# Warming summer temperatures are rapidly restructuring North American bumble bee communities

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### 18 Abstract

- 19 A rapidly warming climate is a primary force driving changes in biodiversity worldwide.
- 20 The impact of warming temperatures on insect communities is of particular interest given
- 21 their importance for ecosystem function and service provision and the uncertainty around
- 22 whether insect communities can keep pace with the rate of increasing temperatures. We
- 23 use a long-term dataset on bumble bee species occurrence along with summer maximum
- temperature trends across North America to characterize community-level responses to
- 25 recent climate warming. We examined responses using the community temperature index
- (CTI) a measure of the balance of cool- and warm-adapted species within local
   communities. Starting in 2010, bumble bee average CTI across North America has rapidly
- 27 communities. Starting in 2010, bumble bee average CTT across North America has rapidly
   28 increased after a period of slight increase from 1989 to the late 2000s. This increase is
- 29 strongly associated with recent increases in maximum summer temperatures. The increase
- 30 in CTI is spatially extensive, but the areas exhibiting the largest increase include mid to
- 31 high latitudes as well as low and high elevations areas relatively shielded from other
- 32 intensive global changes (e.g., land-use). On average, bumble bee CTI has increased 0.99°C
- 33 from 1989 to 2018, a change of similar magnitude to the increase in maximum summer
- 34 temperatures. This shift has been driven by the rapid loss of cold-adapted species and an
- 35 increase in warm-adapted species within bumble bee communities across North American
- 36 ecosystems. Despite evidence that the spatial velocity of community change is keeping pace
- 37 with temperatures, this is a direct result of the decline of cool-adapted species. Our results
- 38 provide strong evidence of the pervasive impacts posed to insect communities by
- 39 temperature increases in the last 30 years.

### 40 Introduction

- 41 Climate change is driving profound changes in animal occurrence and community
- 42 composition worldwide. Long-term increases in average temperature as well as increases
- 43 in acute, extreme weather events (e.g., heat waves) have been linked to both positive
- 44 (Kammerer et al., 2021; Crossley et al., 2021) and negative outcomes for biodiversity
- 45 (Kammerer et al., 2021; Oliver et al., 2016; Outhwaite et al., 2022; Sirois-Delisle & Kerr,
- 46 2018). Regardless of the direction of such outcomes, a rapidly changing climate has the
- 47 potential to fundamentally alter biological processes, including ecosystem services that
- 48 maintain biodiversity and support global agricultural production (Johnson et al., 2023;
- 49 Settele et al., 2016).
- 50 Insect responses to climate change are of specific interest given the growing
- 51 documentation of declines in a variety of taxa and regions (Halsch et al., 2021; Raven &
- 52 Wagner, 2021). Although several anthropogenic drivers of global change are at play
- 53 (Goulson et al., 2015; Hemberger et al., 2021), a changing climate is particularly menacing
- 54 given the number of potential direct and indirect impacts it has on insects and its capacity
- to be a force-multiplier, interacting with other factors to exacerbate changes in insect
- 56 populations (Hoover et al., 2012; Forrest et al., 2018; Kenna et al., 2023). Like many global
- 57 change drivers, rapidly increasing temperatures may favor some species while leading to
- 58 local extirpations of others. Though temperatures above the critical limits of species (e.g.,
- 59  $CT_{max}$ ; Oyen et al., 2018) are unlikely, the extent to which climate warming has contributed
- 60 to local shifts in insect abundance and community structure remains mostly unknown. This
- 61 knowledge gap places our understanding of a host of ecological processes and services in
- 62 limbo.
- 63 Even among the most studied insect taxa there is debate about the extent, severity, and
- 64 direction of effects associated with climate change. Bumble bees are a prime example with
- some studies revealing extensive declines (Soroye et al., 2020; but see Guzman et al., 2021)
- and others suggesting resilience and relative stability (Guzman et al., 2021; Maebe et al.,
- 67 2021) or mixed patterns of decline and increases over time (Jackson et al., 2022). Most
- 68 current approaches examining the long-term influence of climate on bumble bees use
- 69 occupancy models to relate changes in species occurrence to trends in climate, such as
- increases in temperature and changes in precipitation (Janousek et al., 2023). Although this
- 71 method yields valuable insights, it can be challenging to align the framework with the
- incidental and imperfect occurrence data that abounds in large-scale insect databases,
- 73 making model outcomes sensitive to occupancy assumptions (Guzman et al., 2021).
- 74 Moreover, the occupancy approach framework does not explicitly capture any
- 75 physiological mechanisms driving species responses to warming temperatures. As such, a
- 76 more thorough understanding of where/when insects are most impacted by climate change
- 77 requires exploring alternative analytical methods that better tie climatic changes to
- 78 estimates of insect physiological temperature preferences and limits.
- 79 We characterize bumble bee community responses to recent climate warming at the
- 80 continental scale by examining changes in the community temperature index (CTI), a
- 81 physiologically-informed metric of community responses to climate based on the
- 82 composition of cool- and warm-adapted species. This metric can be used to assess the rate
- 83 of change in community composition based on historical species temperature preferences

- 84 (species temperature index, STI), as well as the spatial velocity of community changes
- 85 (Devictor et al., 2008, 2012). When examined over time associated with temperature, CTI
- 86 can help determine whether species are keeping pace with the velocity of temperature
- trends (i.e., an increase in warm-adapted species and a loss of cool-adapted species in
- rapidly warming areas; Fourcade et al. 2019), or whether communities are accruing
- 89 "climate debts", as rising temperatures outpace species turnover (Devictor et al., 2012).
- 90 Using 50 years of records from the Bumble bees of North America database (Richardson
- 91 2023), we test for changes in bumble bee communities using CTI across North America by
- 92 quantifying the association between changes in CTI with trends in maximum summer
- 93 temperature. Specifically, we wanted to address the following questions: (1) is there
- evidence of an increase in bumble bee CTI over time? (2) are changes in CTI associated
- 95 with increases in summer temperatures? (3) are CTI changes greater in areas particularly
- 96 vulnerable to a changing climate (e.g., higher latitudes and elevations)? and (4) is a loss of
- 97 cool-adapted or an increase in warm-adapted species driving the observed changes in CTI?
- 98 We predicted a steady increase in bumble bee CTI in accordance with documented
- 99 increases in average maximum summer temperatures over the past century and that
- 100 changes would be more dramatic at higher latitudes and elevations. We also expected that
- 101 a host of common, warm-adapted species that have increased in occurrence over the past
- 102 several decades would be the strongest drivers of change in CTI across the continent.

### 103 Methods

### 104 North American bumble bee occurrence and community data

- 105 We used occurrence records for 59 species of North American bumble bees from the
- 106 bumble bees of North America database (BBNA; Richardson 2023). This database
- 107 composes 781,280 records from 1805-2020 from a variety of sources (e.g., natural history
- 108 collections, research studies, citizen science programs). To match the temporal range of
- available climate data, we used bumble bee records collected between 1960 and 2018.
- 110 Because the database consists of an amalgam of sources, we took several steps to account
- 111 for known biases (Bartomeus et al., 2019; Gotelli et al., 2021). The species and community
- 112 temperature indices at large scales of our analysis are robust to imprecision in the
- underlying distributional data (Devictor et al., 2008); nonetheless we filtered the original dataset to include only complete records (i.e., identified to species, containing complete
- 115 coordinates) and unique collection events (distinct combinations of species, date,
- 116 coordinates, and observer; Figure 1A). This step helps to minimize the bias associated with
- 117 unequal sampling efforts and differential data collection methods across all observers.
- 118 Moreover, we conducted a range of sensitivity analyses (see below) to determine whether
- 119 our results were robust given our assumptions and methodological decisions.

### 120 Is there evidence of an increase in bumble bee CTI over time?

- 121 Calculating the CTI first requires us to determine the species temperature index (STI; the
- 122 historical average summertime temperature experienced over a species' approximate
- 123 range; Figure 1B) for all species present within a given community. For this calculation, we
- 124 used summer maximum monthly temperature as the bulk of bumble bee records are
- 125 collected during this period corresponding the peak flight for most North American species.
- 126 To calculate the STI, we used a subset of bee occurrence records from 1960-2000 to extract
- 127 historical summertime temperatures at these locations from a global climate database. The
- 128 purpose of extracting these records is to delineate the approximate range of each species.
- Because the range estimates for North American bumble bees are largely based on records
- from this dataset (Williams et al. 2014), we are confident that these data capture the range of almost all included species. Next, using the `raster` package (Hijmans, 2023), we
- 131 of annost an included species. Next, using the Taster package (Fijinans, 2023), we132 calculated the historical average maximum summer temperatures at the specific location
- 133 (i.e., raster pixel) of each occurrence record for that species from the WorldClim version 2.1
- historical climate database at 30 arc-second ( $\sim 1 \text{ km}^2$ ) resolution (Fick & Hijmans, 2017) by
- 135 averaging the maximum monthly temperature for summer months (defined here as June-
- 136 September) for a historical period of 1970 to 2000. We used this historical dataset as it is
- 137 the highest resolution historical temperature raster available through WorldClim. We then
- 138 used this raster to extract mean summer maximum temperature values using our bumble
- 139 bee occurrence records. For each species, we then calculated the mean of the extracted
- 140 values to determine the STI estimate.
- 141 Use of the CTI framework required us to assign bee occurrence records to communities to
- 142 calculate CTI values for given locations/times (Devictor et al., 2008; Figure 1C). To do this,
- 143 we created a hexagonal grid across North America at a broad spatial scale (50 km
- hexagonal grid resolution, center to side: ~ 6600 km2) to act as stand-in "community"
- boundaries. We chose a 50 km resolution to ensure we would capture sufficient records

146 within each grid cell to robustly estimate the broad spatiotemporal trend of CTI (Jackson et

- 147 al., 2022). Although these species assemblages are considerably larger than the scale of an
- ecological community, the analysis is ultimately agnostic to this point, and it does not affect
- our specific questions. We refer to them as communities/CTI to maintain consistency withthe existing literature. Also, because we used occurrence records from a variety of sources
- 150 the existing literature. Also, because we used occurrence records from a variety of sources 151 whose spatial locations varied over time, using fixed sampling locations to delineate
- 152 communities (e.g., Prince and Zuckerberg 2014) was not possible. To determine if the
- 153 resolution of our grid cells impacted our results, we also conducted our analyses using 25
- 154 and 100 km center-to-side hexagonal grid cells. Using these grids, we assigned bumble bee
- 155 occurrence records to each cell to create quasi-communities, requiring each cell to contain
- 156 at least 2 species for a given year to calculate CTI. We used hexagonal grid cells to minimize
- 157 possible edge effects and provide a better fit across the curvature of the earth at large
- 158 spatial scales (e.g., continental; Birch et al. 2007).
- 159 Using STI values, we then calculated CTI within each grid cell where at least 2 species
- 160 records were present in the grid using the full set of bumble bee occurrence records from
- 161 1989-2018 (Figure 1D). We were limited to using only CTI calculations from 1989 onward
- as 1989 was the first year for which we could calculate a 30-year moving average summer
- 163 temperature anomaly (see below). We calculated CTI using two different methods, first
- 164 using occurrence records for species *i* occurring within a given community (grid cell) *j*

165 Equation 1: Occurrence 
$$CTI_j = \frac{\sum_{i=1}^{n} STI_{i,j}}{n}$$

- 166 and then using abundance weighted estimates of species within each community:
- 167 Equation 2: Abundance weighted  $CTI_j = \frac{\sum_{i=1}^{n} a_{i,j} \times STI_{i,j}}{\sum_{i=1}^{n} a_{i,j}}$
- 168 where  $a_{i,j}$  is the abundance of species *i* at site *j*, and *n* is the total number of species within
- a grid cell (Princé & Zuckerberg, 2015). In our case, the true abundance is not known, but
- 170 we use the total number of individuals of species i within the community of bees at site j,
- 171 the grid cell, as a proxy of abundance. These two approaches, though similar, estimate the
- 172 two mechanisms of change in CTI. Using occurrence records (Equation 1) allowed us to test
- shifts in CTI due to changes in occurrence (i.e., immigration/extirpation), while calculating
- 174 CTI using abundance weighting (Equation 2) allowed us to understand shifts in CTI as a
- 175 function of changes in local relative abundance (i.e., species becoming more common/rare
- 176 within a given community).

### 177 Are changes in CTI associated with increases in summer temperatures?

- 178 To determine long-term warming trends across North America, we used WorldClim
- 179 gridded historical monthly weather data from 1961-2018 for our defined summer months
- 180 (Fick & Hijmans, 2017). First, we averaged the maximum monthly temperature for each
- 181 year. Second, we extracted the mean maximum temperature within each of the bumble bee
- 182 community grid cells (Figure 1E). This procedure created a time series of the average
- 183 maximum summer temperature for each year/grid cell from 1961-2018. Third, we
- 184 calculated the average maximum summer temperature for a historical period from 1961-
- 185 2000 for each grid cell; this is our baseline, and we refer to it as the temperature "normal".

- 186 Last, we calculated the summer maximum temperature anomaly (defined here as the
- 187 deviation from long-term normal) and averaged these using 3 moving-window scales of 3,
- 188 10, and 30 years to capture metrics of relatively short-, medium-, and long-term changes in
- 189 maximum summer temperatures, respectively. To illustrate the estimated trends in
- 190 maximum summertime temperatures, we calculated the change in our 3 scales of
- anomalies by subtracting the 1989 average anomaly (first possible year to calculate 30-
- 192 year average) from the 2018 average anomaly for each grid cell. This meant that, when
- 193 modeling change in CTI as a function of temperature anomalies, modeled from 1989-2018
- 194 (i.e., where all measures had maximum temperature anomaly values for all three scales (3-,
- 195 10-, and 30-year).
- 196 We used generalized additive models (GAM) to quantify trends in CTI over space and time
- and determine whether changes in CTI were related to short-, medium-, and long-term
- 198 trends in temperature anomalies (Figure 1F). Generalized additive models provide a highly
- 199 flexible computational framework to account for variable trends in spatiotemporal
- 200 processes (Pedersen et al., 2019) and are especially well-suited for the analysis of
- 201 potentially complex time series and can readily identify periods of significant change
- 202 (Simpson, 2018).
- For each measure of CTI (occurrence and abundance-weighted), we fitted a GAM to model the effects of spatial location (latitude, longitude, and elevation), long-term trend (year), short-, medium-, and long-term estimates of rising temperatures (3, 10, and 30-year
- summertime maximum temperature anomalies). For the remainder of this manuscript, werefer to this GAM as the global model.

208 Equation 3: 
$$CTI_j \sim s(lat, long) + s(year) + s(elevation) + ti(lat, long, year, elevation)$$
  
209  $+ s(eco \ region, bs = "re") + s(\overline{T}_{max 3}) + s(\overline{T}_{max 10}) + s(\overline{T}_{max 30})$ 

We fit the model using the mgcv package in R (Wood, 2011). The goal of each model was to 210 identify the spatiotemporal trend in CTI along with its statistical association with changes 211 212 in summertime maximum temperatures. Because of the differences in geography, land-use, 213 and climate across North America, we included a 2-dimensional smooth of latitude and 214 longitude, and we allowed the estimated temporal trend in CTI to vary according to spatial 215 location by including a tensor product interaction of latitude, longitude, elevation, and year (Pedersen et al., 2019; Equation 3). We also included a random effect smooth of ecological 216 217 region (Omernik, 1987) to further account for variation in the response of CTI associated 218 with common biophysical characteristics within ecological regions, such as commonalities in vegetation and other climate variables (e.g., precipitation). We used the Level 1 219 220 ecoregions defined and maintained by the US EPA (Omernik and Griffith. 2014). We included smooths of 3-, 10-, and 30-year summertime maximum temperature anomalies to 221 determine whether changes in CTI were correlated with trends in warming maximum 222 223 summer temperatures. Including three different anomaly scales allowed us to coarsely 224 estimate the temporal scale of temperature change to which bumble bee communities 225 respond most strongly to. This model was fit to CTI estimates from 1989-2018 as 1989 was the first year for which 30-year temperature anomalies could be calculated for each grid 226 227 cell. We tested the model for spatial and temporal autocorrelation in the residuals. For 228 spatial autocorrelation, we tested simulated residuals with a Moran's I test using the 229 DHARMa package (Hartig, 2022). For temporal autocorrelation, we visually examined the

- autocorrelation function using scaled, simulated residuals, finding no evidence of
- 231 problematic residual correlation.
- 232 To visualize the change in CTI over time, we generated CTI predictions across the spatial
- and temporal extents of our dataset using the global model for each grid cell. We then
- determined the change in CTI from 1989-2018 by subtracting the modeled CTI estimate for
- 235 1989 from that of 2018 for each grid cell. To visualize model uncertainty, we calculated the
- average standard error of global model predictions for each grid cell from 1989-2018. We
  visualized the effect of the three moving-average temperature anomalies on CTI by plotting
- 237 visualized the effect of the time moving-average temperature anomalies on CTT by plotting 238 the partial effects (prediction of CTI as a function of temperature holding other variables
- are at their mean value) of each anomaly from the global model using the gratia (Simpson,
- 240 2023) package.

#### Are CTI changes greater in areas particularly vulnerable to a changing climate (e.g., higher latitudes and elevations)?

- 243 To determine whether CTI changes were most drastic (i.e., greater slope in fitted GAM) in
- areas known to be experiencing accelerated climatic changes, we examined the rate of
- change in the slope (i.e., derivative) of our fitted model smooth (Figure S1). To do this, we
- first fitted a GAM to CTI predictions with a single smooth of year to create a spatially
- 247 explicit, estimated trend of CTI for each grid cell. Then, for each grid cell's fitted GAM year
- smooth, we extracted the first derivative with respect to time (1990-2018) using the
- derivatives() function from the gratia package (Simpson, 2023). For elevation and latitude,
- we calculated the mean derivative value for each grid cell (i.e., the average rate of change of the CTI of a grid cell from 1989-2018) and then plotted this against the mean elevation and
- 251 the C11 of a grid cell from 1989-2018) and then plotted this against the mean eleval 252 latitude of the grid cell. We visualized the relationship with a GAM fit using the
- 253 geom\_smooth() function in the ggplot (Wickham et al., 2019) package. To determine
- whether CTI changes were consistent or have accelerated over time, we calculated the
- 255 derivative values for the year smooth for each grid cell and plotted these values against the
- 256 year. Like elevation, this relationship was visualized with a simple GAM fit.

### 257 Which species are driving any observed changes in CTI?

- Although quantifying the trend in CTI provides evidence for whether communities are
- being restructured in response to a changing climate, the procedure does not identify the
- 260 mechanism driving changes. For example, for areas where CTI is observed to increase, is it
- an increase in warm-adapted or a decrease in cool-adapted driving the change? To address
- this, we modeled the trend in the relative abundance of cool- and warm-adapted species to
- 263 generalize the mechanism underlying the observed changes in CTI across North America.
- First, we assigned species as either cool- or warm-adapted within each grid cell by
- comparing species STI values against the average of STI values (i.e., the CTI) across all
- 266 species present within the cell. STI values above the community mean were assigned
- 267 "warm-adapted" while those below were assigned "cool-adapted". This approach allowed
- species identified as cool- or -warm adapted to change based on the location and
- 269 community composition, which we felt was more realistic than assigning cool- or warm-
- adapted based on a range-wide assessment given that some species have extensive ranges
- that cover large swaths of North America. Next, we calculated the relative abundance of
- each species in each grid cell within 3 temporal bins (each with  $\sim$  equal numbers of

- 273 observations). For each species, we then fit a binomial GLMM with a fixed effect of the year
- bin and a random effect of ecological region to estimate the temporal trend in relative
- abundance across all grid cells where the species was present. Modeling relative
- abundance in this way is a useful proxy to determine approximate trends in species
- occurrence and abundance (Hemberger et al., 2021, Gotelli et al., 2021). Next, we extracted
  the model coefficient for the year bin term and combined these estimates with species
- 278 the model coefficient for the year bin term and combined these estimates with species 279 occurrence records that were assigned as either cool- or warm-adapted. We then fit a
- subsequent GLM that predicted the estimated relative abundance trend as a function of
- species thermal group (cool- or warm adapted) and its interaction with the latitude of the
- grid cell of the observation and an error propagation term. Because this model was fit using
- 283 coefficients from the species-specific relative abundance trend models, we included the
- error propagation term (coefficient standard error / coefficient estimate) to help account
- for the uncertainty in the underlying coefficient estimates. This model allowed us to
- determine whether the rate of change in relative abundance among the two species groupsvaried from South to North given known trends in increasing temperatures across
- 288 latitudes.

### 289 Model validation

- 290 We performed cross validation on our global model using testing data that was filtered out
- 291 of the full BBNA database. These collection events, while not "unique" (i.e., not necessarily
- fully independent given our strict definition), were still valid records that could be used to
- 293 calculate the CTI for any given location. Upon calculating the CTI for grid cells using these
- records, we compared the values against predictions from the global model by using the
- coefficient of determination (R2), root mean square error (RMSE) and mean absolute error(MAE).
- 297 Despite the vast number of individual occurrence records within our dataset, there were
- 298 many grid cells that did not contain species occurrence data for fitting the model. Given
- that we explicitly model CTI over space, we presented our results above using predictions
- 300 within all grid cells given the strength of our global model fits. However, we also assessed
- 301 the results when using model predicted values of CTI only for grid cells containing
- 302 occurrence data. This approach was primarily meant to provide conservative estimates of
- 303 CTI changes, particularly where in space (i.e., latitude, elevation) and time changes have
- 304 been the largest.
- 305 We conducted all data wrangling, GIS operations, modeling, and visualization using R (R
- 306 Core Team, 2017) using the aforementioned and following packages: tidyverse (Wickham
- 307 et al., 2019), raster (Hijmans, 2023), sf (Pebesma, 2018), performance (Lüdecke et al.,
- 308 2021), janitor (Firke, 2021), paletteer (Hvitfeldt, 2021), exactextractr (Daniel Baston,
- 309 2022), foreach (Microsoft & Weston, 2022), and data.table (Dowle & Srinivasan, 2023)
- 310 packages.

### 311 **Results**

## Bumble bee community temperature index has increased across a majority of North America

- From 1989-2018 bumble bee CTI increased substantially across most of North America.
- 315 Overall, CTI increased on average 0.99 ± 1.98 °C (mean ± SD). The magnitude of change in
- CTI was spatially variable, ranging from a decrease of 6.30 °C to an increase of 7.99 °C
- 317 (Figure 2A). The predictions were most certain across the coterminous United States
- 318 where there is a high density of bumble bee records and less certain in the most northern
- 319 grid cells of our study region in the high Tundra and Queen Elizabeth Islands as well as in
- the tropical wet forests of Mexico (Figure 2B). The spatial trends of the increase in CTI
   were nearly identical between occurrence and abundance-weighted CTI; however, changes
- 322 in occurrence CTI were marginally smaller ( $0.78 \pm 1.75$  °C). The global model, which
- 323 quantified the change in CTI as a function of space, time, and changes in short-, medium-,
- and long-term temperature increases, explained a substantial portion of the deviance in
- both the abundance-weighted (Table S1; 86.0%,  $adj-R^2 = 0.849$ ) and occurrence models
- 326 (Table S1; 86.3%, adj-R<sup>2</sup> = 0.851).
- 327 The results of our analysis were consistent irrespective of the grid scale used in
- 328 aggregating communities (Figure S2; Table S2). The exception was in areas of British
- Columbia and Alaska where a highly concentrated spatial pattern of bumble bee records
- likely led to a predicted decrease in CTI in grid cells when aggregated at the 50 and 25 km
- 331 grid scale. Aggregating at the largest scale (100 km center-to-side hexagonal grid) revealed
- the most wide-spread increases in CTI, with nearly all grid cells exhibiting an increase in
- 333 CTI from 1989 to 2018.
- Our models performed well when cross-validated using withheld data from the BBNA
- 335 database (Figure S3). Coefficient of determination (R<sup>2</sup>) values ranged from 0.79-0.81; root
- 336 mean squared error (RMSE) ranged from 1.22-1.31; and mean absolute error (MAE)
- ranged from 0.91-0.96. In addition, our model performance was consistent across the three
- tested grid scales. Predictions were most accurate for CTI values ranging from 23-28°C
- 339 which corresponded to the regions where the bulk of the occurrence records were
- 340 collected. Prediction accuracy was most variable among cool regions in the north and sub-
- 341 arctic (CTI < 23°C).

### 342 Shifts in CTI are strongly related to long-term increases in summer temperature

- 343 Summertime maximum temperatures have increased by 1989-2018 (Fig. 2C-E), with
- increases most apparent at 10- (0.630 ± 0.405 °C) and 30-year average anomalies (0.969 ±
- 345 0.342°C; Figure 1D, E; Figure S4). Increases in the 30-year summertime maximum
- 346 temperature anomaly showed a strong statistical association with increases in bumble bee
- 347 CTI (Figure 2C; F = 4.561, p = 0.002). Increases in the 30-year temperature anomaly
- 348 between 0-0.5°C had no impact on CTI. However, increases of over 0.5°C were associated
- 349 with a rapid increase of up to 1°C in bumble bee CTI (partial effect due solely to 30-year
- 350 temperature anomaly). Beyond a 1°C change in the 30-year temperature anomaly the
- 351 changes in CTI rapidly increase, with gains of 1 to 6.8°C. The relationship of CTI with short
- 352 term, 3-year moving average shifts in summer temperature anomalies, while statistically
- 353 supported, was weak and variable over the range of the anomalies (Figure 3A; F = 2.584, p

354 = 0.032). There was no statistically supported relationship between the 10-year average
anomaly and bumble bee CTI (Figure 3B; F = 0.064, p = 0.802).

## 356 CTI is increasing fastest at low and high elevations, high latitudes, and more recent 357 years

- 358 We examined patterns in the rate of change in CTI across the continent to determine where
- and when the most extreme changes in CTI were occurring and whether these areas
- 360 overlapped with areas known to be heavily impacted by a warming climate (Janousek et al.,
- 361 2023). The rate of change in CTI was greatest at low (< 800 m) and high elevations (> 2000
- m; Figure 5A) and increased with increasing latitude (Figure 5B). CTI increases predicted
- 363 at high elevations also had greater uncertainty than those at low elevations, due in part to a
- higher concentration of occurrence records at lower elevations. Moreover, the rate of
   change in CTI has increased from 1989-2018, with CTI increasing most rapidly after 201
- change in CTI has increased from 1989-2018, with CTI increasing most rapidly after 2010
  (Figure 5C). These results varied slightly when analyzed with predictions from only grid
- 367 cells containing occurrence records, with changes in CTI being greatest at high elevations
- 368 (Figure S5A; > 2000 m) and mid-high latitudes (Figure S5B; 35 60°). The temporal
- 369 patterns of the rate of change were largely similar but were positive only from 2003 and
- beyond (Figure S5C), confirming the accelerating rate of CTI change from 2010 onward
- 371 that is exhibited when using predictions from all grid cells (Figure 5C).

#### 372 CTI changes driven by loss of cool-adapted and increase in warm-adapted species

- 373 In the model predicting the temporal trend coefficient for species relative abundance, there
- 374 was a significant interaction between species thermal niche, latitude, and propagated error
- 375 ( $\chi^2 = 14.53$ , *p* < 0.001, Table S3). The relative abundance of cool-adapted species has
- declined across North America, with the rate of decrease (i.e., binomial model coefficient
- 377 estimates) lowest at high latitudes (Figure 6). In contrast, the relative abundance of warm-
- adapted species has increased across all areas south of ~50° latitude. Beyond this 50°
- 379 parallel, warm-adapted species are also decreasing in relative abundance. This general
- trend, a consistent loss of cold-adapted species and increase in warm-adapted species
- across most latitudes was broadly consistent across ecological regions (Figure S6).

### 382 **Discussion**

383 We documented significant, rapid spatially extensive shifts in the thermal composition of

North American bumble bee communities in response to long-term increases in summer
 temperatures. Over the last 29 years across the continent, bumble bee community

386 assemblages increasingly consist of fewer cool-adapted and more warm-adapted species

- 387 with resultant increases in the community temperature index, a measure of the balance of
- 388 warm- and cool-adapted species, most pronounced at mid- to high latitudes, and high
- elevations in the American Rockies, Intermountain West, and central Mexico. We also
- document an alarming trend suggesting that above 50°N, both cool- and warm-adapted
- species are declining in relative abundance, indicating that warming temperatures are
   outpacing the capacity of bumble bee species to respond or adapt (Kerr et al., 2015). The
- 392 community temperature index increased according to both occurrence and abundance-
- weighted indices, suggesting that shifts in both local abundance (i.e., loss of cool-adapted
- 395 species) and broader changes in species occurrence (i.e., range shifts) underlie the
- 396 observed changes in community composition. Our model results are consistent with
- 397 occupancy patterns that reveal both winners and losers among bumble bee species in
- 398 response to anthropogenic climate change (Jackson et al., 2022). Overall, our work
- 399 provides strong evidence of the pervasive impacts a warming planet has for insect
- 400 biodiversity, particularly for historically cool-adapted species, and identifies regions of
- 401 concern where anthropogenic climate warming is rapidly restructuring the communities of
- 402 an ecologically important group of insects.
- 403 An increase in species turnover within biological communities is a logical consequence of a
- 404 rapidly warming climate (Tingley & Beissinger, 2013). Similar shifts in community
- 405 composition have been observed in bird communities in response to both warming
- 406 summer (Devictor et al., 2008, 2012) and winter (Princé & Zuckerberg, 2015)
- 407 temperatures. Because insects are ectotherms, temperature-induced shifts in range and
- 408 abundance may be even more pronounced. Indeed, large changes in insect CTI have been
- 409 observed for both bumble bees (Fourcade et al., 2019) and butterflies (Devictor et al.,
- 410 2012); however, trends in CTI are often not explicitly tied to spatial and temporal patterns
- 411 of warming temperatures. Our results explicitly link these two phenomena revealing a
- 412 clear statistical relationship between increases in CTI and long-term increases in maximum
- summer temperatures across North America. Areas experiencing a 30-year temperature
   anomaly of greater than or equal to 0.5°C strongly associated with a rapid increase in
- anomaly of greater than or equal to 0.5°C strongly associated with a rapid increase in
  bumble bee CTI (Figure 2; dark orange and red areas). It is worth noting that the historica
- bumble bee CTI (Figure 2; dark orange and red areas). It is worth noting that the historical
  baseline period we choose for calculating species STI values is due to the availability of
- 410 baseline period we choose for calculating species 511 values is due to the availability of417 interpolated climate data. Choosing an earlier baseline period could reveal different
- 418 patterns in community change, however this is unlikely given the stability of summer
- 419 temperatures relative to the dramatic increases observed in recent decades.
- 420 The frontline of species' responses to climate have tended to be at high latitudes. Northern
- 421 regions have experienced rapid increases in temperature leading to pronounced
- 422 phenological shifts across taxa (Parmesan, 2007). Our results support this trend, finding
- 423 greatest rates of bumble bee CTI change at higher latitudes and high elevation. The bumble
- 424 bee species in these locations tend to have narrower ranges and be cold-adapted, traits
- 425 identical to other insect taxa that have exhibited declines due to climate (Engelhardt et al.,

426 2022; Halsch et al., 2021; Neff et al., 2022). Alarmingly, our results found that even warm-

- 427 adapted species are struggling to respond to the pace of warming temperatures at higher
- 428 latitudes (CITE). We found that both cool- and warm-adapted bumble bee species north of
- 429 50°N have exhibited significant declines in relative abundance. This result supports
- 430 previous work describing the limited capacity of bumble bees to track their northern range
- limits in accordance with warming temperatures (Kerr et al., 2015). Though additional
  confirmation is needed, our results suggest that northern bumble bee communities are in
- 432 communities are in 433 crisis, with significant species turnover and declines in abundance that may threaten the
- 434 persistence of populations in the coming decades.
- 435 Rapidly increasing CTI at high elevations suggests that cold-adapted species are being
- 436 displaced by warm-adapted, low-elevation species. This phenomenon has been observed in
- the US Rocky Mountains where bumble bee communities are increasingly dominated by
- 438 low-elevation species using high-elevation habitats as a thermal refugia (Miller-Struttmann
- 439 et al., 2022). An upslope range expansion appears to be a common response of bumble bee
- 440 communities to warming temperatures rather than expansions of northern ranges (Kerr et
- al., 2015; Sirois-Delisle & Kerr, 2018). Despite the rapid changes observed at higher
- latitudes, biological communities in southern latitudes and lower elevations are not
- 443 protected from a changing climate (Dillon et al., 2010), and we documented some shifts in
- 444 CTI in central Mexico and at low elevations. That said, if species lost from communities
- 445 have STI values comparable to those species remaining, shifts in CTI and community
- 446 composition may be effectively masked, highlighting a limitation of our approach.
- 447 An increase in CTI could be the result of two mechanisms. First, shifts in the occurrence of 448 bees within a community (i.e., immigration/extirpation of warm-/cool-adapted species via 449 range expansion/contraction) and second, changes in the local abundance of warm-/cool-450 adapted species. We found evidence supporting both mechanisms by modeling occurrence 451 and abundance-weighted measures of CTI. Shifts in local relative abundance align with 452 existing research (Cameron et al., 2011; J. Hemberger et al., 2021); however, substantial 453 range expansion of warm-adapted bumble bees has not been described (Kerr et al., 2015) 454 and may be unlikely given bumble bee dispersal capacities (Fijen, 2021). That said, select 455 species of bumble bees may be capable of long-distance dispersal (Fijen, 2021), and 456 significant range shifts in other insect taxa have been observed (Warren et al., 2001; 457 Hinckling.2005). Regardless, our thermal niche analysis revealed there are a host of warm-458 adapted species whose relative abundance is increasing significantly. This result indicates 459 that certain species are sensitive to and more capable of effectively tracking/adapting to 460 ideal climatic conditions (Maebe et al., 2021). Indeed, several bumble bee species have exhibited both range increases (Looney et al., 2019; e.g., B. impatiens Palmier et al., 2019) 461 462 and increases in local abundance. However, other species (e.g., B. occidentalis) are not able 463 to successfully track warming and are likely to suffer substantial reductions in range as a result (Janousek et al., 2023). Moreover, our analysis found that, north of the US/Canadian 464 465 border, even warm-adapted species are at risk, with negative trends in species' relative 466 abundance. Such contrasts highlight the species-specific nature of bumble bee responses to 467 a rapidly changing climate (Jackson et al., 2022; Whitehorn et al., 2022). Additional 468 research is needed detailing species-specific responses to warming conditions – focusing 469 on identifying the physiological and evolutionary mechanisms that drive species' plasticity 470 to changing environmental conditions.

471 An increase in the occurrence and abundance of warm adapted species does suggest a 472 physiological/climate preference mechanism is at play (i.e., direct effect). Several studies 473 document significant, direct effects of warming on insect pollinators (CaraDonna et al., 474 2018; Hemberger et al., 2023; Kenna et al., 2021), however indirect effects mediated 475 through biotic interactions may be just as if not more important (Ockendon et al., 2014, but 476 see Iler et al., 2021). In the context of our study, such indirect effects imply that shifts in 477 bumble bee community composition are occurring in part in response to climate-induced 478 changes in the resource landscape (i.e., indirect effects). For example, warming climates 479 can widen the temporal availability of resources due to earlier snowmelts, which in turn 480 lead to an increase in bumble bee abundance (Ogilvie et al., 2017). Warming may also 481 create phenological mismatches that reduce available forage for bees (Pyke et al. 2016, but 482 see Bartomeus et al., 2011). Similarly, an increase in hot, dry summer conditions can 483 significantly reduce floral resources and the bumble bees that depend on them (Iserbyt & 484 Rasmont, 2013), and similar patterns have been observed for butterflies (Crossley et al., 485 2021). Unfavorable conditions, often a result of extreme weather events such as heat waves 486 and droughts, can create resource bottlenecks that have the potential to lead to population 487 declines and local extirpation (Maron et al., 2015). Heat waves, for example, are expected to increase significantly in the coming century (Lopez et al., 2018; Meehl & Tebaldi, 2004; 488 489 Thompson et al., 2022). Because our study could not differentiate between direct and 490 indirect pathways, parsing their relative impacts on bumble bees and other taxa is a critical research need. In the meantime, supporting bumble bees in the face of both direct and 491 492 indirect effects may be accomplished by maintaining climate refugia, such as heterogeneity 493 in vegetation structure, that can provide microclimatic respite from temperature extremes 494 to bees (Pincebourde & Woods, 2020) and other taxa (e.g., birds, Kim et al., 2022) in addition to increasing spatial/temporal resource continuity to minimize negative indirect 495 496 effects (Maron et al., 2015).

- 497 Given the spatiotemporal extent of our study, it is likely that warming summer
- temperatures and the temperature profile of a given bumble bee assemblage may co-vary
- 499 with other, known factors of bumble bee community composition and occurrence. For
- example, losses in certain species across their range may be linked to disease (Szabo et al.,
  2012). Additionally, at large-scales, a loss of suitable habitat via land-use intensification
- 501 2012). Additionally, at large-scales, a loss of suitable habitat via land-use intensification 502 and change is also of concern. However, when examined together with shifts in land-use,
- 503 climatic variables (and their associated indirect effects) tend to have as much or more
- 504 power to explain long-term species trends than land-use or resource availability in bumble
- 505 bees (Kerr et al. 2015) and other wild bee species (Duchenne et al., 2020). Moreover, the
- areas of greatest increase in CTI are in areas removed from the most significant effects of
- 507 land-use change (e.g., high latitudes and elevations; Halsch et al., 2021). Regardless,
- 508 managing habitat offers a critical tool that can be used to mitigate the impacts of a changing
- climate (Kim et al., 2022; Oliver et al., 2016; Oliver et al., 2015; Outhwaite et al., 2022).

### 510 **Conclusions**

- 511 Climate change is poised to have significant, cross-scale impacts on insect behavior,
- 512 populations, and communities (Halsch et al., 2021; Høye et al., 2021; Lehmann et al., 2020;
- Raven & Wagner, 2021). In this paper, we document a substantial shift in the functional
- 514 composition of bumble bee communities with respect to climate that is tied to a long-term

- 515 increase of summer temperatures in North America. Due to changes in both occurrence and
- abundance, several species appear to be tracking climate warming, however we found that
- 517 cold-adapted species appear lack the adaptive capacity to cope with rapidly climbing
- 518 temperatures and are being lost from bumble bee communities across the continent.
- 519 Although the exact mechanisms of these community-level shifts remain unknown (i.e.,
- 520 direct vs. indirect effect of warming), our work adds to a growing body of evidence that
- 521 suggests climate change is having a significant, negative impact on many species with
- unknown consequences for ecosystems. It is critical that we focus on designing adaptation
- 523 measures, such as climate refugia and climate-focused habitat conservation, to help combat 524 the ongoing direct and indirect impacts a rapidly warming planet threatens. However, such
- 6124 the ongoing direct and indirect inpacts a rapidly warning planet threatens. However, su
   625 efforts will only be successful in conjunction with substantial decreases in emissions
- 525 efforts will only be successful in conjunction with substantial decreases in effissions 526 (Oliver et al. 2015) on eccential solution to sefery and the planet's biodiversity for
- 526 (Oliver et al., 2015) an essential solution to safeguard the planet's biodiversity for
- 527 generations to come.

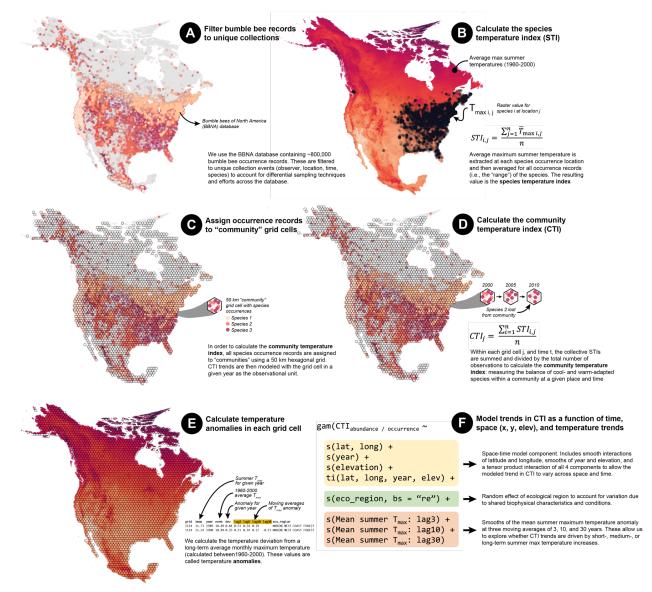
### 528 Acknowledgements

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- 530 project was funded by a USDA NIFA Postdoctoral Fellowship to JH (Award No. 2020-
- 531 67034-31944)

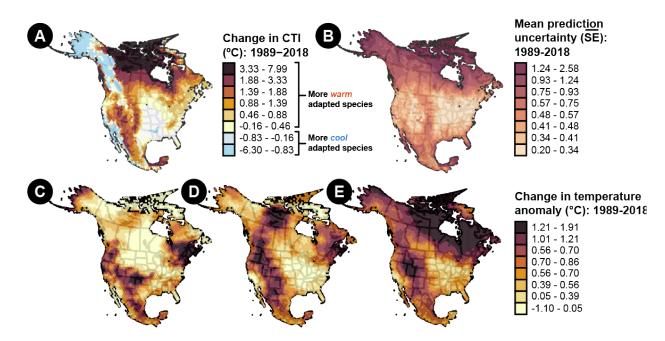
### 532 Data access

- 533 All data and R code for analyses, figures, and manuscript are available on FigShare
- 534 (https://figshare.com/s/59aa9a55ab5bafb65901) and will be made public upon
- 535 publication.

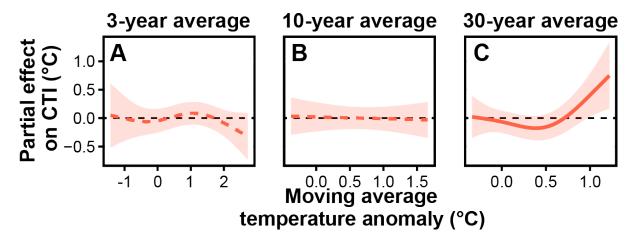
### 536 Figures and Tables



- **Figure 1:** Conceptual figure of data cleaning (A), STI calculation (B), community
- 539 assignment (C), CTI calculation (D), temperature anomaly calculations (E) and modeling
- 540 procedures used in our analyses (F).



- 541
- 542 **Figure 2:** (A) Extrapolated spatial projection of the estimated change in community
- 543 temperature index from 1990-2018 across North America. Differences in CTI were
- calculated for each grid cell by subtracting the model predicted CTIt = 1989 from predicted
- 545 CTIt = 2018. (B) Spatial projection of the mean uncertainty estimates across years from
- 546 1989-2018. (C) Spatial projection of the change in the 3-year, 10-year (D) and 30-year (E)
  547 average temperature anomaly. Differences were calculated by subtracting the 1989
- 548 anomaly from the 2018 anomaly for each grid cell. Hexagonal grid cells are 100 km from
- 549 side to side (~8600 km2).



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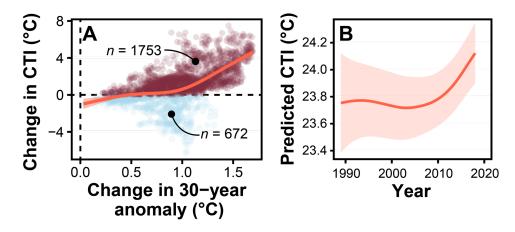
**Figure 3:** Generalized additive model partial plots (i.e., marginal effects) show the model

552 predicted effect of (A) 3, (B) 10, and (C) 30-year moving average temperature anomalies on

the community temperature index. Positive values on the y-axes indicate an increase in CTI,

while positive values on the x-axes indicate an increase in the average temperature relative

to the long-term average. Solid line indicates strong evidence of a relationship.



556

557 **Figure 4:** A significant increase in bumble bee CTI is strongly associated with long-term

warming and has accelerated in the last 15 years. (A) Biplot of change in 30-year

temperature anomaly and change in bumble bee CTI for each grid cell across North

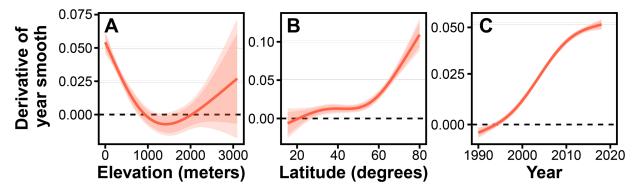
America. Trendline is a GAM fit including the 95% confidence interval. Dashed lines

indicate no change in anomaly or CTI for the X and Y axes, respectively. (B) Model

solution estimated temporal trend in CTI across North America. Yearly predictions are calculated

from the global model for each grid cell, and the trend within each region is illustrated with

a GAM fit including the 95% confidence interval.





**Figure 5:** Estimates of the rate of change in CTI over time across (A) elevation, (B) latitude, and (C) year. Yearly predictions of CTI are calculated from the global model for each grid

cell using a generalized additive model with a single smooth of year to determine the

temporal trend in CTI within the grid cell. For each fitted smooth (except for the year, C),

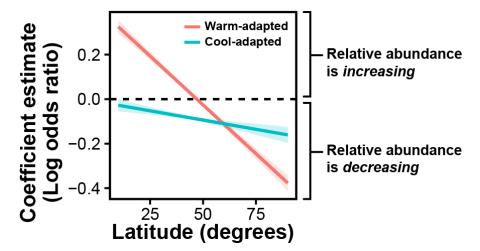
570 we then calculated the mean derivative across its range (1989-2018) for each grid cell. We

then plotted these derivative estimates against elevation and latitude to explore, across the

572 extent of North America, where the rate CTI change is greatest. We visualized the

relationships (red lines) using a simple GAM. Model fits include the 95% confidence

574 interval. # References

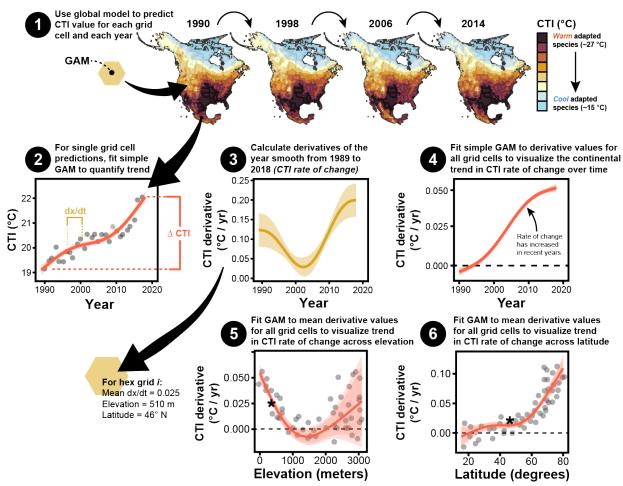


576 577 **Figure 6:** Marginal effect plots describing predicted coefficients for the temporal trend in

- 578 warm- and cool-adapted species relative abundance across latitude in North America (±
- 579 95% CI). Values above zero indicate an increase in relative abundance from 1989-2018,
- 580 while values below zero indicate a decrease.

581 Supplementary Materials

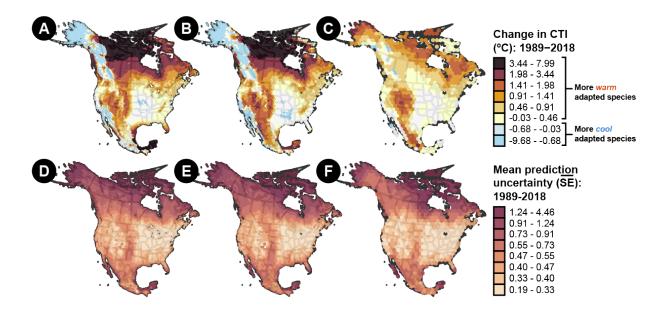




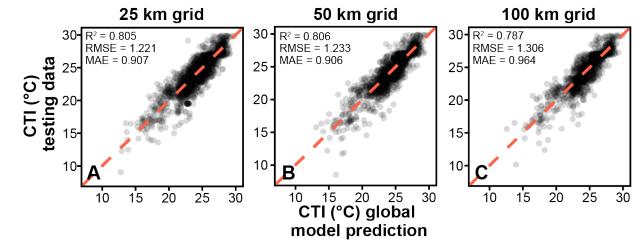
583

584 Figure S1: Conceptual diagram of the derivative calculations conducted to determine 585 whether the rate of increase (i.e., derivative) of bumble bee CTI has remained steady or accelerated over space and time. (1) We use the global model to predict the CTI in each grid 586 587 cell for each year of the study, from 1989-2018. (2) For each grid cell, we fit a GAM through the predicted points to visualize and quantify the trend in CTI from 1989-2018. From these 588 589 data, we also calculated the change in CTI from 1989-2018 (change in CTI) which is plotted 590 in Fig. 1A. The overall change, however, tells us nothing of the functional form of the 591 relationship between CTI and time, elevation, etc. To address this, we calculated the first 592 derivative across the fitted smooth to determine how the rate of change in CTI varied 593 across time, elevation, and latitude (Fig. 2). (3) For each grid cell's fitted GAM, we 594 calculated the derivative of the year smooth at a range of values between 1989-2018. In 595 this example, because CTI is increasing throughout the entire study period, the derivative is > 0 at all years. (4) We then took the derivative estimates for all grid cells and fit a GAM to 596 597 visualize the trend between the derivative and time. For elevation (5) and latitude (6), we 598 first averaged the derivative value from 1989-2018 to determine the mean slope for each 599 grid cell before plotting that against the mean elevation and latitude of each grid cell and visualizing the relationship with a GAM. Transparent points are illustrative (not actual 600

- 601 values) of individual hex grid derivative values across the range of elevation and latitude.
- 602 The black star represents a hypothetical mean derivative value from the example plot in
- 603 (3) to illustrate how mean derivative values are used to assess the trend.

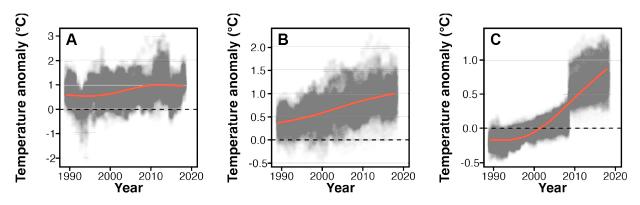


- **Figure S2:** Predicted change in bumble bee CTI across North America between 1989-2018
- at three different spatial resolutions of hexagonal grid (distance indicates side-to-side): (A)
- 50 km; (B) 100 km; (C) 200 km; along with the mean prediction uncertainty at the sameresolutions.



**Figure S3:** Abundance-weighted global model cross validation results at three different

- 611 scales of (A) 25 km, (B) 50 km, and (C) 100 km center-to-edge hexagonal grids. Cross
- 612 validation metrics are given in the top left of each panel including coefficient of
- 613 determination (R2), root mean squared error (RMSE), and mean absolute error (MAE).





614 615 616 **Figure S4:** Trend in summer (June – September) maximum temperature anomalies at (A) 3-year, (B) 10-year, and (C) 30-year moving averages. Transparent points are raw values

617 and red lines are GAM trendlines.

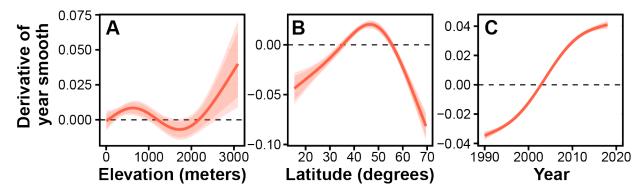




Figure S5: Estimates of the rate of change in CTI over time across (A) elevation, (B)
latitude, and (C) year using predictions only from grid cells containing occurrence records

621 (conservative approach). Yearly predictions are calculated from the global model for each

622 grid cell using simple generalized additive models with a single smooth of year to

623 determine CTI trend within the grid cell. For each fitted smooth (except for the year, C), we

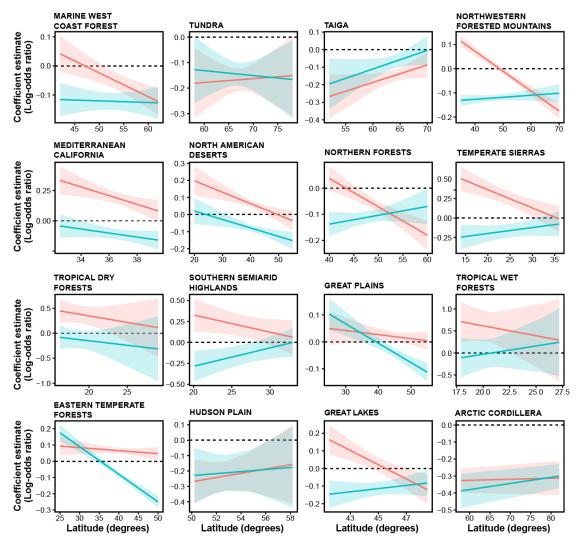
then calculated the mean derivative across its range (1989-2018) for each grid cell. We

625 then plotted these derivative estimates to explore, across the extent of North America,

626 whether increases in CTI were varied with elevation or over time. We calculated

627 predictions (red lines) from a generalized additive model using a thin-plate basis function

and 3 knots for visual purposes only. Estimates include the 95% confidence interval.



630 **Figure S6:** Marginal effect plots describing predicted coefficients for the temporal trend in

631 warm- and cool-adapted species relative abundance across latitude in North America (±

632 95% CI) for each ecological region. Values above zero indicate an increase in relative

633 abundance from 1989-2018, while values below zero indicate a decrease.

**Table S1:** Results from a generalized additive model for CTI using occurrence-only and

635	abundance-weighted records from 1989-2018.
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	Occurrence model			Abundance model				
Model smooth	EDF	F	p value	EDF	F	p value		
Latitude, longitude	282.065	16.778	< 0.001	270.546	18.099	< 0.001		
Year	2.945	4.015	0.004	2.353	3.592	0.022		
Elevation	7.995	23.926	< 0.001	7.615	25.782	< 0.001		
Latitude, longitude, elevation, year	100.149	3.782	< 0.001	97.853	2.996	< 0.001		
Ecological region	10.526	2.388	< 0.001	9.855	1.644	< 0.001		
Mean Tmax 3-year MA	2.968	2.584	0.032	2.827	2.500	0.039		
Mean Tmax 10-year MA	1.002	0.064	0.803	1.001	0.475	0.491		
Mean Tmax 30-year MA	2.967	4.561	0.002	3.377	6.712	< 0.001		
Model n	5273			5273				
Deviance explained	0.860			0.863				
R-squared (adjusted)	0.849			0.851				

\*EDF : estimated degrees of freedom (i.e., smooth wiggliness)

- **Table S2:** Results from a generalized additive model for CTI using occurrence-only records at three different spatial resolutions (community grid scale) at 25, 50, and 100 km from
- 1989-2018.

	25 km grid scale		50 km grid scale			100 km grid scale			
Model smooth	EDF	F	p value	EDF	F	p value	EDF	F	p value
Latitude, longitude	294.224	18.839	< 0.001	270.546	18.099	< 0.001	181.650	17.769	< 0.001
Year	3.653	5.293	< 0.001	2.353	3.592	0.022	2.538	1.609	0.164
Elevation	9.360	34.646	< 0.001	7.615	25.782	< 0.001	1.000	52.860	< 0.001
Latitude, longitude, elevation, year	106.844	4.744	< 0.001	97.853	2.996	< 0.001	51.223	2.558	< 0.001
Ecological region	10.889	4.824	< 0.001	9.855	1.644	< 0.001	11.650	4.843	< 0.001
Mean Tmax 3-year MA	2.881	3.999	0.008	2.827	2.500	0.039	1.000	0.041	0.840
Mean Tmax 10-year MA	1.000	0.061	0.805	1.001	0.475	0.491	3.511	5.344	0.004
Mean Tmax 30-year MA	2.919	5.177	0.001	3.377	6.712	< 0.001	3.029	4.430	0.003
Model n	7582			5273			3078		
Deviance explained	0.840			0.860			0.862		
R-squared (adjusted)	0.831			0.849			0.850		

\*EDF :estimated degrees of freedom (i.e., smooth wiggliness) 

**Table S3:** Type III Wald Chi-squared effect test results from a generalized linear model for the temporal trend in species relative abundance for two thermal niche groups: cool- and 

warm-adapted.

			Degrees of	f
Model term	Model coefficient	χ2	freedom	<i>p</i> -value
Thermal niche	0.497	506.26	1	< 0.001
Latitude	-0.002	461.18	1	< 0.001
Error propagation	-0.005	20.14	1	< 0.001
Thermal niche * latitude	-0.008	191.57	1	< 0.001
Thermal Niche * error	-0.059	67.62	1	< 0.001
Latitude * error	0.0004	54.51	1	< 0.001
Thermal niche * latitude * error	0.001	14.53	1	< 0.001

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