (11)$$

225 M_t and its stratified variants are still standardised measures that quantify the size of
226 heterogeneity relative to the magnitude of meta-analytic mean effect, providing intuitive
227 interpretation. For example, $\sigma_{total} = 0$ leads to $M_{total} = 0$, indicating the population mean
228 effect is fully generalisable, and replicable across different contexts (see a case study in
229 **Model additional source of heterogeneity**). On the other hand, M_{total} and its stratified
230 variants are truncated at one, which overcomes the issue of CV_{total} when the magnitude of
231 meta-analytic mean effect $|\mu|$ approaches zero. Note that there is another mean- and variance-
232 scaled metric, M_{total}^2 , where σ_{total} and $|\mu|$ are replaced by their squared values (**Methods**).
233 CV_{total} , M_{total} and M_{total}^2 can be all be easily stratified using multilevel meta-analytic
234 models (**Model additional source of heterogeneity**).

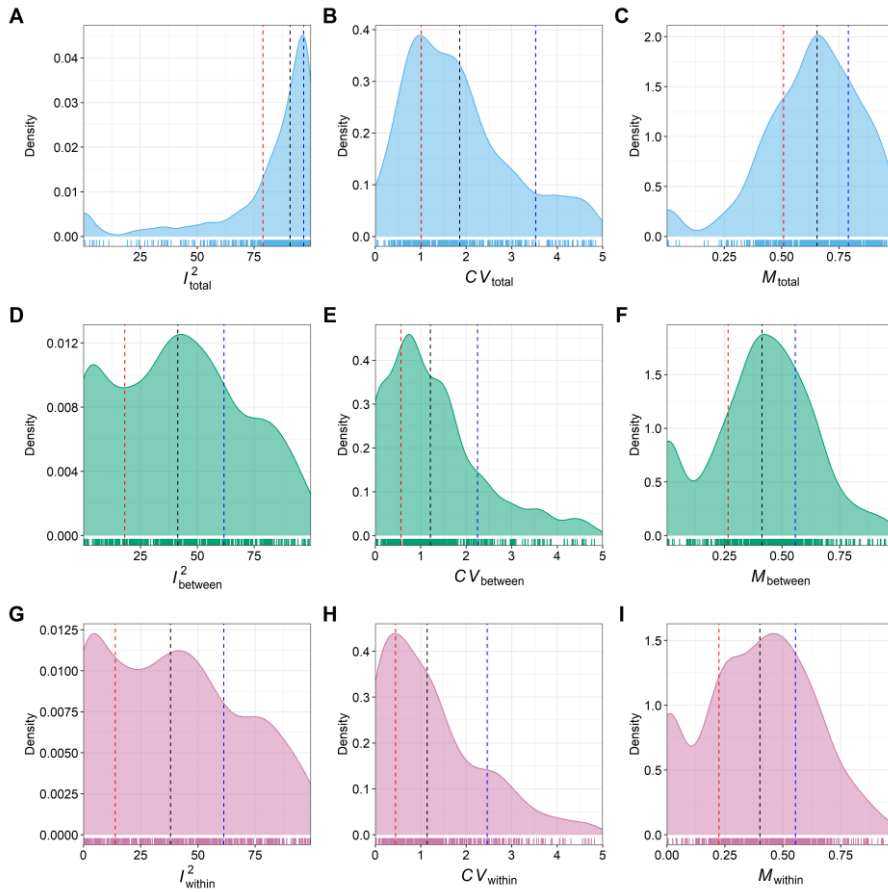
235

236 **Empirical patterns of heterogeneity in ecology and evolution**

237 To evaluate empirical patterns in heterogeneity among meta-analytic studies in ecology and
238 evolution, we applied multilevel meta-analytic models (Equation 1) to 512 published meta-

239 analyses^{18,28}. For each meta-analysis, we quantified and stratified heterogeneity using I_{total}^2 ,
240 CV_{total} , M_{total} . For I_{total}^2 , the 25th, 50th, and 75th percentiles corresponded to 79%, 91%,
241 and 97% I_{total}^2 , respectively (Fig. 3), rather than conventional thresholds for interpreting I^2 ,
242 which typically categorize heterogeneity as small, moderate, or high at 25%, 50%, and 75%
243 I_{total}^2 , respectively¹⁰. This also means, on average, variation in true effect sizes σ^2 was ten
244 times as large as typical sampling error variance ($\frac{\sigma^2}{\bar{v}} = \frac{I^2}{(1-I^2)} = 10$; see Figs. S1 and S2 for
245 empirical distributions of σ^2 and \bar{v}) and 91% of them can be attributed to the ‘true’ biological
246 or methodological differences in research contexts, and thus are theoretically explainable
247 using appropriate predictors.

248
249 While I_{total}^2 displayed a left-skewed and single-modal distribution, its stratified counterparts,
250 $I_{between}^2$ and I_{within}^2 , demonstrated a right-skewed distribution with multi-modal patterns.
251 There was no consistent trend suggesting one type of stratified heterogeneity consistently
252 outweighed the other across the 512 meta-analyses (Fig. 3). Intriguingly, 47% (242 out of
253 512) of the meta-analyses exhibited smaller between-study level heterogeneity than within-
254 study level heterogeneity ($I_{between}^2 < I_{within}^2$; Fig. 4). Within this subset of meta-analyses, the
255 median values for I_{total}^2 , $I_{between}^2$ and I_{within}^2 were 95%, 21%, and 63%, respectively. It
256 highlights a key finding often overlooked by traditional heterogeneity quantification
257 practices: findings from many meta-analyses with high total heterogeneity can still be
258 generalized at the between-study study level. Such generalization is achievable when
259 replication is defined as the testing of the null hypothesis at the between-study level, and
260 when within-study methodological and biological variations can be adequately accounted for
261 (i.e., within-lab heterogenization²⁹) because some meta-analyses have relatively low
262 heterogeneity at the between-study study level.



263

264 **Fig. 3:**

265 The distribution of heterogeneity estimates derived from 512 meta-analyses was systematically

266 assessed using pluralistic measures and stratified across different strata. Total heterogeneity measures

267 (A – C): I^2_{total} , CV_{total} and M_{total} . Between-study heterogeneity measures (D – E): $I^2_{between}$,

268 $CV_{between}$ and $M_{between}$. Within-study heterogeneity measures (G – I): I^2_{within} , CV_{within} and

269 M_{within} . Three dashed lines correspond to the 25th, 50th, and 75th percentiles, respectively. In panels

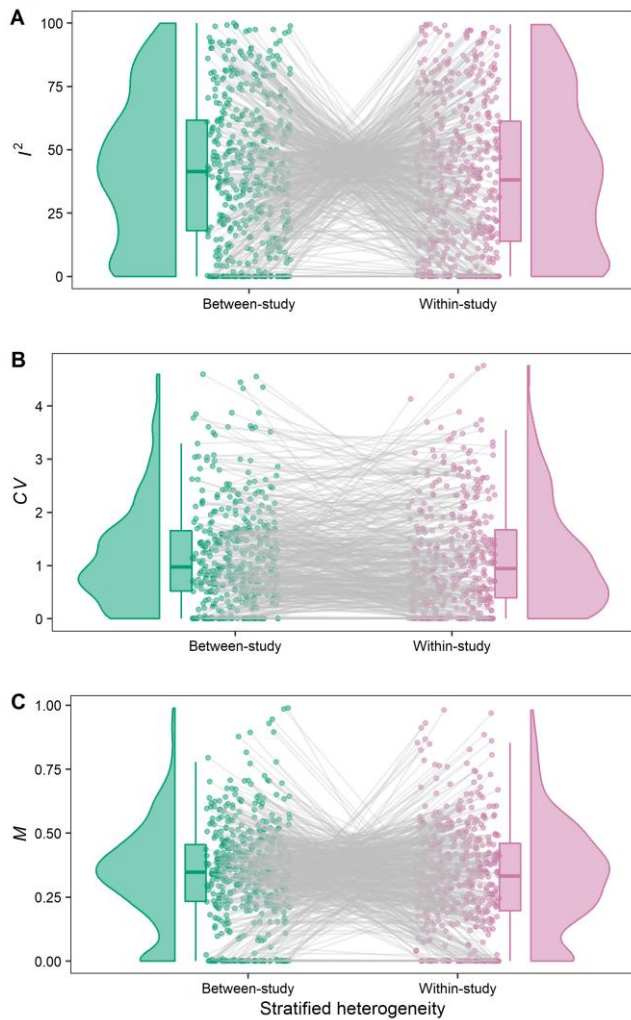
270 B, E, and H, the CV was truncated at five for figure clarity, as very large CV values can be challenging

271 to interpret when the meta-analytic mean effect is small. For example, the maximum CV observed in

272 the 512 meta-analyses was 106, which was inflated by a small meta-analytic mean effect of 0.03. For
273 the figures without truncation, please refer to Figure S3.

274

275 When the CV_{total} metric was used to quantify heterogeneity, the calculated 25th, 50th, and
276 75th percentiles of CV_{total} values were 1.0, 1.8, and 3.5, respectively (Fig. 3). This means
277 that the standard deviation (in this case, heterogeneity) was, on average, nearly twice that of
278 the meta-analytic mean effect. The distributions of both CV_{total} and its stratified versions,
279 $CV_{between}$, and CV_{within} , displayed a right-skewed pattern with a single-mode. In contrast, the
280 distribution of M_t exhibited a more symmetrical pattern, with the 25th, 50th, and 75th
281 percentiles of M_{total} values being 0.5, 0.6, and 0.8, respectively (Fig. 3), albeit with a minor
282 peak around zero. Notably, stratification analysis revealed that $M_{between}$ and M_{within} had
283 patterns similar to those observed for $CV_{between}$ and CV_{within} . This similarity is expected as
284 they can be mathematically transformed into one another using equations $M_{total} =$
285 $CV_{total}/(1 + CV_{total})$ and $logit(M_{total}) = \log(CV_{total})$. The median values for both
286 CV_{total} and M_{total} across the 512 meta-analyses signify a high amount of heterogeneity,
287 thereby warranting a thorough exploration into the drivers influencing such context
288 dependence. However, stratification of M_{total} also suggests that meta-analyses with high
289 heterogeneity can possess a considerable likelihood of generality at the between-study level,
290 given low $M_{between}$ (as we pointed out above with I^2). On average, there was a median
291 $M_{between} = 0.3$ (SD is 41% of the meta-analytic mean effect) observed in 47% of the meta-
292 analyses (242/512) with smaller $M_{between}$ values compared to M_{within} values (Fig. 4).



293

294 **Fig. 4:**

295 Paired comparison of stratified heterogeneity estimates derived 512 meta-analyses for three

296 heterogeneity metrics (A) I^2 , (B) coefficient of variation, CV and (C) M . Heterogeneity was stratified

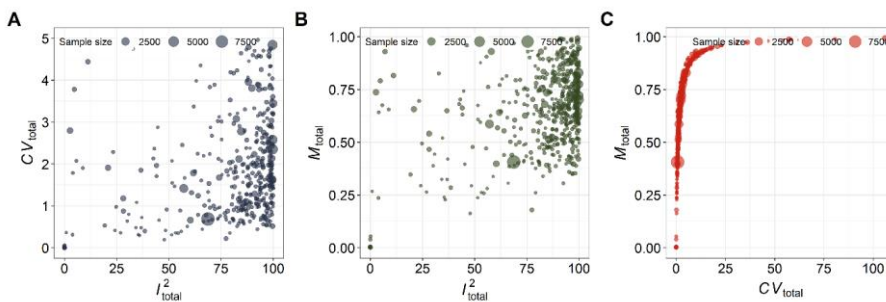
297 at both 'between-study' and 'within-study' levels (x-axes). Each point represents an estimate from

298 each meta-analysis. For panel B, CV has been truncated at five for figure clarity. For the full figures

299 without truncation, please refer to Figure S4. For other details see Fig. 3.

300

301 We found only moderate agreement between heterogeneity measured as I^2 and the
 302 alternatives (CV_{total} : $r_{\text{spearman}} = 0.32$, 95% CI = [0.24, 0.40], M_{total} : $r_{\text{spearman}} = 0.33$, 95% CI =
 303 [0.25, 0.41]; Fig. 5). In cases of meta-analyses with I^2 larger than 75% or smaller than 25%
 304 (identified as large and small heterogeneity by conventional benchmarks¹⁰), the disagreement
 305 between I^2 and CV , as well as I^2 and M , became even more pronounced (Fig. S5 – S7). In
 306 contrast, a near-perfect agreement was observed between CV_{total} and M_{total} , as expected
 307 ($r_{\text{spearman}} = 1$, 95% CI = [0.99, 1]; Fig. 5). Therefore, cross-meta-analysis (meta-scientific)
 308 evidence suggests that the heterogeneity source measure I^2 is not consistent with the
 309 magnitude measures (CV_{total} and M_{total}) for ecological and evolutionary data. We also found
 310 that out of the 512 meta-analyses featuring medium to large I^2_{total} values (>50% based on
 311 conventional guidelines), 80 had small CV_{total} (Fig. 5), indicating that more than 20% of the
 312 large I^2_{total} values were caused by small sampling errors rather than larger amount of
 313 heterogeneity. These findings emphasize the importance of considering multiple metrics to
 314 obtain a holistic understanding of heterogeneity in meta-analyses (see **A pluralistic**
 315 **framework**).



316
 317 **Fig. 5:**
 318 Disagreement (or agreement) between different heterogeneity metrics. For other details see Fig. 3.
 319 The Spearman correlation estimates (r_{spearman}) were: 0.32, 95% CI = [0.24, 0.40] for I^2_{total} and CV_{total} ,
 320 0.33, 95% CI = [0.25, 0.41] for I^2_{total} and M_{total} , and 1, 95% CI = [0.99, 1] for M_{total} and CV_{total} .

321

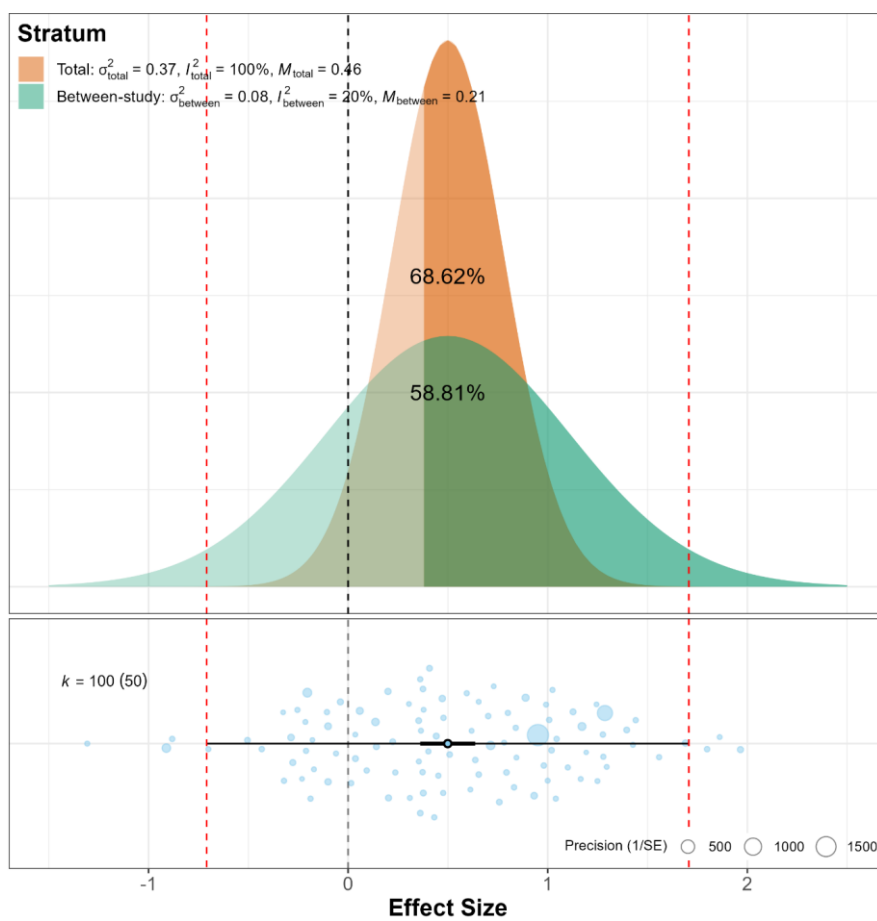
322 **Prediction intervals and predictive distributions: visualising heterogeneity**

323 Prediction intervals (PIs) are underreported but insightful in meta-analytic heterogeneity and

324 generality. Surveys have shown that less than one per cent (1/102) of such studies reported

325 PIs³⁰). PIs are derived from the definition of σ_t^2 and provide a range within which a future

326 effect size is predicted to fall with a certain probability¹⁴, often 95% (Fig. 6).



327

328 **Fig. 6:**

329 Example of how prediction intervals (PIs) combined with ‘prediction distributions’ (PDs) can be used

330 to understand effect size heterogeneity and generality. Effect size data are simulated assuming two

331 effect sizes were collected from a total of $n = 50$ studies, ($k = 100$), with a $\sigma_{between}^2 = 0.1$, $\sigma_{within}^2 =$
332 0.6 and an overall meta-analytic mean, μ , of 0.5
333 (https://yefeng0920.github.io/heterogeneity_guide/). Red dashed lines are the upper and lower
334 95% PI, black dashed line the 'null' effect. The orchard plot ³⁰ displays the overall meta-analytic
335 mean, 95% confidence interval (CI) and 95% PI. The PDs were constructed using t -distribution with k
336 -1 degrees of freedom, μ as location parameter, and total or between-study variance along with
337 sampling variance of around μ as scale parameter (see Equation 11). The percentage of effect sizes
338 beyond a given threshold (i.e., the lower 95% CI) are provided.

339
340 For example, consider a conservation intervention with a mean effect size (SMD) of -0.5 and
341 95% PI of [-0.2 to -0.8]. This indicates that 95% of future interventions implemented in are
342 predicted to decrease the conservation outcomes of interest by between 0.2 to 0.8 standard
343 deviations. Unlike the point estimate of heterogeneity, such as σ_t^2 , PIs offer an interval to
344 inform the extent to which the focal effect can be generalized ³¹. Under Equation 1, 95% PIs
345 can be computed by ⁷:

$$346 \quad 95\%PI = \mu \pm t_{0.975} \sqrt{\sigma_{between}^2 + \sigma_{within}^2 + SE[\mu]^2}, (11)$$

347 where $t_{0.975}$ denotes the 97.5th percentile of a t -distribution (with $k-1$ degrees of freedom ³²,
348 where k is the number of sample size), and $SE[\mu]$ denotes the standard error of the mean
349 effect μ .

350
351 Despite their usefulness, PIs can create the illusion that all effect sizes within the upper and
352 lower intervals are equally likely (Fig. 6; see also ³³). Therefore, statisticians have
353 emphasised the importance of visualising the probability density to accurately capture the
354 likelihood of each effect size within the intervals ^{34,35}. By considering the entire distribution
355 of true effects while accounting for statistical noise, the predictive distribution (PD) offers a

356 more holistic measure of heterogeneity and generality. In the Bayesian framework, PDs,
357 known as posterior distributions, are a natural part of the process, but even frequentist
358 approaches can adopt PDs (sometimes referred to as “empirical Bayes”) to achieve similar
359 aims. An advantage of the PD is its ability to calculate the probability that a true effect size
360 exceeds a biologically or practically meaningful threshold although determining such a
361 threshold usually requires domain-specific knowledge and expertise. The proportion of true
362 effect sizes above a specific threshold could serve as a measure of evidence strength and
363 generality¹⁶. Consider a case that 69% of effect sizes representing the efficacy of a
364 conservation intervention are predicted to surpass a threshold value representing a practically
365 significant effect (Fig.6, where we assumed the lower confidence limit representing the
366 threshold). If assuming similar configurations of study contexts in the sampled future cases,
367 we can infer that the intervention will achieve this benefit in 69% of future cases, with strong
368 implications for policymaking.

369

370 **Modelling additional sources of heterogeneity**

371 In ecological and evolutionary datasets, complexity often arises from the inclusion of diverse
372 species, temporal, and spatial variations³. Such complexity offers a unique opportunity for
373 further disentangling heterogeneity. This can be achieved by embracing a flexible random-
374 effects structure within the multilevel meta-analytic framework^{7,9}. To illustrate this, we will
375 show the principles of how to partition heterogeneity in datasets featuring multiple species
376 (similar principles can be applied to those involving different temporal and spatial contexts).
377 In the case of datasets encompassing multiple species, incorporating species-relevant
378 random-effects terms into Equation 1 would lead to the phylogenetic multilevel meta-analytic
379 model^{5,36}:

$$380 \quad ES_{[i]} = \mu + u_{species[k]} + u_{phylogeny[k]} + u_{between[j]} + u_{within[i]} + e_{[i]}, (12)$$

381 where $u_{s[k]}$ denotes the non-phylogenetic species random effect, which follows a normal
 382 distribution with mean zero and variance $\sigma_{species}^2$; $u_{phylogeny[k]}$ denotes the phylogenetic
 383 species random effect, which follows a multivariate normal distribution with mean zero and
 384 variance-covariance matrix $\sigma_{phylogeny}^2 \mathbf{A}$ (where $\sigma_{phylogeny}^2$ is the phylogenetic species
 385 variance, and \mathbf{A} is phylogenetic correlation matrix based on the distance between species on a
 386 molecular-based phylogenetic tree).

387
 388 With Equation 12 in hand, the total variance can be stratified at the phylogenetic and non-
 389 phylogenetic species level ($\sigma_{phylogeny}^2$ and $\sigma_{species}^2$). Such stratification allows for the
 390 assessment of the generality of a focal effect within these strata, as illustrated in the empirical
 391 example below. Phylogenetic and non-phylogenetic species-level heterogeneity can be
 392 measured using $I_{phylogeny}^2$ and $I_{species}^2$, respectively ⁵:

$$393 \quad I_{phylogeny}^2 = \frac{\sigma_{phylogeny}^2}{\sigma_{phylogeny}^2 + \sigma_{species}^2 + \sigma_{between}^2 + \sigma_{within}^2 + \bar{v}}, \quad (13)$$

$$394 \quad I_{species}^2 = \frac{\sigma_{species}^2}{\sigma_{phylogeny}^2 + \sigma_{species}^2 + \sigma_{between}^2 + \sigma_{within}^2 + \bar{v}}, \quad (14)$$

395 We derive the alternative stratified version of measures as follows:

$$396 \quad CV_{phylogeny} = \frac{\sigma_{phylogeny}}{|\mu|}, \quad (15)$$

$$397 \quad CV_{species} = \frac{\sigma_{species}}{|\mu|}, \quad (16)$$

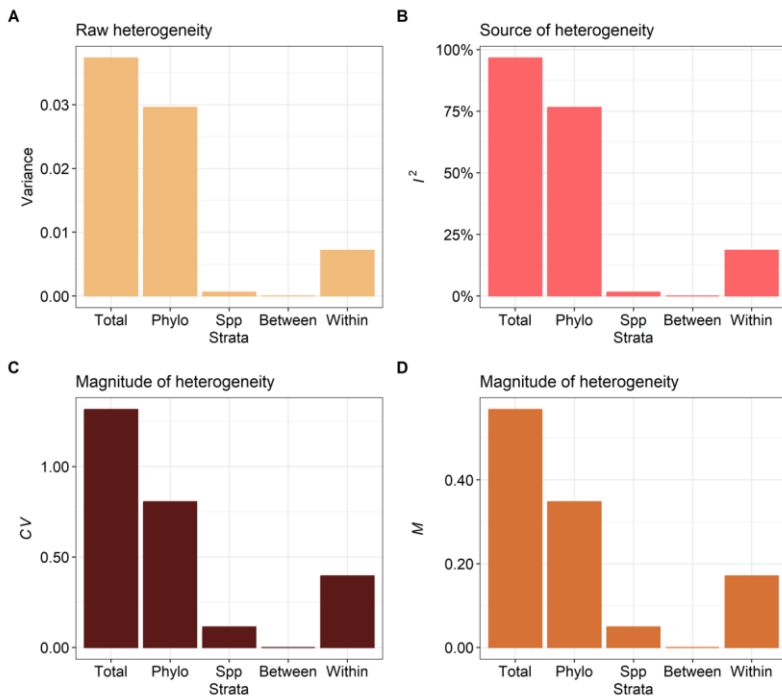
$$398 \quad M_{phylogeny} = \frac{\sigma_{phylogeny}}{\sigma_{phylogeny} + \sigma_{species} + \sigma_{between} + \sigma_{within} + |\mu|}, \quad (17)$$

$$399 \quad M_{species} = \frac{\sigma_{species}}{\sigma_{phylogeny} + \sigma_{species} + \sigma_{between} + \sigma_{within} + |\mu|}, \quad (18)$$

400 Furthermore, the predictive distribution also can be stratified at phylogenetic and non-
401 phylogenetic species-level, which provides a visual means to assess the heterogeneity and
402 generality at these strata.

403

404 To illustrate the insights gained through these extended measures, we present an empirical
405 example. We re-analysed a phylogenetic meta-analysis originally conducted by Risely et al.
406 ³⁷. Our focus centres on a subset of this analysis, specifically examining the impact of
407 infection status on the cost (e.g., movement capacity) of migratory animals. The data and
408 code for replicating all calculations can be found at
409 https://yefeng0920.github.io/heterogeneity_guide/. Our re-analysis yielded three
410 observations. Firstly, $I_{total}^2 = 97\%$ exceeded the 75th percentile of the empirically derived
411 heterogeneity distribution (Fig. 7 and Table S1). This suggests a high amount of
412 heterogeneity according to the conventional benchmarks ¹⁰. However, when we employed
413 magnitude metrics to measure heterogeneity, they fell below between the 25th and 50th
414 percentiles of the empirically derived heterogeneity distribution ($CV_{total} = 1.3$ and $M_{total} =$
415 0.6). This discrepancy was attributed to the small typical sampling variance \bar{v} , which was
416 found to be 0.001 in this case, underscoring I_{total}^2 's limitation of relying on \bar{v} to capture
417 relative magnitude of heterogeneity. On the other hand, we emphasise that the proper
418 interpretation of I_{total}^2 is to use it to indicate the source of heterogeneity rather than the
419 magnitude, as it represents the variance of the true effect in the context of the variance of the
420 observed effect. For example, $I_{total}^2 = 97\%$ suggests a heterogeneity can explain most (97%)
421 of the variability in effect size (only 3% is explained by the sampling variance, or the
422 heterogeneity is 32 times larger than that of statistical noise).



423

424 **Fig. 7:**

425 Heterogeneity quantification and stratification for multiple metrics. (A) The heterogeneity is
 426 quantified using raw variance, (B) source measure I^2 , (C) magnitude measure CV , and (D) magnitude
 427 measure M , and stratified at phylogenetic (Phylo), non-phylogenetic (Spp), between-study (Between),
 428 and within-study (Within) levels. The source measure I^2 sometimes aligns well with the raw variance,
 429 as observed in this example (A and B). However, we note that I^2 values can be challenging to
 430 interpret as the magnitude of heterogeneity, especially when the typical sampling error variance is
 431 extremely small or large. This challenge is often encountered with certain effect size measures, such
 432 as the log coefficient of variation ratio (lnCVR), as demonstrated in a real example at

433 https://yefeng0920.github.io/heterogeneity_guide/.

434

435 Secondly, the estimated mean effect was highly likely to be generalizable and replicable at
436 the between-study- and species-context, if controlling for within-study experimental contexts
437 (e.g., age, sex, outcomes). This is indicated by the stratification analysis that between-study
438 level heterogeneity was extremely low, despite a large heterogeneity according to
439 conventional benchmarks¹⁰. Traditional meta-analytic practices would overlook these
440 valuable insights, potentially leading to erroneous conclusions. For example, random-effects
441 meta-analysis shows that this dataset has high study-level heterogeneity ($I^2_{total} = 96\%$; Fig. 5
442 and Table S1). However, this amount of heterogeneity was not attributable to the study level
443 but, rather, was mainly explained by the phylogenetic signal ($I^2_{phylogeny} = 76\%$). The
444 stratified version of PD also provided a clearer visual clue that the phylogenetic signal was
445 the primary source of heterogeneity (Fig. 7).

446

447 **A pluralistic framework**

448 Given that different measures offer distinct insights into heterogeneity and generality (Table
449 1), we propose adopting a pluralistic framework to comprehensively assess heterogeneity in
450 ecological and evolutionary meta-analyses. Our recommendations are threefold:

451 (1) Employing multilevel meta-analytic framework: Provided data allow, we strongly
452 advocate for the use of a multilevel meta-analytic framework (Equation 1), as
453 opposed to random-effects meta-analysis, for the modelling and stratification of
454 heterogeneity. Additional random effects can be incorporated into Equation 1 as
455 needed to further dissect heterogeneity. For example, the application of the
456 phylogenetic multilevel meta-analytic model (Equation 12) allows for the
457 disentanglement of species-specific heterogeneity.

458 (2) Quantification and stratification of pluralistic heterogeneity measures: We recommend
459 transparently reporting all variance components, including typical sampling error

460 variances in the main text, supplementary tables, or figures (Figs. 6 and 7 and Table
461 1). As such, pluralistic metrics can be computed using the formula above. I^2 , M (with
462 CV being derivable from M), and their stratified versions should be reported as the
463 default measures. PI or PD should also be reported to provide a visual identification
464 of the heterogeneity information. These measures provide complementary
465 information, for example, the source, magnitude, and visual clue of heterogeneity
466 (examples see **Table 1**). We also provide parametric bootstrapping solutions to
467 estimate the uncertainty (e.g., 95%CI) for each of the measures.

468 (3) Check the model parameter identifiability: When models incorporate many random
469 effects, issues of parameter identifiability may arise, wherein unique variance
470 estimates that maximize the likelihood function may not exist (see **Method**)³⁹.
471 Therefore, we recommend assessing whether variance components are all identifiable
472 through means such as checking profile likelihood, before proceeding with
473 heterogeneity quantification and stratification.

474 (4) Carefully interpret heterogeneity measures: It is crucial to interpret both total and
475 stratified heterogeneity to evaluate variation in effect sizes, aiding in the examination
476 of general rules in the fields of ecology and evolution (see a case study in **Modelling
477 additional sources of heterogeneity**). However, neither the conventional benchmarks
478 (25, 50, and 75% as small, moderate and high heterogeneity¹⁰) nor those of
479 empirically derived distributions (Table 1 and Fig. 3) are currently suitable for
480 informing interpretation. Nevertheless, the empirically derived distribution can be
481 employed to interpret heterogeneity within the context of existing ecological and
482 evolutionary meta-analyses.

483

484 We argue that ecologists and evolutionary biologists should treat heterogeneity and the meta-
485 analytic mean effect size with equal importance. We provide a user-friendly tutorial equipped
486 with a set of R functions to streamline the qualification, stratification, and interpretation of
487 heterogeneity https://yefeng0920.github.io/heterogeneity_guide/, empowering ecologists and
488 evolutionary biologists to discern generality.
489

490 Table 1

491 Summary of heterogeneity measures, their stratified counterparts, and empirically derived benchmark values. SMD denotes standardised mean
 492 difference. lnRR denotes log response ratio. Zr denotes Fisher's r-to-z transformed correlation coefficient. 2-by-2 table denotes often
 493 dichotomous (binary) effect size measures, such as log odds ratio, log risk ratio. Uncommon measures represent less frequently used effect size
 494 measures, such as raw mean difference and regression coefficients.

Types	Metrics	Interpretation and examples	Empirically derived benchmark ¹
Test statistic	Q	Null-hypothesis test. Statistical test of heterogeneity in effect sizes.	Not applicable
Unstandardisation	σ^2	Absolute magnitude measure of heterogeneity. Variance (square of standard deviation) of the meta-analytic mean effect (σ_{total}^2) and its stratification at between- and within-study contexts ($\sigma_{between}^2$ and σ_{within}^2).	25th, 50th, and 75th percentiles (Fig. S1): 0.54, 1.25, 3.03 for SMD; 0.11, 0.27, 0.57 for lnRR; 0.06, 0.12, 0.25 for Zr ; 1.04, 1.20, 2.51 for the 2-by-2 table; 0.01, 0.04, 0.27 for uncommon measures. The percentiles of typical sampling variance \bar{v} are reported at Fig. S2.
Variance-standardization	I^2	Heterogeneity source measure. Proportion of variance not due to statistical noise. It measures the source of heterogeneity. For example, $\sigma_{total}^2 = 95\%$ denotes that 95% of variation is the result of nuisance heterogeneity (i.e., differences in contexts). $\sigma_{between}^2 = 80\%$ and $\sigma_{within}^2 = 15\%$ indicate differences in between-study contexts dominate the heterogeneity, pointing towards between-study level predictors as the likely drivers of context-dependent variation.	25th, 50th, and 75th percentiles (Fig. 3): 79%, 91%, 97% for overall; 78%, 89%, 96% for SMD; 88%, 95%, 99% for lnRR; 73%, 87%, 95% for Zr ; 71%, 73%, 89% for the 2-by-2 table; 74%, 91%, 98% for uncommon measures.

Commented [SN1]: can we put a disclaimer that the spread could be underestimated - these values could be underestimated if we have publication bias - this is especially so for CV and M

Should discuss with Shinichi

Commented [YY2R1]: Good point

Mean-standardization	<i>CV</i>	Heterogeneity magnitude measure. Variance expressed as the proportion of the mean effect. It is the measure of the magnitude of heterogeneity in the context of mean effect. For example, $CV_{total} = 1.5$, $CV_{between} = 0.8$, and $CV_{within} = 0.5$ denote that total, between- and within-study variance are 150, 80, and 50% of the mean effect.	25th, 50th, and 75th percentiles (Fig. 3): 1.0, 1.8, 3.5 for overall; 1.1, 2.0, 3.9 for SMD; 1.2, 1.9, 3.5 for lnRR; 0.8, 1.7, 2.9 for Zr; 1.2, 2.2, 2.7 for the 2-by-2 table; 0.7, 1.1, 1.3 for uncommon measures.
Variance-mean-standardization	<i>M</i>	Heterogeneity magnitude measure. Variance expressed as the proportion of the mean effect and a transformation of <i>CV</i> designed with better properties. It is the measure of the magnitude of heterogeneity in the context of mean effect. The interpretation can be eased by back-transformation with $M_{total} = CV_{total}/(1 + CV_{total})$. For example, $CV_{total} = 0.6$, $CV_{between} = 0.5$, and $CV_{within} = 0.4$ denote that total, between- and within-study variance are 150, 100, and 67% of the mean effect.	25th, 50th, and 75th percentiles (Fig. 3): 0.5, 0.7, 0.8 for overall; 0.5, 0.7, 0.8 for SMD; 0.5, 0.7, 0.8 for lnRR; 0.5, 0.6, 0.8 for Zr; 0.5, 0.7, 0.7 for the 2-by-2 table; 0.4, 0.5, 0.6 for uncommon measures.
Visual metric	PI & PD	Heterogeneity visual measure. A plausible interval where a new effect size is predicted to fall with a specified level of probability. It can be used to visually diagnose the heterogeneity and generality of the mean effect. For example, a 95% prediction interval (PI) of [-0.2 to -0.8] indicates that 95% range of future effect sizes are expected in studies with similar contexts. The whole predictive distribution (PD) can be used to derive the probability of a newly observed effect being above a biologically meaningful threshold.	Not applicable

495 ¹The distributions and percentiles could be underestimated if publication bias existed.

497 **Methods**

498 **Meta-analysis database**

499 The ecological and evolutionary database used in this study were originally compiled by
500 Costello¹⁸, O'Dea¹⁷, and their colleges. They conducted a systematic search for meta-
501 analysis papers published in ecological journals, including those from the Ecological Society
502 of America and journals of the British Ecological Society. Additionally, they supplemented
503 the database with high-profile journals, such as Nature, and Science. Their systematic search
504 yielded 522 meta-analysis datasets. We dropped meta-analysis datasets that could not achieve
505 convergence when fitted to the multilevel model. Convergence could not be reached for ten
506 meta-analysis datasets, even after adjusting key parameters of the iterative methods to
507 maximize the log likelihood function (see below for details). Therefore, our database
508 contained 512 meta-analysis datasets encompassing 17,770 primary studies and 109,495
509 effect size estimates. On average, each meta-analysis dataset included 240 effect size
510 estimates sourced from 40 studies, with median values of 64 and 23, respectively.

511

512 **Stratification of hierarchical meta-analytic data**

513 In this section, we elucidate the theoretical background behind employing a three-level meta-
514 analytic approach to stratify datasets characterized by three-level hierarchical structure as
515 outlined above. Note that the stratification of heterogeneity can be further extended to data
516 structures with more than four strata as necessary (see a case study in **Model additional**
517 **source heterogeneity**). In the first-stage modelling procedure, the true (population) effect
518 size $\mu_{between[j]}$ of j -th study is modelled using a normal distribution with expectation μ and
519 variance $\sigma_{between}^2$, where μ is the population mean effect or overall effect and $\sigma_{between}^2$
520 denotes the extent to which $\mu_{between[j]}$ deviates from the overall effect μ ^{24,40}. Moving to the
521 second-stage modelling procedure, the i -th effect size $\mu_{within[i]}$ within j -th study is modelling

522 using a normal distribution with expectation $\mu_{between[j]}$ and variance σ_{within}^2 , where σ_{within}^2
523 represents the extent to which within-study effect $\mu_{within[i]}$ deviates from between-study
524 effect $\mu_{between[j]}$ ^{24,40}. In the third-stage modelling procedure, the effect size estimate $ES_{[i]}$ of
525 $\mu_{within[i]}$ is modelled using a normal distribution with expectation $\mu_{within[i]}$ and sampling
526 error variance $v_{[i]}$. This multilevel modelling framework provides a general way to
527 decompose the variance of effect sizes into different strata, for example between- and within-
528 study levels.

529
530 From the implementation perspective, effect size estimate $ES_{[i]}$ is not sequentially modelled
531 through the three-stage process but rather directly modelled from the overarching distribution
532 with an expectation μ and variance-covariance matrix VCV ^{24,40}:

$$533 \begin{bmatrix} \sigma_{between}^2 + \sigma_{within}^2 + v_{[1]} & \cdots & \sigma_{between}^2 \\ \vdots & \ddots & \vdots \\ \sigma_{between}^2 & \cdots & \sigma_{between}^2 + \sigma_{within}^2 + v_{[k]} \end{bmatrix}, (19)$$

534 The meta-analytic model specified with the variance-covariance matrix VCV is referred to as
535 the multilevel meta-analytic model (Equation 1). VCV can be reparametrized as a compound
536 symmetry random-effects structure within the framework of multivariate meta-analytic model
537 ^{40,41}.

$$538 \begin{bmatrix} \sigma_{total}^2 + v_{[1]} & \cdots & \rho\sigma_{total}^2 \\ \vdots & \ddots & \vdots \\ \rho\sigma_{total}^2 & \cdots & \sigma_{total}^2 + v_{[k]} \end{bmatrix}, (20)$$

539 where $\sigma_{total}^2 = \sigma_{between}^2 + \sigma_{within}^2$ is the total variance in effect sizes and $\rho =$
540 $\sigma_{between}^2 / \sigma_{total}^2$ denotes intraclass correlation coefficient. We used the *rma.mv()* function
541 from the *metafor* package⁴² to fit all 512 meta-analysis datasets to the three-level meta-
542 analytic model (Equation 1). We employed restricted maximum likelihood (REML) as the
543 variance estimator and the quasi-Newton method as the optimizer to maximize the likelihood

544 function over variance estimation ($\sigma_{between}^2$ and σ_{within}^2), with a threshold of 10^{-8} , a step
545 length of 1, and a maximum iteration limit of 1000. All models successfully converged under
546 these settings. We confirmed the identifiability of variance estimation ($\sigma_{between}^2$ and σ_{within}^2)
547 by checking their likelihood profiles. The R code for model fitting can be accessed at the
548 https://github.com/Yefeng0920/heterogeneity_ecoevo.

549

550 **Extended heterogeneity metrics**

551 In addition to CV_{total} , M_{total} , and their stratified counterparts (Equations 6 – 11), we
552 introduce two related heterogeneity measures. CV_{total} has a potential shortcoming that it is
553 not numerically equivalent to the sum of heterogeneity at between- and within-study levels
554 ($CV_{total} \neq CV_{between} + CV_{within}$). This is because the total standard deviation σ_t is not equal
555 to the sum deviations at each stratum ($\sigma_{total} \neq \sigma_{between} + \sigma_{within}$). To address the numerical
556 difference, we propose CV_{total}^2 , an analogue to CV_{total} :

$$557 \quad CV_{total}^2 = \frac{\sigma_{total}^2}{\mu^2}, (21)$$

558 Similarly, we propose between-study level and within-study level variants ($CV_{between}^2$ and
559 CV_{within}^2):

$$560 \quad CV_{between}^2 = \frac{\sigma_{between}^2}{\mu^2}, (22)$$

$$561 \quad CV_{within}^2 = \frac{\sigma_{within}^2}{\mu^2}, (23)$$

562 Following the same principle, M_{total}^2 can be obtained ¹¹:

$$563 \quad M_{total}^2 = \frac{\sigma_{total}^2}{\sigma_{total}^2 + \mu^2}, (24)$$

564 We further propose between-study level (M_{total}^2) and within-study level (M_{total}^2) counterparts
565 as:

566
$$M_{between}^2 = \frac{\sigma_{between}^2}{\sigma_{total}^2 + \mu^2}, (25)$$

567
$$M_{within}^2 = \frac{\sigma_{within}^2}{\sigma_{total}^2 + \mu^2}, (26)$$

568 M_{total}^2 and its stratified variants ($M_{between}^2$ and M_{within}^2) are re-scaling of CV_{total}^2 and its
569 stratified variants ($CV_{between}^2$ and CV_{within}^2). Therefore, they can be converted into each other
570 using simple mathematical relationships, such as $M_{total}^2^{-1} = CV_{total}^2^{-1} + 1$ or
571 $\text{logit}(M_{total}^2) = \log(CV_{total}^2)$.

572 **Data availability**

573 The data needed to reproduce the analyses and figures are archived GitHub repository
574 https://github.com/Yefeng0920/heterogeneity_ecoevo/tree/main, and will be deposited at
575 Zenodo after acceptance.

576 **Code availability**

577 The scripts needed to reproduce the analyses and figures are archived GitHub repository
578 https://github.com/Yefeng0920/heterogeneity_ecoevo/tree/main, and will be deposited at
579 Zenodo after acceptance.

580

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675

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681

682 **Author contributions**

683 YY: Conceptualization; data curation; formal analysis; investigation; methodology; software;
684 visualization; writing – original draft; writing – review and editing. DWAN: Software;
685 visualization; writing – review and editing. RS: Writing – review and editing. AMS: Writing
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690

691 **Competing interests**

692 All authors declare no competing interests.

693

694 **Additional information**

695 Supplementary materials will be available at the online version.