

1 **Beyond single-species models: leveraging multispecies forecasts to navigate the dynamics of**
2 **ecological predictability**

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23 **Authorship:**

24 NJC, SKM and EPW designed the study; SKM and GMY oversaw field data collection; HS, JLS and EPW
25 designed cyberinfrastructure to store and access data; NJC and KANKK performed statistical analyses
26 and visualizations; NJC drafted the manuscript and all authors contributed to revisions.

27

28 **Data availability statement:**

29 All data from the Portal Project is publicly archived to Zenodo (Ernest et al. 2023) and is available in
30 processed forms using the portalr R package (Christensen et al. 2019b). R code to reproduce
31 analyses is provided in Supplementary materials and in a GitHub repository
32 (https://github.com/nicholasjclark/portal_VAR). R code will be permanently archived on Zenodo on
33 acceptance of the manuscript.

34 **ABSTRACT**

35 Forecasting the responses of natural populations to environmental change is a key priority in the
36 management of ecological systems. This is challenging because the dynamics of multispecies
37 ecological communities are influenced by many factors. Populations can exhibit complex, nonlinear
38 responses to environmental change, often over multiple temporal lags. In addition, biotic
39 interactions, and other sources of multi-species dependence, are major contributors to patterns of
40 population variation. Theory suggests that near-term ecological forecasts of population abundances
41 can be improved by modelling these dependencies, but empirical support for this idea is lacking. We
42 test whether models that learn from multiple species, both to estimate nonlinear environmental
43 effects and temporal interactions, improve ecological forecasts for a semi-arid rodent community.
44 Using Dynamic Generalized Additive Models, we analyze monthly captures for nine rodents over 25
45 years. Model comparisons provide strong evidence that multi-species dependencies improve both
46 hindcast and forecast performance, as models that captured these effects gave superior predictions
47 than models that ignored them. We show changes in abundance for some species can have delayed,
48 nonlinear effects on others, and that lagged effects of temperature and vegetation greenness are
49 key drivers of change. Our findings highlight that multivariate models are useful not only to improve
50 near-term ecological forecasts but also to ask targeted questions about community dynamics.

51

52 **KEYWORDS**

53 Biotic interactions, Community dynamics, Ecological forecasting, Generalized additive model, Stan,
54 State-Space model

55 INTRODUCTION

56 Predicting the impacts of environmental change on ecosystem function and biodiversity is a global
57 challenge (Clark et al. 2001, Intergovernmental Science - Policy Platform on Biodiversity and
58 Ecosystem Services 2019, Fredston et al. 2023). Explicit predictions are needed to guide ecological
59 management decisions, inform monitoring programs, and perform scenario planning (Lindenmayer
60 et al. 2012, Tulloch et al. 2020). This has led to a growing emphasis on the importance of near-term
61 ecological forecasting to encourage greater reliance on ecological time series data, and on suitable
62 models that can handle the complexities of these data, to generate quantitative forecasts that can
63 be harnessed to guide management decisions (Dietze et al. 2018, Lewis et al. 2023, Karunarathna et
64 al. 2024). The applications of ecological forecasting are broad, including the prediction of soil
65 microbiome compositions (Averill et al. 2021), carbon cycle dynamics (Dietze et al. 2014) and
66 species' population dynamics (Ward et al. 2014, White et al. 2019, Johnson-Bice et al. 2021).

67 Forecasts for species population dynamics are especially crucial for conservation planning,
68 stock assessments and other ecological management priorities. However, these forecasts typically
69 focus on only a single species at a time (Quinn 2003, Simonis et al. 2021, Lewis et al. 2022) or on
70 aggregate measures such as species richness, biomass or diversity (Algar et al. 2009, Tonkin et al.
71 2017, Clark et al. 2020). This is problematic because key applications of population dynamics
72 forecasts, including changes in ecosystem function and biodiversity loss, are rarely single-species
73 issues (Lindenmayer et al. 2012, Greenville et al. 2016). In addition, because species differ in their
74 niche requirements, ecosystems containing multiple species of interest may require managers to
75 balance competing needs not only between human and ecosystem requirements, but also among
76 different species (e.g., Romañach et al. 2022). Finally, species population dynamics are known to be
77 related to one another due to both direct interactions between species (e.g., competition) and
78 because species respond to shared environmental drivers (Volterra 1931, Warton et al. 2015,
79 Ovaskainen et al. 2017). These associations between the dynamics of different species has resulted

80 in extensive research into multivariate population dynamics models (Ives et al. 2003, Ward et al.
81 2010, Bunin 2017, Ward et al. 2022, Paniw et al. 2023). Leveraging these associations could
82 potentially result in more accurate forecasts from multispecies models and better-informed scenario
83 planning, including approaches to predicting the impacts of species extinctions or the potential
84 spread of invasive species (Ibáñez et al. 2009).

85 However, despite the potential advantages of multispecies dynamic models, their
86 implementation is still rare in ecological applications in general and in population forecasting
87 specifically. A recent review of 178 near-term ecological forecast applications, with targets ranging
88 from wildlife population trajectories to fisheries stocks and algal bloom forecasting, found that only
89 10 (5.6%) used multivariate models to generate and evaluate forecasts (Lewis et al. 2022). This
90 finding is in line with an earlier review of population dynamics models for informing marine reserve
91 design, which found that only 1 of 34 studies considered multi-species dynamics (Gerber et al. 2003).

92 The rarity of multispecies population dynamic forecasting is likely due in part to the
93 increased computational complexity and statistical knowledge needed to formulate multivariate
94 population dynamic models that incorporate real world complexities in ecological data (Karp et al.
95 2023). Forecasting the abundances of multiple species is particularly difficult, for several reasons.
96 Many biological and physiological processes influence population dynamics (Quinn 2003, Hampton
97 et al. 2013), and species often exhibit complex responses to external drivers (including non-linear
98 responses and lags; Cárdenas et al. 2021, Karunarathna et al. 2024). Moreover, temporal
99 autocorrelation is often prevalent in abundance time series data (due to population processes; Ives
100 et al. 2010), which can be difficult to address in ecological models. Finally, because monitoring
101 wildlife is challenging, data complexities (e.g., irregular sampling intervals, observation errors,
102 missing samples, and overdispersed discrete counts with meaningful lower and/or upper bounds)
103 bring additional challenges into an already complicated modelling environment (Clark and Wells
104 2023). In combination, these issues often make population time-series data unsuitable for traditional

105 modelling approaches such as regression or simple time series models. Managers may also have
106 differing needs for forecasts, ranging from predicting the most accurate near-term population sizes
107 to exploring potential responses to differing management scenarios (Clark et al. 2001, Lindenmayer
108 et al. 2012, Moustahfid et al. 2021, Lewis et al. 2023).

109 One area of ecological modelling that has embraced multi-species approaches is Joint
110 Species Distribution Models (JSDMs), which leverage spatial patterns to predict the distribution of
111 species in space and time (Clark et al. 2016, Thorson et al. 2016, Norberg et al. 2019, Tobler et al.
112 2019, Powell-Romero et al. 2023). While many of these models only consider spatial data, recent
113 advances have included time-series structures that can learn multispecies dependencies (Ovaskainen
114 et al. 2017, Abrego et al. 2021, Ruiz-Moreno et al. 2024). While forecasting multispecies population
115 dynamics remains challenging, these types of models have the potential to provide valuable insights
116 for forecast applications. Theory and experimental evidence support the idea that learning from
117 multiple species should improve population forecasts. For example, a recent experimental study
118 induced changes in the abundance of competitors, resulting in altered species interactions that
119 impacted the accuracy of single-species forecasts (Dumandan et al. 2024). Other work has shown
120 that incorporating other species – either by including lagged observations of other species as
121 predictors in single-species models (Abrego et al. 2021, Daugaard et al. 2022) or by building
122 temporal JSDMs with multispecies autoregressive terms (Mutshinda et al. 2009, Hampton et al.
123 2013, Ovaskainen et al. 2017, Ruiz-Moreno et al. 2024) – improves the accuracy of ecological
124 predictions. But despite these findings, the broader use of multispecies forecasts as an ecological
125 application remains unexplored. Validation of multispecies forecasts, and comparisons against
126 forecasts from simpler single species models, have generally been limited to either in-sample
127 predictive measures (Sandal et al. 2022, Ruiz-Moreno et al. 2024) or one-step ahead correlation
128 measures (Ovaskainen et al. 2017, Abrego et al. 2021). We are not aware of any studies that
129 compare single species vs multispecies forecasts beyond a single time step. This is problematic
130 because most forecast applications typically require predicting multiple time steps into the future to

131 assess near-term management needs or responses to likely future scenarios (i.e. loss of important
132 species, shifts in important drivers). Moreover, most multispecies time series models fail to
133 incorporate one or more of the many important real-world complexities – observation errors,
134 missing values, non-linear responses to environmental drivers, and latent temporal dynamics – that
135 plague real-world forecasting applications (Royle and Nichols 2003, Holmes et al. 2014, Daugaard et
136 al. 2022, Clark and Wells 2023). This combination of a limited exploration of the utility of
137 multispecies models for ecological time series applications and the need to incorporate more
138 complex modeling structures constitutes a major gap in our ability to tackle realistic forecasting
139 applications.

140 Here we evaluate whether models that incorporate multi-species relationships can improve
141 near-term population forecasts using data from a long-term ecological monitoring study where there
142 is evidence of both direct biotic interactions between species (Heske et al. 1994, Ernest and Brown
143 2001, Lima et al. 2008, Bledsoe and Ernest 2019, Christensen et al. 2019a) and shared responses to
144 environmental factors (Christensen et al. 2018). Using the framework of Dynamic Generalized
145 Additive Models developed by Clark and Wells (2023), we build a series of models that learn species'
146 shared environmental responses and temporal dependencies to make inference about
147 environmental and biotic factors that relate to community dynamics. Our models highlight how
148 several key challenges can be tackled when modelling the dynamics of multiple species, including
149 how to estimate environmental effects that change nonlinearly over increasing lags, how to capture
150 unobserved temporal autocorrelation, and how to estimate lagged temporal dependencies among
151 species. We then test whether the incorporation of these biotic dependence structures improves
152 forecasts compared to simpler single-species models over multiple near-term timescales (up to 12
153 months) using penalized in-sample performance criteria and out-of-sample forecast metrics. Finally,
154 we demonstrate how these models can be used to perform perturbation experiments for assessing
155 community responses to shifts in key species abundances and to changes in environmental drivers
156 with shared species responses. Because these multi-species dynamic models integrate both species

157 interactions and complex environmental dependencies, our study shows that they can provide a
158 deeper understanding of ecological dynamics while generating more accurate forecasts and
159 predictions for scenario planning. These models are broadly applicable to time-series data, providing
160 a versatile tool for conducting time-series based forecasting to meet the wide-ranging needs of both
161 basic and applied research.

162

163 **MATERIALS AND METHODS**

164 We first describe the study system to outline why it is suitable for testing whether multi-species
165 models lead to better ecological forecasts compared to single-species models. Second, we describe
166 our full dynamic model, from which we can make inferences about the processes that drive
167 community dynamics. Third, we describe how we compare this model to simpler models in an
168 iterative forecasting exercise to ask whether models that include multi-species dependencies (a)
169 improve in-sample fits to the observed data and (b) provide better out-of-sample near-term
170 predictions.

171

172 **Rodent capture data**

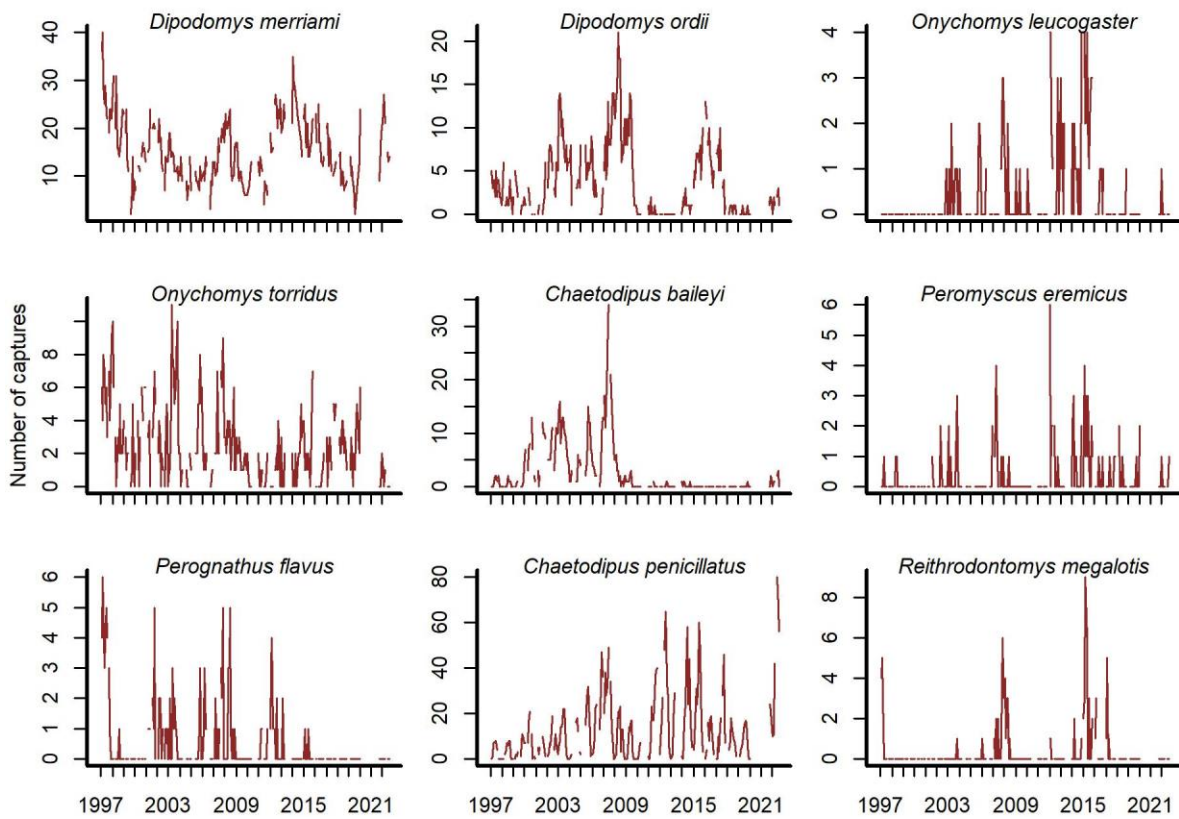
173 Our data come from the Portal Project, a long-term monitoring study of a desert rodent community
174 (Brown 1998, Ernest et al. 2020) that has been undergoing active forecasting since 2016 (White et al.
175 2019). The Portal Project is based in the Chihuahuan Desert near Portal, Arizona. The sampling
176 design includes 24 experimental plots (50m x 50m), each containing a grid of 49 baited traps (Brown
177 1998, Ernest et al. 2020). The design uses three experimental treatments. In control plots (N = 10),
178 holes in the fence are large enough to allow free access for all rodents. Full rodent removal plots (N
179 = 6) have fences with no holes. Kangaroo rat exclosures (N = 8) have fences with holes to allow
180 passage of all rodents except kangaroo rats (*Dipodomys* genus). Investigators close holes during

181 trapping to ensure all captured rodents are residents. Trapping follows the lunar monthly cycle, but
182 weather and other disruptions result in missing observations (~5% on average; Dumandan et al.
183 2023).

184 The Portal dataset exhibits many of the complexities that confront population forecasting.
185 These include observation errors due to imperfect detection, missing samples due to weather or
186 other issues (e.g., global pandemics), and overdispersed discrete counts for many species (20 rodent
187 species) that include large numbers of zeros and upper bounds set by the number of traps.
188 Environmental drivers, including temperature and measures of primary production, exhibit lagged
189 and nonlinear impacts on rodent breeding, activity rates, and population dynamics (Brown and
190 Ernest 2002). Moreover, the rodent species at Portal are known to compete for resources in
191 complex ways, and these biotic interactions are postulated to have important consequences for
192 variation in population dynamics. In other words, the Portal Project exhibits all the complexities that
193 make the ecological forecasting of species populations particularly difficult, making it an ideal real-
194 world test case for exploring whether multi-species models can provide better near-term predictions
195 than single species models.

196 Open-source software exists to access the Portal Project data (Christensen et al. 2019b,
197 Simonis et al. 2022). We used the *portalr* package to extract trapping records from the Portal data
198 (version 3.134.0; downloaded October 2022; <https://doi.org/10.5281/zenodo.7255488>). Our study
199 focused on rodent captures from the long-term control plots for the period December 1996 – August
200 2022. The data has records for 20 rodent species, but some are rarely captured. We excluded
201 species if they were observed in < 10% of trapping sessions. We did this to focus inferences on
202 species with the most influence on community dynamics. Each observation was a vector of total
203 captures on long-term control plots for the nine remaining species (Figure 1).

204



205
206

207 **Figure 1:** Rodent capture data from the Portal Project for the period December 1996 to August 2022.
208 Counts are total captures across long-term control plots. Blanks are missing values.

209

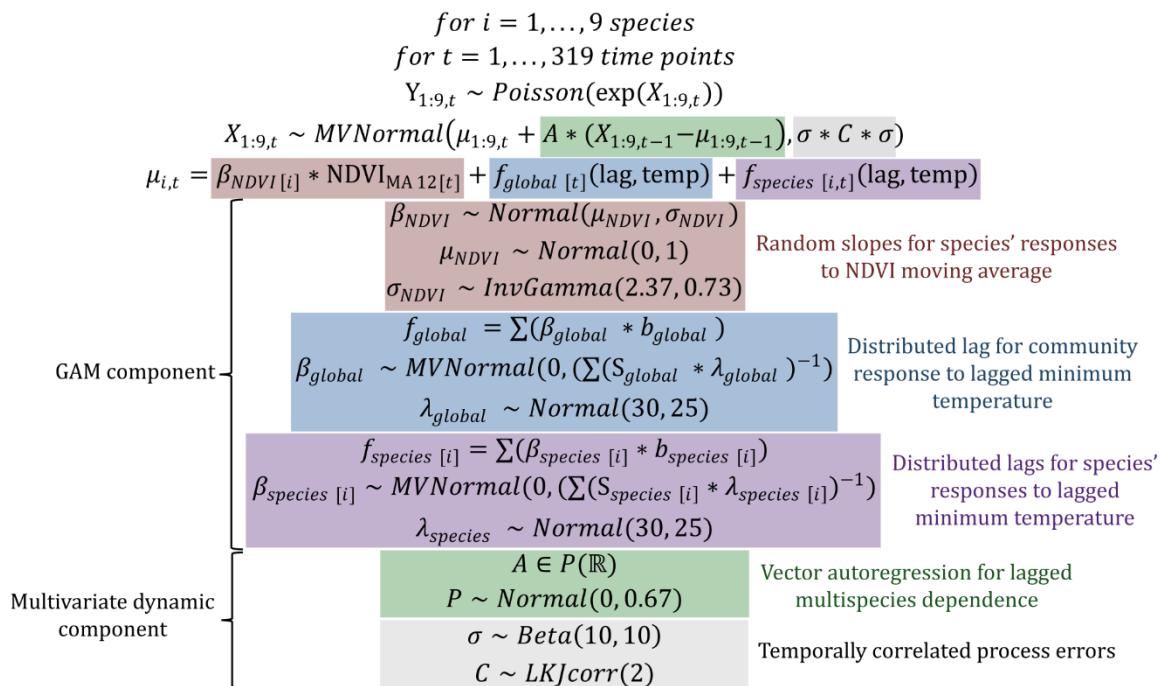
210 **Covariate measurements**

211 Rodent populations at Portal, and the associated number of captures recorded during sampling,
212 depend on environmental conditions that reflect resource availability and seasonal breeding signals.
213 We therefore modelled species' responses to environmental variation using minimum temperature
214 and the Normalized Difference Vegetation Index (NDVI) as covariates. Hourly air temperature (°C) is
215 recorded by an automated weather station, while Landsat images are used to calculate NDVI
216 (accessed from the US Geological Service Earth Resources Observation and Science Center;
217 <https://www.usgs.gov/centers/eros>). Measurements for both covariates were converted to monthly
218 averages. We extracted covariate data from one year before the start of captures (from January
219 1995) so we could calculate lagged and moving average versions. See Ernest et al. (2020) for details.

220

221 **Model description**

222 There were several aspects of the data we needed to consider when designing our model. Total
 223 rodent captures showed both short- and long-term fluctuations (Supplementary Figure S1). Captures
 224 for individual species also undulated over multi-annual cycles and were positively autocorrelated at
 225 lags up to 20 months (Supplementary Figures S2 and S3). To test whether multi-species information
 226 improves model performance, we needed to model these dynamics using a multivariate dependence
 227 structure. Second, we needed to leverage community information to estimate each species' time-
 228 delayed response to variation in vegetation and temperature. Because species' responses to
 229 environmental change in this system are expected to be nonlinear (Brown and Ernest 2002), we
 230 used splines to model these responses. Rodent captures were modelled as *Poisson* observations of
 231 a latent state model that was composed of a hierarchical GAM component (to capture shared
 232 environmental responses) and a multivariate dynamic vector autoregressive component to capture
 233 multispecies dependence. The full description for this model, which we abbreviate to **GAM-VAR**, is
 234 shown in Figure 2.



235

236 **Figure 2:** Model definition and priors for the **GAM-VAR** model. Coloured boxes highlight the five
237 main components of the latent state model (X).

238

239 The GAM component of the model consisted of hierarchical NDVI and minimum
240 temperature effects. The structural forms of these functions were informed by theory and
241 exploration of covariate time series (shown in Supplementary Figures S4-5). We used a 12-month
242 moving average of NDVI ($NDVI_{MA12}$) because we expected rodents to respond gradually to
243 vegetation change. Our model assumed linear effects of $NDVI_{MA12}$, equivalent to a hierarchical
244 slopes model. The partial pooling properties of this model allowed us to regularize weakly informed
245 slopes toward a community average. Responses to temperature were estimated using a hierarchical
246 distributed lag model in which nonlinear effects of minimum temperature varied smoothly with
247 increasing lag. These effects were constructed as tensor products of four cubic basis functions for lag
248 and three thin plate basis functions for minimum temperature. To allow our model to capitalize on
249 multi-species learning, we included a shared community-level response $f_{global}(Mintemp, lag)$ and
250 species-level deviations $f_{species[i]}(Mintemp, lag)$. The sum of these effects allowed each species to
251 show a different temperature response from the wider community, but only if there was
252 information in the data to support such a deviation. We used lags of up to six months in the past.

253 A vector autoregression (VAR) of order 1 captured lagged multispecies dependence, where
254 A was a 9 x 9 matrix of autoregressive coefficients. Diagonal entries of A described density-
255 dependence, or the effect of a species' dynamic process (at time t) on its own lagged values (at $t -$
256 1). Off-diagonals represented cross-dependencies that could provide useful biological insights into
257 interspecific interactions. For example, the entry in $A[2,3]$ described the effect of species 3's
258 dynamic state at time $t - 1$ on the current state estimate for species 2 (at time t). To encourage
259 stability and prevent forecast variance from increasing indefinitely, we enforced stationarity
260 following methods described in Heaps (2023). Briefly, a multistep process was used to map the

261 constrained A matrix to unconstrained partial autocorrelations P . Process errors were allowed to be
262 contemporaneously dependent to capture any unmodelled correlations. Priors for all model
263 components are shown in Figure 2 and described in detail in the accompanying R code.

264

265 **Evaluating whether multi-species dependencies improve prediction performance**

266 We formally tested whether learning from multiple species improved our model's predictions using
267 prediction-based model comparisons. To do so, we estimated a series of benchmark models that
268 acted as natural simplifications of the **GAM-VAR** by eliminating multi-species components in a
269 stepwise manner. The first benchmark model used the same HGAM linear predictor as the **GAM-**
270 **VAR** but replaced the multi-species VAR(1) dynamics with an AR(1) process. This model (called **GAM-**
271 **AR** in subsequent sections) eliminated the covariances and temporal cross-dependencies among
272 species' latent states, allowing us to ask whether the multivariate dynamic component was
273 supported for improving model fit. Next, we further simplified the **GAM-AR** by removing the
274 hierarchical environmental response functions from the linear predictor. This forced the model to
275 learn environmental responses for each species without using information from other species in the
276 data (**GAM-AR no pooling**). The final benchmark, referred to as **AR**, also used independent AR(1)
277 states but removed the GAM component entirely. Because this model only learned from past
278 observations, comparisons against it helped us understand how covariates impacted predictions and
279 inferences. Each candidate model was trained on all observations (through August 2022, $N = 319$
280 timepoints). Models were then compared using Pareto-smoothed importance sampling leave-one-
281 out cross-validation (PSIS-LOO), a method that reweights posterior draws to estimate leave-one-out
282 pointwise prediction accuracy using Estimated Log Predictive Density (ELPD) values (Vehtari et al.
283 2017).

284 To adequately evaluate competing forecast models, it is also necessary to perform out-of-
285 sample validation (Harris et al. 2018, Clark et al. 2022, Lewis et al. 2022). This is particularly

286 important because LOO-CV is designed to ask how models would generalize to new observations
287 within the training window. This metric does not evaluate a time series model's ability to forecast, as
288 information from future timepoints is used to influence predictions for previous time points. To
289 evaluate forecasts in a way that respected the temporal nature of our forecasting exercise, we used
290 exact leave-future-out cross-validation in an iterative expanding window framework. Models were
291 re-trained on the first 273 time points (~22 years), with the subsequent 12 time points (through
292 November 2019; selected to avoid a large sampling gap due to the COVID-19 pandemic) used to
293 evaluate forecasts. This allowed us to gauge how models might perform in a forecast scenario, but it
294 only provided a single comparison. To further scrutinize models, we retrained models on the first 75,
295 115, 154, 194, and 233 observations, and evaluated the subsequent 12 observations in each cross-
296 validation fold. All forecast comparisons used an evenly weighted combination of two proper
297 multivariate scoring rules. We chose the variogram score, which penalizes distributions that do not
298 adequately capture correlations in test observations, and the energy score, which ignores
299 correlations but penalizes forecasts if they are not well-calibrated (Scheuerer and Hamill 2015).

300

301 **Estimation**

302 We estimated posterior distributions with Hamiltonian Monte Carlo in Stan (Carpenter et al. 2017,
303 Stan Development Team 2022), specifically the *cmdstanr* interface (Gabry and Češnovar 2021).
304 Stan's algorithms provide state-of-the-art diagnostics for probabilistic models (Betancourt 2017). For
305 example, Hamiltonian Markov chains diverged when attempting to estimate minimum temperature
306 deviations for some species in the **GAM-VAR**. Our data were not informative enough to learn how,
307 or even if, these species responded to temperature change in ways that differed from the
308 community response. Stan's diagnostics guided us to a model that could be reliably estimated, which
309 included deviation functions for the four most frequently captured species (*D. ordii*, *D. merriami*,
310 *Onychomys torridus* and *C. penicillatus*). Posteriors were processed in R 4.3.1 (R Core Team 2023)

311 with the *mvgam* R package (Clark and Wells 2023). Traceplots, rank normalized split- \hat{R} and effective
312 sample sizes interrogated convergence of four parallel chains. Each chain was run for 500 warmup
313 and 1600 sampling iterations. R code to replicate all analyses and produce Figures is included in the
314 Supplementary materials and will be permanently archived on Zenodo on acceptance.

315

316 RESULTS

317 Modeling relationships among species improves prediction performance

318 Our data included total captures for nine rodent species over 319 time points. All models showed
319 adequate convergence and posterior exploration, and randomized quantile residuals showed no
320 obvious evidence of unmodelled temporal or systematic variation (Supplementary Figures S6 – S7).
321 However, in-sample performances differed, with models that leveraged multi-species information
322 producing higher ELPD scores compared to simpler models (Table 1). This was the case for all
323 stepwise comparisons apart from one: although the **GAM-AR**, which used partial pooling to learn
324 species' environmental responses, was favoured over the simpler **GAM-AR no pooling**, overlapping
325 ELPD standard errors suggested there was still large uncertainty about the magnitude of this
326 difference (Table 1).

327

328 Table 1: Approximate Pareto-smoothed importance sampling leave-one-out cross-validation (PSIS-
329 LOO) was used to compute the Estimated Log Predictive Density (ELPD) of competing models. A
330 higher ELPD indicates a model is expected to generalize better to new data within the training
331 window.

Model	ELPD difference	SE of ELPD difference
GAM-VAR	0.0	0.0

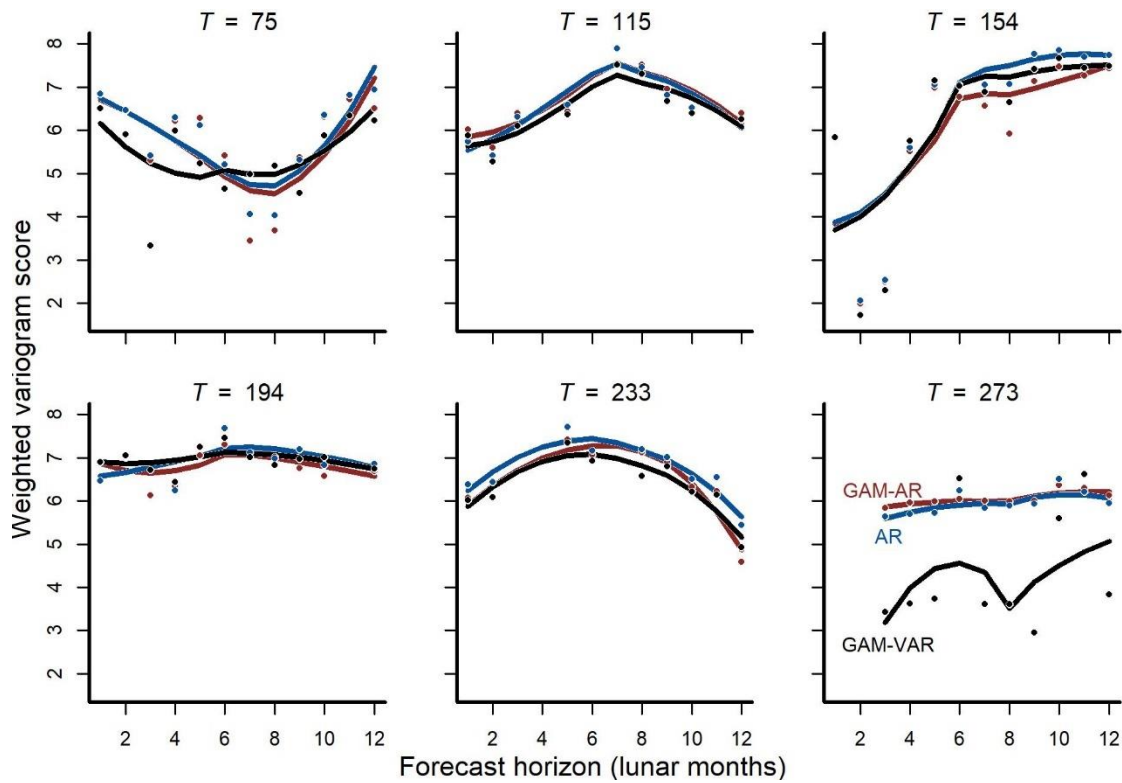
GAM-AR	-15.5	8.3
GAM-AR no pooling	-22.1	7.0
AR	-74.3	12.6

332

333 We also found that forecast performance differed among models, with more complex multi-species
334 models again tending to score higher for forecast performance than simpler models. Forecasts from
335 the multi-species **GAM-VAR** were the most accurate when considering all validation points in
336 aggregate and for 4 / 6 cross-validation folds (Figure 3; Supplementary Figure S8). The **GAM-AR** and
337 **GAM-AR no pooling** models gave similar predictions and effectively tied for second in forecast
338 performance, giving the most accurate forecasts in 2 / 6 cross-validation folds (Figure 3). The
339 simplest **AR** model gave the worst forecasts.

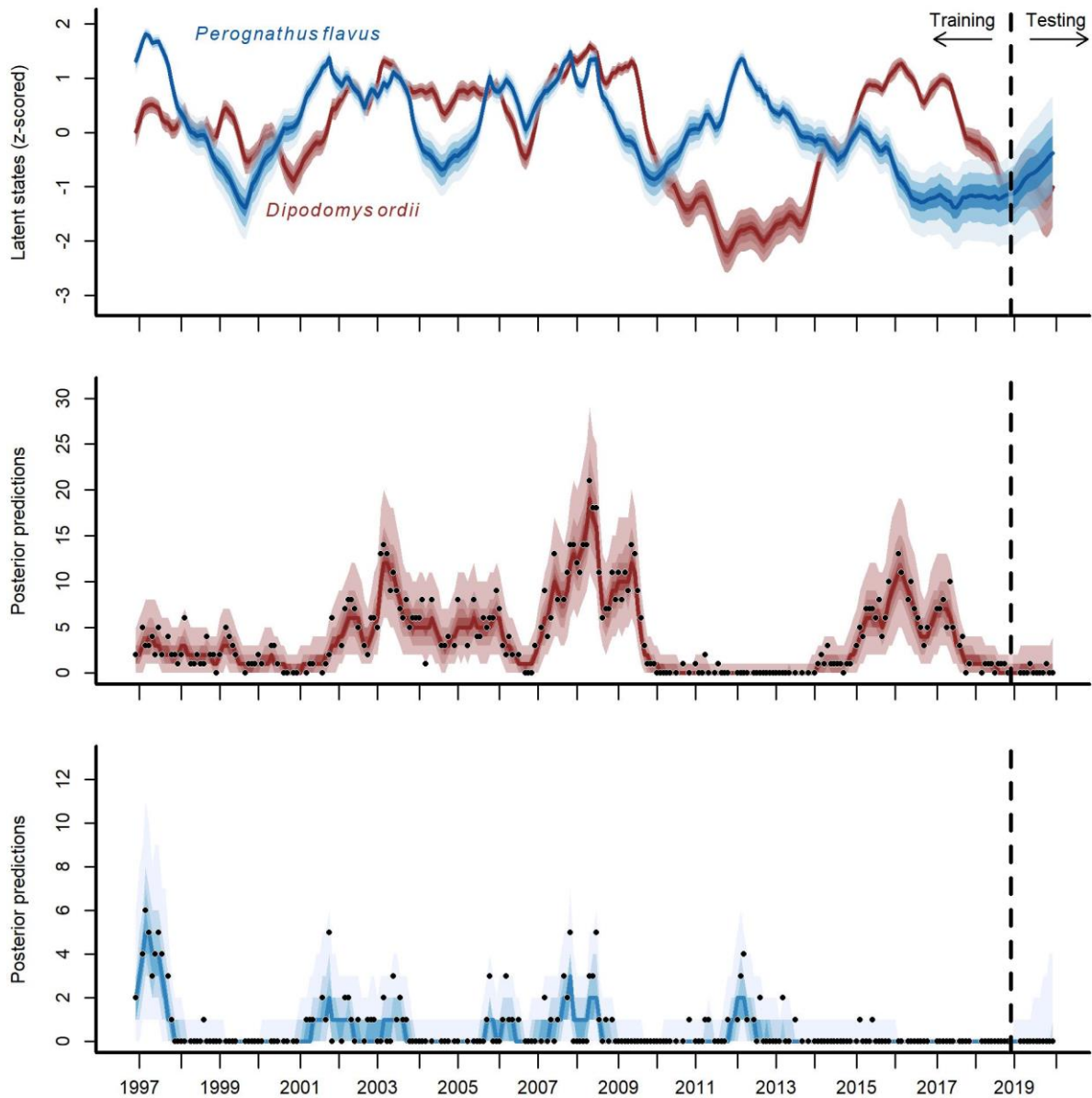
340 The multi-species **GAM-VAR** model estimated large, positive autoregressive coefficients for
341 most species (diagonal entries in Supplementary Figure S9). It also relied strongly on information
342 from multiple species by estimating many non-zero cross-dependence effects (off-diagonal entries in
343 Supplementary Figure S9) and process error correlations (Supplementary Figure S10), which
344 provided structure that the model leveraged to accurately simulate historical dynamics. The model
345 recovered multiple notable transitions observed in the time-series including a major shift in
346 community composition around 2000 following the establishment of Bailey's pocket mouse *C.*
347 *baileyi*, and a second restructuring that happened following a drought in 2008 – 09 (Supplementary
348 Figure S11). It was these multispecies effects that enabled the **GAM-VAR** to produce more accurate
349 forecasts compared to the benchmarks. For example, Ord's kangaroo rat (*D. ordii*) and silky pocket
350 mouse (*P. flavus*) had negative cross-dependencies in the **GAM-VAR**, providing structure that the
351 model used to make predictions (Figure 4). The benchmarks, which ignored this structure, produced
352 smoother, less synchronous trends and wider uncertainties (Supplementary Figure S12). In the

353 following sections, we use simulations to briefly interpret each of the multi-species effects that
 354 allowed the **GAM-VAR** to outperform simpler models.



355
 356 **Figure 3:** Cross-validation forecast performances for three of the competing models (we do not show
 357 metrics for the **GAM-AR no pooling** model as they were not clearly distinguishable from the **GAM-**
 358 **AR** metrics). Y-axis shows the log of the weighted variogram score, a scoring rule that penalizes
 359 multivariate forecasts if they are not well calibrated and do not capture inter-series correlations in
 360 observed data (lower scores are preferred). 12-step ahead predictions were evaluated over a
 361 sequence of six evenly spaced origins. Points show individual forecast scores, with lower scores
 362 indicating a better forecast. Lines show Loess smoothed trend lines. Missing points indicate that
 363 sampling did not occur at the time point for that horizon.

364



365

366 **Figure 4:** Posterior latent state estimates (top panel) and posterior predictions (bottom two panels)

367 from the **GAM-VAR** model for Ord's kangaroo rat (*Dipodomys ordii*; in red) and silky pocket mouse

368 (*Perognathus flavus*; in blue) for the training and testing periods (demarked by the vertical dashed

369 line). State estimates were scaled to unit variance for comparisons. Ribbon shading shows posterior

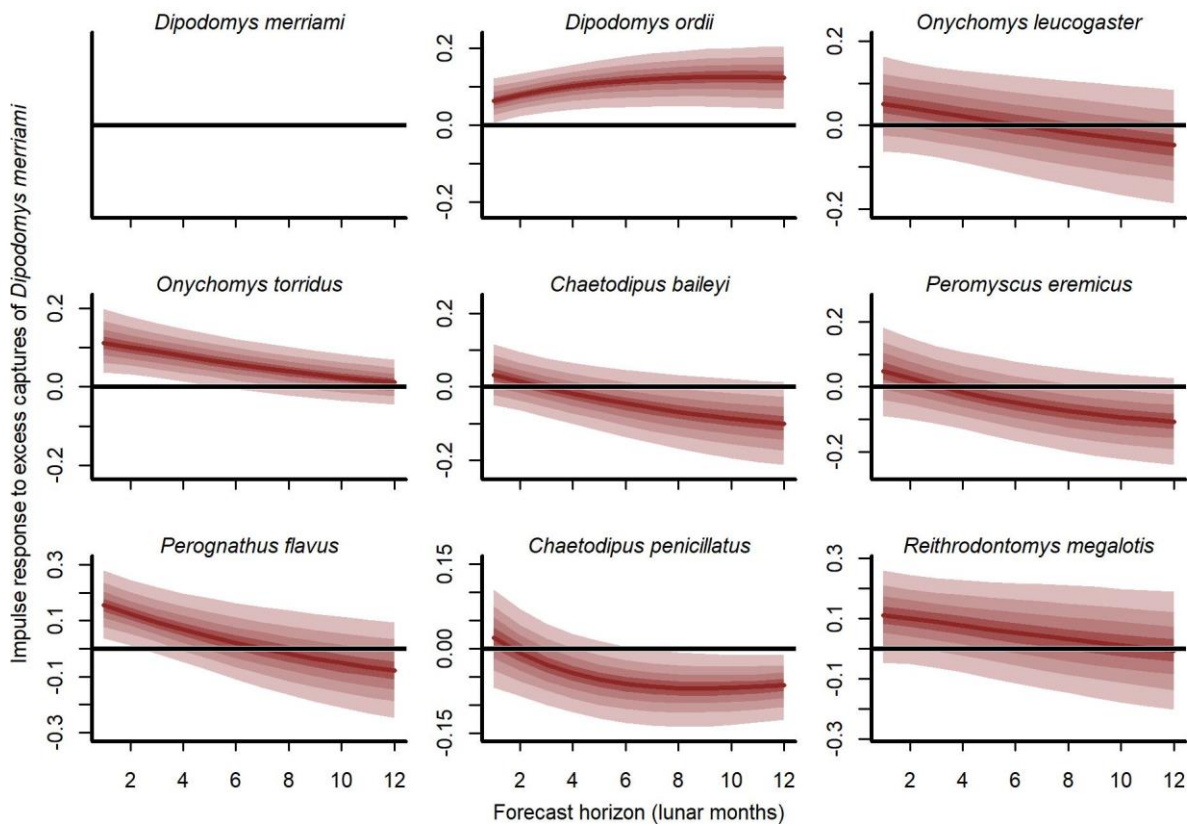
370 empirical quantiles (90th, 60th, 40th and 20th). Dark lines show posterior medians. Points show

371 observations.

372

373 **Modeling relationships among species provides unique insights into community dynamics**

374 Our cross-validation metrics strongly favoured the **GAM-VAR** and suggested that the multivariate
375 dynamic component was a particularly important driver of increased performance. Estimates of
376 process error were larger for the benchmarks than the **GAM-VAR** for nearly all species
377 (Supplementary Figure S13), suggesting this model used additional information from multi-species
378 cross-dependencies to produce better predictions. But interpreting this cross-dependence is difficult
379 because VAR effects provide only a marginal view into the complex network of conditional
380 associations. We used impulse response functions (Lütkepohl 1990) to better understand the model.
381 These functions involve simulating an ‘impulse’ in captures for one species and then evaluating how
382 predicted captures for other species changed over the next six months (Figure 5). Following a
383 simulated impulse of three extra captures for Merriam’s kangaroo rat (*D. merriami*), the model
384 expected some initial increases (due to the correlated process errors) followed by declines in
385 captures for most of the other species (Figure 5). The shapes of these declines varied by species.
386 Captures for the two pocket mouse species (*C. baileyi* and *C. penicillatus*) showed more immediate
387 declines, while the two grasshopper mouse species (*O. leucogaster* and *O. torridus*) declined more
388 gradually (Figure 5). In contrast, the other kangaroo rat species (*D. ordii*) was expected to increase
389 following a *D. merriami* pulse (Figure 5). Different effects were expected when changing the focal
390 species (Supplementary Figure S14)



391

392 **Figure 5:** Expected responses to a simulated pulse in captures of Merriam’s kangaroo rat (*D.*

393 *merriami*). Ribbon plots show how mean captures (μ , on the log scale) are expected to change over

394 the next six months if three additional *D. merriami* individuals are captured. Ribbons show posterior

395 empirical quantiles (90th, 60th, 40th and 20th). Dark red lines show posterior medians.

396

397 **Positive NDVI associations and hierarchical minimum temperature effects**

398 We found broad support for positive $NDVI_{MA12}$ associations (Figure 6). Conditional simulations,

399 which asked how rodents might respond if moved from a relatively dry/brown vegetation state to a

400 relatively moist/green vegetation state, gave higher probability to increased captures in the

401 moist/green scenario for all species. But uncertainties about this effect varied. Greatest increases

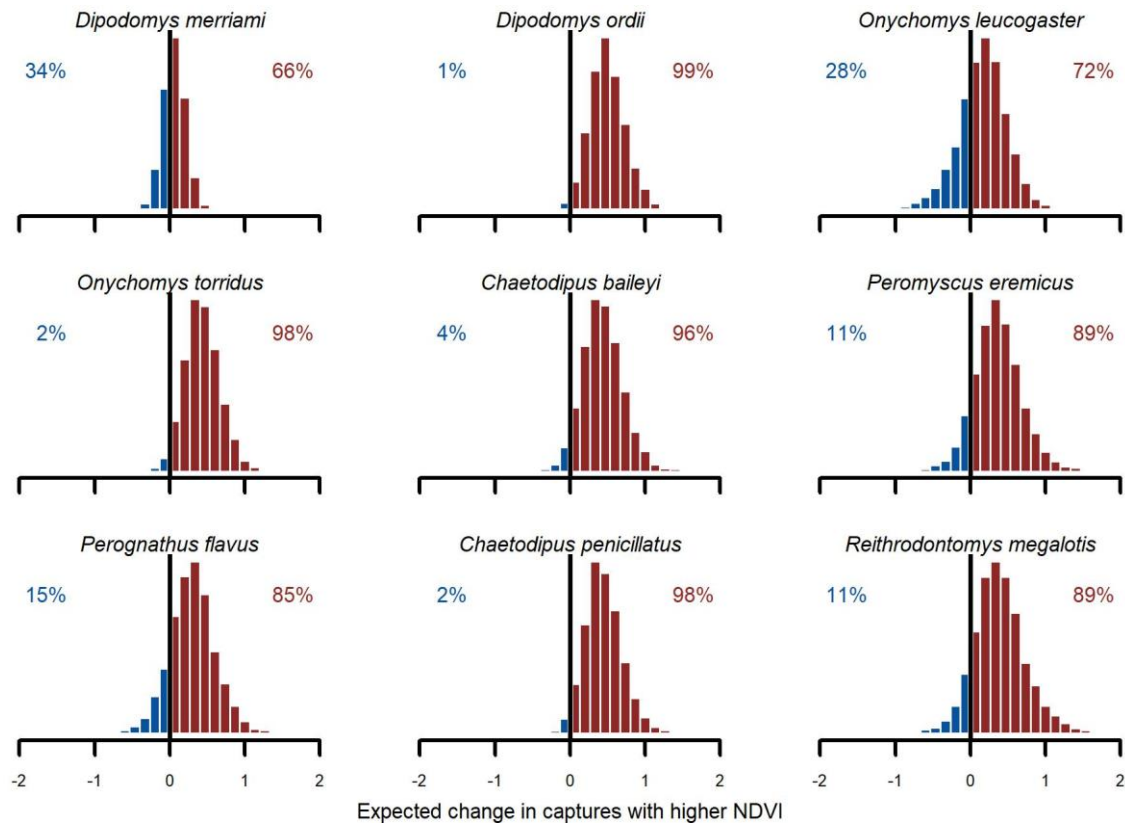
402 were expected for Ord's kangaroo rat (*D. ordii*), Western harvest mouse (*R. megalotis*) and cactus

403 mouse (*Peromyscus eremicus*). The model was less confident about the direction of effect for

404 Northern grasshopper mouse (*O. leucogaster*) and for one of the most dominant species in the

405 study, Meriam's kangaroo rat (*D. merriami*). For these species, the model expected increases in
 406 ~70% of simulations and decreases in ~30% (Figure 6). While primary conclusions were generally
 407 similar when using the **GAM-AR no pooling** model, which did not leverage multi-species learning,
 408 the estimates of these contrasts were much more variable (Figure S15).

409



411 **Figure 6:** Posterior NDVI contrasts from the hierarchical slopes component of the **GAM-VAR** model.
 412 Histograms illustrate how much the expected number of captures, $exp(\mu)$, would change if the z-
 413 scored NDVI 12-month moving average ($NDVI_{MA12}$) changed from a relatively low value (-0.50) to a
 414 relatively high value (0.50). Numbers in each plot indicate the proportion of probability mass at or
 415 below zero (in blue) vs above zero (in red).

416

417 Interpreting minimum temperature distributed lag effects also required simulation. We visualized
 418 1,000 simulated functions for each species using temperatures from 1997 (Figure S15). There was

419 large uncertainty in function shapes for all species except desert pocket mouse (*C. penicillatus*).
420 Captures for this species were expected to increase from May to October and decrease sharply in
421 winter. For seven of the other eight species, the model generally expected more captures in spring
422 (March – May) and fewer in late summer / autumn (July – October). But the shapes of these
423 responses varied. The two kangaroo rats (*D. merriami* and *D. ordii*) had smoother shapes that
424 decreased gradually from mid-summer to mid-winter. But the model expected *D. ordii* captures to
425 peak slightly later (May as opposed to March for *D. merriami*). The Southern grasshopper mouse (*O.*
426 *torridus*) was the only species that was expected to show higher captures in late autumn / early
427 winter (Figure S16). The five species that relied solely on the global function (*O. leucogaster*, *C.*
428 *baileyi*, *P. eremicus*, *P. flavus* and *R. megalotis*) were expected to show tighter spring peaks and
429 autumn troughs. When simulating from the **GAM-AR no pooling** model, the lack of multi-species
430 learning was immediately obvious. There was not enough information to learn nonlinear distributed
431 lag functions for these five species, with the model instead estimating flat functions centred on zero
432 for all five species (Figure S17).

433

434 **DISCUSSION**

435 Understanding and predicting change in species abundances requires models that capture realistic
436 biotic structure and address data complexities to produce near-term ecological forecasts (Hampton
437 et al. 2013, Holmes et al. 2014). Our results show that incorporating relationships between species
438 to estimate their lagged dependence, and to learn their potentially non-linear associations with
439 environmental drivers, yields more accurate in-sample and out-of-sample predictions. In addition to
440 improved quantitative forecasts, incorporating these multi-species complexities provides insights
441 into the dynamics of the system that could be important for scenario planning and other qualitative
442 forecasting approaches. For example, our dynamic VAR process uncovered biotic structure
443 representing a cascading network of relationships within the community. Captures for all species

444 increased with higher NDVI and responded nonlinearly to temperature change, but the shapes and
445 magnitudes of these responses differed across species. Our results show that models that describe
446 biological complexity, both through nonlinear covariate functions and multi-species dependence, are
447 useful both for generating more accurate near-term forecasts and for asking targeted questions
448 about drivers of ecological change (Ives et al. 2003, Greenville et al. 2016, Ovaskainen et al. 2017,
449 Pedersen et al. 2019).

450

451 **Leveraging relationships between species for ecological forecasting**

452 Interactions and dependencies among multiple species are hypothesized to play pivotal roles in the
453 assembly of ecological communities and to broader ecosystem functions (Dobzhansky 1950,
454 Mutshinda et al. 2009, Mayfield and Stouffer 2017, Fecchio et al. 2019). This study shows why
455 models that target multi-species effects in both their environmental responses and their biotic
456 dependencies should also be strongly considered when studying community dynamics. Our
457 approach to constructing hierarchical dynamic GAMs and evaluating forecasts using multivariate
458 proper scoring rules offers a way to quantitatively assess multispecies forecasts and scrutinize their
459 value in real-world ecological forecasting applications. We also demonstrate how inferences from
460 these models provide deeper insights into why they may or may not perform better. For example,
461 the **GAM-VAR**'s process variance estimates were smaller than those from the benchmarks because it
462 used more information from the data. By learning about the relationships between species the
463 model could better capture both shared responses to environmental factors (e.g., a wet year in the
464 desert is good for most species) and direct temporal effects (e.g., competition for seeds). These
465 relationships between species can allow forecasts for less common species to borrow strength from
466 more common species, yielding better hindcasts and forecasts compared to simpler single-species
467 models. But like other multivariate autoregressive models (Ives et al. 2003, Holmes et al. 2014,
468 Hannaford et al. 2023) the VAR parameters of the **GAM-VAR** should not be interpreted as a species

469 interaction matrix, because these relationships can result from multiple sources (i.e., shared
470 environmental responses and direct interactions). While the parameters are not interpretable as
471 direct interactions, this approach does make it possible to gain a more detailed understanding of
472 population dynamics. Conducting simulations from this type of model allows exploring which species
473 have the strongest cascading effects, what changes might we expect if management increases or
474 decreases abundance for target species, and how these effects relate to regime transitions.

475 Our hierarchical modelling approach also makes it possible to partition variance among
476 observation error, environmental responses, and multispecies dependence to guide future efforts to
477 improve ecological forecasting. In our study, forecasts were dominated by uncertainty in the
478 dynamic process model, but using a vector autoregressive process allowed us to dissect this
479 uncertainty in meaningful ways (Lütkepohl 1990, Ives et al. 2003). Simulated responses to sudden
480 impulses in captures were often delayed and nonlinear. Despite the restriction to a VAR of lag of one
481 month, these responses resulted in cascading changes that lasted up to six months. Our model's
482 ability to simulate and dissect community change in this way offers a useful avenue for ecologists to
483 better understand, and expand on, theoretical predictions from both classical and more recent
484 empirical studies that have described strong interspecific interactions in ecological systems (Volterra
485 1931, Ebersole 1977, Mayfield and Stouffer 2017, Dumandan et al. 2024).

486

487 **Learning hierarchical nonlinear effects from community data**

488 Our model captured linear, nonlinear, and lagged responses to environmental and climatic
489 covariates that were informed by data from all species at once. We found positive linear associations
490 between capture rates and a 12-month moving average of NDVI. This response was expected
491 because the rodents at Portal depend on plants for food and other resources (Brown and Ernest
492 2002, Ernest et al. 2020) and NDVI measures vegetation greenness in the landscape. Within this
493 overarching community pattern there were interesting patterns of variation in these responses

494 among species. The strongest positive association was shown by Ord's kangaroo rat (*D. ordii*), a
495 species that field evidence suggests consumes and harvests grasses (Kerley et al. 1997). In contrast,
496 Merriam's kangaroo rat (*D. merriami*) showed weaker associations with NDVI. This species has been
497 predicted to increase in prevalence in the study region with more severe droughts, in part due to a
498 preference for more open foraging habitat with less vegetation (Cárdenas et al. 2021).

499 Distributed lag functions determine the best combination of lags for environment covariates
500 but are not commonly used in ecology (but see Ogle et al. 2015, Wells et al. 2016, Karunarathna et
501 al. 2024). Our study shows how these effects can be learned hierarchically and provides useful
502 insights into delayed responses to temperature change for rodents at Portal. Most species showed
503 higher captures when minimum temperatures were low 3 – 4 months prior, suggesting increases
504 begin during mid to late spring when resources such as seeds become available. But others, such as
505 Merriam's kangaroo rat and Southern grasshopper mouse, showed increases during cooler months
506 in autumn and winter. Asynchronous phenology, where species show different reproductive timing,
507 is sometimes expected in competitive communities (Carter and Rudolf 2022). Analysis of individual
508 reproductive status in different biotic contexts suggests that some species shift their reproductive
509 timing in the presence of strong competitors in the Portal system (Dumandan et al. 2023). Do these
510 competitive forces play a role in seasonal capture variation over the long-term? Comparing
511 temperature responses on control vs experimental plots would be one interesting way to tackle this
512 question.

513 Interestingly, despite the relatively large number of observations our data contained for
514 each species, estimates of environmental responses were still more precise and informative when
515 using hierarchical models (which use partial pooling) as opposed to a no-pooling model that only
516 considers species' effects in isolation. While hierarchical intercepts and slopes are commonly used in
517 ecological models, there has been less emphasis on hierarchical nonlinear functions (but see
518 Pedersen et al. 2019). Open access to new software that makes it easy to construct and estimate

519 these types of functions, such as the *mvgam* R package that we used here (Clark and Wells 2023),
520 should improve their uptake in ecological forecasting exercises.

521 But despite the power of hierarchical environmental effects to improve predictions, we
522 cannot interpret environmental response estimates as directly causal for several reasons. First, we
523 know NDVI is not a perfect measure of changes in seed production. Second, it is likely that changes
524 to NDVI and minimum temperature are both related to other unmeasured environmental drivers
525 that may also influence rodent abundance. Major storms, the El Niño Southern Oscillation and other
526 factors that influence moisture levels can all influence temperature and vegetation change (Sun and
527 Kafatos 2007). These other drivers could act as unmeasured confounds, biasing estimates in a causal
528 inference framework (McElreath 2020).

529

530 **Future directions**

531 Two additional steps would be useful to fully assess the value of multi-species models for ecological
532 forecasting, both in this system and more broadly as an ecological application. First, a more diverse
533 suite of candidate models could be estimated to determine how forecasts could be combined into
534 an ensemble to provide the best predictions in situations where prediction accuracy is the primary
535 goal (Clark et al. 2022, Powell-Romero et al. 2023). This could be especially useful for detecting
536 changes in the system. For example, **GAM-VAR** gave better forecasts in most cross-validation tests,
537 but its performance was slightly worse than the simpler **GAM-AR** when the training window stopped
538 just prior to a major restructuring of rodent abundances that was taking place in response to a
539 drought. Second, determining which models are best for true forecasting requires evaluating
540 forecasts in the presence of uncertainty in future covariate values. In this study we were hindcasting
541 and therefore used the actual observed environmental measurements for the period reserved for
542 model evaluation. Fortunately, the system is undergoing active forecasting involving a suite of
543 simpler models and leveraging actual forecasts for environmental covariates (White et al. 2019,

544 Simonis et al. 2022). A natural next step for this work is to compare the performance of the **GAM-**
545 **VAR** model to simpler models both using hindcasting with observed covariates and when making
546 true forecasts relying on predictions instead of observations for NDVI and minimum temperature.

547 The Portal Project also provides a unique opportunity to disentangle the combined influence
548 of shared environmental responses and direct species interactions in driving observed relationships
549 between species. The site includes a long-term experimental manipulation where kangaroo rats
550 (*Dipodomys* species) are excluded from some plots. Recent work shows that single species
551 forecasting models for *C. baileyi* do not transfer well between the control plots and this
552 experimental manipulation, likely due to the different competitive environment experienced in the
553 absence of the behaviorally dominant kangaroo rats (Dumandan et al. 2023). Multi-species models
554 like the **GAM-VAR** have the potential to transfer better in situations where one or more species are
555 removed from the system by accurately predicting the response of the other species to this removal.
556 Therefore, a key next step in evaluating the potential strengths of our models is to determine if they
557 can more effectively transfer to make accurate predictions on the plots with the experimentally
558 manipulated species composition. More broadly, we hope that the ability to estimate multi-species
559 dependence and species-level variation in nonlinear environmental responses result in more
560 accurate forecasts, inspire new questions, and lead to an improved understanding of the factors that
561 govern ecological community dynamics.

562

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572

573 SUPPORTING INFORMATION CAPTIONS

574 **Figure S1:** Total rodent captures from the Portal Project for the period December 1996 to August
575 2022. Counts represent total captures for nine species across long-term control plots, sampled at
576 each cycle of the lunar moon. Blanks represent missing values.

577 **Figure S2:** Autocorrelation functions of rodent capture time series in the Portal Project. Dashed lines
578 show values beyond which the autocorrelations are considered significantly different from zero.

579 **Figure S3:** Histograms of rodent capture time series in the Portal Project. Counts represent total
580 captures across long-term control plots, sampled at each cycle of the lunar moon.

581 **Figure S4:** Seasonal and Trend decomposition using Loess smoothing (STL) applied to observed
582 minimum temperature time series for the period December 1996 – August 2022. The top panel
583 shows the raw time series. The middle plot shows the estimated long-term trend (calculated using a
584 Loess regression to the de-seasoned time series). The bottom plot shows the time-varying estimate
585 of seasonality (calculated using a Loess regression that smooths across years). STL components were
586 estimated using the *msts()* function in the *forecast* R package (Hyndman and Khandakar 2008).

587 **Figure S5:** Top panel: observed Normalized Difference Vegetation Index (NDVI) time series for the
588 period December 1996 – August 2022, with obvious seasonal fluctuations. Bottom panel: a 12-
589 month moving average that represents smooth, gradual changes in NDVI at the study site.

590 **Figure S6:** Autocorrelation functions of randomized quantile residuals from the **GAM-VAR** model.
591 Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines show
592 posterior medians. Dashed lines show values beyond which the autocorrelations would be
593 considered significantly different from zero in a Frequentist paradigm.

594 **Figure S7:** Normal quantile-quantile plots of randomized quantile residuals from the **GAM-VAR**
595 model. Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark lines show
596 posterior medians.

597 **Figure S8:** Posterior predictions from the **GAM-VAR** model for the training and testing periods
598 (demarcated by the vertical dashed line). Latent state estimates were scaled to unit variance for
599 comparisons. Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark
600 lines show posterior medians. Points show observations.

601 **Figure S9:** Posterior distributions of vector autoregressive coefficients (matrix A). Off-diagonals
602 represent cross-dependencies. For example, the entry in $A[1,2]$ captures the effect of **DO**'s state at
603 time $t - 1$ on the current state estimate for **DM** (at time t). Diagonals (with grey shading) represent
604 autoregressive coefficients (the effect of a species' state at time $t - 1$ on its own state at time t).
605 Colours indicate the proportion of probability mass at or below zero (in blue) vs above zero (in red).
606 **DO**, *Dipodomys merriami*; **DO**, *Dipodomys ordii*; **OL**, *Onychomys leucogaster*; **OT**, *Onychomys*
607 *torridus*; **PB**, *Chaetodipus baileyi*; **PE**, *Peromyscus eremicus*; **PF**, *Perognathus flavus*; **PP**, *Chaetodipus*
608 *penicillatus*; **RM**, *Reithrodontomys megalotis*.

609 **Figure S10:** Posterior distributions for process error correlations (matrix C). Colours indicate the
610 proportion of probability mass at or below zero (in blue) vs above zero (in red). **DO**, *Dipodomys*
611 *merriami*; **DO**, *Dipodomys ordii*; **OL**, *Onychomys leucogaster*; **OT**, *Onychomys torridus*; **PB**,
612 *Chaetodipus baileyi*; **PE**, *Peromyscus eremicus*; **PF**, *Perognathus flavus*; **PP**, *Chaetodipus penicillatus*;
613 **RM**, *Reithrodontomys megalotis*.

614 **Figure S11:** Simulated rodent communities. Using the **GAM-VAR** model's posterior predictive
615 distribution, we simulated communities of 200 individuals at different timepoints to investigate how
616 well the model captured known community transitions. Colours represent different species

617 **Figure S12:** Posterior trend estimates from three competing models for Ord's kangaroo rat
618 (*Dipodomys ordii*; in red) and silky pocket mouse (*Perognathus flavus*; in blue). Trends were scaled to
619 unit variance for comparisons. Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th
620 and 20th). Dark lines show posterior medians.

621 **Figure S13:** Posterior estimates of trend standard deviations from the three competing models.
622 Estimates are the square root of diagonal parameters from the trend covariance matrix (Σ_{VAR}) for
623 the **GAM-VAR** (black), **GAM-AR** (red) and **AR** (blue).

624 **Figure S14:** Expected responses to a pulse in captures of the desert pocket mouse (*Chaetodipus*
625 *penicillatus*). Ribbon plots show how mean captures (μ , on the log scale) are expected to change
626 over the next six months if three additional *C. penicillatus* individuals are captured. Ribbon shading
627 shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines show posterior
628 medians.

629 **Figure S15:** Conditional distributed lag minimum temperature functions from the hierarchical
630 smooth component of the **GAM-VAR** model, using temperatures observed in 1997. All other effects
631 were ignored. Functions for *O. leucogaster*, *C. baileyi*, *P. eremicus*, *P. flavus* and *R. megalotis* were
632 drawn solely from the global function. Functions for other species were the sum of the global
633 function and a species-specific deviation function. Estimates were scaled to unit variance for
634 comparisons. Ribbons show posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines
635 show posterior medians.

636 **Figure S16:** Conditional distributed lag minimum temperature functions from the independent
637 smooth component of the **GAM-AR no pooling** model, using temperatures observed in 1997. All
638 other effects were ignored. Functions for *O. leucogaster*, *C. baileyi*, *P. eremicus*, *P. flavus* and *R.*
639 *megalotis* were drawn solely from the global function. Functions for other species were the sum of
640 the global function and a species-specific deviation function. Estimates were scaled to unit variance
641 for comparisons. Ribbons show posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines
642 show posterior medians.

643

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