

1 **Treating gaps and biases in biodiversity data as a missing data**  
2 **problem**

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4 Diana E. Bowler<sup>1\*</sup>, Robin J. Boyd<sup>1</sup>, Corey T. Callaghan<sup>2</sup>, Robert A. Robinson<sup>3</sup>, Nick J. B.

5 Isaac<sup>1</sup>, Michael J. O. Pocock<sup>1</sup>

6

7 Affiliations

8 1 UK Centre for Ecology & Hydrology, Wallingford, OX10 8BB, UK

9 2 Department of Wildlife Ecology and Conservation, Fort Lauderdale Research and Education  
10 Center, University of Florida, Davie, FL 33314-7719

11 3 British Trust for Ornithology, The Nunnery, Thetford, IP24 2PU, UK

12

13 \* Corresponding author: [diana.e.bowler@gmail.com](mailto:diana.e.bowler@gmail.com)

14

## 15 **Abstract**

16 Big biodiversity datasets have great potential for monitoring and research because of their large  
17 taxonomic, geographic and temporal scope. Such datasets have become especially important for  
18 assessing the temporal change of species' populations and distributions. Gaps in the available  
19 data, however, often hinder drawing large-scale inferences about species' trends. Here, we  
20 conceptualise biodiversity data gaps as a missing data problem, which provides a unifying  
21 framework for the challenges and potential solutions across different types of biodiversity  
22 datasets. We characterise the typical types of data gaps in biodiversity data as different classes of  
23 missing data and then use missing data theory to explore the implications for different research  
24 questions. By using this framework, we show that bias due to data gaps can arise when the  
25 factors affecting sampling and/or data availability overlap with those affecting biodiversity. But  
26 the outcome also depends on the ecological questions, which determines choices around the  
27 analytical approach. We argue that typical approaches to long-term species trend modelling are  
28 especially susceptible to data gaps since such models do not tend to account for the factors that  
29 drive missingness. To identify general solutions, we review empirical studies and use simulation  
30 studies to compare some of the most frequently employed approaches to deal with data gaps,  
31 including subsampling, weighting and imputation. All these methods have the potential to reduce  
32 bias but may come at the cost of increased uncertainty of parameter estimates. Weighting  
33 approaches are arguably the least used so far in ecology and have the potential to reduce both the  
34 bias and variance of parameter estimates. Regardless of the method, the ability to reduce bias  
35 critically depends on knowledge of, and the availability of data on, the factors creating data gaps.  
36 We use our review to outline the necessary considerations when dealing with data gaps at  
37 different stages of the data collection and analysis workflow.

38 Keywords: Biodiversity change; Citizen Science; Ecological Modelling; Macroecology; Spatial

39 bias

40

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## 66 **I. Introduction: uneven sampling of biodiversity**

67 Ecologists have ever-growing access to data on species' occurrence and abundances. Potential  
68 sources of data include long-term citizen-science monitoring schemes (such as the North  
69 American Breeding Bird Survey) (Bled *et al.*, 2013), data aggregators (such as the Global  
70 Biodiversity Information Facility) (Garcia-Rosello *et al.*, 2015), remote-sensing platforms  
71 (Fretwell, Scofield & Phillips, 2017) and synthesis databases (such as BioTIME or the Living  
72 Planet Database) (Dornelas *et al.*, 2014). Since these data cover broad spatial and temporal  
73 scales, they are especially useful for large-scale questions, for instance, about species'  
74 distributions, population and community-level trends, and ecological niches (Chandler *et al.*,  
75 2017; Sullivan *et al.*, 2017; Fink *et al.*, 2020). These data also underpin many biodiversity trend  
76 indicators that are central for national and international conservation policy (Gregory *et al.*,  
77 2005; van Swaay *et al.*, 2008; Fraisl *et al.*, 2020).

78         Despite the impressive volume of data, biodiversity data, regardless of the source, tend to  
79 be filled with gaps and redundancies (Boakes *et al.*, 2010). Data gaps are not necessarily  
80 problematic; indeed, most studies rely on statistical inference to make inferences about a broader  
81 region of interest from a sample. Data gaps, however, can be problematic when they lead to  
82 biases (Boakes *et al.*, 2010; Bled *et al.*, 2013; Amano, Lamming & Sutherland, 2016). Already  
83 many ecologists have raised concerns about the impacts of bias on estimated spatial or temporal  
84 biodiversity patterns (Bayraktarov *et al.*, 2019; Valdez *et al.*, 2023). Developing methods to deal  
85 with data gaps and associated biases within large-scale biodiversity data is an increasingly  
86 important task to make full use of the growing big data sources.

87           Patterns in the availability of biodiversity data can be affected by the original motivations  
88 for, and constraints on, data collection activities. While some data are collected as part of  
89 scientific studies, much of the available data on species' occurrence and abundance are collected  
90 through citizen science initiatives (Chandler *et al.*, 2017). Spatial patterns in data availability  
91 from citizen science have been especially well-studied. Citizen science programs have varying  
92 degrees of protocol and sampling designs (Isaac & Pocock, 2015; Pocock *et al.*, 2017) but more  
93 data are typically collected in accessible areas such as near roads and urban areas (Geldmann *et*  
94 *al.*, 2016). Such biases are not unique to citizen science data, as even data collected during  
95 formal scientific studies have potential sampling biases towards regions undergoing less habitat  
96 change, which may lead to underestimates of biodiversity change (Gonzalez *et al.*, 2016; Forister  
97 *et al.*, 2023; Cardinale *et al.*, 2018). Various solutions have been proposed to deal with these  
98 biases (Hefley *et al.*, 2013; Cretois *et al.*, 2021; Johnston *et al.*, 2020; Ver Hoef *et al.*, 2021), but  
99 there is still a lack of a general framework for ecologists to guide decisions on when and how to  
100 deal with data gaps.

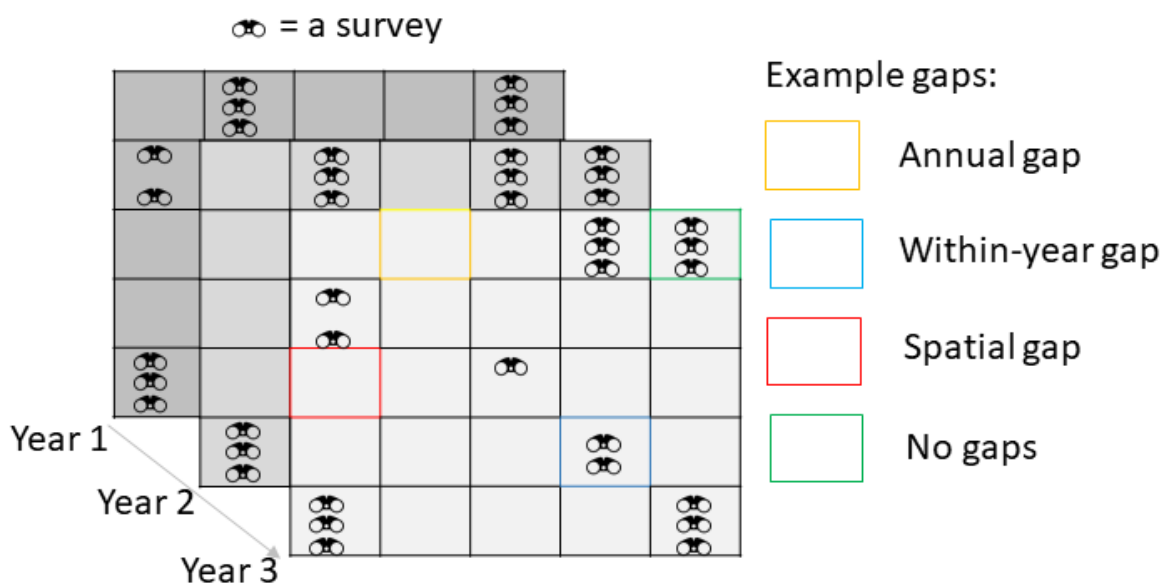
101           Here, we show how using missing data theory (Rubin, 1976) can unify problems  
102 associated with data gaps across different types of biodiversity datasets. Missing data are a  
103 widespread problem crossing disciplines, with a large body of literature on the implications and  
104 possible solutions (Little & Rubin, 2019; Carpenter & Kenward, 2012). We expect that aligning  
105 the generalized problem of missing data, conceptualized within missing data theory, to the  
106 problem of biodiversity data biases discussed above will yield opportunities so far overlooked.  
107 We mostly focus our review on modelling trends in species occupancy or abundance using  
108 monitoring data collected by volunteer citizen scientists, but the ideas transfer to other types of  
109 biodiversity data or questions. We show that bias is not a property of a dataset but rather a

110 property of the use of a dataset for a specific question and target population that are imposed by  
111 the data analyst. We review some commonly used solutions to missing data to highlight potential  
112 approaches that could be considered in biodiversity analyses.

## 113 **II. Classifying data gaps using missing data theory**

### 114 **(1) Biodiversity data gaps**

115 Species occurrence or abundance data can have gaps in different dimensions. We distinguish  
116 between spatial, annual and within-year gaps (Fig. 1). We define spatial gaps as those formed by  
117 sites with no data, and annual gaps as those formed by a lack of data in some years at sites that  
118 have been otherwise sampled. Together, spatial and annual gaps determine the spatial and  
119 temporal coverage of a dataset. Within-year gaps arise when data are lacking in specific seasons  
120 or months, which can be important because most organisms are seasonal and multiple visits are  
121 usually necessary to robustly estimate detection probabilities. Considering why these gaps arise  
122 can help understand their likely impact, for instance, on long-term species trend estimation. Data  
123 gaps are found in different types of monitoring data including highly structured monitoring  
124 schemes with a standardised protocol, such as many national bird survey schemes, as well as  
125 opportunistic monitoring data that are typically an aggregation of heterogeneous observations.  
126 Biodiversity datasets can also have taxonomic gaps (Troudet *et al.*, 2017), but this is outside the  
127 scope of this paper.



128

129 **Fig. 1 Different types of data gaps within biodiversity data.**

130 We imagine a scenario where there are multiple survey visits across sites and years. Visits can be in  
 131 response to a protocol ('structured' data) or opportunistic ('unstructured'), and repeat visits can be by the  
 132 same or multiple recorders. Data gaps, or more generally uneven data availability, can arise due to (a)  
 133 within-year gaps (e.g., blue square, i.e., ordinarily there are three visits, but some sites are only visited  
 134 once or twice in a year); (b) annual gaps (e.g., yellow square, i.e., some sites that are usually sampled are  
 135 entirely unvisited in some years) or (c) spatial gaps (e.g., red square, i.e., some sites within the region of  
 136 interest are never visited across all years). Some sites are well-sampled within and across years and hence  
 137 have no missing data (e.g., green square).

138

139 While both structured and opportunistic monitoring data can be affected by similar data  
 140 gaps (Binley & Bennett, 2023), there are some key differences between these types of  
 141 monitoring data. Moreover, structured schemes themselves vary in the degree of structure and  
 142 standardisation. In structured schemes with a formal spatial sampling design, data gaps include



143 both planned and unplanned gaps. Planned gaps arise because only a sample of sites was ever  
144 intended to be sampled. Unplanned gaps occur because of failure to recruit and retain surveyors  
145 at sites that were intended to be sampled (Zhang *et al.*, 2021; Marsh & Cosentino, 2019). In most  
146 other types of data, gaps are neither planned or unplanned. Some monitoring schemes have  
147 sampling protocols but participants are free to choose their own sampling sites. In fully  
148 opportunistic monitoring schemes, participants make individual decisions about where to sample  
149 and gaps emerge from unevenness in the cumulative sampling effort of all participants. Due to  
150 the high number of participants, and lack of coordination of their effects, sampling effort is  
151 generally more strongly skewed across space and time in opportunistic schemes than in  
152 structured schemes, leading to more pervasive data gaps (Geldmann *et al.*, 2016). Synthesis  
153 databases such as BioTIME and the Living Planet Database, and data aggregators such as GBIF,  
154 are similar in these respects to schemes without a formal spatial sampling design since they  
155 contain data that were independently collected as part of separate studies, without coordinated  
156 efforts.

157         Despite these differences, correlates of data gaps tend to be similar across monitoring  
158 schemes, especially those involving citizen scientists. Spatial gaps often occur in remote areas  
159 because there is a smaller pool of potential participants nearby (Geldmann *et al.*, 2016;  
160 Mandeville, Nilsen & Finstad, 2022). Spatial gaps can also be more common where species have  
161 lower abundance or land cover is perceived to be less attractive for biodiversity and for  
162 surveying e.g., agricultural land (Tulloch *et al.*, 2013; Dambly *et al.*, 2021; Marsh & Cosentino,  
163 2019). Annual gaps can arise due to project turnover or because of external factors (e.g. the 2020  
164 season for most countries was highly compromised by the Covid-19 pandemic). Annual gaps  
165 have also been linked with local land use changes that negatively affected species abundance and

166 the attractiveness of a site for sampling (Zhang *et al.*, 2021; Marsh & Cosentino, 2019). Within-  
167 year data gaps can be caused by periods of inclement weather (Zimney & Smart, 2022; Diekert  
168 *et al.*, 2023) or vary seasonally e.g., missing surveys for butterflies are more common at start and  
169 end of the main flight period (Dennis *et al.*, 2016), while bird sampling can be higher during  
170 their migration periods (La Sorte & Somveille, 2020).

171

## 172 **(2) Classes of missing data**

173 Within the classic missing data theory, there are three classes of missing data (Missing  
174 Completely at Random, Missing at Random, Missing Not at Random), defined below, each with  
175 different consequences for bias (Table 1) (Rubin, 1976; Nakagawa & Freckleton, 2008; Little &  
176 Rubin, 2019). These classes vary in their missing data mechanism, which describes the  
177 relationship between the probability of missing data (or sampling effort in the monitoring  
178 context) and the values of other variables. Hefley *et al.* (2013) already proposed viewing spatial  
179 biases in presence-only data as a form of missing data. Here, we extend it more broadly across  
180 different types of biodiversity data.

181         Within the context of biodiversity data, missingness can be regarded as Missing  
182 Completely at Random (MCAR) if the factors affecting biodiversity sampling, and causing  
183 missingness, are independent of those affecting biodiversity (Table 1). Under MCAR, the  
184 observed data are effectively a random sample of the whole population, and the values of the  
185 variable of interest are similar in sampled and non-sampled sites or times. For instance, if site  
186 selection is driven by human accessibility, but species distribution is primarily driven by climate,  
187 and if accessibility and climate are not correlated, then spatial data gaps would be MCAR.

188 Within-year gaps associated with weekdays (Evans & Day, 2002; Courter *et al.*, 2013), or annual

189 gaps associated with project turnover, are also examples likely to cause MCAR data gaps since  
190 such gaps are probably not associated with biodiversity patterns (Table 1). In this case, missing  
191 data could reduce the precision of parameters estimates through reduced sample size, but not  
192 increase the bias.

193         When the factors affecting sampling are the same as, or correlated with, those affecting  
194 biodiversity, the missing data mechanism can either be Missing at Random (MAR) or Missing  
195 Not at Random (MNAR). For instance, if road density affects both sampling probability and  
196 species abundance, then spatial gaps are not MCAR. Road density might affect sampling  
197 probability directly (e.g., if people are more often looking for wildlife along roads) or indirectly  
198 (e.g., if road density affects species detectability); in either case, road density influences data  
199 gaps. Similarly, habitat degradation could reduce both species abundance and observer retention  
200 to continue sampling at a site, creating an annual data gap that is MAR or MNAR (Table 1). In  
201 these cases, there are systematic differences in the biodiversity quantity of interest between  
202 sampled and non-sampled sites or times (Table 1).

203         To borrow from an infamous quote, if we regard data gaps as “unknowns”, then MAR  
204 can be thought of as “known unknowns” while MNAR are “unknown unknowns”. The “known”  
205 needed for MAR is knowledge and availability of data on the shared covariates affecting  
206 sampling and biodiversity. If complete data for shared covariates are available and included in  
207 the analysis, then the missing data mechanism is MAR. Hence, despite its name, MAR does not  
208 mean that sampling effort is randomly distributed in the landscape. Rather, it means that the  
209 covariates affecting sampling are known and that there is available covariate data to fully explain  
210 the differences between sampled and non-sampled potential data. If any of the relevant factors  
211 affecting sampling and biodiversity are unknown, or not modelled, the missing data mechanism

212 becomes MNAR (Table 1). Hence, decisions of the analyst can determine whether a data gap is  
213 MNAR or MAR (discussed more fully in section III). MNAR may also arise when missingness  
214 is dependent on the value of biodiversity itself, i.e., if sampling effort directly depends on species  
215 occurrence or abundance.

216           Statistical tests can only partly indicate which missing data class is most likely (Little,  
217 1988). Analysis of relationships between data availability and observed covariates can point  
218 towards MAR if some relationships are significant. But a lack of any association, or an  
219 incomplete explanation of data gaps, could reflect MCAR or MNAR. Because MNAR is  
220 associated with unavailable data, it cannot be directly tested. Concerns about whether  
221 missingness in the biodiversity data is directly associated with its values could be explored if  
222 there is a related variable that is fully available (Wu, 2022). We argue that MCAR is unlikely in  
223 most biodiversity data since unplanned data gaps can affect even the most structured monitoring  
224 schemes.

225

226 **Table 1 Missing data mechanisms in biodiversity data, including examples and implications**

Mechanism	Typical meaning	Meaning in the context of biodiversity data	Examples	Typical implications
Missing completely at random (MCAR)	Missingness is independent of observed and unobserved variables.	Sampling is independent of any covariates, or covariates that affect sampling probability are independent of those affecting biodiversity	Within-year: Weekday gaps Annual/Spatial: Gaps caused by the completion of a fixed-term project or retirement of a participant	Ignorable
Missing at random (MAR)	Missingness is associated with observed data but not any unobserved variables	Covariates that affect sampling probability are shared with those affecting biodiversity, but data are available on all these covariates	within-year: Season (day of year) Annual: Urban development Spatial: Accessibility	Ignorable if the model includes all relevant covariates
Missing not at random (MNAR)	Missingness depends on unobserved variables or the missing values itself	CS sampling varies with biodiversity value or an unknown or unavailable covariate affects sampling and biodiversity	within-year/annual/spatial: unknown factors causing variation in species activity/abundance that are also correlated with sampling effort	Non-ignorable - the missing data mechanism needs to be modelled

227 **III. Implications of missingness for ecological questions**

228 Missing data (i.e., data gaps) themselves do not necessarily have strong impacts on the results of  
 229 biodiversity modelling, but can depend on the specific question and parameter of interest  
 230 (Bartlett, Harel & Carpenter, 2015; Collins, Schafer & Kam, 2001; Little *et al.*, 2022). Viewing  
 231 data gaps as a form of missing data can help decide whether a particular data gap matters. As we

232 note above, data gaps that are MCAR do not cause bias, but data gaps in biodiversity data are  
233 unlikely to be wholly MCAR. For a data gap to be MAR rather than MNAR can depend on the  
234 ecological question being pursued by an analyst. This is because the ‘missing at random’  
235 assumption of MAR is conditional on controlling for covariates affecting sampling probability,  
236 which means that these covariates are known, reflected in available data and included in the  
237 analysis (Fig. 2) (Conn, Thorson & Johnson, 2017; Hefley *et al.*, 2013). Different ecological  
238 questions will lead to different decisions about which variables to include in an analysis. Hence,  
239 data gaps of the same dataset might be MAR under some questions but MNAR under others. To  
240 illustrate these potential differences, we contrast two typical questions asked with biodiversity  
241 data.

242

### 243 **(1) Understanding the roles of environmental drivers on species’ distributions**

244 Monitoring data are often used to understand the environmental factors explaining species  
245 distribution patterns. The implications of missing data for species distribution models have been  
246 considered in terms of niche truncation. Niche truncation happens when a dataset only contains  
247 occurrence data from part of the geographic range of a species, which usually also means that the  
248 dataset only covers part of the ecological/environmental space that is suitable for the species  
249 (Chevalier *et al.*, 2022; Albert *et al.*, 2010; Guo *et al.*, 2023). These studies show that the  
250 implications of niche truncation depend on the functional form of the relationship between the  
251 associated covariate and the species response (Chevalier *et al.*, 2022) and whether occurrence  
252 data are presence/absence or presence-only (Baker *et al.*, 2022).

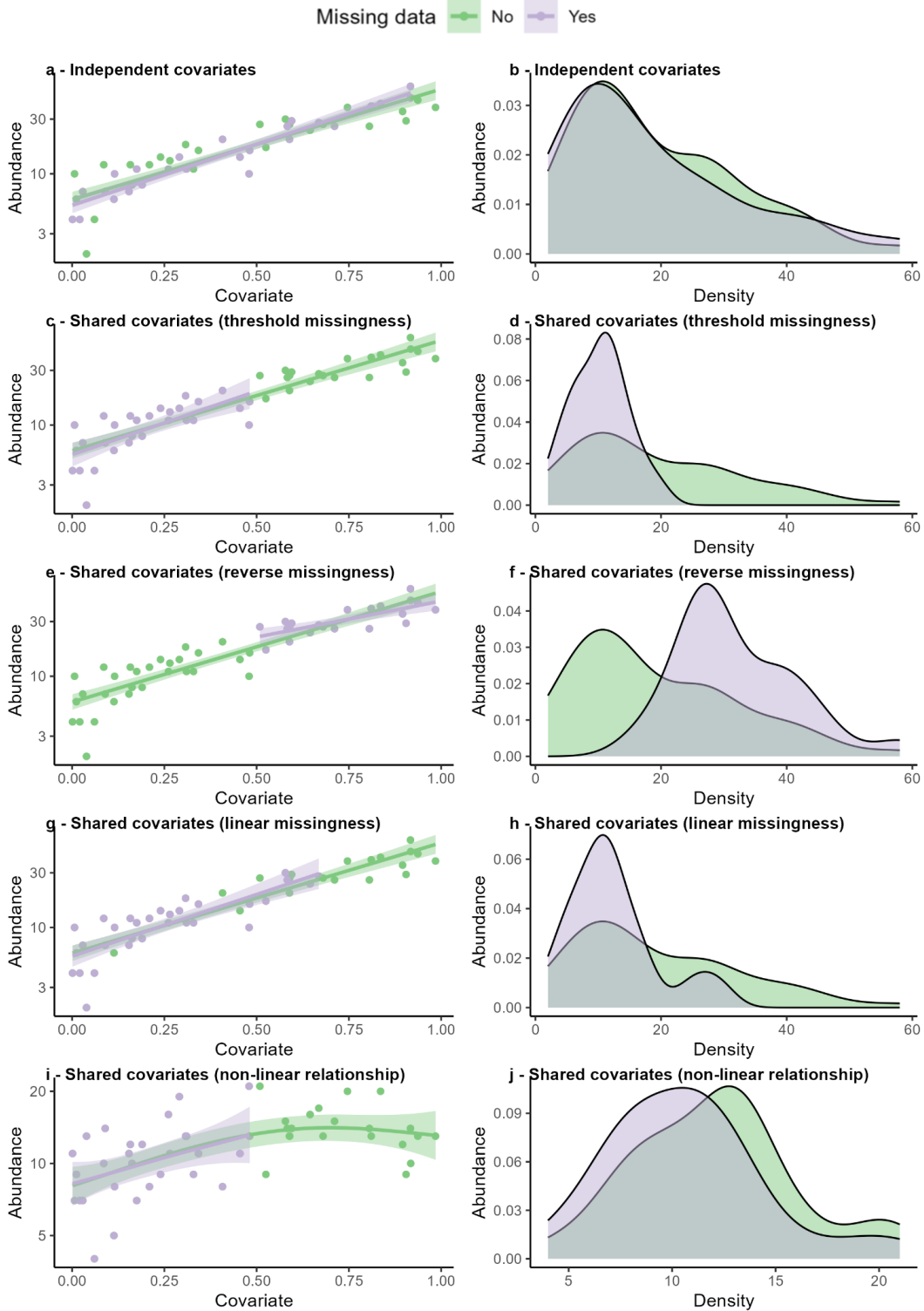
253 We begin considering the scenario when presence-absence data are available. In this case,  
254 if there is a simple linear relationship, missing data do not necessarily cause bias in the estimated

255 effect of the covariate on biodiversity, even when missingness depends on the same covariate  
256 (Fig. 1a, c, e, g) (Collins *et al.*, 2001). For instance, we could estimate the effect of elevation on  
257 species occurrence, if it is linearly related, even if elevation is also associated with data gaps.  
258 This is because the relationship between the covariate and species occurrence can be estimated  
259 without bias using data over a restricted range of covariate values. This is shown in e.g., Fig. 1c -  
260 the same relationship is found with a full dataset (green) or a restricted dataset with data gaps  
261 (purple). Missing data can, however, cause problems when the underlying relationship between  
262 the covariate and species occurrence is non-linear. In this case, data gaps can hinder estimating  
263 the true form of the relationship (see Fig. 2i - a curved relationship in fit with the full dataset but  
264 a simple positive linear relationship with the restricted dataset). The relationship that is fit using  
265 the restricted dataset will critically depend on which portion of the covariate range is sampled.  
266 Since many ecological associations show some non-linearity, or context-dependencies such that  
267 relationships depend on the value of other variables (Spake *et al.*, 2023), we expect this issue is  
268 likely to be widespread in species distribution models.

269 We now consider the alternative scenario of fitting a distribution model with presence-  
270 only data. In this case, any data gaps could represent a lack of sampling or a lack of true species  
271 occurrence. This creates an inherent identifiability challenge for any model seeking to separate  
272 the processes affecting sampling from the true ecological processes affecting species  
273 distributions with presence-only data (Hefley *et al.*, 2013; Baker *et al.*, 2022). Many methods  
274 have been developed to generate pseudo-absences (Barbet-Massin *et al.*, 2012; Hertzog, Besnard  
275 & Jay-Robert, 2014), but such models are still usually more prone to biases when there are  
276 shared covariates affecting sampling and species occurrence (Baker *et al.*, 2022). More recent

277 approaches to modelling presence-only data, by integrating them with any available presence-  
278 absence data (Fithian *et al.*, 2015), may help minimise some of these biases.

279



280



281 **Fig. 2 The impacts of different missing data mechanisms on regression (left) and sample**  
282 **distributions (right).**

283 We use a hypothetical dataset to highlight different missing data mechanisms. In (a) and (b), the covariate  
284 affecting sampling probability is *independent* from the covariate affecting species abundance. In this case,  
285 both the estimated effect of the covariate (e.g., in a linear regression, shown in a by the solid line) and the  
286 sample distribution (b) are similar in a dataset with (purple) and without (green) missing data. (i.e.,  
287 missingness is MCAR). In (c) and (d), the covariate affecting sampling probability is the *same as or*  
288 *correlated with* the covariate affecting species abundance - in this case, data are missing when the  
289 covariate is above average (i.e., threshold missingness). The estimated effect of the covariate is the same  
290 in the dataset with and without missing values (shown in c) but the sampling distribution is different (d).  
291 In (e) and (f), the missingness pattern is reversed compared to (c) and (d) (i.e., data are missing when the  
292 covariate is below average), but we can similarly retrieve the same unbiased covariate effect (e) even  
293 though there is greater mean abundance in the dataset with missing values (f). In (g) and (h), the covariate  
294 affecting sampling probability is the *same as or correlated with* the covariate affecting species abundance  
295 - in this case, the probability of missing data increases with the value of the covariate (i.e., linear  
296 missingness). Again, the estimated effect of the covariate is the same (shown in g) but the sampling  
297 distribution is different (h). In (i) and (j), the covariate affecting sampling probability is the *same as or*  
298 *correlated with* the covariate affecting species abundance; additionally, the true relationship between the  
299 covariate and species abundance is non-linear and data are missing when the covariate is above average.  
300 The mechanism is now MNAR since the model cannot be correctly specified with the observed data.

301

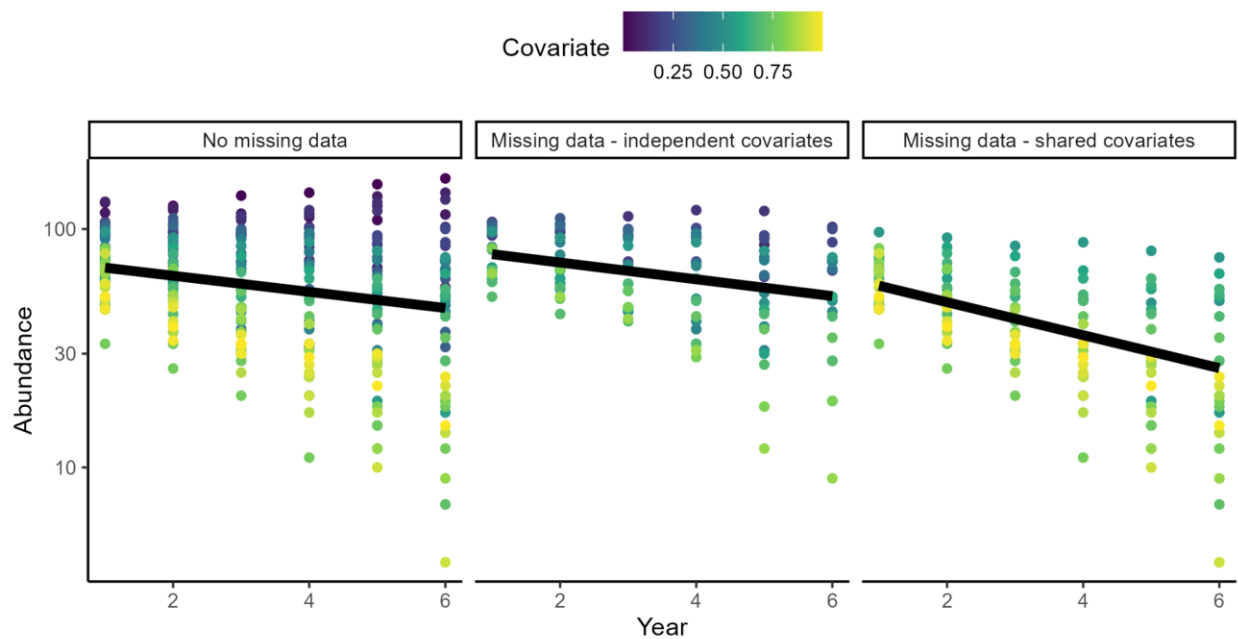
302 **(2) Estimating trends in species abundances**

303 Models to estimate species' trends tend to be descriptive: spatial variation is modelled by  
304 including site identity (as a fixed or random term) while any temporal trend is modelled as a  
305 simple year effect (either as a linear, spline or a categorical term) (Amano *et al.*, 2012; Bled *et*

306 *al.*, 2013). Drivers of the trend are not explicitly modelled when the goal is to simply estimate  
307 the mean trend over time. Broader inferences about the trend estimated by such models are based  
308 on the assumed representativeness of the sample, or prior knowledge of the inclusion  
309 probabilities of sampling units (see design weights discussed in section IV 2). Basing inference  
310 from the sampling design is the most traditional approach to surveys (Smith, 1976) and the  
311 approach typically taken by official governmental surveys (van den Brakel & Bethlehem, 2008).  
312 This approach has the advantage of avoiding complex assumptions in the statistical analysis  
313 (Buckland *et al.*, 2012) and is perhaps also easier to analyse and communicate to stakeholders  
314 and laypersons.

315         Simple trend models may, however, lead to biased trend estimates for biodiversity when  
316 data gaps are not MCAR. We illustrate this in a simple simulation in which site-level species  
317 trends were assumed to depend on a site-level covariate e.g., urban cover (Fig. 3). We assumed  
318 sites were sampled either with a probability affected by an independent covariate (Fig. 3 middle  
319 panel) or with a probability affected by the same site-level covariate affecting species trends  
320 (Fig. 3 right panel), a scenario already identified as a pitfall in some monitoring schemes  
321 (Buckland & Johnston, 2017). We estimated trends using a simple mixed effect model including  
322 site and year. This shows that when the site-level covariate affected both sampling effort and  
323 species' trends, the trends were biased, but site-level trends were unbiased when an independent  
324 covariate affected sampling. In real world situations, many factors will influence the trend of a  
325 species, but this toy simulation highlights the potential for bias caused by shared covariates.  
326 Since the specific covariates affecting sampling effort and biodiversity trends are not considered  
327 in the typical forms of analysis for trend modelling, trend analyses are liable to be affected by  
328 MNAR, whereas by including appropriate covariates (where possible), the data gaps become

329 MAR instead and trends will be unbiased. Without conditioning on the covariates involved, trend  
 330 estimates might be overestimated if missing data are more common in static regions where  
 331 species trends are more stable; but underestimated if missing data are more common in dynamic  
 332 regions where species trends more strongly deviate from zero (Fig. 3) (Bowler *et al.*, 2022;  
 333 Buckland & Johnston, 2017).



334

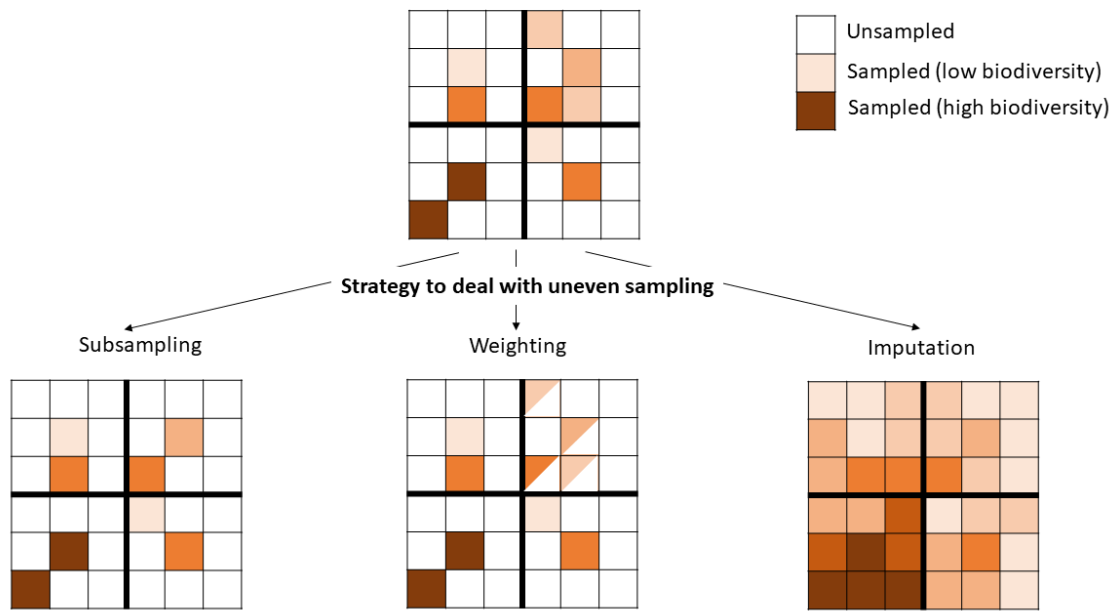
335 **Fig. 3 The impacts of different missing data mechanisms on trend modelling**

336 We use a hypothetical scenario in which a mean trend model is fit to datasets that vary in their missing  
 337 data mechanism. We assumed a scenario of 50 sites that varied in an environmental covariate affecting  
 338 species trends (trends were stable or even increasing at low values of the covariate and declining at  
 339 increasingly high values of the covariate). When missing data was independent (i.e., a MCAR pattern -  
 340 the covariate affecting sampling probability was a different and uncorrelated covariate), the overall mean  
 341 trend (estimated by the year effect in a generalized linear mixed effect model that also included a site  
 342 random effect) was similar with (middle panel) and without (left panel) missing data. By contrast, when  
 343 the same covariate affected both species' trends and sampling probability, leading to less sampling in sites

344 with low values of the covariate (notice there are fewer blue points in the right panel - a MNAR pattern),  
345 the overall mean trend was downward biased with missing data (right panel) compared to the scenario of  
346 no missing data (left panel).

## 347 **IV. Missing data solutions**

348 A broad range of methods to deal with missing data have been used in ecology (Hossie, Gobin &  
349 Murray, 2021; Nakagawa & Freckleton, 2008; Lopucki *et al.*, 2022). Many solutions are  
350 particularly relevant when data are missing in both response and predictor variables. Here, we  
351 focus on the typical scenario in biodiversity modelling of missing data only in the response  
352 variable (i.e., in the biodiversity data) since typical predictors tend to have no or few gaps (e.g.,  
353 site identity or environmental data from remote sensing). We organise solutions into three groups  
354 - subsampling, weighting and imputation (Fig. 4), which have been tested to varying degrees  
355 already with both structured and unstructured biodiversity data (Table 2). Most solutions to deal  
356 with missing data are only appropriate for MCAR or MAR missingness. MNAR is the most  
357 challenging missing data mechanism to deal with in statistical modelling so, we deal with  
358 MNAR in a later section.



359

360 **Fig. 4 Visualisation of contrasting approaches to deal with data gaps.**

361 We focus on spatial gaps to illustrate the possible approaches, but the ideas apply to other types of data  
 362 gaps (Fig. 1). (top) the landscape is divided into four quarters (e.g., representing different habitats or  
 363 geographic regions). One quarter (top right quarter) has been sampled more (4 sampling sites) than the  
 364 others (2 sampling sites). Solutions: Random subsampling (bottom left): two sites are randomly  
 365 subsampled from the oversampled quarter to create a dataset with an even sampling coverage across  
 366 quarters. Weighting (bottom middle): data from the oversampled quarter is downweighted in the  
 367 statistical model so data from all quarters similarly influence the modelled results. Imputation (bottom  
 368 right): missing values at unsampled sites are imputed based on the spatial pattern in the data and/or  
 369 environmental covariates, and summary parameters are calculated based on both predictions at sampled  
 370 and unsampled sites. In subsampling and weights, the aim is to improve the representativeness of the  
 371 sample for statistical inference at the population-level. In imputation, the aim is to directly predict  
 372 population-level values.

373

374 **Table 2 Example applications of the solutions to deal with data gaps within biodiversity data.**

Type of data	Typical approaches:
<b>gaps</b>	
Within-year	Sometimes imputed e.g., spline terms to smooth over seasonal variation in sampling times during the flight period of butterflies (Dennis <i>et al.</i> , 2016)
Annual	Sometimes imputed e.g., general linear models to impute annual gaps based on mean site and year effects, optionally allowing for habitat effects e.g., used in TRIM abundance indices, (Lehikoinen <i>et al.</i> , 2016)
Spatial	Often ignored, but occasionally weighting by geographic regions (Bled <i>et al.</i> , 2013) or imputed (Breivik <i>et al.</i> , 2021), or reduced by subsampling (Johnston <i>et al.</i> , 2021).

375

### 376 **(1) Subsampling**

377 The ‘Big Data Paradox’ highlights that there can be trade-offs between dataset size and dataset  
 378 quality (Bradley *et al.*, 2021; Meng, 2018). Small datasets can be preferable to large datasets, if  
 379 they are more representative and less heterogeneous than a larger dataset (Bayraktarov *et al.*,  
 380 2019). Based on such thinking, some studies have proposed to ‘reverse engineer’ structure in  
 381 biodiversity data by filtering data (Rapacciuolo, Young & Johnson, 2021). Part of this reverse  
 382 engineering has attempted to deal with spatial biases; for instance, by spatially subsampling data  
 383 to reduce the unevenness of sampling effort across the landscape (Steen *et al.*, 2021; Matutini *et*  
 384 *al.*, 2021; Steen, Elphick & Tingley, 2019; Boria *et al.*, 2014; Robinson *et al.*, 2020). This has  
 385 been tested on, for instance, the semi-structured data compiled by eBird (Johnston *et al.*, 2020).  
 386 Some have also applied this approach to reduce temporal skews in sampling effort (Hof &  
 387 Bright, 2016; Zbinden *et al.*, 2014), although not always successfully (Callcutt, Croft & Smith,  
 388 2018). Subsampling can also be used to balance the amount of data across a single or multi-  
 389 dimensional environmental gradient; essentially stratified sampling of the original sample

390 (Meng, 2022; Nunez-Penichet *et al.*, 2022). Recent class balancing approaches have been  
391 developed to ensure that important observations, especially for rare species, are not lost during  
392 the subsampling process (Robinson *et al.*, 2020; Steen *et al.*, 2021; Gaul *et al.*, 2022).

393

## 394 **(2) Weighting**

395 Weighting is a common practice in survey analysis, especially in the social sciences (Li *et al.*,  
396 2013; Seaman & White, 2013; Raghunathan, 2004). Weighting can serve different purposes,  
397 including reducing the impact of confounding variables when the goal is to estimate the causal  
398 effect of an intervention. But weighting can also be used to deal with missing data that is not  
399 MCAR. For instance, weighting can be used to reduce selection bias caused by participant  
400 nonresponse in surveys (Seaman & White, 2013), but it is less often used to account for data  
401 gaps in biodiversity data (Boyd, Powney & Pescott, 2023a; Aubry & Francesiaz, 2022).

402 Different types of weights have been used in the analysis of biodiversity data: (1) design  
403 weights; (2) estimated non-response weights (or sampling weights) and (3) population weights.  
404 Each form of weighing is intended to improve sample representativeness of some target  
405 population but vary in terms of whether the weights derive from the sampling design and the  
406 dimension of representativeness under consideration. Design weights are based on the study  
407 sampling design and assumed to be known with certainty, and hence are only relevant for  
408 structured schemes with a sampling design. For instance, in many national bird breeding  
409 schemes, the design weights are based on the geographic strata that underlie a random stratified  
410 study design (Buckland *et al.*, 2012). Non-response weights are used to account for unplanned  
411 missing data in structured schemes (Frair *et al.*, 2004) or variation in sampling effort in  
412 unstructured schemes (Johnston *et al.*, 2020; Hefley *et al.*, 2013), which means that are not

413 known with certainty and must be estimated. Population weights are used to ensure the sample is  
414 representative of the full distribution/population of a species and are typically assumed to be  
415 known. Population weights are used in the calculation of supranational/international indicators  
416 (e.g., farmland or woodland bird indicators (Gregory *et al.*, 2005)) in which national estimates  
417 are combined by giving greater weight to regions that harbour a larger proportion of the species  
418 total population.

419         Non-response weights are usually the most difficult to include since they are not known *a*  
420 *priori* and need to be estimated. Predictive models (e.g., random forest models) have been used  
421 to predict the probability that a site is sampled based on the set of covariates (e.g., land cover or  
422 climate, or accessibility) available across all sampled and unsampled sites, with the inverse of  
423 these probabilities used as weights (Little *et al.*, 2022; Johnston *et al.*, 2020). Alternatively,  
424 poststratification (for categorical covariates), or more generalized calibration approaches  
425 (allowing both continuous and categorical covariates), can be used, which adjust the weight  
426 given to each data point until the joint or marginal distributions of covariate values in the  
427 observed sample matches those for the population (Boyd, Stewart & Pescott, 2023b). In both  
428 cases, weighting can cause problems when there are regions within the target population with  
429 close to zero probability of being sampled, which could lead to some data points having  
430 extremely large weights. In this case, weights may need to be redefined e.g., by coarsening the  
431 covariates used to define the weights, or by truncating weight values so that extreme weights are  
432 not produced (Battaglia, Hoaglin & Frankel, 2009). Poststratification can be preceded by multi-  
433 level regression (for so-called “Mr P” analysis) for partial pooling of information across strata  
434 before poststratification of the model predictions, which may be especially useful when some  
435 strata contain few data points (Gelman, 2007).



436           The most appropriate approach is likely to be question- and taxon-specific, varying with  
437 how much the species range extends across the region of interest. For example, it would usually  
438 not be important to upweight under-sampled regions where a species is rare, or even absent,  
439 when estimating trends in its total population size. If, however, the goal is to estimate trends in  
440 the average site-level population trend, then it would be important to up-weight data from under-  
441 sampled regions, even from where the species is rare. For instance, in the UK bat monitoring  
442 scheme, data are weighted to allow for the different sampling rates across England, Scotland and  
443 Wales in proportion to the ratio of non-upland area to number of sites surveyed for the relevant  
444 country (Bat Conservation Trust, 2023). However, this weighting is not applied to range  
445 restricted species, such as the serotine bat, *Eptesicus serotinus* that is only found in southern  
446 England.

447

### 448 **(3) Imputation**

449 Imputation involves replacing missing values in a dataset with plausible estimates. A range of  
450 imputation procedures have been developed, which can fill gaps in both response and predictor  
451 variables (Carpenter & Kenward, 2012). Imputation is probably the most flexible and widely  
452 used approach to account for missing data across ecology and beyond. In biodiversity modelling,  
453 missing values are more often concentrated in the response variable (i.e., the biodiversity value),  
454 hence imputation here can be equated with making model predictions at unsampled sites and  
455 times.

456           Imputation is already in use in biodiversity trend monitoring, especially to account for  
457 within-year and annual data gaps (Table 3). Early approaches used chain indices or route  
458 regression (Ter Braak *et al.*, 1992) or the Underhill index, using an expectation-maximisation

459 algorithm (Underhill & PrysJones, 1994) designed for waterbirds (Rehfishch *et al.*, 2003). A range  
460 of further model-based approaches have been developed that fill data gaps using mean effects of  
461 site and year, e.g., to fill annual gaps using TRIM/birdSTATs, commonly used for bird indices  
462 (Lehikoinen *et al.*, 2016); or using splines e.g., to fill seasonal gaps in butterfly data (Schmucki  
463 *et al.*, 2016; Dennis *et al.*, 2016) or using ecological covariates (Dakki *et al.*, 2021). A Bayesian  
464 framework is especially useful for dealing with missing values in the response since they are  
465 naturally imputed with a full probability distribution during model fitting. Bayesian occupancy-  
466 detection models have been used to analyse opportunistic species observations from citizen  
467 science, with annual data gaps imputed before the predicted annual proportion of occupied sites  
468 is calculated (Outhwaite *et al.*, 2019). The flexibility of Bayesian models means they could also  
469 incorporate expert knowledge as priors as a way to help fill data gaps (Johnson *et al.*, 2023).

470         While imputation is already used to deal with annual and within-year gaps, it has been  
471 less often used to deal with spatial gaps when the focus is mean trend modelling of species'  
472 abundances or occurrences. An exception is studies of changes in species' range sizes, which use  
473 distribution models to predict the full distribution of a species at multiple time points, before  
474 change is assessed (Grattarola, Bowler & Keil, 2023). Monitoring schemes with large spatial  
475 coverage are also beginning to use distribution or abundance models to predict spatio-temporal  
476 patterns of abundance change across whole countries (e.g., [eBird maps](#) and [BTO maps](#)). In these  
477 cases, regression models fit to the available data make predictions at unsampled sites based on  
478 the effects of environmental covariates and/or spatial structure (Bush *et al.*, 2017; Ver Hoef *et al.*,  
479 2021; Breivik *et al.*, 2021). Geostatistical methods also offer a range of interpolation  
480 methods for spatial data, including kriging, which are especially useful when there is a strong  
481 spatial pattern in the data (Ballesteros-Mejia *et al.*, 2013; Kreft & Jetz, 2007; Lin *et al.*, 2008).

482 **V. Pro and cons of each solution**

483 **Table 3 Summary of the pros and cons of each approach to deal with missing data in biodiversity**  
 484 **monitoring**

Solution	Pros	Cons
Subsampling	<ul style="list-style-type: none"> <li>- arguably the simplest approach, especially for spatial gaps</li> <li>- already a routine feature of many species distribution modelling protocols</li> <li>- aligns with rarefaction approaches used in community ecology</li> </ul>	<ul style="list-style-type: none"> <li>- could mean excluding a large amount of data, which may be unacceptable for citizen science and engaging/retaining volunteers</li> <li>- most protocols focus on a single dimension (e.g., filtering by geographic region)</li> <li>- more complex to implement when gaps are multi-dimensional or temporally varying</li> </ul>
Weighting	<ul style="list-style-type: none"> <li>- standard practise to deal with sample unrepresentativeness in other disciplines, especially social sciences</li> </ul>	<ul style="list-style-type: none"> <li>- poorly understood in ecology</li> <li>- diverse range of possible weighting techniques (Valliant, 2020; Boyd <i>et al.</i>, 2023b) but little ecological guidance available to help selection</li> </ul>
Imputation	<ul style="list-style-type: none"> <li>- suitable approach if missing data are within the environmental covariates as well as within the biodiversity response</li> <li>- offers the promise to generate the continuous space-time data cubes of the Essential Biodiversity Variable framework (Kissling <i>et al.</i>, 2018; Jetz <i>et al.</i>, 2019).</li> </ul>	<ul style="list-style-type: none"> <li>- becomes inefficient as missingness increases, e.g., when the number of unsampled locations/times is large</li> <li>- requires a good understanding of the ecological system to predict the missing biodiversity values</li> </ul>

485

486 All of the approaches have the potential to reduce the bias in parameter estimates but

487 differ in complexity, scope and typical practice (Table 3) (Little *et al.*, 2022; Collins *et al.*,

488 2001). Moreover, while we separated the methods into three categories for convenience, their  
489 distinctions are not absolute. For instance, subsampling essentially assigns those population units  
490 included in the subsample a weight of 1 and the remainder a weight of 0. Often, but not always,  
491 the reduction in bias due to application of the above solutions comes at a cost of increasing  
492 parameter uncertainty: the classic bias-variance trade-off (Hefley *et al.*, 2013). This is because  
493 subsampling directly reduces the sample size; weighting reduces the effective sample size; and  
494 imputation adds uncertainties via predictions at unsampled points. But this trade-off does not  
495 always apply; for instance, poststratification can lead to the dual benefits of reduced bias and  
496 increased precision depending on the choice of covariates (Little & Vartivarian, 2005).

497         Covariates used to account for data gaps are often called ‘auxiliary variables’ (Little *et*  
498 *al.*, 2022), which are typically not of central interest to the scientific questions but are included in  
499 one or more of the analysis steps for subsampling, weighting or imputing. The general  
500 recommendation from the missing data theory and survey sampling literature is to be generous  
501 when deciding which covariates to use to adjust for data gaps, considering covariates relating to  
502 the missingness (i.e., sampling effort in the context of biodiversity data gaps) to reduce bias and  
503 those related to the biodiversity outcome to reduce the variance (Collins *et al.*, 2001; Caughey *et*  
504 *al.*, 2020). It is worth noting, however, that selecting auxiliary variables on a purely correlative  
505 basis can increase bias in some circumstances (Thoemmes & Rose, 2014), and a safer strategy is  
506 to select them on theoretical grounds (Mohan & Pearl, 2021). When auxiliary variables are  
507 related to both the biodiversity outcome and the pattern of missingness, weighting approaches  
508 can reduce bias and improve precision (Little & Vartivarian, 2005). The success of any of the  
509 solutions, hence, critically depends on the choice of auxiliary variables (Little *et al.*, 2022). A  
510 recent study testing the use of weighting approaches to account for spatial biases in a reasonably

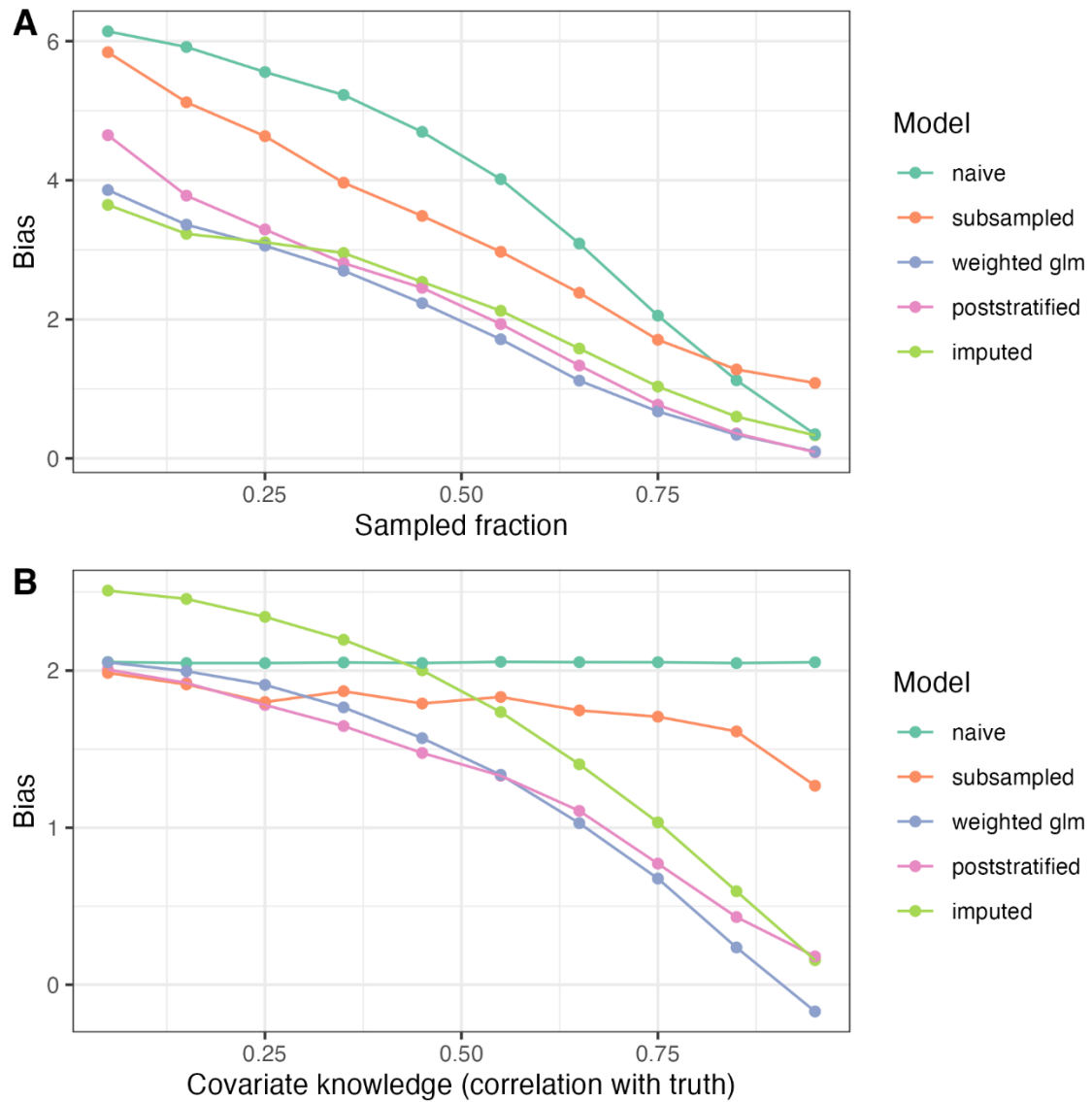
511 well-understood ecological system found that the selected auxiliary variables had only limited  
512 success in mitigating bias (Boyd *et al.*, 2023b).

513         We illustrate some of these challenges and the application of each potential solution with  
514 a toy example of an abundance dataset with missing values (Fig. 5). We simulated a landscape in  
515 which a covariate (let's say representing 'habitat quality') affected both species abundance and  
516 the likelihood of a site being sampled. The analysis aimed to estimate the mean abundance of the  
517 species across all sites in the landscape. We varied the total fraction of sites that were sampled  
518 and the degree of knowledge available on the covariate affecting sampling/species (modelled as  
519 the correlation between the covariate involved in the data generation process and the covariate  
520 available to the modeller). We compared subsampling, weighting and imputation, which all used  
521 the available covariate data for adjustment. For subsampling, we subsampled one site at random  
522 at each habitat quality value. For weighing, we compared two approaches: fitting a weighted  
523 regression model using model-robust sandwich variance estimators or using a poststratification  
524 approach. For imputation, we fit a Bayesian model using JAGS in which NA values were  
525 inserted to represent the missing response data.

526         The results show that all methods do better than a naive approach that did not attempt to  
527 account for missingness in the estimation of the mean abundance (Fig. 5). Subsampling  
528 performed the worst, while weighting and imputation performed similarly. Poststratification  
529 tends to perform less well with a lower sampling fraction i.e., when the number of missing  
530 values was high (Fig. 5A), because the sample did not always contain all the habitat quality  
531 values found in the population and the weighing could not account for entirely unsampled  
532 regions. All models performed less well at the available covariate became a weaker proxy of the  
533 true driving covariate (Fig. 5B). In further simulations, we found that imputation performed

534 poorer when there were additional covariates affecting species abundance and these covariates  
535 were not modelled, highlighting the importance of understanding the ecological system for  
536 imputation (Fig. S1). We do not intend this simulation to be exhaustive - rather to highlight the  
537 potential ways in which the availability of data and degree of knowledge about the factors  
538 causing bias and the availability of covariate data affects any attempts to account for missing  
539 data.

540 We point the reader towards some useful R packages and functions in the Supporting  
541 Information (Table S1).



542

543 **Fig. 5 The ability of missing data solutions to adjust for bias in biodiversity data.**

544 We assumed a landscape of 400 cells and that a covariate affected both species abundance and the likelihood of a  
 545 cell being sampled. In A: we vary the fraction of the cells that were sampled. In B: we vary the correlation between  
 546 the true covariate and the covariate available for analysis, as measure of the available knowledge (correlation of 1 =  
 547 perfect covariate and knowledge). The models to estimate the parameter of interest (mean abundance) were: naive  
 548 (no correction); subsampled (cells were subsampled along the covariate gradient), weighted (two methods: weighted  
 549 glm using the svyglm function, and weighted by poststratification, using postStratify, both in the survey package)

550 and imputed (using JAGS to impute NAs in the response). Points show the mean bias (difference between model  
551 prediction and truth) across 100 independent runs.

552

## 553 **VI. Dealing with Missing Not at Random**

554 Dealing with Missing Not at Random (MNAR) is more challenging than dealing with the other  
555 data mechanisms (Little & Rubin, 2019). In this case, missingness is directly associated with  
556 unavailable data, which could be either the missing biodiversity values or missing covariate data  
557 that are not known to be important or are not measured/measurable. This makes MNAR  
558 especially difficult to diagnose (but see Conn et al. (2017) for suggestions) and model, since  
559 auxiliary variables are not available. MNAR can arise through a number of mechanisms in  
560 biodiversity monitoring data.

561 MNAR can be an outcome of preferential sampling - more intense sampling effort where  
562 the species is expected (Diggle, Menezes & Su, 2010; McClure & Rolek, 2023) - which leads to  
563 more missing values in places where the species is rare or absent. Preferential sampling can  
564 arise, for instance, if observers visit a location to specifically observe a species that others have  
565 observed there before (Laney *et al.*, 2021; Pennino *et al.*, 2019). Preferential sampling can also  
566 be a planned sampling strategy (Alessi *et al.*, 2023). For rare species, preferential sampling can  
567 be optimal when the goal is to estimate species detection probability and account for imperfect  
568 detection, since sufficient observations of the species can only be achieved by sampling where  
569 they are more common (Specht *et al.*, 2017). Similarly, it can be optimal to expend greater  
570 sampling effort where the species is common if the goal is to estimate trends in the total  
571 population size, since regions where the species is scarce are less important for the overall trend.



572 For organisms associated with specific habitats, such as wetland species or colonial seabirds,  
573 dedicated structured monitoring schemes target their habitats (McClure & Rolek, 2023). In such  
574 schemes, missing data outside of these core habitats are not considered part of the target  
575 population.

576 Typical approaches to modelling data allowing for MNAR are selection models  
577 (Heckman, 1979) and pattern-mixture models (Herzog and Rubin, 1983). Both model the joint  
578 distribution of the data and the data availability, but differ in how these processes are  
579 decomposed. Both also require making strong assumptions about the missing data mechanism,  
580 but can be useful to explore the consequences of plausible options as a sensitivity analysis  
581 (Little, 1995). In the ecological literature, preferential sampling has been modelled using marked  
582 point process models, which jointly model the sampling intensity (the points), the biodiversity  
583 value at those points (the marks) and the dependence between them (Conn *et al.*, 2017; Pennino  
584 *et al.*, 2019; Laxton *et al.*, 2023). Another approach to inference in a NMAR scenario is to use  
585 instrumental variables i.e., variables that affect the probability of sampling/data availability but  
586 are independent of the biodiversity variable of interest (Tchetgen & Wirth, 2017; Bailey, 2023).  
587 The challenge, however, is to identify such variables.

## 588 **VII. General guidelines for dealing with biodiversity data**

### 589 **gaps**

590 Our review highlights the potential value of ‘missing data thinking’ when analysing biodiversity  
591 data. We argue that MCAR data gaps are unlikely in most biodiversity data contexts, which  
592 means that researchers will need to consider whether and how they deal with data gaps in their

593 analysis. While it is premature to make very specific guidelines, we summarise here some of the  
594 considerations needed when dealing with data gaps in biodiversity data at different stages of data  
595 collection, analysis and reporting.

## 596 **(1) Study design**

597 For new monitoring schemes, planned data gaps that deviate from MCAR (i.e., a random  
598 sample) can be seen as opportunities rather than challenges since solutions are available to deal  
599 with missing data. Intentionally missing some data has been proposed for ethical or practical  
600 reasons in some study designs e.g., (Noble & Nakagawa, 2021; Herrera, 2019). In citizen  
601 science, planned data gaps could help increase uptake and avoid participant fatigue, especially  
602 caused by collecting difficult data. For instance, the UK Breeding Bird Survey includes an  
603 ‘upland rovers’ component in which the standard protocol is modified to allow for fewer visits to  
604 remote sites (Darvill *et al.*, 2020). Alternative study designs, such as wave missingness (Little &  
605 Rhemtulla, 2013) or a rotating panel design (Nielsen *et al.*, 2009) may increase the sustainability  
606 of long-term monitoring for some taxa or regions with few willing participants. But such an  
607 approach has to balance the cost of increased study design complexity and potential implications  
608 for the range of questions that can be addressed.

609 For existing monitoring schemes, data gaps may be filled, where possible, by promoting  
610 data collection in certain areas. Within citizen science projects, there is evidence that participants  
611 can be nudged to collect more data in regions identified as sampling priorities (Callaghan *et al.*,  
612 2023; Callaghan *et al.*, 2019). Previous studies have identified sampling priorities in different  
613 ways; for instance, based on the expected influence of a data point (Callaghan *et al.*, 2019) or  
614 predictions based on species distribution models (Chiffard *et al.*, 2020). Since data collected by

615 monitoring schemes are often collected for multiple purposes, the challenge is identifying the  
616 common set of sampling priorities.

617         For synthesis studies compiling data from independent studies, data mobilisation efforts  
618 may be tailored to improve sample representativeness of the target population, by expending  
619 more effort to under-sampled units. This could be informed by exploring the transferability of  
620 model predictions across spatial or temporal units based on currently available data (Spake *et al.*,  
621 2022). Regions with high transferability may represent appropriate sampling strata to guide  
622 mobilisation efforts. Moreover, these sampling strata may inform the adjustment for data gaps in  
623 subsequent modelling of the population mean.

## 624 **(2) Evaluating and reporting missingness**

625 Developing a causal model (e.g., using a DAG) of the factors affecting sampling probability and  
626 biodiversity can be useful first step to identify auxiliary variables for adjusting data gaps –  
627 variables linked to both sampling probability and biodiversity are those creating bias (Mohan &  
628 Pearl, 2021). As far as possible, data should then be collected on the covariates that are likely to  
629 explain missingness. Statistical models can be used to test whether covariates that are associated  
630 with missingness are also associated with biodiversity patterns, though of course this is only  
631 possible in the sampled data. Unplanned missingness in structured schemes could be investigated  
632 by disseminating follow-up surveys to participants to determine their reasons for missed surveys.  
633 Follow-on data collection, e.g., with paid surveys, in regions or times of missing data could also  
634 help understand whether there are fundamental differences in biodiversity patterns between the  
635 original dataset and the extended dataset.

636         Missingness, and how it is dealt with, tends to be insufficiently reported in biodiversity  
637 trend analyses. Some reporting frameworks for missing data have been developed for other

638 disciplines (Lee *et al.*, 2021) but are in their early stages in ecology (Boyd *et al.*, 2022). At a  
639 minimum, we propose that missingness can be reported in terms of the proportion of sampling  
640 units that are spatial, annual and within-year gaps, and the number of unplanned gaps for  
641 structured monitoring schemes (Fig. 1). Visualizations of the distributions of covariates in  
642 sampled and non-sampled times/sites could also effectively highlight key systematic differences.

### 643 **(3) Modelling to account for data gaps**

644 The impact of data gaps depends on multiple factors: whether the factors affecting missingness  
645 are independent of the factors affecting biodiversity and biodiversity itself; the ecological  
646 questions being asked and which covariates are available and included in the analysis. Because  
647 of this, potential impacts of missingness have to be considered for each species-question-dataset  
648 combination. A dataset *per se* is not biased. Subsampling, weighting and imputation all have the  
649 potential to reduce bias caused by data gaps. Many, but not all, solutions will navigate the bias-  
650 variance trade-off. Weighting is probably the most under-used in ecology and could be applied  
651 more often, especially to account for spatial gaps when the goal is estimating mean abundance or  
652 abundance trends. Imputation methods offer the potential to fill in spatio-temporal gaps to  
653 generate the space-time data cubes of the Essential Biodiversity Framework (Kissling *et al.*,  
654 2018), but its success is dependent on the ability to model the variation in the biodiversity  
655 response. Since available covariates are likely to be only partly successful in reducing bias,  
656 sensitivity analysis could be help explore how different assumptions of missingness change the  
657 model coefficients and predictions, and the uncertainties of them (Little, 1995; Leurent *et al.*,  
658 2018). For some contexts, it might be more statistically efficient and ecologically interpretable to  
659 redefine the target region of interest to a region with fewer data gaps.

## 660 **VIII. Conclusions**

661 (1) Biodiversity datasets containing information on species' occurrences and abundances are  
662 rapidly growing in size, but data gaps are not necessarily closing. Nonetheless, big  
663 biodiversity datasets are invaluable for a broad range of basic and applied questions, and  
664 increasingly for policy-relevant questions about the status and trends of biodiversity at  
665 large-scales. Heterogeneity in sampling efforts - whether by citizen scientists or scientists  
666 - creates different types of data gaps in the available data. Such data gaps are among the  
667 biggest hindrances to making use of these growing data sources for large-scale  
668 inferences.

669 (2) We show how 'missing data thinking' can help decide whether a data gap is problematic  
670 in a given context and provides directions on possible solutions. We show that an  
671 important determinant of bias is whether factors affecting sampling effort are correlated  
672 with those affecting biodiversity: shared covariates affecting sampling effort and  
673 biodiversity have the potential to lead to biased analyses if not taken into account.

674 (3) Multiple approaches are available to account for missing data but they depend on  
675 knowledge and availability of covariates associated with missingness. A lack of training  
676 for ecologists in commonly employed approaches in other disciplines has meant there are  
677 few standard practices in ecology to deal with gaps. We highlight multiple methods that  
678 are ripe for comparison across different ecological problems.

679 (4) At the same, statistical solutions can only go so far, closing data gaps with more  
680 coordinated data collection across monitoring stakeholders is also important as we move  
681 forwards.

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## 691 **X. Supporting Information**

692 **Table S1** Selected R tools that can help with missing data problems and their potential  
693 application for use in biodiversity research.

694 **Fig. S1** Extended analysis of missing data solutions when additional covariates affect the  
695 biodiversity response.

696 R script for the example solution simulations (Fig. 5) are here:

697 <https://github.com/bowlerbear/dataGaps>

698

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## Supporting Information

**Table S1 Selected R tools that can help with missing data problems and their potential application for use in biodiversity research.**

R packages	Applications	Useful functions
<i>Exploring missing data</i>		
<a href="#">naniar</a>	visualizing/exploring the missing data pattern	<i>mcar_test</i> – Little's missing completely at random (MCAR) test <i>vis_miss</i> – plot the missing data for all variables
<a href="#">occAssess</a>	measure of the potential for bias in taxonomic, temporal, spatial, and environmental dimensions	<i>assessEnvBias</i> - assess whether data are sampled from a representative portion of environmental space in the spatial domain of interest <i>assessSpatialBias</i> – assess whether data resemble a random distribution in the geographic space of interest for inference <i>assessSpatialCov</i> – assess whether a representative portion of the spatial domain of interest has been sampled and whether the same portion of geographic space has been sampled over time
<a href="#">sambias</a>	a Bayesian approach to estimate how sampling rates vary as a function of proximity to one or multiple bias factors	<i>calculate_bias</i> - calculating the bias effect of sampling bias due to geographic structures, such as the vicinity to cities, airports, rivers and roads
<i>Subsampling</i>		
base	Base R functions	<i>sample</i> - sample data with predefined inclusion probabilities specified with the prob argument
<a href="#">sampling</a>	draw random samples using different sampling schemes	<i>balancedcluster</i> – selects a balanced cluster sample according to defined auxiliary variables <i>strata</i> - stratified sampling with unequal probabilities.

<a href="#">spatialEco</a>	spatial data manipulation and modelling	<i>stratified.random</i> - creates a stratified random sample of an sp class object <i>stratified.distance</i> - draws a minimum, and optional maximum constrained, distance sub-sampling
<a href="#">spThin</a>	Spatial thinning of species occurrence records	<i>thin</i> - returns a dataset with the maximum number of records for a given thinning distance
<a href="#">terra</a>	spatial data manipulation and processing	<i>spatSample</i> – sample a SpatRaster, SpatVector or SpatExtent objcy
<i>Imputation</i>		
<a href="#">agTrend</a>	modelling regional trends with missing data	<i>mcmc.aggregate</i> - a zero-inflated, nonparameteric model with a definable observation model, augmenting missing values before calculating regional abundances
<a href="#">INLA/</a> <a href="#">inlabru</a>	fitting Bayesian models, especially useful for spatial models via its spatial mesh	<i>inla/bru</i> - fit a Bayesian model using Integrated Nested Laplace approximation <i>predict</i> – draw predictions from the fitted model, where the prediction data frame can be a SpatialPointsDataFrame object
<a href="#">LORI</a>	imputation of missing count data	<i>lori</i> – impute missing count data using a large covariate set, including interactions, with a LASSO penalty
<a href="#">mice</a>	multiple imputation by chained equations	<i>mice</i> – multiple imputation method that will generate plausible values for any missing data – in the response and in any covariates
<a href="#">Rjags</a> <a href="#">JAGS</a> <a href="#">nimble</a>	fitting Bayesian models allowing for missing values in the response	<i>Jags/runMCMC</i> - fitting Bayesian models allowing for imputation of missing values in the response during model fitting (options available for missing values in covariates too)
<a href="#">rtrim</a>	functions to calculate annual indices and trends of abundances	fit a GLM imputing missing values based on mean site and year effects, with optional covariates
<i>Weighting</i>		



<a href="#">survey</a> <a href="#">srvyr</a>	range of functions for analysis of data from complex surveys, including fitting models with weights	<i>Svyglm</i> – generalized linear models with survey weights <i>postStratify</i> – function for post-stratification to match the joint distribution of the variables of the population
<a href="#">svrep</a>	Analysis of replicate/bostrapped survey weights	<i>svyby_repwts</i> – compare estimates from different sets of weights
<a href="#">twang</a>	functions to estimate propensity scores and weights	<i>ps</i> - gradient boosted trees to predict non-response from covariates <i>bal.table</i> – compare covariate values between sample and population

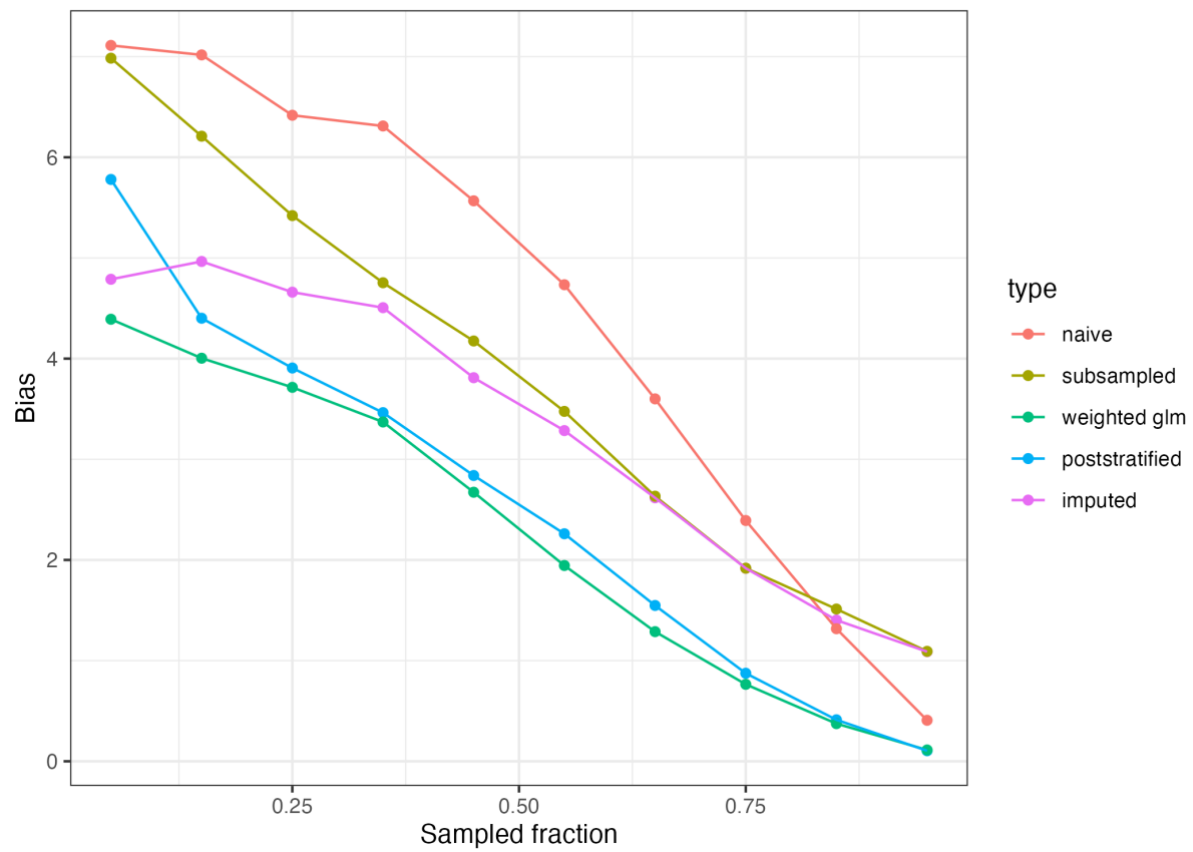


Fig. S1 The ability of missing data solutions to adjust for bias in biodiversity data.

We assumed a landscape of 400 cells and that a covariate affected both species abundance and the likelihood of a cell being sampled. We vary the fraction of the cells that were sampled. In contrast to Fig 5A (main text), we assumed that the species abundance was affected by an additional covariate that did not affect sampling; this variable was not included in any of the analysis. The models to estimate the parameter of interest (mean abundance) were: naive (no correction); subsampled (cells were subsampled along the covariate gradient to reduce the sampling bias), weighted (two methods: weighted glm using the svyglm function, and weighted by poststratification, using postStratify, both in the survey package) and imputed (using JAGS to impute NAs in the response). Points show the mean bias (difference between model prediction and truth) across 100 independent runs.