

1 **Sex-specific associations between social behaviour, its predictability and fitness in a wild**  
2 **lizard**

3 **Authors:** Barbara Class<sup>1\*</sup>, Kasha Strickland<sup>2</sup>, Dominique Potvin<sup>3</sup>, Nicola Jackson<sup>4</sup>, Shinichi  
4 Nakagawa<sup>5</sup>, Celine Frere<sup>4</sup>

5 <sup>1</sup> Department of Biology, Ludwig-Maximilians-Universität München, Großhaderner Straße 2, 82152  
6 Planegg-Martinsried

7 <sup>2</sup> Institute of Ecology and Evolution, School of Biological Science, University of Edinburgh, Edinburgh,  
8 EH9 3FL, UK

9 <sup>3</sup> School of Science, Technology and Engineering, University of the Sunshine Coast, Petrie, QLD, 4502,  
10 Australia

11 <sup>4</sup> School of the Environment, The University of Queensland, St Lucia, QLD 4067, Australia

12 <sup>5</sup> Evolution & Ecology Research Centre, Centre for Ecosystem Science, and School of Biological, Earth  
13 and Environmental Sciences, University of New South Wales, Sydney, NSW, 2052, Australia

14 \*Correspondance: [barbara.a.class@gmail.com](mailto:barbara.a.class@gmail.com)

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16

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33

34 **Abstract**

35 Social environments impose a number of constraints on individuals' behaviour. These constraints have  
36 been hypothesized to generate behavioural variation among individuals, social responsiveness, and  
37 within-individual behavioural consistency (also termed 'predictability'). In particular, the social niche  
38 specialization hypothesis posits that higher levels of competition associated with higher population  
39 density should increase among-individual behavioural variation and individual predictability, as a way  
40 to reduce conflicts. Being predictable should hence have fitness benefits in group-living animals.  
41 However, to date, empirical studies on fitness consequences of behavioural predictability remain scarce.  
42 In this study, we investigated the associations between social behaviour, its predictability and fitness in  
43 the eastern water dragon (*Intellagama lesueurii*), a wild gregarious lizard. Since this species is sexually  
44 dimorphic, we examined these patterns both between sexes and among individuals. Although females  
45 were more sociable than males, there was no evidence for sex-differences in among-individual variation  
46 or predictability. However, females exhibited positive associations between social behaviour, its  
47 predictability and survival while males only exhibited a positive association between mean social  
48 behaviour and fitness. These findings hence partly support predictions from the social niche  
49 specialization hypothesis and suggest that the function of social predictability may be sex-dependent.

50

## 51 **Introduction**

52 For many organisms, group living implies repeatedly interacting with the same social partners over time,  
53 which often requires dealing with conflict. For sociality to evolve, individuals must therefore adapt ways  
54 to minimise conflict and to maximise received fitness benefits. Behaviourally, individuals can do this in  
55 multiple ways. First, they can partition their behavioural (and/or ecological) niche (i.e., adopt different  
56 roles) such that conflicts do not arise as frequently (Bergmüller & Taborsky, 2010). Second, they may  
57 learn to respond to a partner's behaviour by altering their own behaviour (i.e., be socially responsive)  
58 (Johnstone, 2001). Third, they may become "predictable" in their behaviour (i.e., behave consistently),  
59 allowing conflict to be avoided by allowing social partners to learn an optimal behavioural response  
60 (Wolf et al., 2011). These principles could explain why we often observe dominance hierarchies, spatio-  
61 temporal segregation or cooperation in group-living animals. The implication, therefore, is that for social  
62 evolution to occur, there needs to be both heritable variation in, and selection for social behaviours,  
63 social responsiveness and behavioural predictability within populations. Such variation is expected to  
64 be maintained by natural selection, via frequency-dependent mechanisms (Wolf & McNamara, 2013).

65 Heritable social behaviours should manifest as consistent behavioural differences among individuals. In  
66 the past two decades, animal personality research has uncovered a variety of ways in which individuals  
67 consistently differ in their behaviour. Importantly, consistent differences may not only concern average  
68 levels of repeatedly expressed behaviours (e.g., some individuals are bolder than others, Bell et al.,  
69 2009), but also the way individuals plastically adjust their behaviour to environmental conditions (e.g.,  
70 some individuals are more plastic than others, Dingemanse et al., 2010) and their behavioural  
71 predictability (e.g., some individuals behave more consistently than others, Stamps et al., 2012).

72 Although less studied than among-individual variation in behaviours and plasticity, among-individual  
73 variation in predictability has been reported in an increasing number of species (see Mitchell et al.,  
74 2021), has been shown to be determined by additive genetic effects (Martin et al., 2017; Prentice et al.,  
75 2020) and may have fitness benefits. For instance, behaving unpredictably has been suggested to  
76 decrease predation risks (Briffa, 2013; Humphries & Driver, 1970; Richardson et al., 2018), and shown  
77 to increase fitness when foraging under unpredictable environmental conditions (Patrick et al., 2021).

78 In contrast, and as mentioned earlier, higher predictability should be favoured in social contexts that  
79 involve repeated interactions (Johnstone, 2001; Wolf et al., 2011). Individual variation in social  
80 behaviour and social responsiveness has been widely studied (Aplin et al., 2015; Chervet et al., 2011;  
81 Watson et al., 2018) and social behaviour has been shown to correlate with fitness (Brent et al., 2013;  
82 Delmé et al., 2023; Kohn, 2017; Silk, 2007). At the same time, social behaviours and responsiveness  
83 have also been shown to be heritable (Bailey & Desjonquères, 2022; Godoy et al., 2022; Lea et al.,  
84 2010). In contrast, predictability has seldom been studied in social traits and the association between  
85 social predictability and fitness has, to our knowledge, never been reported.

86 In the present study, we take advantage of a unique behavioural dataset collected in wild eastern water  
87 dragons (*Intellagama lesueurii*) to estimate individual predictability in social behaviour and investigate  
88 its correlation with individual mean social behaviour and fitness. Here, we focus on social distance, a  
89 proxy for social tolerance, which refers to non-aggressive physical proximity between individuals and  
90 represents a fundamental first step for the evolution of affiliative behaviour (Smith & Ivins, 1984). The  
91 eastern water dragon is a territorial and long-lived agamid lizard, which exhibits long-term social  
92 associations and avoidances (Strickland et al., 2017). In this sexually dimorphic species, sexes differ in  
93 their spatial and social behaviour. In particular, females are generally more sociable than males and  
94 occur in smaller territories which they largely share with other individuals (Baird et al., 2012). Previous  
95 research has also shown that dragons vary consistently in social behaviours (e.g., degree, social  
96 tendency, choosiness, centrality, density, Delmé et al., 2023; Strickland & Frère, 2018) and in their  
97 social plasticity (e.g., how they adjust their social tendency to local density, Strickland & Frère, 2019).

98 In particular, we test predictions from the social niche specialization hypothesis (Bergmüller &  
99 Taborsky, 2010) between and within sexes. This hypothesis, which is derived from ecological niche  
100 theory, posits that individuals within a population should adopt different behavioural niches to reduce  
101 the costs of intraspecific competition, which drives among-individual variation. Under higher densities  
102 or when competition intensifies, populations are hence expected to exhibit more among-individual  
103 variation and individuals to be more predictable. In dragons, individuals of both sexes compete for  
104 resources (e.g., food, shelter and basking sites) and the costs of such competition increase with density

105 (Baxter-Gilbert & Whiting, 2019). Males show alternative mating strategies associated with body size,  
106 territory exclusiveness, and agonistic displays (i.e., territorial to satellite, Baird et al., 2012; Piza-Roca  
107 et al., 2020 but see Ball et al., 2023) due competition for access to females. This additional source of  
108 intra-sexual competition in males may drive a higher among-individual variation and a higher  
109 predictability in their social behaviour compared to females. Alternatively, male mating strategies may  
110 be unrelated to social tolerance, and we could find the opposite pattern between sexes. Indeed, females  
111 are more often found in the vicinity of each other (Baird et al., 2012; Strickland & Frère, 2019), which  
112 may drive a higher level of competition for resources. The latter would be supported by a recent finding  
113 that females are slightly more repeatable than males in several social traits (i.e. social tendency and  
114 association strength in same-sex interactions and density Delmé et al., 2023).

115 Predictions from the social niche specialization hypothesis could also be extended to the individual level  
116 as social environments often vary within populations. That is, we would expect more socially tolerant  
117 individuals, which consistently occur in denser environments (be it by choice or not, see Webber &  
118 Vander Wal, 2018), to be more predictable than individuals that are less socially tolerant and occur in  
119 sparser environments. If predictability is an adaptation to reduce the costs of social conflicts, we would  
120 expect it to be positively associated with fitness. Alternatively, unpredictability could be favoured if it  
121 reflects individuals being more opportunistic and better at appropriately adjusting their behaviour to  
122 every situation (Briffa & Lane, 2017). Using double-hierarchical generalized linear models and >30000  
123 observations made over 4 years in >300 individuals, we provide rare estimates of predictability in social  
124 behaviour and its association with mean social behaviour and fitness in the wild.

## 125 **Methods**

### 126 *Study species*

127 The eastern water dragon is a long-lived (lifespan of 28–40 years, Harlow and Harlow 1997, Griffiths  
128 2006) agamid lizard exhibiting male-biased sexual dimorphism (Thompson, 1993), polyandry (Frère et  
129 al., 2015) and alternative mating strategies (Baird et al., 2012). In this species, individuals hold  
130 territories, which vary in size and exclusiveness between females, territorial males and non-territorial

131 males (Baird et al., 2012; Gardiner et al., 2014). While females' territories are smaller and largely  
132 overlap with each other, male territories generally overlap with females' and either exclude (for  
133 territorial males) or overlap with other male's (non-territorial males) (Baird et al., 2012). In this species,  
134 spatial proximity between individuals has been used as a proxy for affiliative interactions. Previous  
135 research has established that individuals remaining within 1.85m of each other without exhibiting  
136 agonistic behaviours (e.g., head bobbing, arm waiving, tail slapping) express social tolerance towards  
137 each other (Strickland et al., 2014). However, this distance represents a small percentile of all social  
138 distances and individuals can adjust their social behaviour depending on their location within their entire  
139 home range (which are larger than 4200m<sup>2</sup> for males and 2700m<sup>2</sup> for females, Piza-Roca et al., 2018).  
140 Therefore, variation in social tolerance likely encompasses a larger range of distances. Previous research  
141 in this population has also shown that dragons form non-random associations with other individuals  
142 (e.g., as preferences and avoidances, see Strickland et al., 2014, 2017) and that sexes differ in their  
143 average levels of sociality. Indeed, females are generally more social than males and form stronger  
144 associations with females and males than males do with each other (Strickland et al., 2014). Sexes differ  
145 in their repeatability, males being more repeatable in social network metrics (Strickland & Frère, 2018)  
146 and females being more repeatable in density (Delmé et al., 2023). Sexes were also found to differ in  
147 their social plasticity (Strickland & Frère, 2019). Indeed, while both sexes increase social tendency and  
148 non-random associations in response to local density, females plastically respond to their social  
149 environments (e.g., sex ratio, number of avoidances) more than males (Strickland et al., 2018; Strickland  
150 & Frère, 2019). Finally, recent research has shown a positive association between reproductive success  
151 and social behaviour (degree and social tendency) towards opposite-sex partners in males, while no such  
152 association was found in females (Delmé et al., 2023).

### 153 *Data collection*

154 Behavioural data was collected as part of an ongoing long-term study (started in 2010) of a wild urban  
155 population of eastern water dragons at the Roma Street Parkland, Brisbane, Australia (27° 27' 46" S,  
156 153° 10' 11" E). This population, which has an average estimated size of 336 individuals, has been  
157 monitored through frequent behavioural surveys (most of them from September to April when

158 individuals are most active). As with many other long-term studies, data was collected by a number of  
159 researchers (including authors of this study) and students and was used in independent research projects.  
160 Behavioural surveys were conducted once or twice a day (am and pm) on average three times a week.  
161 Observers walked along a defined transect of the parkland, which covers 85% of the population  
162 (Strickland et al., 2014), and collected data for all the individuals encountered. Individuals in this  
163 population are recognizable thanks to their unique facial scale and colouration patterns (Gardiner et al.,  
164 2014). Head profile photographs of each individual were taken upon encounter (using a Canon EOS 600  
165 digital camera) to allow their later identification using the I3S Manta software package (Van Tienhoven  
166 et al., 2007). Each individual's behaviour was also recorded (e.g., aggression or mating behaviours),  
167 along with its GPS coordinates (using a Garmin eTrex 10 handheld device, with a 3m resolution). Sex  
168 determination was based on sexual dichromatism and dimorphism present in the species, males being  
169 larger, with wider jaws, and presenting red colouration on their chest (Thompson, 1993). The behavioural  
170 data used in this study was collected between August 2016 and August 2020. During this period,  
171 individuals had on average at least 41 observations per field season. Data was collected under animal  
172 ethics approval through Animal Ethics Committee of the University of the Sunshine Coast (ANA1487)  
173 and Scientific Purposes Permit (WISP17696616).

#### 174 *Response variable*

175 For each observation of each individual, we used the minimum distance (in m) to other individuals  
176 present within the same survey as a proxy for social tolerance such that the higher the distance, the less  
177 socially tolerant an individual was. Importantly, we excluded observations during which individuals  
178 were engaged in agonistic or mating interactions from the analyzed dataset. This measure is hereafter  
179 referred to as "social distance". Previous work on social behaviour in this population focused on the  
180 occurrence of social proximity between individuals, to capture social tolerance or gregariousness  
181 (Strickland et al., 2014). However, the distribution of such sociality measure (binary) did not allow  
182 fitting DHGLMS, which is why we chose a distance variable. Note that social distance was perfectly  
183 correlated ( $r=-1$ ) to this previously used binary index and to the number of individuals present within  
184 9.25 m (the average dragon territory size, see Strickland & Frère, 2019) at the among-individual level



185 (Text S1, Tables S1-S4). Hence, these measures capture the same aspect of an individual's social  
186 environment (Webber & Vander Wal, 2018). Individual variation in social environments can be driven  
187 by individuals establishing their territories in habitats that have varying densities or by seeking or  
188 avoiding the proximity of other individuals. We note that both processes cannot be disentangled because  
189 dragons, which are free to express habitat choice, can disperse to any suitable area of this urban park.

#### 190 *Parentage analyses*

191 To determine reproductive success, we conducted annual sampling between 2010 and 2020. We caught  
192 all possible adult individuals using a lassoing technique and collected tail tip tissue or blood through  
193 caudal venepuncture (Littleford-Colquhoun et al., 2017). We further measured snout-vent length and  
194 captured head profile images to match individuals to behavioural data using the previously described  
195 facial scale identification method (Gardiner et al., 2014). We extracted DNA from blood and tissue using  
196 DNeasy Blood and Tissue extraction kits (Qiagen) as per manufactures instructions. The extracted DNA  
197 was then sequenced at Diversity Arrays Technology, Canberra, using DArTcap™ technology (Feutry et  
198 al., 2020). This resulted in a total of 6,425 SNP loci prior to filtering, across 1285 individuals for which  
199 775 were adults with behavioural survey data. As per Delmé et al., 2023, we restricted homozygote  
200 genotype calls to those with a minimum read depth of 5 to account for false homozygote genotype calls.  
201 Using the dartR package (Mijangos et al., 2022), SNPs were further filtered for an individual call rate  
202 of 80% or greater, a locus call rate of 99% or greater, removal of monomorphic loci, the proportion of  
203 technical replicates assay pairs of at least 99%, and only one locus per short read tag was retained. We  
204 used plink to filter for linkage disequilibrium with locus pairs exhibiting an  $r^2 \geq 0.7$  removed from the  
205 dataset (Chang et al., 2015). To determine the optimal minor allele frequency (MAF) filtering we  
206 conducted a sensitivity analysis by using a range of MAF values (0.3-0.4). Each subsequent dataset was  
207 then used to assign parentage using the R package Sequoia (Huisman, 2017). We selected the MAF  
208 filtering parameters by the highest assignment rate of known mothers and lowest number of  
209 misassignments to known mothers (Jackson et al., 2019) and cross validated these using relatedness  
210 estimates (See Delmé et al., 2023, for which parentage data herein overlaps). The best performing  
211 filtering used a MAF  $\geq 0.43$  producing a resultant data set of 179 high quality loci were then used to

212 assign parentage with the R package, Sequoia (Huisman, 2017). From the 775 adult individuals included  
213 in the parentage assignment, 77 dams and 82 sires were assigned at least one offspring. A total of 535  
214 offspring assignments were made to 239 dams and 296 sires.

#### 215 *Fitness metrics*

216 We defined three fitness metrics: lifetime reproductive success, reproductive success, and survival.  
217 Lifetime reproductive success was the total number of adult offspring produced by individuals during  
218 their entire lifetime. This metric was hence only available for individuals assumed to be dead (not seen  
219 during the 2022-2023 field season). In this rather closed population (Littleford-Colquhoun et al., 2017),  
220 99% of consecutive sightings of individuals occurred within 27 days and 97.6% of multi-season  
221 sightings of individuals occurred in consecutive field seasons. Lifetime reproductive success was  
222 corrected (using fixed effects, see section below) for the year of appearance to account for the fact that  
223 individuals that appeared later during the study (particularly after 2017) were less likely to have  
224 offspring that would be caught as adults (i.e., at least 2-3 years old) before the end of 2020 (end of the  
225 sampling, see above). For similar reasons, a new recruit can be assigned to a parent several years after  
226 its parent died. We tested whether the year of disappearance was associated with a decline in lifetime  
227 reproductive success but found no evidence for such association. Lifetime reproductive success data  
228 was available for 125 females and 134 males.

229 Our second fitness proxy, age-corrected reproductive success, was the total number of adult offspring  
230 of individuals (regardless of whether they were dead or still alive), corrected by the number of years  
231 individuals had been in the population and year of appearance (using fixed effects, see section below).  
232 This proxy therefore measures the average number of offspring an individual has at a given age.  
233 Reproductive success data was available for 177 males and 188 females.

234 Our third fitness proxy, survival, was the number of years an individual had been seen in the population  
235 (data from March 2023) and was corrected for its year of appearance and detectability (the median time  
236 between its successive observations). This metric therefore indicated whether for a given year of

237 appearance and detectability level, an individual lived longer than another. Survival data was available  
238 for 180 males and 196 females and included individuals that were still alive and that were not genotyped.  
239 In this dataset, 10 individuals severely infected with the recently emerged fungal pathogen *Nannizziopsis*  
240 *barbatae* (Peterson et al., 2020) were euthanized between 2018 and 2021. These individuals were  
241 excluded from the lifetime reproductive success and survival analyses. Although a substantial (>30%)  
242 proportion of the population now exhibits symptoms concordant with *N.barbatae* infection, most cases  
243 are relatively mild and a previous study did not detect any impact of the disease on dragons' social  
244 behaviour (Tacey et al., 2023). These individuals therefore remained in the social behaviour analyses.

#### 245 *Statistical analyses*

246 The recent introduction of double-hierarchical generalized linear models (DHGLMs, Cleasby et al.,  
247 2015) in ecological research and the availability of statistical tools (e.g., brms, Bürkner, 2017) have  
248 allowed the investigation of increasingly complex aspects of behavioural variation such as the  
249 relationship between average levels of behaviours and their predictability (Hertel et al., 2020). Here, we  
250 performed our statistical analyses in R (R Core Team, 2022) using the package brms (Bürkner, 2017).  
251 We fitted all models in males and females separately. A DHGLM is a mixed model which includes a  
252 mean model and a dispersion model. While the mean model of a DHGLM allows estimating how fixed  
253 and random effects impact the response (here social distance), the dispersion model fits residuals from  
254 the former as a function of the same or different fixed and random effects to test how these effects impact  
255 deviations in the response relative to the population mean. Among-individual variation in the mean  
256 model hence captures individual differences in the average value of the response while among-  
257 individual variation in the dispersion model captures individual differences in the average value of  
258 deviations (i.e., predictability). Importantly, these deviations include random “noise” but also plastic  
259 responses to unknown external factors (Stamps et al., 2012; Westneat et al., 2015). The higher these  
260 deviations, also termed intra-individual variation (IIV) are, the more unpredictable or plastic an  
261 individual is.

262 Fixed effects in both models included: i) a measure of body size (snout-vent length, measured during  
263 annual captures, Littleford-Colquhoun et al., 2017) to account for individual differences due to age  
264 and/or dominance (Baird et al., 2012; Piza-Roca et al., 2020), ii) quadratic seasonal effects (day since  
265 the 1<sup>st</sup> of August), iii) field season (4 years, factorial) and iv) time of the day (morning vs. afternoon  
266 surveys). Random effects in both models included individual identity (ID) to estimate long-term  
267 individual variation in intercepts, a within-season individual component (a combination of ID and Field  
268 Season, Araya-Ajoy et al., 2015) to account for consistent individual variation driven by short-term  
269 environmental factors, and date of observation. The parameters of interest for this study were therefore  
270 the long-term individual variance in observed behaviour (mean model), in deviations (dispersion model)  
271 and their correlation. A similar model was run for the total population dataset, and included sex as a  
272 fixed effect to test for sex differences in mean sociability and predictability.

273 In all models, social distance was fitted using a lognormal distribution. However, the distribution of  
274 social distances included zeros (8% of the records), most of which were probably not true zeros (two  
275 individuals touching) but distances lower than the detection threshold (the minimum non-zero distance  
276 recorded), here 1.64 m, which is roughly two dragon lengths. To be able to fit a lognormal distribution  
277 on our model, we therefore, needed to add a constant to these zeros, the value of which was determined  
278 using simulations (Text S2). These simulations showed that a constant equal to 80% of our detection  
279 threshold combined with left-censoring allowed retrieving unbiased parameter estimates (Figures S1-  
280 S2). We note that zeros represent a small proportion of our data and that our detection threshold of  
281 1.64m is lower than the distance previously established for social tolerance (1.85m) in this study system  
282 (Strickland et al., 2014). Therefore, most variation in social behaviour in this study species is expected  
283 to occur at a larger scale.

284 All models used default uninformative priors, 4 chains and thinning intervals of 4. Univariate models  
285 were run for 5000 iterations (1000 warmups). More informative priors were also fitted and yielded  
286 similar results (Table S5). All models had an effective sample size above 1000 and model performance  
287 was assessed based on R-hat values being under 1.01, and visualizing trace plots and posterior predictive  
288 plots. Model fit was also assessed by calculating  $R^2$ , which was 0.19 for males and 0.21 for females.

289 Raw-scale estimates of individual repeatability ( $R_p$ ) and coefficients of variation (CV, the ratio of  
290 standard deviation on mean) were computed for each sex. These two metrics, which are variance-  
291 standardized for the former and mean-standardized for the latter, facilitate between-studies comparisons  
292 and meta-analyses. Their calculation followed the procedures described in O’Dea et al., 2022, which  
293 involve computing the variance of fixed effects, and back-transforming variance estimates to the original  
294 data scale for the mean and dispersion models. Finally, to assess sex-bias in variability we computed the  
295 natural logarithm of the ratio of male and female CVs (lnCVR, see Nakagawa et al., 2015; Zajitschek et  
296 al., 2020), positive values indicating a male bias.

297 To investigate how social distance and its predictability were associated with fitness, our three fitness  
298 measures were fitted as a function of individual-specific values derived from the model (best linear  
299 unbiased predictors, BLUPs). More specifically, each fitness measure was either fitted as a function of  
300 individual values in mean social distance or individual values for IIV in social distance. We modelled  
301 reproductive success responses with zero-inflated Poisson distributions (using the `pscl` package,  
302 Jackman, 2010; Zeileis et al., 2008), and survival with a Cox proportional-hazards model (using the R  
303 package `survival`, Therneau, 2024). These models included additional fixed effects such as year of  
304 appearance, number of years in the population (for reproductive success) and detectability (for survival).  
305 Each model was run for each individual posterior value (i.e., 4000 times) and we stored each coefficients  
306 to later calculate their median and 95% credible intervals. Caution has been raised against the use of  
307 BLUPs in subsequent analyses when such procedure does not consider uncertainty around each BLUP,  
308 which can result in spurious results (Hadfield et al., 2010). However, we here used their entire posterior  
309 distribution and checked using simulations (Text S3, Figure S3) that this approach generated little to no  
310 bias for estimating selection when fitness has a zero-inflated Poisson distribution and the phenotype is  
311 moderately repeatable. Individual-level correlations between traits and fitness have been shown to  
312 estimate selection more accurately than do phenotypic correlations (Dingemanse et al., 2021). While  
313 multi-response models would have been preferable for these analyses, we could not satisfactorily  
314 implement a multivariate DHGLM for traits that do not have the same number of repeats (fitness being  
315 measured once), and different distributions (number of offspring being zero-inflated) in the package

316 brms. This prompted our decision to use a 2-step approach, after carefully verifying that this approach  
317 would yield unbiased results.

## 318 **Results**

### 319 *Summary statistics*

320 The entire dataset included 30475 records from 189 males and 197 females. Individuals were observed  
321 on average 41 times per field season. Approximately 46% of all individuals were recorded for more than  
322 one field season. Regarding sex-specific trait distributions, males seemed to be less social than females.  
323 Indeed, the raw social distance was on average 7.1 m (sd=6.8) in males and 5.9 m (sd = 6.4) in females.  
324 As a result, females experienced higher conspecific densities than males ( $6.5 \times 10^{-3}$  [ $6.1 \times 10^{-3}$ ;  $6.9 \times 10^{-3}$ ]  
325 and  $5.6 \times 10^{-3}$  [ $5.2 \times 10^{-3}$ ;  $6.0 \times 10^{-3}$ ] individuals per m<sup>2</sup>, respectively, Text S4). Social distance also seemed  
326 to be slightly more variable phenotypically in females than in males (lnCVR = -0.11). Regarding fitness  
327 proxies, the reproductive success of individuals known to be dead was similar between sexes (1.01 for  
328 males and 1.04 for females), although males exhibited more variability than females (sd=2.55 for males  
329 and 1.63 for females, lnCVR =0.48). In the entire population, males and females also had on average  
330 the same number of offspring (0.87 for males and 0.85 for females), and males also exhibited more  
331 variability than females (sd=2.34 for males and 1.46 for females, lnCVR = 0.44). In contrast, males had  
332 an average survival of 3.39 years (sd=2.38), which was lower than females' survival (mean =3.80,  
333 sd=2.47) but variability in survival was quite similar between the sexes (lnCVR = 0.08).

### 334 *Individual variation in mean social distance*

335 Neither body size nor time of the day influenced social distance, which increased within field seasons  
336 in females and decreased across years in both sexes (Table 1). Females were hence less socially tolerant  
337 towards the end of a field season but both sexes were increasingly tolerant across years. The model  
338 pooling both sexes confirmed that males and females statistically differed in their mean social distance,  
339 males having higher values (indicative of lower social tolerance) than females (0.16 [0.07; 0.25], Table  
340 2, Figure 1). In both sexes, we detected long-term (across years) among-individual variation in mean  
341 social distance (sd = 0.28 [0.24; 0.33] for males and 0.34 [0.29; 0.40] for females). The point estimate

342 of among-individual variance was larger in females than in males although evidence for this female bias  
343 was weak ( $\ln\text{CVR} = -0.19 [-0.48; 0.09]$ ). Long-term individual effects were greater contributors to  
344 variation in social distance than date ( $\text{sd}=0.12 [0.10; 0.14]$  for both) and within-season individual effects  
345 ( $\text{sd} = 0.17 [0.13; 0.20]$  for males and  $0.20 [0.17; 0.24]$  for females). While social distance had a rather  
346 low long-term repeatability in both sexes, females tended to be slightly more repeatable than males ( $R$   
347  $= 0.18 [0.14; 0.24]$  for females and  $0.13 [0.09; 0.17]$  for males, Table 1).

#### 348 *Individual variation in predictability of social distance*

349 Similar to mean social distance, predictability in social distance was not associated with body size or  
350 age but individuals were increasingly predictable (decrease in IIV) within and between seasons (Table  
351 1). The model pooling both sexes did not show any statistical difference between sexes in average  
352 predictability (difference= $-0.01 [-0.04; 0.03]$ , Table 2, Figure 1). Males and females both exhibited long-  
353 term among-individual variation in predictability ( $\text{sd}=0.08 [0.05; 0.11]$  in males and  $0.07 [0.04; 0.09]$   
354 in females). This long-term individual component was a slightly greater contributor to variation in  
355 predictability than date ( $\text{sd}=0.02 [0.00; 0.05]$  for males and  $0.07 [0.05; 0.09]$  for females) and short-term  
356 individual effects ( $\text{sd}=0.06 [0.01; 0.09]$  for males and  $0.04 [0.00; 0.07]$  for females). There was however  
357 no evidence that among-individual variation in predictability differed between sexes (Table 1,  $\ln\text{CVR} =$   
358  $0.23 [-0.34; 0.74]$ ) and the repeatability of this intra-individual component (i.e., its relative importance  
359 compared to the total phenotypic variance for this trait) was very low (1%) in both sexes. This is because  
360 the dispersion model explained a very small proportion (3%) of the residual variance from the mean  
361 model (Table S6). As a result, most of the variance in social distance was left unassigned (76% for  
362 males, 68% for females). Finally, we found within females that individuals with higher social distances  
363 were less predictable (higher intra-individual variation, or IIV) as evidenced by a strong positive  
364 correlation between mean social distance and IIV ( $0.79 [0.49; 0.98]$ ) while males did not exhibit any  
365 association between social distance and IIV ( $0.00 [-0.29; 0.31]$ , Table 1, Figure 1).

366

367

368 **Table 1.** Estimates (median and 95% credible intervals) from double hierarchical mixed models fitting  
369 social distance in males and females separately. Raw-scale repeatability (R) and coefficient of variation  
370 (CV) were calculated for the individual component in both the mean and the dispersion parts of the  
371 model. The model directly estimated the correlation between mean and intra-individual variance in  
372 social distance (Cor mean-dispersion). In the random effects, the Individual\_Season component  
373 estimates within-year among-individual variation. Fixed effects for which the 95% CI did not include  
374 zero are printed in bold.

		Males	Females
<b>Fixed effects</b>			
Mean model	Intercept	1.78 [ 1.69; 1.86]	1.65 [ 1.56; 1.75]
	Body size	0.01 [-0.02; 0.05]	0.02 [-0.02; 0.06]
	Time of day (pm)	-0.01 [-0.04; 0.02]	-0.02 [-0.05; 0.01]
	Season day	0.00 [-0.02; 0.02]	<b>0.05 [ 0.03; 0.07]</b>
	Season day <sup>2</sup>	0.02 [ 0.00; 0.04]	<b>-0.03 [-0.05;-0.01]</b>
	Season 2017-2018	<b>-0.10 [-0.19;-0.01]</b>	<b>-0.12 [-0.21;-0.02]</b>
	Season 2018-2019	<b>-0.14 [-0.23;-0.05]</b>	<b>-0.15 [-0.24;-0.05]</b>
	Season 2019-2020	<b>-0.21 [-0.32;-0.10]</b>	<b>-0.23 [-0.34;-0.11]</b>
Dispersion model	Intercept	-0.14 [-0.18;-0.10]	-0.17 [-0.21;-0.13]
	Body size	0.02 [ 0.00; 0.03]	0.01 [-0.01; 0.03]
	Time of day (pm)	-0.02 [-0.04; 0.01]	0.02 [-0.01; 0.04]
	Season day	<b>-0.03 [-0.04;-0.01]</b>	<b>-0.03 [-0.04;-0.01]</b>
	Season day <sup>2</sup>	<b>0.04 [ 0.03; 0.06]</b>	<b>0.04 [ 0.02; 0.05]</b>
	Season 2017-2018	<b>-0.12 [-0.16;-0.08]</b>	<b>-0.11 [-0.15;-0.06]</b>
	Season 2018-2019	<b>-0.14 [-0.18;-0.09]</b>	<b>-0.12 [-0.16;-0.07]</b>
	Season 2019-2020	<b>-0.17 [-0.22;-0.12]</b>	<b>-0.18 [-0.23;-0.12]</b>
<b>Random effects (sd)</b>			
Mean model	Date	0.12 [0.10; 0.14]	0.12 [0.10; 0.14]
	Individual	0.28 [0.24; 0.33]	0.34 [0.29; 0.40]
	Individual_Season	0.17 [0.13; 0.20]	0.20 [0.17; 0.24]
	R	0.13 [0.09; 0.17]	0.18 [0.14; 0.24]
	CV	0.35 [0.28; 0.42]	0.42 [0.34; 0.51]
Dispersion model	Date	0.02 [0.00; 0.05]	0.07 [0.05; 0.09]
	Individual	0.08 [0.05; 0.11]	0.07 [0.04; 0.09]
	Individual_Season	0.06 [0.01; 0.09]	0.04 [0.00; 0.07]
	Cor mean-dispersion	0.00 [-0.29; 0.31]	0.79 [0.49; 0.98]
	R	0.01 [0.00;0.02]	0.01 [0.00;0.01]
	CV	0.51 [0.32; 0.68]	0.41 [0.27; 0.55]

375



376 **Table 2.** Estimates (median and 95% credible intervals) from double hierarchical mixed models fitting  
377 social distance in both sexes. Sex was fitted as a fixed effect in the mean and dispersion models to test  
378 for sex-differences in social distance and its predictability. The model directly estimated the correlation  
379 between mean and intra-individual variance in social distance (Cor mean-dispersion). In the random  
380 effects, the ‘Individual\_Season’ component estimates within-year among-individual variation. Fixed  
381 effects for which the 95% CI did not include zero are printed in bold.

		Estimates
<b>Fixed effects</b>		
Mean model	Intercept	1.65 [ 1.57; 1.74]
	Body size	0.02 [-0.02; 0.05]
	Time of day (pm)	-0.02 [-0.03; 0.00]
	Sex (male)	<b>0.12 [ 0.03; 0.21]</b>
	Season day	<b>0.05 [ 0.03; 0.07]</b>
	Season day <sup>2</sup>	-0.02 [-0.05; 0.00]
	<b>Sex (male) : Season day</b>	<b>-0.05 [-0.07;-0.03]</b>
	<b>Sex (male) : Season day<sup>2</sup></b>	<b>0.05 [ 0.02; 0.04]</b>
	Season 2017-2018	<b>-0.10 [-0.17;-0.03]</b>
	Season 2018-2019	<b>-0.14 [-0.21;-0.07]</b>
Season 2019-2020	<b>-0.22 [-0.30;-0.13]</b>	
Dispersion model	Intercept	-0.15 [-0.19;-0.11]
	Body size	0.02 [ 0.00; 0.04]
	Time of day (pm)	0.00 [-0.02; 0.02]
	Sex (male)	0.00 [-0.04; 0.03]
	Season day	<b>-0.03 [-0.04;-0.02]</b>
	Season day <sup>2</sup>	<b>0.04 [ 0.03; 0.05]</b>
	Season 2017-2018	<b>-0.12 [-0.15;-0.08]</b>
	Season 2018-2019	<b>-0.12 [-0.16;-0.09]</b>
Season 2019-2020	<b>-0.18 [-0.22;-0.14]</b>	
<b>Random effects (sd)</b>		
Mean model	Date	0.13 [ 0.11; 0.14]
	Individual	0.31 [ 0.28; 0.35]
	Individual_Season	0.18 [ 0.16; 0.21]
Dispersion model	Date	0.07 [ 0.05; 0.08]
	Individual	0.07 [ 0.05; 0.09]
	Individual_Season	0.05 [ 0.02; 0.08]
	Cor mean-dispersion	0.38 [ 0.15; 0.61]

382

383

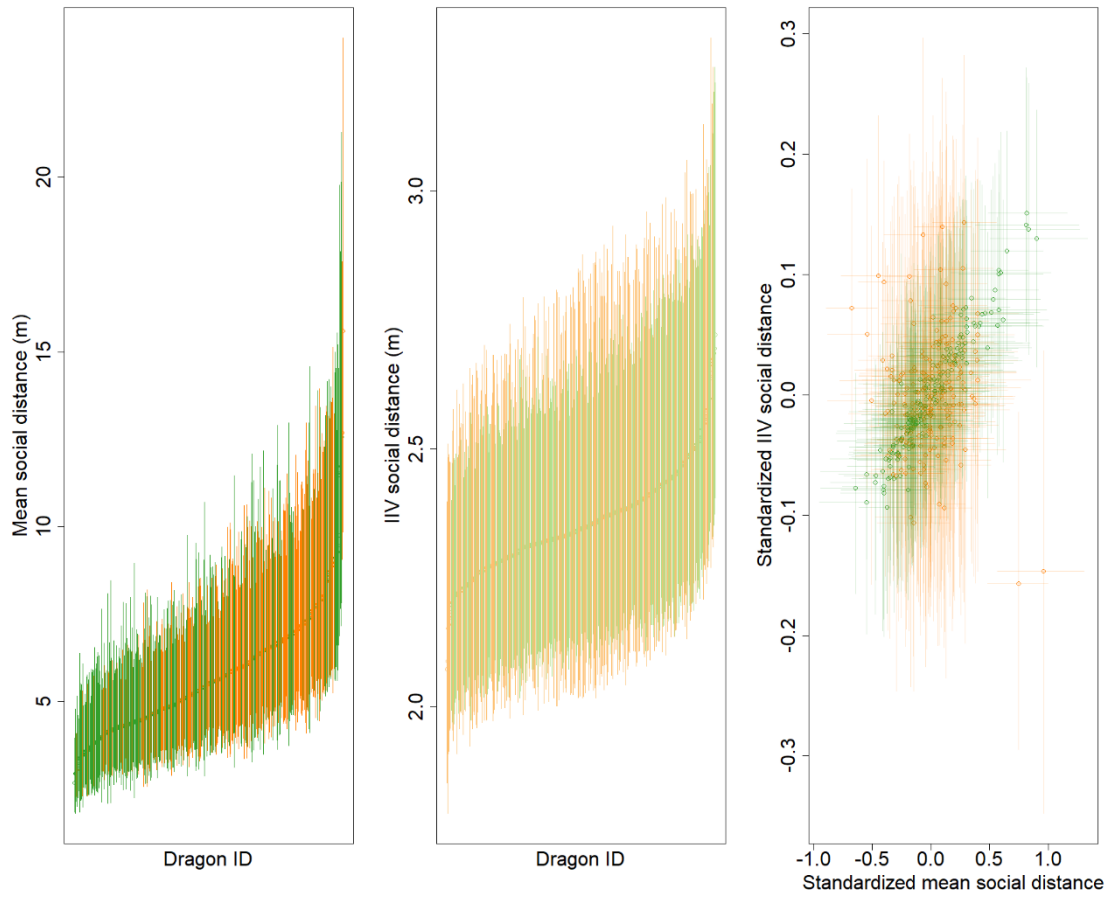
384 *Links with fitness*

385 We found evidence for sex-specific associations between social distance and our fitness proxies. First,  
386 lifetime reproductive success was negatively associated with mean social distance in males but not in  
387 females (Figure 2, 3). Males that were less socially tolerant produced fewer offspring in their lifetime (-  
388 0.13 [-0.29; 0.00]) and had a higher probability of siring no offspring (0.27 [0.00; 0.54], Figure 2). As a  
389 result, more social males were predicted to have 1.3 offspring and less social males to have 0.2 offspring  
390 on average in their lifetime (Figure 3). In contrast, there was weak evidence for unpredictable males to  
391 have more offspring in their lifetime (0.27 [-0.01; 0.55]) and no evidence for a relationship between  
392 predictability and the probability of siring no offspring (0.20 [-0.22; 0.69], Figure 2). As a result, the  
393 overall relationship between IIV and lifetime reproductive success in males (0.2 offspring for  
394 predictable and 0.9 offspring for unpredictable males) did not differ from zero (Figure 3). This  
395 relationship was of similar in magnitude in females (0.3 offspring for predictable and 1.1 offspring for  
396 unpredictable females, Figure 3).

397 Second, the relationship between age-corrected reproductive success and mean social distance was also  
398 negative in males, although the magnitude and level of evidence of this relationship were both weaker  
399 (Figure 2, 3). Indeed, there was no evidence that less socially tolerant males produced fewer offspring  
400 at a given age (-0.09 [-0.29; 0.07]), but these males had higher probabilities of siring no offspring (0.34  
401 [0.08; 0.60], Figure 2). As a result, more social males were predicted to have 1.1 offspring and less  
402 social males to have 0.7 offspring on average at age 6 (Figure 3). In contrast, we found no evidence that  
403 reproductive success was associated with mean social tolerance in females and social predictability in  
404 both sexes (Figure 2, 3).

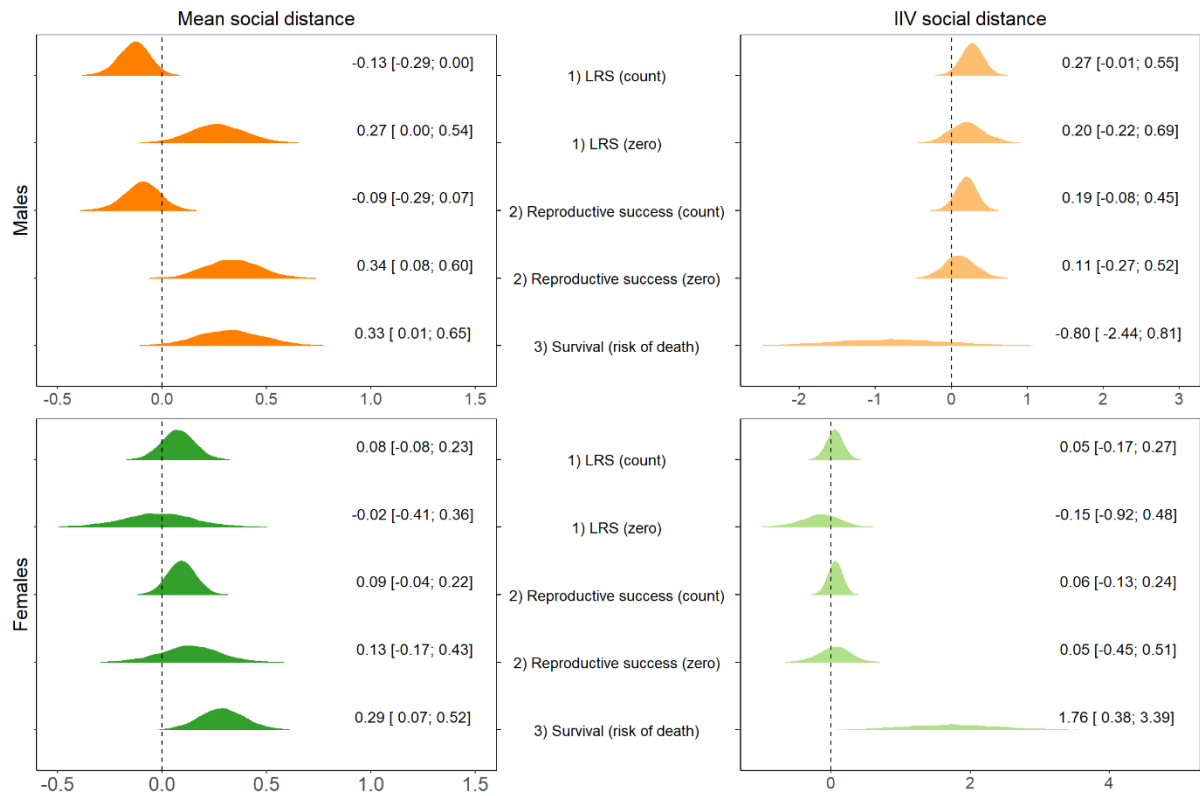
405 Third, we found that less socially tolerant individuals of both sexes had shorter lifespans (Figure 2, 4).  
406 Indeed, the effect of social distance on the risk of death was positive and similar in magnitude in both  
407 sexes (0.33 [0.01; 0.65] in males, 0.29 [0.07; 0.52] in females, Figure 2). As a result, more socially  
408 tolerant individuals lived on average longer (Q1, Figure 4; 3.93 [3.84; 4.02] for males, 4.53 [4.45; 4.61]  
409 for females) than less socially tolerant individuals (Q3, Figure 4; 3.75 [3.65; 3.84] for males, 4.34 [4.26;  
410 4.40] for females). In contrast, we found the association between predictability and survival to be

411 positive in females (1.76 [0.38; 3.39]) and null in males (-0.80 [-2.44; 0.81], Figure 2). As a result, more  
412 predictable females lived longer on average (Q1, Figure 4; 4.55 [4.45; 4.66]) than less predictable  
413 females (Q3, Figure 4; 4.32 [4.22; 4.40]) while the mean survival of males did not differ between  
414 predictable (3.77 [3.63; 3.91]) and unpredictable males (3.91 [3.75; 4.07]).



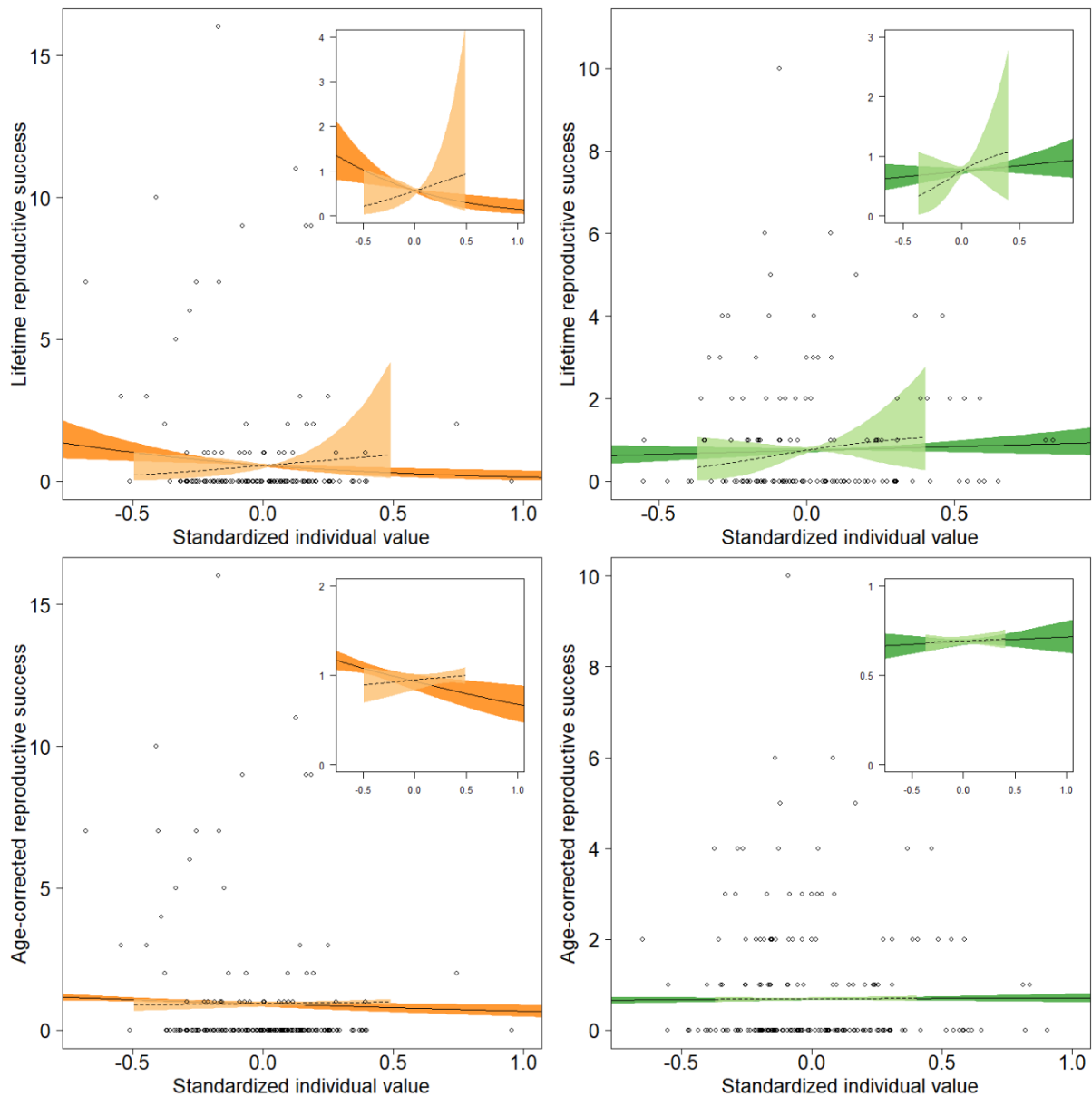
415

416 **Figure 1:** Posteriors of mean social distance, intra-individual variation (IIV) in social distance, and their  
 417 relationship in 386 eastern water dragons (197 females in green, 189 males in orange). Each point  
 418 corresponds to a posterior median and each vertical or horizontal segment denotes 95% credible  
 419 intervals. In the left and middle panels, estimates were scaled back to the raw data scale (distance  
 420 measured in meters units).



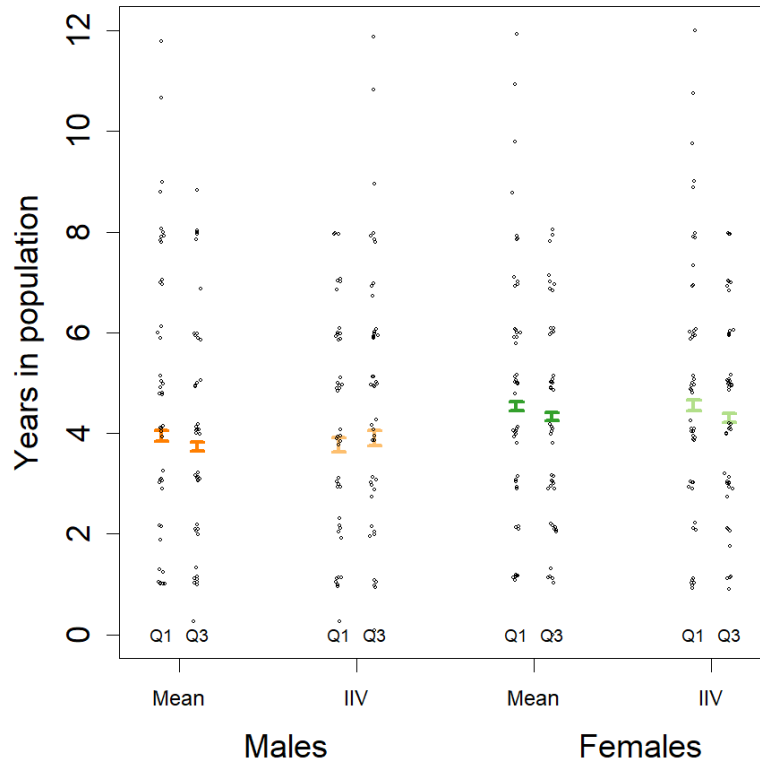
421

422 **Figure 2:** Full posterior distributions of the effects of individual social distance (mean: left, dark shades,  
 423 intra-individual variation (IIV): right, light shades) on fitness measures in males (top, orange) and  
 424 females (bottom, green). Coefficients were obtained by fitting each fitness proxy as a function of  
 425 individual posteriors. For each fitness component (1: lifetime reproductive success (here LRS), 2: age-  
 426 corrected reproductive success (here reproductive success), 3: Survival), we detail whether positive  
 427 coefficients increase offspring number (count), the probability of having no offspring (zero) or decrease  
 428 survival (increased risk of death). Posterior medians and 95% credible intervals are printed next to each  
 429 distribution.



430

431 **Figure 3:** Predicted lifetime reproductive success (top) and age-corrected reproductive success at a  
 432 given age (here, 6 years, bottom) as a function of individual values for mean (dark shades) and intra-  
 433 individual variation (light shades) in social distance in males (orange, left) and females (green, right).  
 434 The black lines and shaded areas depict the median estimates and 95% credible intervals of the  
 435 distribution of coefficients derived from zero-inflated Poisson models fitting each individual posterior.  
 436 A close-up view of each of these relationships is printed in the top right corner of each plot. Raw fitness  
 437 measures as a function of standardized mean social distance are depicted by open circles.



438

439 **Figure 4:** Predicted mean survival (95% credible interval, CI) in males (orange) and females (green) for  
 440 the first and third quartiles (Q1 and Q3, respectively) of mean social distance (dark shades) and intra-  
 441 individual variation (IIV) in social distance (light shades). These estimates were obtained by fitting the  
 442 number of years individuals were seen in the population as a function of individual posteriors in a Cox  
 443 proportional-hazards model. Observed survival when individuals' values are  $\leq Q1$  and  $\geq Q3$  are depicted  
 444 by open circles.

445

## 446 **Discussion**

447 This study investigated for the first time the links between social behaviour, its predictability and fitness.  
448 We demonstrated, in eastern water dragons, that i) females were more social than males but both sexes  
449 exhibited similar among-individual variation and predictability in their social behaviour; ii) individuals  
450 varied in their average social behaviour and predictability; iii) more social females were also more  
451 predictable in their social behaviour; iv) more social males had a higher fitness (survival and  
452 reproductive success) while more social and predictable females lived longer lives. These findings  
453 provide partial support for the social niche specialization hypothesis, which posits that higher density  
454 environments should favour niche partitioning and behavioural predictability. We detail below how our  
455 results align or differ from predictions of the social niche specialization hypothesis and discuss their  
456 broader implications.

### 457 *Sex differences in means, variance and predictability*

458 Sexual dimorphism arises because sexes are subject to different selective pressures and can manifest not  
459 only in the average value of a trait, but also in the variability of the trait among and within individuals  
460 (Poissant et al., 2010; Zajitschek et al., 2020). In this study, we found that female dragons are on average  
461 more social than males, which is in line with previous research (Strickland & Frère, 2019). However,  
462 evidence for a female bias in among-individual variation was weak and there was no evidence for sex-  
463 differences in predictability. These findings therefore did not clearly support the social niche  
464 specialization hypothesis, which predicted that the more social sex, here females, would be more  
465 variable and more predictable.

466 Various mechanisms have been proposed to explain sex-differences in trait variance (Zajitschek et al.,  
467 2020). For instance, the ‘greater male variability hypothesis’, predicts condition-dependent sexual traits  
468 to vary more among males than among females (Cuervo & Møller, 1999, 2001; Pomiankowski &  
469 Møller, 1997) as a result of sexual selection, and was mainly supported for morphological traits. In this  
470 population, males indeed appeared to be more variable in their reproductive success than females, which  
471 likely results from intra-sexual competition for mating opportunities (Bateman, 1948; Wade, 1979). In



472 contrast, there was no evidence for social behaviour to be more variable in males, probably because our  
473 studied trait was not associated with body size, an important contributor to reproductive success in males  
474 (Piza-Roca et al., 2020).

475 Alternatively, the ‘estrus-mediated variability hypothesis’ predicts that females are more variable than  
476 males due to experiencing different stages of their reproductive cycle, and is predominantly supported  
477 by evidence using physiological traits (Zajitschek et al., 2020). In eastern water dragons, reproduction  
478 is seasonal and mainly occurs during the first half of the field season (i.e., between September and  
479 December, Thompson, 1993). Whether female dragons vary in their seasonal hormonal changes and  
480 which hormones influence dragons’ social behaviour remain unknown. Females likely experience  
481 higher levels of sexual harassment and higher competition over resources during the reproductive  
482 period, which may explain their seasonal change in social distance. However, previous research did not  
483 find females to vary in their seasonal social plasticity (Strickland & Frère, 2019).

484 Finally, sex-differences in predictability opposite to predictions of the social niche specialization  
485 hypothesis could have been the result of sexual selection. For instance, females could have preferred  
486 males that were more behaviourally predictable, while unpredictability in females could have helped  
487 them avoid sexual coercion from males. However, we did not find evidence for sex-differences in  
488 predictability and our estimates of sex-specific associations between predictability and fitness were in  
489 opposite direction to this prediction.

#### 490 *Mean-predictability relationships*

491 Previous studies have reported associations between mean behaviour (e.g., boldness, aggression,  
492 sociability, movement) and its predictability in other vertebrates (Hertel et al., 2020, 2021; Jolles et al.,  
493 2019; Mitchell et al., 2016; O’Dea et al., 2022). Our findings align with one of these studies (O’Dea et  
494 al., 2022) in that more social females were also more predictable, although one should keep in mind that  
495 relationships between mean behaviours and their predictability likely vary between populations and  
496 traits (Mitchell et al., 2021). Here, we showed that this correlation could also differ between sexes and  
497 we propose an explanation below.

498 First of all, the magnitude of the mean-predictability correlation found in females was higher than all  
499 other mean-predictability correlations ever reported. One may argue that such a mean-variance  
500 relationship can mathematically arise with distance data as its distribution is likely non-normal (bound  
501 to zero and right-skewed). However, simulations showed that our model accurately estimated the mean-  
502 predictability correlation for a trait that had a similar distribution as the studied trait (Text S2, Figure  
503 S1). More importantly, this bias would not explain the strikingly different mean-predictability  
504 correlations between sexes. Therefore, these results are probably better explained by the eastern water  
505 dragon's social structure.

506 While the positive mean-predictability correlation found in females aligns with predictions from the  
507 social niche specialization hypothesis, the absence of correlation in males does not necessarily rule it  
508 out and may provide interesting nuance. An intuitive mechanism underlying the social niche  
509 specialization hypothesis is the social enforcement of predictability via density. Previous research in this  
510 population indeed showed that density not only increases interaction opportunities but also the number  
511 of non-random associations in both sexes (Strickland et al., 2018). By definition, non-random  
512 associations imply a certain level of social predictability, and because females experience higher  
513 densities than males (Text S4), predictability would thus be more enforced in females than in males.  
514 Whether this explains the magnitude of the sex-difference in mean-predictability correlations is however  
515 unclear. One may speculate that female-female interactions, which are the most frequent in this species  
516 may be the main driver of social niches. In contrast, the benefits of predictability in female-male or  
517 male-male interactions may vary (but see below). Recent work in this population (Delmé et al., 2023)  
518 showed a higher degree and social tendency towards the opposite sex to be positively associated with  
519 reproductive success in males. Future research investigating the sex-specific links between social  
520 predictability and fitness would provide interesting new insights on when the social niche specialization  
521 hypothesis may apply.

522 Despite the strong mean-predictability relationship found in females, individual variation in mean and  
523 predictability represented small to very small proportions of the total variance in social behaviour. This  
524 means that most of the phenotypic variation, which consisted of intra-individual variation, was due to

525 factors that other than intrinsic properties of individuals. Such variance partitioning reflects the high  
526 lability of our studied trait. Indeed, social distance depends on the respective locations of an individual  
527 and its social environment, which are influenced by variation in resources and risks. Such environmental  
528 variation may be particularly unpredictable in this highly frequented urban park and may have hence  
529 been the main contributor of intra-individual variation in social distance. This explanation contrasts with  
530 behaviours being generally more repeatable in the field than in the lab (Bell et al., 2009), although the  
531 influence of habitat on the repeatability of intra-individual variation has not been studied yet.

### 532 *Association with fitness*

533 While, the fitness benefits of social behaviours have been investigated and reported in several species  
534 (Brent et al., 2013; Frère et al., 2010; Kohn, 2017; Silk, 2007) including this study system (Delmé et al.,  
535 2023), our study goes one step further by investigating the relationships between different fitness  
536 components and social predictability. This study is also one of the rare studies connecting behavioural  
537 predictability and fitness in the wild (Cain et al., 2023; Patrick et al., 2021).

538 For males, we found that more social individuals lived longer lives and produced more offspring, which  
539 resulted in a higher lifetime reproductive success and suggested that mean social tolerance may be under  
540 directional selection. This positive relationship between males' social behaviour and reproductive  
541 success was consistent with Delmé et al. (2023) and was likely driven by increased mating opportunities.  
542 In dragons, males mainly interact with females (Baird et al., 2012; Strickland et al., 2014) and shorter  
543 social distances may hence signal stronger social bonds with potential mates. As directional selection is  
544 expected to deplete additive genetic variance (Fisher, 1930), which contributes to among-individual  
545 variance, we would have expected this sex-specific selective regime to drive a lower among-individual  
546 variance in males than females. Evidence for this was, however, weak. In contrast, the association  
547 between predictability and fitness in males was rather mixed, which may explain its zero correlation  
548 with mean social behaviour. Indeed, predictability was not associated with survival or the probability to  
549 have offspring but unpredictable individuals that had offspring had more of them. Given the species'  
550 biology, we can speculate that reproductive success and survival in males mainly result from interactions  
551 with females and males, respectively. While being unpredictable did not seem to provide benefits in

552 male-male contests (contrary to predictions by Briffa & Lane, 2017), which is known to be particularly  
553 costly (Baxter-Gilbert & Whiting, 2019), it may increase mating success (e.g., via sexual coercion) in  
554 males that have a greater access to females, namely territorial males (Baird et al., 2012). On the other  
555 hand, the null correlation between predictability and survival could signal that social predictability and  
556 male mating strategies are independent.

557 For females, we found no evidence that social behaviour and its predictability were associated with  
558 reproductive success which was also consistent with Delmé et al. (2023) but contrasted with numerous  
559 studies in mammals (Silk, 2007). Because breeding systems may affect the adaptive value of females'  
560 social behaviour (Silk, 2007), our findings might be explained by the absence of parental care in dragons,  
561 which precludes any form of cooperation over offspring rearing. While the positive association between  
562 mean social behaviour and survival was similar to that in males, we also found such positive association  
563 for social predictability in females. These results aligned with predictions from the social niche  
564 specialization hypothesis and may have been driven by correlational selection on mean and  
565 predictability of social behaviour. Alternatively, the very high mean-predictability correlation may have  
566 driven their positive association with survival. Interestingly, the survival advantage of being more social  
567 did not translate into a higher lifetime reproductive success in females. This could be because social  
568 behaviour was only associated with one fitness component in females (as opposed to both fitness  
569 components in males) and because females exhibited less variance in lifetime reproductive success than  
570 males.

571 In both sexes, we found a positive association between social behaviour and survival which could be  
572 due to social tolerance decreasing the costs of competition (e.g., fights for resources or territories,  
573 Haunhorst et al., 2017), sexual harassment of females by males (Fox, 2002), or increasing social  
574 information about the location of food sources (Aplin et al., 2012) or predators (Beauchamp, 2010).  
575 Individuals that are more socially tolerant (and females that are more predictable) may also occupy a  
576 different position in the social network, which has repeatedly been found to correlate with longevity  
577 (Barocas et al., 2011; Holt-Lunstad et al., 2010; Silk et al., 2010; Vander Wal et al., 2015). Alternatively,  
578 a higher social tolerance could be found in dominant individuals, which pay low costs of social

579 proximity to other subordinate individuals. However, here, social behaviour was not associated with  
580 individual size, a predictor of dominance in this species (Piza-Roca et al., 2020).

581 Altogether, our findings that social behaviour, its predictability and survival were all positively  
582 correlated in females provided partial support for the social niche specialization hypothesis in eastern  
583 water dragons. This study also provided new insights into sex-specific functions of social behaviour and  
584 suggested that in dragons, not only mean social behaviour but also its predictability have different fitness  
585 consequences between sexes. Assuming that mean and predictability in social behaviour are heritable  
586 and genetically correlated across sexes (Connallon & Clark, 2014; Kaufmann et al., 2023), one could  
587 speculate sex-specific selection to be a mechanism maintaining among-individual variation in these  
588 traits (Schuett et al., 2010). We are hence yet to determine whether social behaviour and its predictability  
589 are heritable and genetically correlated between sexes in eastern water dragons.

## 590 **Conclusion**

591 Our study demonstrated that social behaviour and its predictability covary with each other and with  
592 fitness in a sex-specific way in a wild lizard. While sexes did not clearly differ in their among-individual  
593 variation or predictability, females exhibited a strong positive association between social behaviour, its  
594 predictability, and survival. In males, we found mean social behaviour to be positively associated with  
595 fitness but not with predictability, which did not have clear fitness benefits. These findings hence partly  
596 supported predictions from the social niche specialization hypothesis in that social enforcement of  
597 behavioural predictability may only have occurred in the more social sex. We also provided rare  
598 empirical estimates of behavioural predictability and of its association with fitness in a wild animal. This  
599 study therefore highlights the evolutionary relevance of behavioural predictability and aims to promote  
600 further research on the links between social behaviour, animal personality and predictability.

## 601 **Data and Code accessibility**

602 Data and R code used for this study are available on OSF (<https://osf.io/3y6s7/>)

## 603 **Authors Contributions**

604 BC,CF, SN conceived the original idea of the study. BC performed the statistical analyses and wrote the  
605 manuscript, with input from all other authors. NJ performed the parentage analyses. BC,CF,NJ,KS  
606 contributed to data collection. All authors provided comments on the manuscript and agreed on the final  
607 version to be submitted for publication.

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## 616 **Conflict of interest disclosure**

617 Authors declare no conflict of interest

## 618 **References**

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