

Sex-specific associations between social behaviour, its predictability and fitness in a wild lizard

Authors: Barbara Class^{1*}, Kasha Strickland², Dominique Potvin³, Nicola Jackson⁴, Shinichi Nakagawa⁵, Celine Frere⁴

¹ Department of Biology, Ludwig-Maximilians-Universität München, Großhaderner Straße 2, 82152 Planegg-Martinsried, Germany

² Institute of Ecology and Evolution, School of Biological Science, University of Edinburgh, Edinburgh, EH9 3FL, UK

³ School of Science, Technology and Engineering, University of the Sunshine Coast, Petrie, QLD, 4502, Australia

⁴ School of the Environment, The University of Queensland, St Lucia, QLD 4067, Australia

⁵ Evolution & Ecology Research Centre, Centre for Ecosystem Science, and School of Biological, Earth and Environmental Sciences, University of New South Wales, Sydney, NSW, 2052, Australia

*Correspondance: class@biologie.uni-muenchen.de

Keywords: social behaviour, predictability, fitness, agamid, wild population, sexual dimorphism

1 **Abstract**

2 Social environments impose a number of constraints on individuals' behaviour. These constraints have
3 been hypothesized to generate behavioural variation among individuals, social responsiveness, and
4 within-individual behavioural consistency (also termed 'predictability'). In particular, the social niche
5 hypothesis posits that higher levels of competition associated with higher population density should
6 increase among-individual behavioural variation and individual predictability, as a way to reduce
7 conflicts. Being predictable should hence have fitness benefits in group-living animals. However, to
8 date, empirical studies on fitness consequences of behavioural predictability remain scarce. In this study,
9 we investigated the associations between predictability of social behaviour and fitness in the eastern
10 water dragon (*Intellagama lesueurii*), a wild gregarious lizard. Since this species is sexually dimorphic,
11 we examined these patterns both between sexes and among individuals. Females, which were more
12 sociable than males, exhibited slightly higher levels of among-individual variation but similar social
13 predictabilities as males. Females that were more sociable were also more predictable while no
14 association was found between social behaviour and predictability in males. Finally, social behaviour
15 and its predictability showed sex-specific associations with survival and reproductive success. These
16 findings partly support predictions from the social niche hypothesis and highlight the importance of
17 behavioural predictability in social evolution.

18

19 **Introduction**

20 For many organisms, group living implies repeatedly interacting with the same social partners over time,
21 which often requires dealing with conflict. For sociality to evolve, individuals must therefore adapt ways
22 to minimise conflict in order to maximise received fitness benefits. Behaviourally, individuals can do this
23 in multiple ways. First, they can partition their behavioural (and/or ecological) niche (i.e. adopt different
24 roles) such that conflicts do not arise as frequently (Bergmüller & Taborsky, 2010). Second, they may
25 learn to respond to a partner's behaviour by altering their own behaviour (i.e., be socially responsive)
26 (Johnstone, 2001). Third, they may become “predictable” in their behaviour (i.e. behave consistently),
27 allowing conflict to be avoided by allowing social partners to learn an optimal behavioural response
28 (Wolf et al., 2011). These principles could explain why we often observe dominance hierarchies, spatio-
29 temporal segregation or cooperation in group-living animals. The implication, therefore, is that for social
30 evolution to occur, there needs to be both heritable variation in, and selection for social behaviours,
31 social responsiveness and behavioural predictability within populations. Such variation is expected to
32 be maintained by natural selection, via frequency-dependent mechanisms (Wolf & McNamara, 2013).

33 Heritable social behaviours should manifest as consistent behavioural differences among individuals. In
34 the past two decades, animal personality research has uncovered a variety of ways in which individuals
35 consistently differ in their behaviour. Importantly, consistent differences may not only concern average
36 levels of repeatedly expressed behaviours (e.g. some individuals are bolder than others, Bell et al., 2009),
37 but also the way individuals plastically adjust their behaviour to environmental conditions (e.g. some
38 individuals are more plastic than others, Dingemanse et al., 2010) and their behavioural predictability
39 (e.g. some individuals behave more consistently than others, Stamps et al., 2012). Although less studied
40 than among-individual variation in behaviours and plasticity, among-individual variation in
41 predictability has been reported in an increasing number of species (see Mitchell et al., 2021), has been
42 shown to be determined by additive genetic effects (Martin et al., 2017; Prentice et al., 2020) and may
43 have fitness benefits. For instance, behaving unpredictably has been suggested to decrease predation
44 risks (Briffa, 2013; Humphries & Driver, 1970; Richardson et al., 2018), and shown to increase fitness
45 when foraging under unpredictable environmental conditions (Patrick et al., 2021). In contrast, and as

46 mentioned earlier, higher predictability should be favoured in social contexts that involve repeated
47 interactions (Johnstone, 2001; Wolf et al., 2011). Individual variation in social behaviour and social
48 responsiveness has been widely studied (Aplin et al., 2015; Chervet et al., 2011; Watson et al., 2018)
49 and, in some species, social behaviour has been shown to correlate with fitness (Brent et al., 2013; Delmé
50 et al., 2023; Kohn, 2017; Silk, 2007). At the same time, social behaviours and responsiveness have also
51 been shown to be heritable both in the wild and in the lab (Bailey & Desjonquères, 2022; Godoy et al.,
52 2022; Lea et al., 2010). In contrast, predictability has seldom been studied in social traits and the
53 association between social traits' predictability and fitness has, to our knowledge, never been reported.
54 In the present study, we take advantage of a unique behavioural dataset collected in wild eastern water
55 dragons (*Intellagama lesueurii*) to estimate individual predictability in social behaviour and investigate
56 its correlation with individual mean social behaviour and fitness. Here, we focus on social tolerance,
57 which refers to non-aggressive physical proximity between individuals and represents a fundamental
58 first step for the evolution of affiliative behaviour (Smith & Ivins, 1984). The eastern water dragon is a
59 gregarious and long-lived agamid lizard that exhibits long-term social associations and avoidances
60 (Strickland et al., 2017). Females are polyandrous and males exhibit mating strategies varying from
61 territorial to satellite behaviours (Baird et al., 2012; Ball et al., 2023; Frère et al., 2015). In this sexually
62 dimorphic species, sexes differ in the repeatability and plasticity of their social behaviours (Strickland
63 & Frère, 2018, 2019) and plastic responses in social behaviour consistently vary between individuals
64 (Strickland et al., 2021).

65 In particular, we test predictions from the social niche hypothesis (Bergmüller & Taborsky, 2010)
66 between and within sexes. This hypothesis, which is derived from ecological niche theory, posits that
67 individuals within a population should adopt different behavioural niches to reduce the costs of
68 intraspecific competition, which drives among-individual variation. Under higher densities or when
69 competition intensifies, populations are hence expected to exhibit more among-individual variation and
70 individuals to be more predictable. In dragons, males show various mating strategies due to a strong
71 intra-sexual competition, which may drive a higher among-individual variation and a higher
72 predictability in their social behaviour compared to females. Alternatively, and because mating

73 strategies may be unrelated to social tolerance, we could also find the opposite pattern as females occur
74 at higher densities than males (Strickland & Frère, 2019) and may compete more for resources. The
75 latter would be supported by a recent study suggesting that females are more repeatable than males in a
76 range of social behaviours (Delmé et al., 2023). Predictions from the social niche hypothesis could also
77 be extended to the individual level as social environments often vary within populations. That is, we
78 would expect more socially tolerant individuals, which consistently occur in denser environments (be it
79 by choice or not, see Webber & Vander Wal, 2018), to be more predictable than individuals that are less
80 socially tolerant and occur in sparser environments. If predictability is an adaptation to reduce the costs
81 of social conflicts, we would expect it to be positively associated with fitness. Alternatively,
82 unpredictability could be favoured if it reflects individuals being more opportunistic and better at
83 appropriately adjusting their behaviour to every situation (Briffa & Lane, 2017).

84 The recent introduction of double-hierarchical generalized linear models (DHGLMs, Cleasby et al.,
85 2015) in ecological research and the availability of statistical tools (e.g. brms, Bürkner, 2017) have
86 allowed the investigation of increasingly complex aspects of behavioural variation such as the
87 relationship between average levels of behaviours and their predictability (Hertel et al., 2020). This
88 approach notably requires substantial sample sizes and amounts of repeated measures, which remain
89 rare in behavioural ecology studies (Mitchell et al., 2021). Using DHGLMs and >30000 observations
90 made over 4 years in >300 individuals, we provide rare estimates of predictability in social behaviour
91 and its association with fitness in the wild.

92 **Methods**

93 *Data collection*

94 Behavioural data was collected as part of an ongoing long-term study (started in 2010) of a wild
95 urban population of Eastern water dragons at Roma Street Parkland, Brisbane, Australia (27° 27' 46" S,
96 153° 10' 11" E). This population, which has an average estimated size of 336 individuals, has been
97 monitored through frequent behavioural surveys (most of them from September to April when
98 individuals are most active). Behavioural surveys were conducted by researchers, once or twice a day

99 (am and pm) on average three times a week. Observers walked along a defined transect of the parkland,
100 which covers 85% of the population (Strickland et al., 2014), and collected data for all the individuals
101 encountered. Individuals in this population are recognizable thanks to their unique facial scale and
102 colouration patterns (Gardiner et al., 2014). Head profile photographs of each individual were taken
103 upon encounter (using a Canon EOS 600 digital camera) to allow their later identification using the I3S
104 Manta software package (Van Tienhoven et al., 2007). Each individual's behaviour was also recorded
105 (e.g. aggression or mating behaviours), along with its GPS coordinates (using a Garmin eTrex 10
106 handheld device). Sex was determined based on sexual dichromatism and dimorphism present in the
107 species, males being larger, with wider jaws, and presenting red colouration on their chest (Thompson,
108 1993). The behavioural data used in this study was collected between August 2016 and August 2020.
109 During this period, individuals had on average at least 41 observations per field season. Data was
110 collected under animal ethics approval through Animal Ethics Committee of the University of the
111 Sunshine Coast (ANA1487) and Scientific Purposes Permit (WISP17696616).

112 *Response variable*

113 For each observation of each individual, the minimum distance (in m) to other individuals present within
114 the same survey was used as a proxy for social tolerance such that the higher the distance, the less
115 socially tolerant an individual was. Importantly, observations during which individuals were engaged in
116 agonistic or mating interactions were excluded from the analyzed dataset. This measure is hereafter
117 referred to as "Social distance". Previous work on social behaviour in this population focused on the
118 occurrence of social proximity between individuals, to capture social tolerance or gregariousness
119 (Strickland et al., 2014). However, the distribution of such sociality measure (binary) did not allow
120 fitting DHGLMS, which is why we chose a distance variable. Note that social distance was perfectly
121 correlated ($r=-1$) to this previously used binary index and to the number of individuals present within
122 9.25 m (the average dragon territory size, see Strickland & Frère, 2019) at the among-individual level
123 (Text S1, Tables S1-S4). Hence, these measures capture the same aspect of an individual's social
124 environment (Webber & Vander Wal, 2018). Individual variation in social environments can be driven
125 by individuals establishing their territories in habitats that have varying densities or by seeking or

126 avoiding the proximity of other individuals. We note that both processes cannot be disentangled because
127 dragons, which are free to express habitat choice, can disperse to any suitable area of this urban park.

128 *Parentage analyses*

129 To determine reproductive success, additional annual sampling was conducted to complement
130 behavioural survey data. Annual sampling data used in this study were collected between 2010 and
131 2020. This involved catching all possible adult individuals using a lassoing technique and collecting tail
132 tip tissue or blood through caudal venepuncture (Littleford-Colquhoun et al., 2017). We further
133 measured snout-vent length and captured head profile images to match individuals to behavioural data
134 using described facial scale method identification method. DNA was extracted from blood and tissue
135 using DNeasy Blood and Tissue extraction kits (Qiagen) as per manufactures instructions and sequenced
136 using DArTcap™ technology through Diversity Arrays Technology, Canberra (Feutry et al., 2020). This
137 resulted in a total of 6,425 SNPs prior to filtering, across 1285 individuals for which 775 were adults
138 with behavioural survey data. As per Delmé et al., 2023, we filtered homozygote SNPs for a read depth
139 of 5, and using the *dartR* package SNPs were further filtered for an individual call rate of 80% or greater,
140 a SNP call rate of 99% or greater, removal of monomorphic loci, the proportion of technical replicates
141 assay pairs of at least 99%, only one SNP per short read tag was retained. We used *plink* to filter for
142 linkage disequilibrium with SNP pairs exhibiting an $r^2 \geq 0.7$ removed from the dataset (Chang et al.,
143 2015). To determine the optimal minor allele frequency (MAF) filtering we conducted a sensitivity
144 analysis by using a range of MAF values (0.3-0.4). Each subsequent SNP dataset was then used to assign
145 parentage using the R package Sequoia (Huisman, 2017). We selected the MAF filtering parameters by
146 the highest assignment rate of known mothers and lowest number of misassignments to known mothers
147 (Jackson et al., 2019) and cross validated these using relatedness estimates (See Delmé et al., 2023, for
148 which parentage data herein overlaps). The best performing filtering used a MAF ≥ 0.43 producing a
149 resultant data set of 179 high quality SNPs were then used to assign parentage with the R package,
150 Sequoia (Huisman, 2017). From the 775 adult individuals included in the parentage assignment, 77 dams
151 and 82 sires were assigned at least one offspring. A total of 535 offspring assignments were made to
152 239 dams and 296 sires.

154 We defined three fitness metrics: lifetime reproductive success, survival, and reproductive rate. Lifetime
155 reproductive success (LRS) was the total number of adult offspring of individuals that were assumed to
156 be dead (not seen during the 2022-2023 field season). In this rather closed population (Littleford-
157 Colquhoun et al., 2017), 95% of individuals known to be alive had consecutive sightings that were
158 lower than 160 days and 99% were resighted within 327 days. LRS was corrected for the year of
159 appearance (hereafter “cohort”) to account for the fact that individuals that appeared later during the
160 study (particularly after 2017) were less likely to have offspring that would be caught as adults (i.e. at
161 least 2-3 years old) before the end of 2020 (end of the sampling, see above). For similar reasons, a new
162 recruit can be assigned to a parent several years after its parent died. We tested whether the year of
163 disappearance was associated with a decline in LRS. In particular, we expected that individuals last seen
164 in field seasons 2020-2021 and 2021-2022 had lower LRS than individuals last seen earlier during this
165 study because their offspring may not have been sampled yet. However, this did not seem to be the case
166 as no negative association was found between year of disappearance and LRS. Lifetime reproductive
167 success data was available for 125 females and 134 males.

168 Our second fitness proxy, reproductive rate, was the total number of adult offspring of individuals
169 (regardless of whether they were dead or still alive), corrected by the number of years individuals had
170 been in the population. Reproductive rate data was available for 177 males and 188 females.

171 Our third fitness proxy, survival, was the number of years an individual had been seen in the population
172 (data from March 2023) and was corrected for its cohort and detectability (the median time between its
173 successive measurements). This metric therefore indicated whether for a given year of appearance and
174 detectability level, an individual lived longer than another. Survival data was available for 180 males
175 and 196 females and included individuals that were still alive and that were not genotyped.

176 In this dataset, 10 individuals severely infected with the recently emerged fungal pathogen *N. barbatae*
177 (Peterson et al., 2020) were euthanized between 2018 and 2021. These individuals were excluded from
178 the LRS and survival analyses. Although a substantial (>30%) proportion of the population now exhibits

179 symptoms concordant with *N.barbatae* infection, most cases are relatively mild and a previous study
180 did not detect any impact of the disease on dragons' social behaviour (Tacey et al., 2023). These
181 individuals therefore remained in the social behaviour analyses.

182 *Statistical analyses*

183 Statistical analyses were performed in R (R Core Team, 2022) using the package 'brms' (Bürkner,
184 2017). We fitted all models in males and females separately. DHGLM is a mixed model which includes
185 a mean model and a dispersion model. While the mean model of a DHGLM allows estimating how fixed
186 and random effects impact the response (here social distance), the dispersion model fits residuals from
187 the former as a function of the same or different fixed and random effects to test how these effects impact
188 deviations in the response relative to the population mean. Among-individual variation in the mean
189 model hence captures individual differences in the average value of the response while among-
190 individual variation in the dispersion model captures individual differences in the average value of
191 deviations (i.e. predictability). Importantly, these deviations include random "noise" but also plastic
192 responses to unknown external factors. The higher these deviations, also termed intra-individual
193 variation (IIV) are, the more unpredictable or plastic an individual is.

194 Fixed effects in both models included: i) a measure of body size (snout-vent length, measured during
195 annual captures, Littleford-Colquhoun et al., 2017) to account for individual differences due to age
196 and/or dominance, ii) quadratic seasonal effects (day since the 1st of August), iii) field season (4 years,
197 factorial) and iv) time of the day (morning vs. afternoon surveys). Random effects in both models
198 included individual identity (ID) to estimate long-term individual variation in intercepts, a within-season
199 individual component (a combination of ID and Field Season, Araya-Ajoy et al., 2015) to account for
200 consistent individual variation driven by short-term environmental factors, and date of observation. The
201 parameters of interest for this study were therefore the long-term individual variance in observed
202 behaviour (mean model), in deviations (dispersion model) and their correlation. A similar model was
203 run for the total population dataset, and included sex as a fixed effect to test for sex differences in mean
204 sociability and predictability.

205 In all models, social distance was fitted using a lognormal distribution. However, the distribution of
206 social distances included zeros (8% of the records), most of which were probably not true zeros (two
207 individuals touching) but distances lower than the detection threshold (the minimum non-zero distance
208 recorded), here 1.64 m, which is roughly two dragon lengths. To be able to fit a lognormal distribution
209 on our model, we therefore, needed to add a constant to these zeros, the value of which was determined
210 using simulations (Text S2). These simulations showed that a constant equal to 80% of our detection
211 threshold combined with left-censoring allowed retrieving unbiased parameter estimates (Figures S1-
212 S2).

213 All models used default uninformative priors, 4 chains and thinning intervals of 4. Univariate models
214 were run for 5000 iterations (1000 warmups). More informative priors were also fitted and yielded
215 similar results (Table S5). All models had an effective sample size above 1000 and model performance
216 was assessed based on R-hat values being under 1.01, and visualizing trace plots and posterior predictive
217 plots. Model fit was also assessed by calculating R^2 , which was 0.19 for males and 0.21 for females.
218 Raw-scale estimates of individual repeatability (R_p) and coefficients of variation (CV, the ratio of
219 standard deviation on mean) were computed for each sex. These two metrics, which are variance-
220 standardized for the former and mean-standardized for the latter, facilitate between-studies comparisons
221 and meta-analyses. Their calculation followed the procedures described in O’Dea et al., 2022, which
222 involve computing the variance of fixed effects, and back-transforming variance estimates to the original
223 data scale for the mean and dispersion models. Finally, to assess sex-bias in variability we computed the
224 natural logarithm of the ratio of male and female CVs (lnCVR, see Nakagawa et al., 2015; Zajitschek et
225 al., 2020), positive values indicating a male bias.

226 To investigate how social distance and its predictability were associated with fitness, our three fitness
227 measures were fitted as a function of individual-specific values derived from the model (best linear
228 unbiased predictors, BLUPs). This was done using three generalized linear models fitting reproductive
229 rate and LRS with zero-inflated Poisson distributions (using the pscl package, Jackman, 2010; Zeileis
230 et al., 2008), and survival with a Poisson distribution. Coefficients from each model and each posterior
231 were stored and we calculated their distribution (median and 95% credible intervals). Caution has been

232 raised against the use of BLUPs in subsequent analyses because such procedure does not consider
233 uncertainty around each BLUP, which can result in spurious results (Hadfield et al., 2010). However,
234 we here used their entire posterior distribution and checked using simulations (Text S3, Figure S3) that
235 this approach generated little to no bias for estimating selection when fitness has a zero-inflated Poisson
236 distribution and the phenotype is moderately repeatable.

237 **Results**

238 *Summary statistics*

239 The entire dataset included 30475 records from 189 males and 197 females. Individuals were observed
240 on average 41 times per field season. Approximately 46% of all individuals were recorded for more than
241 one field season. Regarding sex-specific trait distributions, males were on average less social than
242 females. Indeed, social distance was on average 7.1 m (sd=6.8) in males and 5.9 m (sd=6.4) in females.
243 Social distance was overall slightly more variable in females than in males (lnCV= -0.11). Regarding
244 fitness proxies, the reproductive success of individuals known to be dead was similar between sexes
245 (1.01 for males and 1.04 for females), although males exhibited more variability than females (sd=2.55
246 for males and 1.63 for females, lnCV=0.48). In the entire population, males and females also had on
247 average the same number of offspring (0.87 for males and 0.85 for females), and males also exhibited
248 more variability than females (sd=2.34 for males and 1.46 for females, lnCV=0.44). In contrast, males
249 had an average survival of 3.39 years (sd=2.38), which was lower than females' survival (mean =3.80,
250 sd=2.47) but variability in survival was quite similar between the sexes (lnCV=0.08).

251 *Individual variation in mean social distance*

252 Neither body size nor time of the day influenced social distance, which mainly varied temporally within
253 and between field seasons (Table 1). The model pooling both sexes showed that males and females
254 differed in their mean social distance, males having higher values (indicative of lower social tolerance)
255 than females (0.16 [0.07; 0.25], Table 2). In both sexes, we detected long-term (across years) among-
256 individual variation in mean social distance (sd=0.28 [0.24; 0.33] for males and 0.34 [0.29; 0.40] for
257 females). These long-term individual effects were greater contributors to variation in social distance

258 than date (sd=0.12 [0.10; 0.14] for both) and within-season individual effects (sd=0.17 [0.13; 0.20] for
259 males and 0.20 [0.17; 0.24] for females). While social distance had a rather low long-term repeatability
260 in both sexes, females were slightly more repeatable than males (R= 0.18 [0.14; 0.24] for females and
261 0.13 [0.09; 0.17] for males, Table 1). This was possibly due to a higher among-individual variance in
262 females compared to males (lnCVR= -0.19 [-0.48; 0.09]) although this female bias did not statistically
263 differ from zero.

264 *Individual variation in predictability of social distance*

265 Similar to mean social distance, predictability in social distance varied between and within seasons but
266 was not associated with body size or age (Table 1). The model pooling both sexes did not show any
267 statistical difference between sexes in average predictability (difference=-0.01 [-0.04; 0.03], Table 2).
268 Males and females both exhibited long-term among-individual variation in predictability (sd=0.08 [0.05;
269 0.11] in males and 0.07 [0.04; 0.09] in females). This long-term individual component was a slightly
270 greater contributor to variation in predictability than date (sd=0.02 [0.00; 0.05] for males and 0.07 [0.05;
271 0.09] for females) and short-term individual effects (sd=0.06 [0.01; 0.09] for males and 0.04 [0.00; 0.07]
272 for females). Furthermore, males seemed more variable in their predictability than females but such
273 difference did not differ statistically from zero (Table 1, lnCVR=0.23 [-0.34; 0.74]). The repeatability
274 of this intra-individual component (i.e. its relative importance compared to the total phenotypic variance
275 for this trait) was however very low (1%) in both sexes. This is because the dispersion model explained
276 a very small proportion (3%) of the residual variance from the mean model (Table S6). As a result, most
277 of the variance in social distance was left unassigned (76% for males, 68% for females). Finally, we
278 found within females that individuals with higher social distances were less predictable (higher intra-
279 individual variation, or IIV) as evidenced by a strong positive correlation between mean social distance
280 and IIV (0.79 [0.49; 0.98]) while males did not exhibit any association between social distance and IIV
281 (0.00 [-0.29; 0.31], Table 1).

282

283

284 **Table 1.** Estimates (median and 95% credible intervals) from double hierarchical mixed models fitting
 285 social distance in males and females separately. Raw-scale repeatability (R) and coefficient of variation
 286 (CV) were calculated for the individual component in both the mean and the dispersion parts of the
 287 model. Fixed effects for which the 95% CI did not include zero are printed in bold.

		Males	Females
Fixed effects			
Mean model	Intercept	1.78 [1.69; 1.86]	1.65 [1.56; 1.75]
	Body size	0.01 [-0.02; 0.05]	0.02 [-0.02; 0.06]
	Time of day (pm)	-0.01 [-0.04; 0.02]	-0.02 [-0.05; 0.01]
	Season day	0.00 [-0.02; 0.02]	0.05 [0.03; 0.07]
	Season day ²	0.02 [0.00; 0.04]	-0.03 [-0.05;-0.01]
	Season 2017-2018	-0.10 [-0.19;-0.01]	-0.12 [-0.21;-0.02]
	Season 2018-2019	-0.14 [-0.23;-0.05]	-0.15 [-0.24;-0.05]
	Season 2019-2020	-0.21 [-0.32;-0.10]	-0.23 [-0.34;-0.11]
Dispersion model	Intercept	-0.14 [-0.18;-0.10]	-0.17 [-0.21;-0.13]
	Body size	0.02 [0.00; 0.03]	0.01 [-0.01; 0.03]
	Time of day (pm)	-0.02 [-0.04; 0.01]	0.02 [-0.01; 0.04]
	Season day	-0.03 [-0.04;-0.01]	-0.03 [-0.04;-0.01]
	Season day ²	0.04 [0.03; 0.06]	0.04 [0.02; 0.05]
	Season 2017-2018	-0.12 [-0.16;-0.08]	-0.11 [-0.15;-0.06]
	Season 2018-2019	-0.14 [-0.18;-0.09]	-0.12 [-0.16;-0.07]
	Season 2019-2020	-0.17 [-0.22;-0.12]	-0.18 [-0.23;-0.12]
Random effects (sd)			
Mean model	Date	0.12 [0.10; 0.14]	0.12 [0.10; 0.14]
	Individual	0.28 [0.24; 0.33]	0.34 [0.29; 0.40]
	Individual_Season	0.17 [0.13; 0.20]	0.20 [0.17; 0.24]
	R	0.13 [0.09; 0.17]	0.18 [0.14; 0.24]
	CV	0.35 [0.28; 0.42]	0.42 [0.34; 0.51]
Dispersion model	Date	0.02 [0.00; 0.05]	0.07 [0.05; 0.09]
	Individual	0.08 [0.05; 0.11]	0.07 [0.04; 0.09]
	Individual_Season	0.06 [0.01; 0.09]	0.04 [0.00; 0.07]
	Cor mean-dispersion	0.00 [-0.29; 0.31]	0.79 [0.49; 0.98]
	R	0.01 [0.00;0.02]	0.01 [0.00;0.01]
	CV	0.51 [0.32; 0.68]	0.41 [0.27; 0.55]

288

289 **Table 2.** Estimates (median and 95% credible intervals) from double hierarchical mixed models fitting
 290 social distance in both sexes. Sex was fitted as a fixed effect in the mean and dispersion models to test
 291 for sex-differences in social distance and its predictability. Fixed effects for which the 95% CI did not
 292 include zero are printed in bold.

		Estimates
Fixed effects		
Mean model	Intercept	1.64 [1.55; 1.72]
	Body size	0.02 [-0.01; 0.05]
	Time of day (pm)	-0.02 [-0.03; 0.00]
	Sex (male)	0.16 [0.07; 0.25]
	Season day	0.03 [0.01; 0.04]
	Season day ²	0.00 [-0.02; 0.02]
	Season 2017-2018	-0.10 [-0.18;-0.03]
	Season 2018-2019	-0.14 [-0.21;-0.07]
	Season 2019-2020	-0.22 [-0.31;-0.13]
Dispersion model	Intercept	-0.15 [-0.19;-0.11]
	Body size	0.02 [0.00; 0.04]
	Time of day (pm)	0.00 [-0.02; 0.02]
	Sex (male)	-0.01 [-0.04; 0.03]
	Season day	-0.03 [-0.04;-0.02]
	Season day ²	0.04 [0.03; 0.05]
	Season 2017-2018	-0.12 [-0.15;-0.08]
	Season 2018-2019	-0.13 [-0.16;-0.09]
	Season 2019-2020	-0.18 [-0.22;-0.14]
Random effects (sd)		
Mean model	Date	0.13 [0.11; 0.14]
	Individual	0.31 [0.28; 0.35]
	Individual_Season	0.18 [0.16; 0.21]
Dispersion model	Date	0.07 [0.05; 0.08]
	Individual	0.08 [0.06; 0.09]
	Individual_Season	0.05 [0.02; 0.07]
	Cor mean-dispersion	0.38 [0.15; 0.59]

293

294

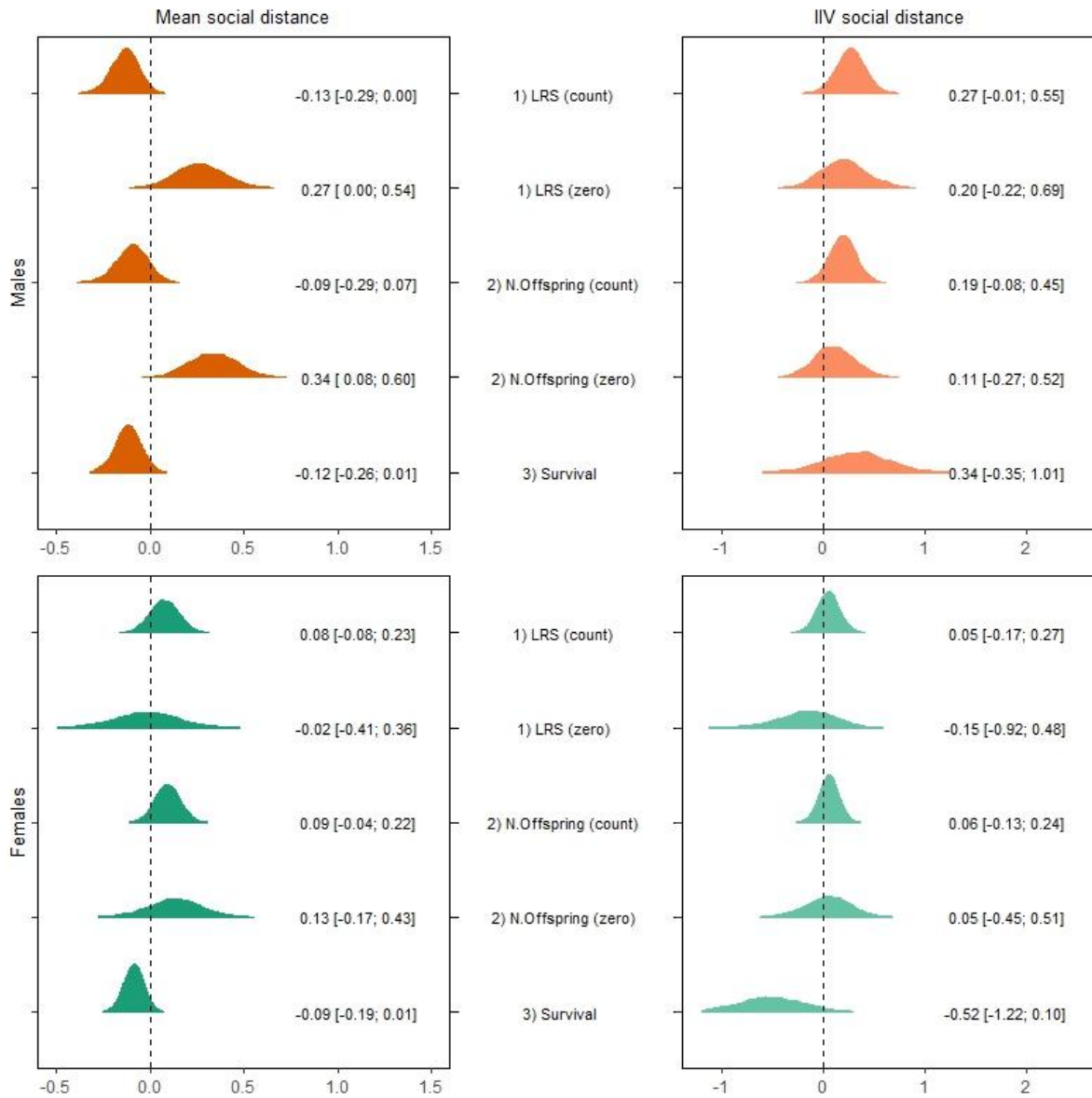
295 *Links with fitness*

296 In both sexes, we found evidence for associations between social distance and our fitness proxies. First,
297 LRS was negatively associated with mean social distance in males (Figure 1, 2). Indeed, males that
298 were more socially distant produced fewer offspring in their lifetime (-0.13 [-0.29;0.00]) and had a
299 higher probability of siring no offspring (0.27 [0.00; 0.54]). In contrast, there was a tendency for
300 unpredictable males to have more offspring in their lifetime (0.27 [-0.01; 0.55]) but no clear relationship
301 between unpredictability and the probability of siring no offspring (0.20 [-0.22; 0.69]). As a result, the
302 relationship between unpredictability and LRS appeared weakly positive in males (Figure 2) and similar
303 in magnitude to relationships found between social behaviour and LRS in females, which did not differ
304 from zero (Figure 1, 2).

305 Second, reproductive rate was negatively associated with mean social distance in males (Figure 1, 3).
306 While there was no evidence that more socially distant males produced fewer offspring at a given age
307 (-0.09 [-0.29; 0.07]), these males had higher probabilities of siring no offspring (0.34 [0.08; 0.60]). In
308 contrast, we found no evidence that reproductive rate was associated with predictability in males, mean
309 social tolerance in females and predictability in females (Figure 1, 3).

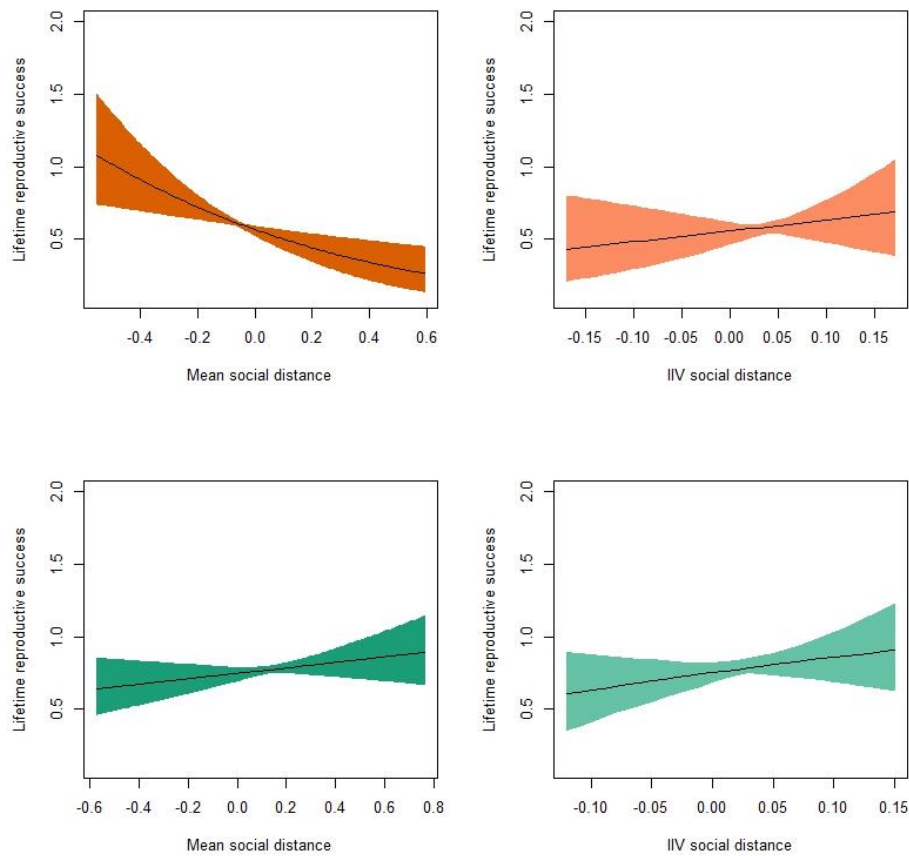
310 Third, we found a tendency for more socially distant individuals to have shorter lifespans in males (-
311 0.12 [-0.26; 0.01]) and females (-0.09 [-0.19; 0.01], Figure 1, 4). In contrast, the association between
312 unpredictability and survival showed different signs between the sexes, being positive in males (0.34 [-
313 0.35; 1.01]), and negative in females (-0.52 [-1.22; 0.10]). These estimates, however, did not statistically
314 differ from zero in either sex.

315



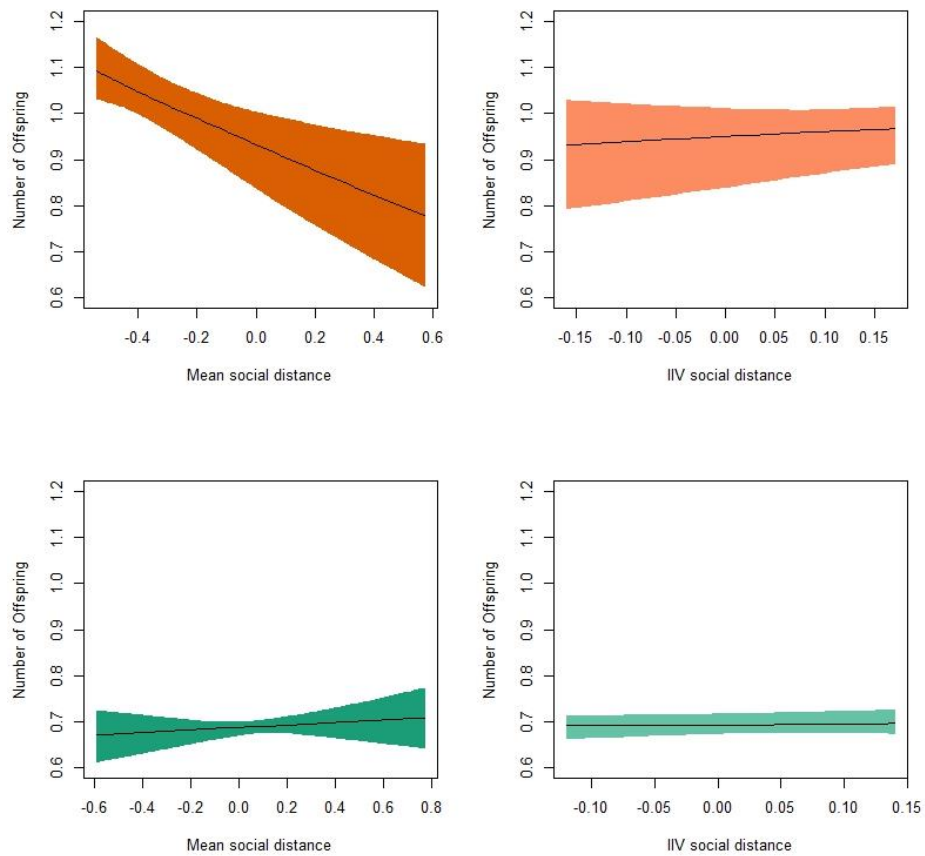
316

317 **Figure 1:** Posterior distributions of the effects of individual social distance (mean on the left, IIV on the
 318 right) on fitness in each sex (males in top row, females in bottom row). These estimates were obtained
 319 by fitting each fitness proxy as a function of posterior distributions of individual values derived from a
 320 DHGLM. Posterior median and 95% credible intervals are printed next to each distribution.



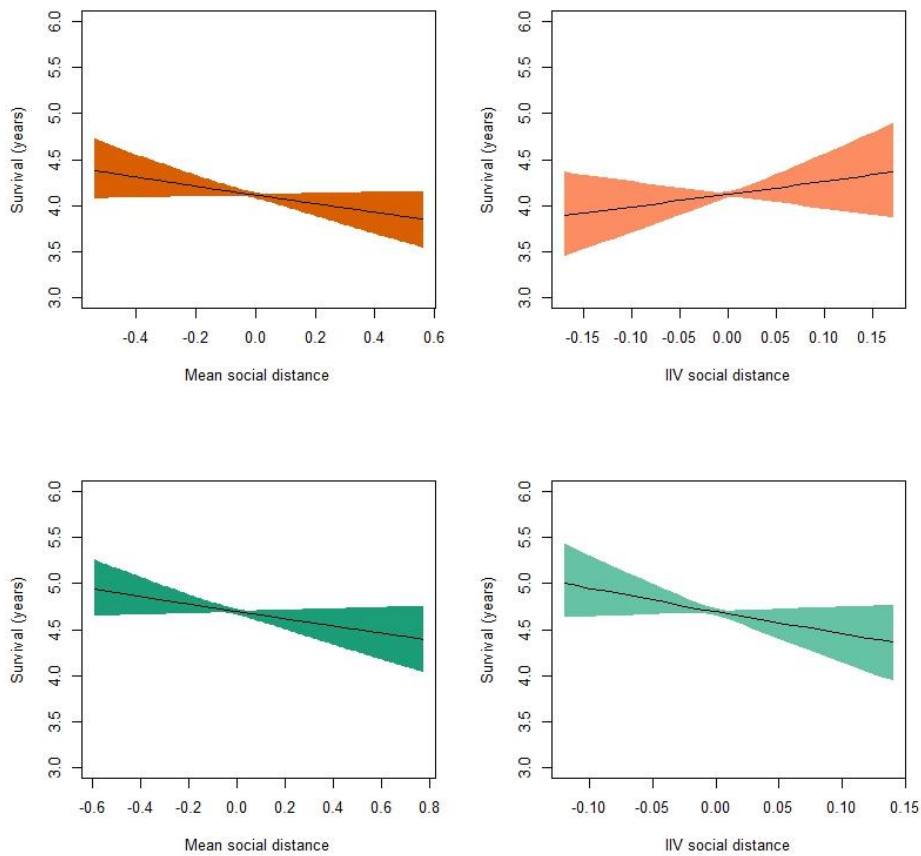
321

322 **Figure 2:** Predicted lifetime reproductive success as a function of individual mean social distance (left,
 323 dark colours) and IIV (right, light colours) in males (top row, orange) and females (bottom row, green).
 324 The black lines and the shaded areas depict the median estimates and 95% credible intervals of the
 325 distribution of coefficients derived from zero-inflated Poisson models fitting each individual posterior.



326

327 **Figure 3:** Predicted reproductive rate as a function of individual mean social distance (left, dark colours)
 328 and IIV (right, light colours) in males (top row, orange) and females (bottom row, green). The black
 329 lines and the shaded areas depict the median estimates and 95% credible intervals of the distribution of
 330 coefficients derived from zero-inflated Poisson models fitting each individual posterior.



331

332 **Figure 4:** Predicted survival as a function of individual mean social distance (left, dark colours) and IIV
 333 (right, light colours) in males (top row, orange) and females (bottom row, green). The black lines and
 334 the shaded areas depict the median estimates and 95% credible intervals of the distribution of
 335 coefficients derived from Poisson models fitting each individual posterior.

336

337 **Discussion**

338 This study not only provided the first estimates of predictability of social behaviour, but also investigated
339 for the first time its link with social behaviour and fitness in a wild animal. We demonstrated, in eastern
340 water dragons, that i) females were more social and slightly more variable in their social behaviour than
341 males, but both sexes were similarly predictable in their social behaviour; ii) in both sexes, individuals
342 varied both in their average and in the predictability of their social behaviour; iii) more social females
343 were also more predictable in their social behaviour; iv) more social males had a higher fitness (survival
344 and reproductive success) while more social (and predictable) females lived longer lives. These findings
345 provide partial support for the social niche hypothesis, which posits that higher density environments
346 should favour niche partitioning and behavioural predictability. We detail below how our results align
347 or differ from predictions of the social niche hypothesis and discuss their broader implications.

348 *Sex differences in means, variance and predictability*

349 Sexual dimorphism arises because sexes are subject to different selective pressures. Thus, especially for
350 behavioural traits, sexual dimorphism can manifest not only in the strength or average level of a trait,
351 but also in the variability of the trait among individuals (Zajitschek et al., 2020). While our finding that
352 male dragons are less social than females is in line with earlier studies in this system (Strickland & Frère,
353 2019), our results do not align with the ‘greater male variability hypothesis’, which predicts condition-
354 dependent sexual traits to vary more among males than among females (Cuervo & Møller, 1999, 2001;
355 Pomiankowski & Møller, 1997). In this population, males indeed appeared to be more variable in their
356 reproductive success than females, which likely results from competition for mating opportunities. In
357 contrast, evidence that males are more variable than females for behavioural traits to date remains
358 equivocal (Wyman & Rowe, 2014; Zajitschek et al., 2020), which may be because behaviours are not
359 necessarily body condition-dependent. This is seen in our study, whereby social behaviour was not
360 associated with body size, an important contributor to fitness in males. In contrast to the greater male
361 variability hypothesis, the social niche hypothesis would predict the more social sex, here females, to
362 be more variable and more predictable. Here, we found weak support for this prediction, as females
363 were slightly more variable than males in their social behaviour but we did not find evidence for sex-

364 differences in predictability. Similar sex-differences in among-individual variance and repeatability of
365 social behaviour, were previously found in this population (Delmé et al., 2023), in bighorn sheep (*Ovis*
366 *canadensis*, Vander Wal et al., 2015) and in chimpanzees (*Pan troglodytes*, Tkaczynski et al., 2020),
367 and could be caused by sex-differences in selection rather than a greater social niche partitioning in
368 females (see below).

369 *Mean-predictability relationships*

370 Previous laboratory studies of behavioural predictability under experimental conditions have found
371 associations between mean behaviours and their predictabilities. For example, Mitchell et al. (2016)
372 found that more active individual guppies (*Poecilia reticulata*) were also more predictable, and Jolles
373 et al. (2019) found that bolder three-spined sticklebacks (*Gasterosteus aculeatus*) were also more
374 predictable. An extensive study by O’Dea et al. (2022) in zebrafish (*Danio rerio*) also found that more
375 active and more social individuals were more predictable, that neophobic and aggressive individuals
376 were less predictable, and that boldness and its predictability did not correlate. Our social tolerance
377 results align with this latter study in that more social females were also more predictable, although one
378 should keep in mind that relationships between mean behaviours and their predictability likely vary
379 between populations and traits (Mitchell et al., 2021). Here, we showed that this correlation could also
380 differ between the sexes in a sexually dimorphic species, potentially because of sex-specific social
381 constraints (see below).

382 A few studies have also quantified mean-predictability relationships in wild animals (Hertel et al., 2020,
383 2021; Martin et al., 2017). For example, previous studies in bears (*Ursus arctos*) and elephants
384 (*Loxodonta africanus*) found that individuals that moved large daily distances were less predictable than
385 individuals that moved shorter distances (Hertel et al., 2020, 2021). Correlations between distances
386 moved and predictability in these distances were also relatively strong, although not as strong as the
387 correlation we found for social behaviour in female dragons. One may argue that such a mean-variance
388 relationship can mathematically arise with distance data as its distribution is likely non-normal (bound
389 to zero and right-skewed). However, this would then have also driven a positive mean-predictability
390 correlation in males, which we did not find. Furthermore, simulations showed that our model accurately

391 estimated the mean-predictability correlation for a trait that had a similar distribution as social distance
392 (Text S2, Figure S1). We also note that contrary to experimental studies performed in controlled
393 conditions, predictability in the wild may also include plasticity to unknown environmental factors
394 (Stamps et al., 2012; Westneat et al., 2015) and thus reflects how generally plastic vs. constrained an
395 individual's behaviour is. Our results therefore suggest that the strong positive correlations between
396 mean behaviour and its predictability are caused by female-specific constraints on social behaviour. This
397 individual-level association between social behaviour and predictability would align with the social
398 niche predictions even though we expected to find this pattern in both sexes. The fact that this association
399 only occurred in females however does not contradict our prediction that individuals in the more social
400 sex may exhibit stronger constraints on their social behaviour. While our analyses of fitness could not
401 detect such constraints on adult females' social behaviour (see below), we cannot rule out selection on
402 social behaviour at earlier life stages (Nuñez et al., 2015).

403 *Association with fitness*

404 While the fitness benefits of social behaviours have been investigated and reported in several species
405 (Brent et al., 2013; Frère et al., 2010; Kohn, 2017; Silk, 2007) and a previous study reported sex-specific
406 selection on social behaviours in this study system (Delmé et al., 2023), our study goes one step further
407 by simultaneously investigating different components of fitness and incorporating predictability.
408 Importantly, this study is also one of the rare studies connecting behavioural predictability and fitness
409 in the wild (Cain et al., 2023; Patrick et al., 2021).

410 For males, we found that more social individuals lived slightly longer lives and produced more offspring,
411 which resulted in a higher lifetime reproductive success and suggested that mean social tolerance may
412 be under directional selection. This positive relationship between males' social behaviour and
413 reproductive success was consistent with Delmé et al. (2023) and was likely driven by social associations
414 with females, which increased mating opportunities for more social males. As directional selection is
415 expected to deplete additive genetic variation of traits, the finding that males had slightly lower among-
416 individual variation than females could therefore be a consequence of such sex-specific selective regime.
417 Interestingly, unpredictability seemed to be associated with a higher lifetime reproductive success in

418 that unpredictable males that had offspring had more of them than predictable males. The magnitude of
419 the relationship between predictability and reproductive rate was similar in magnitude to that with LRS,
420 while predictability was clearly not associated with survival in males. These results would suggest
421 reproductive benefits of unpredictability for males, although power to detect such effects was likely
422 limited in our dataset. Because unpredictability includes unmeasured plasticity, its positive association
423 with reproduction could be caused by males adjusting their behaviour more appropriately to social
424 contexts, which may provide them with increased mating opportunities. These benefits of
425 unpredictability were however masked when considering total reproductive outputs (including failures
426 to reproduce). The reason why predictability did not associate with the probability to produce offspring
427 in males remains unclear.

428 For females, we found no evidence that social behaviour and its predictability were associated with
429 reproductive rate which was also consistent with Delmé et al. (2023) but contrasted with numerous
430 studies in mammals (Silk, 2007). Because breeding systems may affect the adaptive value of females'
431 social behaviour (Silk, 2007), our findings might be explained by the absence of parental care in dragons,
432 which precludes any form of cooperation over offspring rearing. As for males, we found that more social
433 (and predictable) females had a slightly higher survival. We note that the very high correlation between
434 mean and predictability was probably causing both aspects of females' social behaviour to be similarly
435 associated with survival (Figure 3). This possible survival advantage of being more social however did
436 not seem to translate into a higher lifetime reproductive success in females. This could be because social
437 behaviour was only (and weakly) associated with one fitness component in females (as opposed to both
438 fitness components in males) and because females exhibited less variance in lifetime reproductive
439 success than males.

440 In both sexes, we found a small positive association between social behaviour and survival in both sexes,
441 which could be due to social tolerance decreasing the costs of competition (e.g. fights for resources or
442 territories, Haunhorst et al., 2017), sexual harassment of females by males (Fox, 2002), or increasing
443 social information about the location of food sources (Aplin et al., 2012) or predators (Beauchamp,
444 2010). Individuals that are more socially tolerant (and females that are more predictable) may also

445 occupy a different position in the social network, which has repeatedly been found to correlate with
446 longevity (Barocas et al., 2011; Holt-Lunstad et al., 2010; Silk et al., 2010; Vander Wal et al., 2015).
447 Alternatively, a higher social tolerance could be found in dominant individuals, which pay low costs of
448 social proximity to other subordinate individuals. However, here, social behaviour was not associated
449 with individual size, a predictor of dominance in this species (Piza-Roca et al., 2020). In contrast to
450 mean social behaviour, we did not find behavioural predictability to clearly associate with survival in
451 either sex. In dragons, we expected that more predictable individuals would survive longer due to
452 experiencing lower levels of conflict or, alternatively, that unpredictable individuals would be favoured
453 in contest situations due to being more plastic (Briffa & Lane, 2017). These negative results contrasted
454 with a recent study in barn owls (*Tyto alba*) finding that individuals that were less predictable in their
455 movement behaviour had a higher survival (Cain et al., 2023), and therefore (unsurprisingly) suggest
456 that benefits of behavioural predictability are context-dependent. Additionally, although sex-specific
457 associations between survival and predictability were very uncertain should be interpreted with caution,
458 their difference in signs raises the idea that selection on predictability could also be sex-specific.

459 An important finding of this study is that sexes differed in how their mean social behaviour and, to some
460 extent, its predictability, associated with reproductive success. Sex-differences in reproductive selection
461 and similar survival selection in both sexes was previously found in a bighorn sheep population where
462 social behaviour increased reproductive success in females only and increased survival in both sexes
463 (Vander Wal et al., 2015). Assuming that mean and predictability in social behaviour are heritable and
464 genetically correlated across sexes (Connallon & Clark, 2014; Kaufmann et al., 2023), one could
465 speculate sex-specific reproductive selection to be a mechanism maintaining among-individual variation
466 in these traits (Schuett et al., 2010). We are hence yet to determine whether social behaviour and its
467 predictability are heritable and genetically correlated between sexes in eastern water dragons.

468 **Conclusion**

469 Our study demonstrated that social behaviour and its predictability covary with each other and with
470 fitness in a sex-specific way in a wild lizard. Females, which were more social than males exhibited a
471 slightly higher among-individual variation and a strong positive association between social behaviour

472 and predictability. Although sexes did not differ in the predictability of their social behaviour, this mean-
473 predictability association was absent in males. Mean social behaviour was associated with a higher
474 reproductive success in males and a slightly higher survival in both sexes. Altogether, these findings
475 partly supported predictions from the social niche hypothesis at the sex and individual levels. We also
476 provided rare empirical estimates of behavioural predictability and of its association with fitness in a
477 wild animal. This study therefore highlights the evolutionary relevance of behavioural predictability and
478 will hopefully promote further research on the links between social behaviour, animal personality and
479 predictability.

480 **Authors' contributions**

481 BC,CF, SN conceived the original idea of the study. BC performed the statistical analyses and wrote the
482 manuscript, with input from all other authors. NJ performed the parentage analyses. BC,CF,NJ,KS
483 contributed to data collection. All authors provided comments on the manuscript and agreed on the final
484 version to be submitted for publication.

485 **Funding**

486 This work was supported by Dr Celine Frere's internal funding from the University of the Sunshine
487 Coast and ARC Future Fellowship FT200100192.

488 **Acknowledgements**

489 We are grateful to all the students and volunteers who have collected and processed data for this project
490 over the years. We also would like to thank the staff at Roma Street Parklands for their ongoing support
491 with data collection and Dr Sinta Frère for the development of the dragon database. We also thank Tracy
492 Burkhard for feedback on an early version of this manuscript and Raphael Royauté for statistical advice.

493 **Ethics statement**

494 This research was conducted with the approval of the University of the Sunshine Coast's animal ethics
495 committee (approval numbers: ANA1487, ANA20161, ANS1858).

496

497 **Data accessibility**

498 Data and R code used for this study are available on OSF:

499 https://osf.io/3y6s7/?view_only=512a97f97623411781eacdd81bf2b6cd

500 **Conflict of interest disclosure**

501 Authors declare no conflict of interest

502 **References**

503 Aplin, L. M., Farine, D. R., Morand-Ferron, J., & Sheldon, B. C. (2012). Social networks predict patch
504 discovery in a wild population of songbirds. *Proceedings of the Royal Society B: Biological*
505 *Sciences*, 279(1745), 4199–4205. <https://doi.org/10.1098/rspb.2012.1591>

506 Aplin, L. M., Firth, J. A., Farine, D. R., Voelkl, B., Crates, R. A., Culina, A., Garroway, C. J., Hinde,
507 C. A., Kidd, L. R., Psorakis, I., Milligan, N. D., Radersma, R., Verhelst, B. L., & Sheldon, B.
508 C. (2015). Consistent individual differences in the social phenotypes of wild great tits, *Parus*
509 *major*. *Animal Behaviour*, 108, 117–127. <https://doi.org/10.1016/j.anbehav.2015.07.016>

510 Araya-Ajoy, Y. G., Mathot, K. J., & Dingemanse, N. J. (2015). An approach to estimate short-term,
511 long-term and reaction norm repeatability. *Methods in Ecology and Evolution*, 6(12), 1462–
512 1473. <https://doi.org/10.1111/2041-210X.12430>

513 Bailey, N. W., & Desjonquères, C. (2022). The Indirect Genetic Effect Interaction Coefficient ψ :
514 Theoretically Essential and Empirically Neglected. *Journal of Heredity*, 113(1), 79–90.
515 <https://doi.org/10.1093/jhered/esab056>

516 Baird, T. A., Baird, T. D., & Shine, R. (2012). Aggressive Transition between Alternative Male Social
517 Tactics in a Long-Lived Australian Dragon (*Physignathus lesueurii*) Living at High Density.
518 *PLOS ONE*, 7(8), e41819. <https://doi.org/10.1371/journal.pone.0041819>

519 Ball, S., Potvin, D. A., Class, B., & Frère, C. H. (2023). Agonism does not covary with territoriality in
520 a gregarious reptile. *Behavioral Ecology and Sociobiology*, 77(6), 73.
521 <https://doi.org/10.1007/s00265-023-03348-0>

522 Barocas, A., Ilany, A., Koren, L., Kam, M., & Geffen, E. (2011). Variance in Centrality within Rock
523 Hyrax Social Networks Predicts Adult Longevity. *PLOS ONE*, 6(7), e22375.
524 <https://doi.org/10.1371/journal.pone.0022375>

525 Beauchamp, G. (2010). Relaxed predation risk reduces but does not eliminate sociality in birds.
526 *Biology Letters*, 6(4), 472–474. <https://doi.org/10.1098/rsbl.2009.1063>

527 Bell, A. M., Hankison, S. J., & Laskowski, K. L. (2009). The repeatability of behaviour: A meta-
528 analysis. *Animal Behaviour*, 77(4), 771–783. <https://doi.org/10.1016/j.anbehav.2008.12.022>

529 Bergmüller, R., & Taborsky, M. (2010). Animal personality due to social niche specialisation. *Trends*
530 *in Ecology & Evolution*, 25(9), 504–511. <https://doi.org/10.1016/j.tree.2010.06.012>

531 Brent, L. J. N., Heilbronner, S. R., Horvath, J. E., Gonzalez-Martinez, J., Ruiz-Lambides, A.,
532 Robinson, A. G., Skene, J. H. P., & Platt, M. L. (2013). Genetic origins of social networks in
533 rhesus macaques. *Scientific Reports*, 3(1), 1042. <https://doi.org/10.1038/srep01042>

534 Briffa, M. (2013). Plastic proteans: Reduced predictability in the face of predation risk in hermit crabs.
535 *Biology Letters*, 9(5), 20130592. <https://doi.org/10.1098/rsbl.2013.0592>

536 Briffa, M., & Lane, S. M. (2017). The role of skill in animal contests: A neglected component of
537 fighting ability. *Proceedings of the Royal Society B: Biological Sciences*, 284(1863),
538 20171596. <https://doi.org/10.1098/rspb.2017.1596>

539 Bürkner, P.-C. (2017). brms: An R Package for Bayesian Multilevel Models Using Stan. *Journal of*
540 *Statistical Software*, 80, 1–28. <https://doi.org/10.18637/jss.v080.i01>

541 Cain, S., Solomon, T., Leshem, Y., Toledo, S., Arnon, E., Roulin, A., & Spiegel, O. (2023).
542 Movement predictability of individual barn owls facilitates estimation of home range size and
543 survival. *Movement Ecology*, 11(1), 10. <https://doi.org/10.1186/s40462-022-00366-x>

544 Chang, C. C., Chow, C. C., Tellier, L. C., Vattikuti, S., Purcell, S. M., & Lee, J. J. (2015). Second-
545 generation PLINK: Rising to the challenge of larger and richer datasets. *GigaScience*, 4(1),
546 s13742-015-0047–0048. <https://doi.org/10.1186/s13742-015-0047-8>

547 Chervet, N., Zöttl, M., Schürch, R., Taborsky, M., & Heg, D. (2011). Repeatability and Heritability of
548 Behavioural Types in a Social Cichlid. *International Journal of Evolutionary Biology*, 2011,
549 1–15. <https://doi.org/10.4061/2011/321729>

550 Cleasby, I. R., Nakagawa, S., & Schielzeth, H. (2015). Quantifying the predictability of behaviour:
551 Statistical approaches for the study of between-individual variation in the within-individual
552 variance. *Methods in Ecology and Evolution*, 6(1), 27–37. [https://doi.org/10.1111/2041-](https://doi.org/10.1111/2041-210X.12281)
553 210X.12281

554 Connallon, T., & Clark, A. G. (2014). Balancing Selection in Species with Separate Sexes: Insights
555 from Fisher's Geometric Model. *Genetics*, 197(3), 991–1006.
556 <https://doi.org/10.1534/genetics.114.165605>

557 Cuervo, J. J., & Møller, A. P. (1999). Phenotypic variation and fluctuating asymmetry in sexually
558 dimorphic feather ornaments in relation to sex and mating system. *Biological Journal of the*
559 *Linnean Society*, 68(4), 505–529. <https://doi.org/10.1111/j.1095-8312.1999.tb01186.x>

560 Cuervo, J. J., & Møller, A. P. (2001). Components of phenotypic variation in avian ornamental and
561 non-ornamental feathers. *Evolutionary Ecology*, 15(1), 53–72.
562 <https://doi.org/10.1023/A:1011913804309>

563 Delmé, C., Jackson, N., Class, B., Strickland, K., Potvin, D. A., & Frère, C. H. (2023). Adaptive
564 significance of affiliative behaviour differs between sexes in a wild reptile population.
565 *Proceedings of the Royal Society B: Biological Sciences*, 290(2001), 20230805.
566 <https://doi.org/10.1098/rspb.2023.0805>

567 Dingemanse, N. J., Kazem, A. J. N., Réale, D., & Wright, J. (2010). Behavioural reaction norms:
568 Animal personality meets individual plasticity. *Trends in Ecology & Evolution*, 25(2), 81–89.
569 <https://doi.org/10.1016/j.tree.2009.07.013>

570 Feutry, P., Devloo-Delva, F., Tran Lu Y, A., Mona, S., Gunasekera, R. M., Johnson, G., Pillans, R. D.,
571 Jaccoud, D., Kilian, A., Morgan, D. L., Saunders, T., Bax, N. J., & Kyne, P. M. (2020). One
572 panel to rule them all: DArTcap genotyping for population structure, historical demography,
573 and kinship analyses, and its application to a threatened shark. *Molecular Ecology Resources*,
574 20(6), 1470–1485. <https://doi.org/10.1111/1755-0998.13204>

575 Fox, E. A. (2002). Female tactics to reduce sexual harassment in the Sumatran orangutan (*Pongo*
576 *pygmaeus abelii*). *Behavioral Ecology and Sociobiology*, 52(2), 93–101.
577 <https://doi.org/10.1007/s00265-002-0495-x>

- 578 Frère, C. H., Chandrasoma, D., & Whiting, M. J. (2015). Polyandry in dragon lizards: Inbred paternal
579 genotypes sire fewer offspring. *Ecology and Evolution*, 5(8), 1686–1692.
580 <https://doi.org/10.1002/ece3.1447>
- 581 Frère, C. H., Krützen, M., Mann, J., Connor, R. C., Bejder, L., & Sherwin, W. B. (2010). Social and
582 genetic interactions drive fitness variation in a free-living dolphin population. *Proceedings of*
583 *the National Academy of Sciences*, 107(46), 19949–19954.
584 <https://doi.org/10.1073/pnas.1007997107>
- 585 Gardiner, R. Z., Doran, E., Strickland, K., Carpenter-Bundhoo, L., & Frère, C. (2014). A Face in the
586 Crowd: A Non-Invasive and Cost Effective Photo-Identification Methodology to Understand
587 the Fine Scale Movement of Eastern Water Dragons. *PLoS ONE*, 9(5), e96992.
588 <https://doi.org/10.1371/journal.pone.0096992>
- 589 Godoy, I., Korsten, P., & Perry, S. E. (2022). Genetic, maternal, and environmental influences on
590 sociality in a pedigreed primate population. *Heredity*, 129(4), 203–214.
591 <https://doi.org/10.1038/s41437-022-00558-6>
- 592 Hadfield, J. D., Wilson, A. J., Garant, D., Sheldon, B. C., & Kruuk, L. E. B. (2010). The Misuse of
593 BLUP in Ecology and Evolution. *The American Naturalist*, 175(1), 116–125.
594 <https://doi.org/10.1086/648604>
- 595 Haunhorst, C. B., Heesen, M., Ostner, J., & Schülke, O. (2017). Social bonds with males lower the
596 costs of competition for wild female Assamese macaques. *Animal Behaviour*, 125, 51–60.
597 <https://doi.org/10.1016/j.anbehav.2017.01.008>
- 598 Hertel, A. G., Niemelä, P. T., Dingemanse, N. J., & Mueller, T. (2020). A guide for studying among-
599 individual behavioral variation from movement data in the wild. *Movement Ecology*, 8(1), 30.
600 <https://doi.org/10.1186/s40462-020-00216-8>
- 601 Hertel, A. G., Royauté, R., Zedrosser, A., & Mueller, T. (2021). Biologging reveals individual
602 variation in behavioural predictability in the wild. *Journal of Animal Ecology*, 90(3), 723–737.
603 <https://doi.org/10.1111/1365-2656.13406>

604 Holt-Lunstad, J., Smith, T. B., & Layton, J. B. (2010). Social Relationships and Mortality Risk: A
605 Meta-analytic Review. *PLOS Medicine*, 7(7), e1000316.
606 <https://doi.org/10.1371/journal.pmed.1000316>

607 Huisman, J. (2017). Pedigree reconstruction from SNP data: Parentage assignment, sibship clustering
608 and beyond. *Molecular Ecology Resources*, 17(5), 1009–1024. [https://doi.org/10.1111/1755-](https://doi.org/10.1111/1755-0998.12665)
609 0998.12665

610 Humphries, D. A., & Driver, P. M. (1970). Protean defence by prey animals. *Oecologia*, 5(4), 285–
611 302. <https://doi.org/10.1007/BF00815496>

612 Jackman, S. (2010). pscl: Classes and methods for R. Developed in the Political Science
613 Computational Laboratory, Stanford University. Department of Political Science, Stanford
614 University, Stanford, CA. R package version 1.03.5. <Http://Www.Pscl.Stanford.Edu/>.
615 <https://cir.nii.ac.jp/crid/1572261549901049728>

616 Jackson, N., Cristescu, R. H., Piza-Roca, C., Littleford-Colquhoun, B. L., Strickland, K., & Frère, C.
617 H. (2019). Maternal nesting behaviour in city dragons: A species with temperature-dependent
618 sex determination. *Journal of Urban Ecology*, 5(1), juz005. <https://doi.org/10.1093/jue/juz005>

619 Johnstone, R. A. (2001). Eavesdropping and animal conflict. *Proceedings of the National Academy of*
620 *Sciences of the United States of America*, 98(16), 9177–9180.

621 Jolles, J. W., Briggs, H. D., Araya-Ajoy, Y. G., & Boogert, N. J. (2019). Personality, plasticity and
622 predictability in sticklebacks: Bold fish are less plastic and more predictable than shy fish.
623 *Animal Behaviour*, 154, 193–202. <https://doi.org/10.1016/j.anbehav.2019.06.022>

624 Kaufmann, P., Howie, J. M., & Immonen, E. (2023). Sexually antagonistic selection maintains genetic
625 variance when sexual dimorphism evolves. *Proceedings of the Royal Society B: Biological*
626 *Sciences*, 290(1995), 20222484. <https://doi.org/10.1098/rspb.2022.2484>

627 Kohn, G. M. (2017). Friends give benefits: Autumn social familiarity preferences predict reproductive
628 output. *Animal Behaviour*, 132, 201–208. <https://doi.org/10.1016/j.anbehav.2017.08.013>

629 Lea, A. J., Blumstein, D. T., Wey, T. W., & Martin, J. G. A. (2010). Heritable victimization and the
630 benefits of agonistic relationships. *Proceedings of the National Academy of Sciences of the*
631 *United States of America*, 107(50), 21587–21592. <https://doi.org/10.1073/pnas.1009882107>

632 Littleford-Colquhoun, B. L., Clemente, C., Whiting, M. J., Ortiz-Barrientos, D., & Frère, C. H. (2017).
633 Archipelagos of the Anthropocene: Rapid and extensive differentiation of native terrestrial
634 vertebrates in a single metropolis. *Molecular Ecology*, *26*(9), 2466–2481.
635 <https://doi.org/10.1111/mec.14042>

636 Martin, J. G. A., Pirotta, E., Petelle, M. B., & Blumstein, D. T. (2017). Genetic basis of between-
637 individual and within-individual variance of docility. *Journal of Evolutionary Biology*, *30*(4),
638 796–805. <https://doi.org/10.1111/jeb.13048>

639 Mitchell, D. J., Beckmann, C., & Biro, P. A. (2021). Understanding the unexplained: The magnitude
640 and correlates of individual differences in residual variance. *Ecology and Evolution*, *11*(12),
641 7201–7210. <https://doi.org/10.1002/ece3.7603>

642 Mitchell, D. J., Fanson, B. G., Beckmann, C., & Biro, P. A. (2016). Towards powerful experimental
643 and statistical approaches to study intraindividual variability in labile traits. *Royal Society*
644 *Open Science*, *3*(10), 160352. <https://doi.org/10.1098/rsos.160352>

645 Nakagawa, S., Poulin, R., Mengersen, K., Reinhold, K., Engqvist, L., Lagisz, M., & Senior, A. M.
646 (2015). Meta-analysis of variation: Ecological and evolutionary applications and beyond.
647 *Methods in Ecology and Evolution*, *6*(2), 143–152. <https://doi.org/10.1111/2041-210X.12309>

648 Nuñez, C. M. V., Adelman, J. S., & Rubenstein, D. I. (2015). Sociality increases juvenile survival
649 after a catastrophic event in the feral horse (*Equus caballus*). *Behavioral Ecology*, *26*(1),
650 138–147. <https://doi.org/10.1093/beheco/aru163>

651 O’Dea, R. E., Noble, D. W. A., & Nakagawa, S. (2022). Unifying individual differences in
652 personality, predictability and plasticity: A practical guide. *Methods in Ecology and Evolution*,
653 *13*(2), 278–293. <https://doi.org/10.1111/2041-210X.13755>

654 Patrick, S. C., Martin, J. G. A., Ummenhofer, C. C., Corbeau, A., & Weimerskirch, H. (2021).
655 Albatrosses respond adaptively to climate variability by changing variance in a foraging trait.
656 *Global Change Biology*, *27*(19), 4564–4574. <https://doi.org/10.1111/gcb.15735>

657 Peterson, N. R., Rose, K., Shaw, S., Hyndman, T. H., Sigler, L., Kurtböke, D. İ., Llinas, J., Littleford-
658 Colquhoun, B. L., Cristescu, R., & Frère, C. (2020). Cross-continental emergence of

659 Nannizziopsis barbatae disease may threaten wild Australian lizards. *Scientific Reports*, 10(1),
660 20976. <https://doi.org/10.1038/s41598-020-77865-7>

661 Piza-Roca, C., Schoeman, D., & Frere, C. (2020). Fitness benefits of male dominance behaviours
662 depend on the degree of individual inbreeding in a polyandrous lizard. *Proceedings of the*
663 *Royal Society B: Biological Sciences*, 287(1927), 20200097.
664 <https://doi.org/10.1098/rspb.2020.0097>

665 Pomiankowski, A., & Møller, A. P. (1997). A resolution of the lek paradox. *Proceedings of the Royal*
666 *Society of London. Series B: Biological Sciences*, 260(1357), 21–29.
667 <https://doi.org/10.1098/rspb.1995.0054>

668 Prentice, P. M., Houslay, T. M., Martin, J. G. A., & Wilson, A. J. (2020). Genetic variance for
669 behavioural ‘predictability’ of stress response. *Journal of Evolutionary Biology*, 33(5), 642–
670 652. <https://doi.org/10.1111/jeb.13601>

671 R Core Team. (2022). *R: A Language and Environment for Statistical Computing*. R Foundation for
672 Statistical Computing, Vienna, Austria.

673 Richardson, G., Dickinson, P., Burman, O. H. P., & Pike, T. W. (2018). Unpredictable movement as
674 an anti-predator strategy. *Proceedings of the Royal Society B: Biological Sciences*, 285(1885),
675 20181112. <https://doi.org/10.1098/rspb.2018.1112>

676 Schuett, W., Tregenza, T., & Dall, S. R. X. (2010). Sexual selection and animal personality. *Biological*
677 *Reviews*, 85(2), 217–246. <https://doi.org/10.1111/j.1469-185X.2009.00101.x>

678 Silk, J. B. (2007). The adaptive value of sociality in mammalian groups. *Philosophical Transactions of*
679 *the Royal Society B: Biological Sciences*, 362(1480), 539–559.
680 <https://doi.org/10.1098/rstb.2006.1994>

681 Silk, J. B., Beehner, J. C., Bergman, T. J., Crockford, C., Engh, A. L., Moscovice, L. R., Wittig, R. M.,
682 Seyfarth, R. M., & Cheney, D. L. (2010). Strong and Consistent Social Bonds Enhance the
683 Longevity of Female Baboons. *Current Biology*, 20(15), 1359–1361.
684 <https://doi.org/10.1016/j.cub.2010.05.067>

685 Smith, A. T., & Ivins, B. L. (1984). Spatial Relationships and Social Organization in Adult Pikas: A
686 Facultatively Monogamous Mammal. *Zeitschrift Für Tierpsychologie*, 66(4), 289–308.
687 <https://doi.org/10.1111/j.1439-0310.1984.tb01370.x>

688 Stamps, J. A., Briffa, M., & Biro, P. A. (2012). Unpredictable animals: Individual differences in
689 intraindividual variability (IIV). *Animal Behaviour*, 83(6), 1325–1334.
690 <https://doi.org/10.1016/j.anbehav.2012.02.017>

691 Strickland, K., & Frère, C. H. (2018). Predictable males and unpredictable females: Repeatability of
692 sociability in eastern water dragons. *Behavioral Ecology*, 29(1), 236–243.
693 <https://doi.org/10.1093/beheco/axx148>

694 Strickland, K., & Frère, C. H. (2019). Individual Variation in the Social Plasticity of Water Dragons.
695 *The American Naturalist*, 194(2), 194–206. <https://doi.org/10.1086/704089>

696 Strickland, K., Gardiner, R., Schultz, A. J., & Frère, C. H. (2014). The social life of eastern water
697 dragons: Sex differences, spatial overlap and genetic relatedness. *Animal Behaviour*, 97, 53–
698 61. <https://doi.org/10.1016/j.anbehav.2014.08.009>

699 Strickland, K., Levensgood, A., Foroughirad, V., Mann, J., Krzyszczyk, E., & Frère, C. H. (2017). A
700 framework for the identification of long-term social avoidance in longitudinal datasets. *Royal*
701 *Society Open Science*, 4(8), 170641. <https://doi.org/10.1098/rsos.170641>

702 Strickland, K., Mitchell, D. J., Delmé, C., & Frère, C. H. (2021). Repeatability and heritability of
703 social reaction norms in a wild agamid lizard. *Evolution*, 75(8), 1953–1965.
704 <https://doi.org/10.1111/evo.14298>

705 Tacey, J., Class, B., Delmé, C., Powell, D., & Frère, C. H. (2023). Impacts of fungal disease on dyadic
706 social interactions in a wild agamid lizard. *Animal Behaviour*, 200, 125–136.
707 <https://doi.org/10.1016/j.anbehav.2023.04.002>

708 Thompson, M. B. (1993). Estimate of the population structure of the eastern water dragon,
709 *Physignathus lesueurii* (Reptilia: Agamidae), along riverside habitat. *Wildlife Research*, 20(5),
710 613–619. <https://doi.org/10.1071/wr9930613>

711 Tkaczynski, P. J., Mielke, A., Samuni, L., Preis, A., Wittig, R. M., & Crockford, C. (2020). Long-term
712 repeatability in social behaviour suggests stable social phenotypes in wild chimpanzees. *Royal*
713 *Society Open Science*, 7(8), 200454. <https://doi.org/10.1098/rsos.200454>

714 Van Tienhoven, A. M., Den Hartog, J. E., Reijns, R. A., & Peddemors, V. M. (2007). A computer-
715 aided program for pattern-matching of natural marks on the spotted raggedtooth shark
716 *Carcharias taurus*. *Journal of Applied Ecology*, 44(2), 273–280.
717 <https://doi.org/10.1111/j.1365-2664.2006.01273.x>

718 Vander Wal, E., Festa-Bianchet, M., Réale, D., Coltman, D. W., & Pelletier, F. (2015). Sex-based
719 differences in the adaptive value of social behavior contrasted against morphology and
720 environment. *Ecology*, 96(3), 631–641. <https://doi.org/10.1890/14-1320.1>

721 Watson, S. K., Vale, G. L., Hopper, L. M., Dean, L. G., Kendal, R. L., Price, E. E., Wood, L. A.,
722 Davis, S. J., Schapiro, S. J., Lambeth, S. P., & Whiten, A. (2018). Chimpanzees demonstrate
723 individual differences in social information use. *Animal Cognition*, 21(5), 639–650.
724 <https://doi.org/10.1007/s10071-018-1198-7>

725 Webber, Q. M. R., & Vander Wal, E. (2018). An evolutionary framework outlining the integration of
726 individual social and spatial ecology. *Journal of Animal Ecology*, 87(1), 113–127.
727 <https://doi.org/10.1111/1365-2656.12773>

728 Westneat, D. F., Wright, J., & Dingemanse, N. J. (2015). The biology hidden inside residual within-
729 individual phenotypic variation: The biology of residual phenotypic variance. *Biological*
730 *Reviews*, 90(3), 729–743. <https://doi.org/10.1111/brv.12131>

731 Wolf, M., Doorn, G. S. V., & Weissing, F. J. (2011). On the coevolution of social responsiveness and
732 behavioural consistency. *Proceedings of the Royal Society of London. Series B, Biological*
733 *Sciences*, 278(1704), 440–448. <https://doi.org/10.1098/rspb.2010.1051>

734 Wolf, M., & McNamara, J. M. (2013). Adaptive between-individual differences in social competence.
735 *Trends in Ecology & Evolution*, 28(5), 253–254. <https://doi.org/10.1016/j.tree.2013.01.006>

736 Wyman, M. J., & Rowe, L. (2014). Male Bias in Distributions of Additive Genetic, Residual, and
737 Phenotypic Variances of Shared Traits. *The American Naturalist*, 184(3), 326–337.
738 <https://doi.org/10.1086/677310>

739 Zajitschek, S. R., Zajitschek, F., Bonduriansky, R., Brooks, R. C., Cornwell, W., Falster, D. S., Lagisz,
740 M., Mason, J., Senior, A. M., Noble, D. W., & Nakagawa, S. (2020). Sexual dimorphism in
741 trait variability and its eco-evolutionary and statistical implications. *ELife*, 9, e63170.
742 <https://doi.org/10.7554/eLife.63170>

743 Zeileis, A., Kleiber, C., & Jackman, S. (2008). Regression Models for Count Data in R. *Journal of*
744 *Statistical Software*, 27(8). <https://doi.org/10.18637/jss.v027.i08>

745