

1 **Beyond single-species forecasts: leveraging multispecies models to improve predictions and**
2 **navigate the dynamics of ecological predictability**

3 Nicholas J. Clark^{1,2*}, S. K. Morgan Ernest³, Henry Senyondo³, Juniper L. Simonis^{3,4}, Ethan P. White³,
4 Glenda M. Yenni³, K. A. N. K. Karunaratna^{1,2,5}

5

6 ¹ School of Veterinary Science, Faculty of Science, The University of Queensland, Queensland 4343,
7 Australia

8 ² UQ Spatial Epidemiology Laboratory, School of Veterinary Science, The University of Queensland,
9 Gatton, Queensland 4343, Australia

10 ³ Wildlife Ecology and Conservation, University of Florida, Gainesville, Florida 32611, USA

11 ⁴ DAPPER Stats, 3519 NE 15th Avenue, Suite 467, Portland, Oregon 97212, USA

12 ⁵ Department of Mathematics, Faculty of Science, Eastern University, Sri Lanka

13

14 ***Corresponding author:** Nicholas J Clark, n.clark@uq.edu.au

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23 **Authorship:**

24 NJC, SKM and EPW designed the study; SKM and GMY oversaw field data collection; HS, JLS and EPW
25 designed cyberinfrastructure to store and access data; NJC and KANKK performed statistical analyses
26 and visualizations; NJC drafted the manuscript and all authors contributed to revisions.

27

28 **Data availability statement:**

29 All data from the Portal Project is publicly archived to Zenodo (Ernest et al. 2023) and is available in
30 processed forms using the portalr R package (Christensen et al. 2019b). R code to reproduce
31 analyses is provided in Supplementary materials and in a GitHub repository
32 (https://github.com/nicholasjclark/portal_VAR). R code will be permanently archived on Zenodo on
33 acceptance of the manuscript.

34

35 **ABSTRACT**

36 Forecasting the responses of natural populations to environmental change is a key priority in the
37 management of ecological systems. This is challenging because the dynamics of multispecies
38 ecological communities are influenced by many factors. Populations can exhibit complex, nonlinear
39 responses to environmental change, often over multiple temporal lags. In addition, biotic
40 interactions, and other sources of multi-species dependence, are major contributors to patterns of
41 population variation. Theory suggests that near-term ecological forecasts of population abundances
42 can be improved by modelling these dependencies, but empirical support for this idea is lacking. We
43 test whether models that learn from multiple species, both to estimate nonlinear environmental
44 effects and temporal interactions, improve ecological forecasts for a semi-arid rodent community.
45 Using Dynamic Generalized Additive Models, we analyze monthly captures for nine rodents over 25
46 years. Model comparisons provide strong evidence that multi-species dependencies improve
47 performance, as models that captured these effects gave superior predictions than models that
48 ignored them. We show changes in abundance for some species can have delayed, nonlinear effects
49 on others, and that lagged effects of temperature and vegetation greenness are key drivers of
50 change. Our findings highlight that multivariate models are useful not only to improve near-term
51 ecological forecasts but also to ask targeted questions about community dynamics.

52

53 **KEYWORDS**

54 Biotic interactions, Community dynamics, Ecological forecasting, Generalized additive model, Stan,
55 State-Space model

56 INTRODUCTION

57 Predicting the impacts of environmental change on ecosystem function and biodiversity is a global
58 challenge (Clark et al. 2001, Intergovernmental Science - Policy Platform on Biodiversity and
59 Ecosystem Services 2019, Fredston et al. 2023). Explicit predictions are needed to guide ecological
60 management decisions, inform monitoring programs, and perform scenario planning (Lindenmayer
61 et al. 2012, Tulloch et al. 2020). This has led to a growing emphasis on the importance of near-term
62 ecological forecasting to encourage greater reliance on ecological time series data, and on suitable
63 models that can handle the complexities of these data, to generate quantitative forecasts that can
64 be harnessed to guide management decisions (Dietze et al. 2018, Lewis et al. 2023, Karunarathna et
65 al. 2024). The applications of ecological forecasting are broad, including the prediction of soil
66 microbiome compositions (Averill et al. 2021), carbon cycle dynamics (Dietze et al. 2014) and
67 species' population dynamics (Ward et al. 2014, White et al. 2019, Johnson-Bice et al. 2021).

68 Forecasts for species population dynamics are especially crucial for conservation planning,
69 stock assessments and other ecological management priorities. However, these forecasts typically
70 focus on only a single species at a time (Quinn 2003, Simonis et al. 2021, Lewis et al. 2022) or on
71 aggregate measures such as species richness, biomass or diversity (Algar et al. 2009, Tonkin et al.
72 2017, Clark et al. 2020). This is problematic because key applications of population dynamics
73 forecasts, including changes in ecosystem function and biodiversity loss, are rarely single-species
74 issues (Lindenmayer et al. 2012, Greenville et al. 2016). In addition, because species differ in their
75 niche requirements, ecosystems containing multiple species of interest may require managers to
76 balance competing needs not only between human and ecosystem requirements, but also among
77 different species (e.g., Romañach et al. 2022). Finally, species population dynamics are known to be
78 related to one another due to both direct interactions between species (e.g., competition) and
79 because species respond to shared environmental drivers (Volterra 1931, Warton et al. 2015,
80 Ovaskainen et al. 2017). These associations between the dynamics of different species has resulted

81 in extensive research into multivariate population dynamics models (Ives et al. 2003, Ward et al.
82 2010, Bunin 2017, Ward et al. 2022, Paniw et al. 2023). Leveraging these associations could
83 potentially result in more accurate forecasts from multispecies models and better-informed scenario
84 planning, including approaches to predicting the impacts of species extinctions or the potential
85 spread of invasive species (Ibáñez et al. 2009).

86 However, despite the potential advantages of multispecies dynamic models, their
87 implementation is still rare in ecological applications in general and in population forecasting
88 specifically. A recent review on near-term ecological forecasts found that only XX% used
89 multispecies models to generate and evaluate forecasts (Lewis et al. 2022), in line with an earlier
90 review of population dynamics models for informing marine reserve design which found that only 1
91 of 34 studies considered multi-species dynamics (Gerber et al. 2003). One area of ecological
92 modelling that has embraced multi-species approaches is Joint Species Distribution Models (JSDMs),
93 which leverage spatial patterns to predict the distribution of species in space and time (Clark et al.
94 2016, Thorson et al. 2016, Norberg et al. 2019, Tobler et al. 2019, Powell-Romero et al. 2023). While
95 many of these models only consider spatial data, recent advances have included time-series
96 structures in the form of multivariate autoregressive components (Ovaskainen et al. 2017, Abrego et
97 al. 2021).

98 The rarity of multispecies population dynamic forecasting is likely due in part to the
99 increased computational complexity and statistical knowledge needed to formulate multivariate
100 population dynamic models that incorporate real world complexities in ecological data (Karp et al.
101 2023). Forecasting the abundances of multiple species is particularly difficult, for several reasons.
102 Many biological and physiological processes influence population dynamics (Quinn 2003, Hampton
103 et al. 2013), and species often exhibit complex responses to external drivers (including non-linear
104 responses and lags; Cárdenas et al. 2021, Karunarathna et al. 2024). Moreover, temporal
105 autocorrelation is often prevalent in abundance time series data (due to population processes; Ives

106 et al. 2010), which can be difficult to address in ecological models. Finally, because monitoring
107 wildlife is challenging, data complexities (e.g., irregular sampling intervals, observation errors,
108 missing samples, and overdispersed discrete counts with meaningful lower and/or upper bounds)
109 bring additional challenges into an already complicated modelling environment (Clark and Wells
110 2023). In combination, these issues often make population time-series data unsuitable for traditional
111 modelling approaches such as regression or simple time series models. Managers may also have
112 differing needs for forecasts, ranging from predicting the most accurate near-term population sizes
113 to exploring potential responses to differing management scenarios (Clark et al. 2001, Lindenmayer
114 et al. 2012, Moustahfid et al. 2021, Lewis et al. 2023).

115 While forecasting multispecies population dynamics is challenging, recent work suggests
116 that it has the potential to provide valuable insights for forecast applications. Incorporating other
117 species – either by including lagged observations of other species as predictors in single-species
118 models (Abrego et al. 2021, Daugaard et al. 2022) or by building temporal JSDMs with multispecies
119 autoregressive terms (Hampton et al. 2013, Ovaskainen et al. 2017) can result in improved
120 predictions compared to single-species models. Experimentally induced changes in the abundance of
121 competitors also show that altered species interactions can impact the accuracy of single-species
122 forecasts (Dumandan et al. 2024). While suggestive that multi-species modelling is important, none
123 of these existing approaches incorporates the full suite of important real world complexities – data
124 issues, non-linear responses to environmental drivers, temporal autocorrelation – that are
125 necessary for many ecological forecast applications. Consequently, validation of multispecies
126 forecasts, and comparisons against forecasts from simpler single species models, are rare and have
127 generally been limited to either in-sample predictive measures (Sandal et al. 2022) or one-step
128 ahead correlation measures (Ovaskainen et al. 2017, Abrego et al. 2021). Most forecast applications
129 typically require predicting multiple time steps into the future to assess near-term management
130 needs or responses to likely future scenarios (i.e. loss of important species, shifts in important
131 drivers). The combination of a limited exploration of the utility of multispecies models for ecological

132 time series applications and the need to incorporate more complex modeling structures constitutes
133 a major gap in our ability to tackle realistic forecasting applications.

134 Here we evaluate whether models that incorporate multi-species relationships can improve
135 near-term population forecasts using data from a long-term ecological monitoring study where there
136 is evidence of both direct biotic interactions between species (Heske et al. 1994, Ernest and Brown
137 2001, Lima et al. 2008, Bledsoe and Ernest 2019, Christensen et al. 2019a) and shared responses to
138 environmental factors (Christensen et al. 2018). Using the framework of Dynamic Generalized
139 Additive Models developed by Clark and Wells (2023), we build a series of models that learn species'
140 shared environmental responses and temporal interactions to make inference about environmental
141 and biotic factors that relate to community dynamics. Our models highlight how several key
142 challenges can be tackled when modelling the dynamics of multiple species, including how to
143 estimate environmental effects that change nonlinearly over increasing lags, how to capture
144 temporal autocorrelation, and how to estimate lagged temporal dependencies among species. We
145 then test whether the incorporation of these biotic dependence structures improves forecasts over
146 multiple near-term timescales (up to 12 months) using penalized in-sample performance criteria and
147 out-of-sample forecast metrics. We also demonstrate how these models can be used to perform
148 perturbation experiments for assessing community responses to shifts in key species abundances
149 and to changes in environmental drivers with shared species responses. Because these multi-species
150 dynamic models integrate both species interactions and complex environmental dependencies, we
151 show that they can provide a deeper understanding of ecological dynamics while generating more
152 accurate forecasts and predictions for scenario planning. These models are broadly applicable to
153 time-series data, providing a versatile tool for conducting time-series based forecasting to meet the
154 wide-ranging needs of both basic and applied research.

155

156 **MATERIALS AND METHODS**

157 We first describe the study system to outline why it is suitable for testing whether multi-species
158 models lead to better ecological forecasts compared to single-species models. Second, we describe
159 our full dynamic model, from which we can make inferences about the processes that drive
160 community dynamics. Third, we describe how we compare this model to simpler models in an
161 iterative forecasting exercise to ask whether models that include multi-species dependencies (a)
162 improve in-sample fits to the observed data and (b) provide better out-of-sample near-term
163 predictions.

164

165 **Rodent capture data**

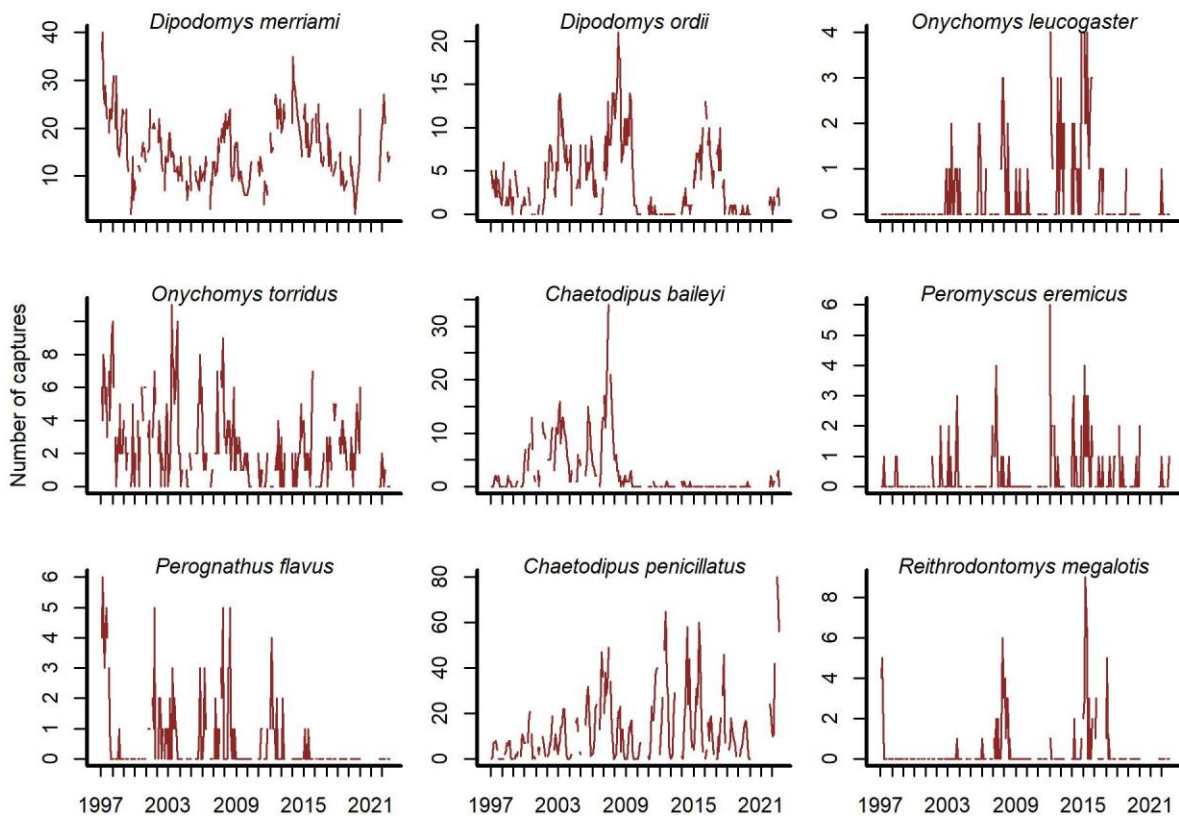
166 Our data come from the Portal Project, a long-term monitoring study of a desert rodent community
167 (Brown 1998, Ernest et al. 2020) that has been undergoing active forecasting since 2016 (White et al.
168 2019). The Portal Project is based in the Chihuahuan Desert near Portal, Arizona. The sampling
169 design includes 24 experimental plots (50m x 50m), each containing a grid of 49 baited traps (Brown
170 1998, Ernest et al. 2020). The design uses three experimental treatments. In control plots (N = 10),
171 holes in the fence are large enough to allow free access for all rodents. Full rodent removal plots (N
172 = 6) have fences with no holes. Kangaroo rat exclosures (N = 8) have fences with holes to allow
173 passage of all rodents except kangaroo rats (*Dipodomys* genus). Investigators close holes during
174 trapping to ensure all captured rodents are residents. Trapping follows the lunar monthly cycle, but
175 weather and other disruptions result in missing observations (~5% on average; Dumandan et al.
176 2023).

177 The Portal dataset exhibits many of the complexities that confront population forecasting.
178 These include observation errors due to imperfect detection, missing samples due to weather or
179 other issues (e.g., global pandemics), and over dispersed discrete counts for many species (20 rodent
180 species) that include large numbers of zeros and upper bounds set by the number of traps.
181 Environmental drivers, including temperature and measures of primary production, exhibit lagged

182 and nonlinear impacts on rodent breeding, activity rates, and population dynamics (Brown and
183 Ernest 2002). Moreover, the rodent species at Portal are known to compete for resources in
184 complex ways, and these biotic interactions are postulated to have important consequences for
185 variation in population dynamics. In other words, the Portal Project exhibits all the complexities that
186 make the ecological forecasting of species populations particularly difficult, making it an ideal real-
187 world test case for exploring whether multi-species models can provide better near-term predictions
188 than single species models.

189 Open-source software exists to access the Portal Project data (Christensen et al. 2019b,
190 Simonis et al. 2022). We used the *portalr package* to extract trapping records from the Portal data
191 (version 3.134.0; downloaded October 2022; <https://doi.org/10.5281/zenodo.7255488>). Our study
192 focused on rodent captures from the long-term control plots for the period December 1996 – August
193 2022. The data has records for 20 rodent species, but some are rarely captured. We excluded
194 species if they were observed in < 10% of trapping sessions. We did this to focus inferences on
195 species with the most influence on community dynamics. Each observation was a vector of total
196 captures on long-term control plots for the nine remaining species (Figure 1).

197



198
199

200 **Figure 1:** Rodent capture data from the Portal Project for the period December 1996 to August 2022.

201 Counts are total captures across long-term control plots. Blanks are missing values.

202

203 **Covariate measurements**

204 Rodent populations at Portal, and the associated number of captures recorded during sampling,

205 depend on environmental conditions that reflect resource availability and seasonal breeding signals.

206 We therefore modelled species' responses to environmental variation using minimum temperature

207 and the Normalized Difference Vegetation Index (NDVI) as covariates. Hourly air temperature (°C) is

208 recorded by an automated weather station, while Landsat images are used to calculate NDVI

209 (accessed from the US Geological Service Earth Resources Observation and Science Center;

210 <https://www.usgs.gov/centers/eros>). Measurements for both covariates were converted to monthly

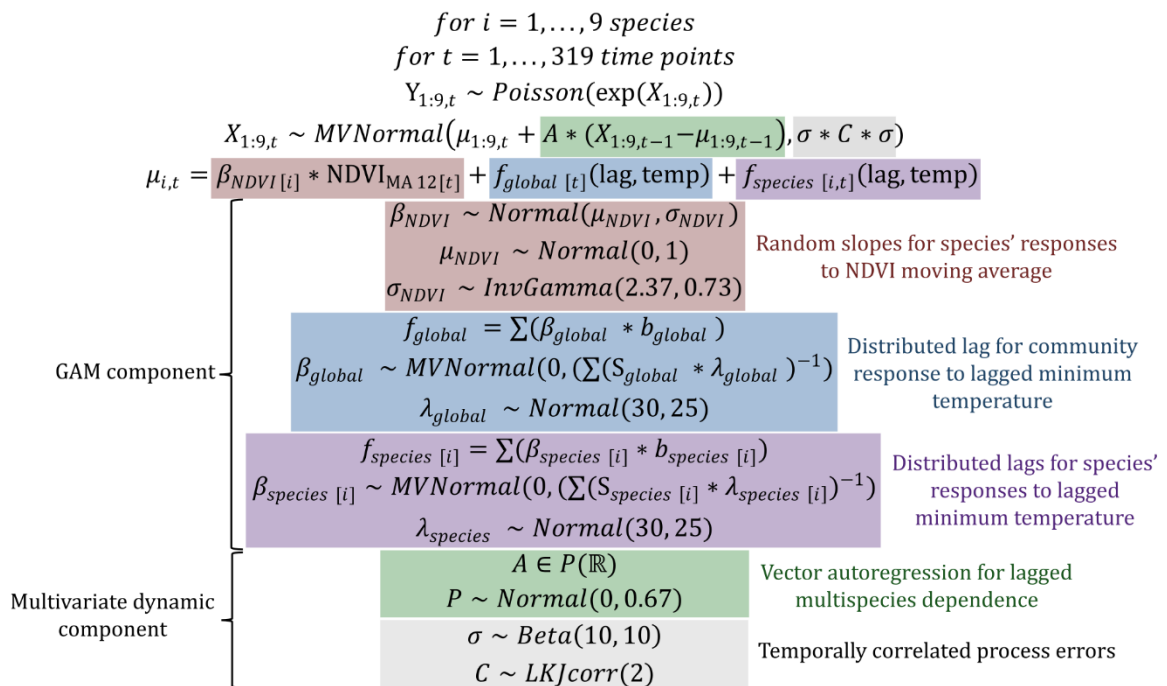
211 averages. We extracted covariate data from one year before the start of captures (from January

212 1995) so we could calculate lagged and moving average versions. See Ernest et al. (2020) for details.

213

214 **Model description**

215 There were several aspects of the data we needed to consider when designing our model. Total
 216 rodent captures showed both short- and long-term fluctuations (Supplementary Figure S1). Captures
 217 for individual species also undulated over multi-annual cycles and were positively autocorrelated at
 218 lags up to 20 months (Supplementary Figures S2 and S3). To test whether multi-species information
 219 improves model performance, we needed to model these dynamics using a multivariate dependence
 220 structure. Second, we needed to leverage community information to estimate each species' time-
 221 delayed response to variation in vegetation and temperature. Because species' responses to
 222 environmental change in this system are expected to be nonlinear (Brown and Ernest 2002), we
 223 used splines to model these responses. Rodent captures were modelled as *Poisson* observations of
 224 a latent state model that was composed of a hierarchical GAM component (to capture shared
 225 environmental responses) and a multivariate dynamic vector autoregressive component to capture
 226 multispecies dependence. The full description for this model, which we abbreviate to **GAM-VAR**, is
 227 shown in Figure 2.



228

229 **Figure 2:** Model definition and priors for the **GAM-VAR** model. Coloured boxes highlight the five
230 main components of the latent state model (X).

231

232 The GAM component of the model consisted of hierarchical NDVI and minimum
233 temperature effects. The structural forms of these functions were informed by theory and
234 exploration of covariate time series (shown in Supplementary Figures S4-5). We used a 12-month
235 moving average of NDVI ($NDVI_{MA12}$) because we expected rodents to respond gradually to
236 vegetation change. Our model assumed linear effects of $NDVI_{MA12}$, equivalent to a hierarchical
237 slopes model. The partial pooling properties of this model allowed us to regularize weakly informed
238 slopes toward a community average. Responses to temperature were estimated using a hierarchical
239 distributed lag model in which nonlinear effects of minimum temperature varied smoothly with
240 increasing lag. These effects were constructed as tensor products of four cubic basis functions for lag
241 and three thin plate basis functions for minimum temperature. To allow our model to capitalize on
242 multi-species learning, we included a shared community-level response $f_{global}(Mintemp, lag)$ and
243 species-level deviations $f_{species[i]}(Mintemp, lag)$. The sum of these effects allowed each species to
244 show a different temperature response from the wider community, but only if there was
245 information in the data to support such a deviation. We used lags of up to six months in the past.

246 A vector autoregression (VAR) of order 1 captured lagged multispecies dependence, where
247 A was a 9 x 9 matrix of autoregressive coefficients. Diagonal entries of A described density-
248 dependence, or the effect of a species' dynamic process (at time t) on its own lagged values (at $t -$
249 1). Off-diagonals represented cross-dependencies that could provide useful biological insights into
250 interspecific interactions. For example, the entry in $A[2,3]$ described the effect of species 3's
251 dynamic state at time $t - 1$ on the current state estimate for species 2 (at time t). To encourage
252 stability and prevent forecast variance from increasing indefinitely, we enforced stationarity
253 following methods described in Heaps (2023). Briefly, a multistep process was used to map the

254 constrained A matrix to unconstrained partial autocorrelations P . Process errors were allowed to be
255 contemporaneously dependent to capture any unmodelled correlations. Priors for all model
256 components are shown in Figure 2 and described in detail in the accompanying R code.

257

258 **Evaluating whether multi-species dependencies improve prediction performance**

259 We formally tested whether learning from multiple species improved our model's predictions using
260 prediction-based model comparisons. To do so, we estimated a series of benchmark models that
261 acted as natural simplifications of the **GAM-VAR** by eliminating multi-species components in a
262 stepwise manner. The first benchmark model used the same HGAM linear predictor as the **GAM-**
263 **VAR** but replaced the multi-species VAR(1) dynamics with an AR(1) process. This model (called **GAM-**
264 **AR** in subsequent sections) eliminated the covariances and temporal cross-dependencies among
265 species' latent states, allowing us to ask whether the multivariate dynamic component was
266 supported for improving model fit. Next, we further simplified the **GAM-AR** by removing the
267 hierarchical environmental response functions from the linear predictor. This forced the model to
268 learn environmental responses for each species without using information from other species in the
269 data (**GAM-AR no pooling**). The final benchmark, referred to as **AR**, also used independent AR(1)
270 states but removed the GAM component entirely. Because this model only learned from past
271 observations, comparisons against it helped us understand how covariates impacted predictions and
272 inferences. Each candidate model was trained on all observations (through August 2022, $N = 319$
273 timepoints). Models were then compared using Pareto-smoothed importance sampling leave-one-
274 out cross-validation (PSIS-LOO), a method that reweights posterior draws to estimate leave-one-out
275 pointwise prediction accuracy using Estimated Log Predictive Density (ELPD) values (Vehtari et al.
276 2017).

277 To adequately evaluate competing forecast models, it is also necessary to perform out-of-
278 sample validation (Harris et al. 2018, Clark et al. 2022, Lewis et al. 2022). This is particularly

279 important because LOO-CV is designed to ask how models would generalize to new observations
280 within the training window. This metric does not evaluate a time series model's ability to forecast, as
281 information from future timepoints is used to influence predictions for previous time points. To
282 evaluate forecasts in a way that respected the temporal nature of our forecasting exercise, we used
283 exact leave-future-out cross-validation in an iterative expanding window framework. Models were
284 re-trained on the first 273 time points (~22 years), with the subsequent 12 time points (through
285 November 2019; selected to avoid a large sampling gap due to the COVID-19 pandemic) used to
286 evaluate forecasts. This allowed us to gauge how models might perform in a forecast scenario, but it
287 only provided a single comparison. To further scrutinize models, we retrained models on the first 75,
288 115, 154, 194, and 233 observations, and evaluated the subsequent 12 observations in each cross-
289 validation fold. All forecast comparisons used an evenly weighted combination of two proper
290 multivariate scoring rules. We chose the variogram score, which penalizes distributions that do not
291 adequately capture correlations in test observations, and the energy score, which ignores
292 correlations but penalizes forecasts if they are not well-calibrated (Scheuerer and Hamill 2015).

293

294 **Estimation**

295 We estimated posterior distributions with Hamiltonian Monte Carlo in Stan (Carpenter et al. 2017,
296 Stan Development Team 2022), specifically the *cmdstanr* interface (Gabry and Češnovar 2021).
297 Stan's algorithms provide state-of-the-art diagnostics for probabilistic models (Betancourt 2017). For
298 example, Hamiltonian Markov chains diverged when attempting to estimate minimum temperature
299 deviations for some species in the **GAM-VAR**. Our data were not informative enough to learn how,
300 or even if, these species responded to temperature change in ways that differed from the
301 community response. Stan's diagnostics guided us to a model that could be reliably estimated, which
302 included deviation functions for the four most frequently captured species (*D. ordii*, *D. merriami*,
303 *Onychomys torridus* and *C. penicillatus*). Posteriors were processed in R 4.3.1 (R Core Team 2023)

304 with the *mvgam* R package (Clark and Wells 2023). Traceplots, rank normalized split- \hat{R} and effective
305 sample sizes interrogated convergence of four parallel chains. Each chain was run for 500 warmup
306 and 1600 sampling iterations. R code to replicate all analyses and produce Figures is included in the
307 Supplementary materials and will be permanently archived on Zenodo on acceptance.

308

309 RESULTS

310 Modeling relationships among species improves prediction performance

311 Our data included total captures for nine rodent species over 319 time points. All models showed
312 adequate convergence and posterior exploration, and randomized quantile residuals showed no
313 obvious evidence of unmodelled temporal or systematic variation (Supplementary Figures S6 – S7).
314 However, in-sample performances differed, with models that leveraged multi-species information
315 producing higher ELPD scores compared to simpler models (Table 1). This was the case for all
316 stepwise comparisons apart from one: although the **GAM-AR**, which used partial pooling to learn
317 species' environmental responses, was favoured over the simpler **GAM-AR no pooling**, overlapping
318 ELPD standard errors suggested there was still large uncertainty about the magnitude of this
319 difference (Table 1).

320

321 Table 1: Approximate Pareto-smoothed importance sampling leave-one-out cross-validation (PSIS-
322 LOO) was used to compute the Estimated Log Predictive Density (ELPD) of competing models. A
323 higher ELPD indicates a model is expected to generalize better to new data within the training
324 window.

Model	ELPD difference	SE of ELPD difference
GAM-VAR	0.0	0.0

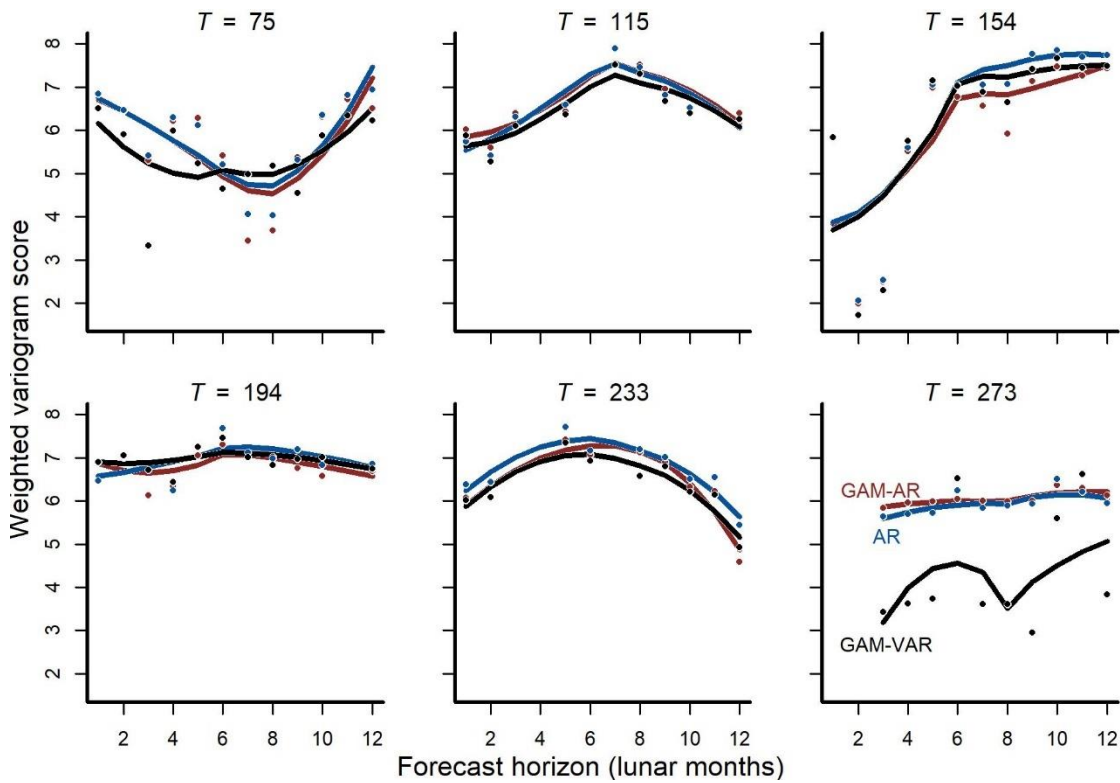
GAM-AR	-15.5	8.3
GAM-AR no pooling	-22.1	7.0
AR	-74.3	12.6

325

326 We also found that forecast performance differed among models, with more complex multi-species
327 models again tending to score higher for forecast performance than simpler models. Forecasts from
328 the multi-species **GAM-VAR** were the most accurate when considering all validation points in
329 aggregate and for 4 / 6 cross-validation folds (Figure 3; Supplementary Figure S8). The **GAM-AR** and
330 **GAM-AR no pooling** models gave similar predictions and effectively tied for second in forecast
331 performance, giving the most accurate forecasts in 2 / 6 cross-validation folds (Figure 3). The
332 simplest **AR** model gave the worst forecasts.

333 The multi-species **GAM-VAR** model estimated large, positive autoregressive coefficients for
334 most species (diagonal entries in Supplementary Figure S9). It also relied strongly on information
335 from multiple species by estimating many non-zero cross-dependence effects (off-diagonal entries in
336 Supplementary Figure S9) and process error correlations (Supplementary Figure S10), which
337 provided structure that the model leveraged to accurately simulate historical dynamics. The model
338 recovered multiple notable transitions observed in the time-series including a major shift in
339 community composition around 2000 following the establishment of Bailey's pocket mouse *C.*
340 *baileyi*, and a second restructuring that happened following a drought in 2008 – 09 (Supplementary
341 Figure S11). It was these multispecies effects that enabled the **GAM-VAR** to produce more accurate
342 forecasts compared to the benchmarks. For example, Ord's kangaroo rat (*D. ordii*) and silky pocket
343 mouse (*P. flavus*) had negative cross-dependencies in the **GAM-VAR**, providing structure that the
344 model used to make predictions (Figure 4). The benchmarks, which ignored this structure, produced
345 smoother, less synchronous trends and wider uncertainties (Supplementary Figure S12). In the

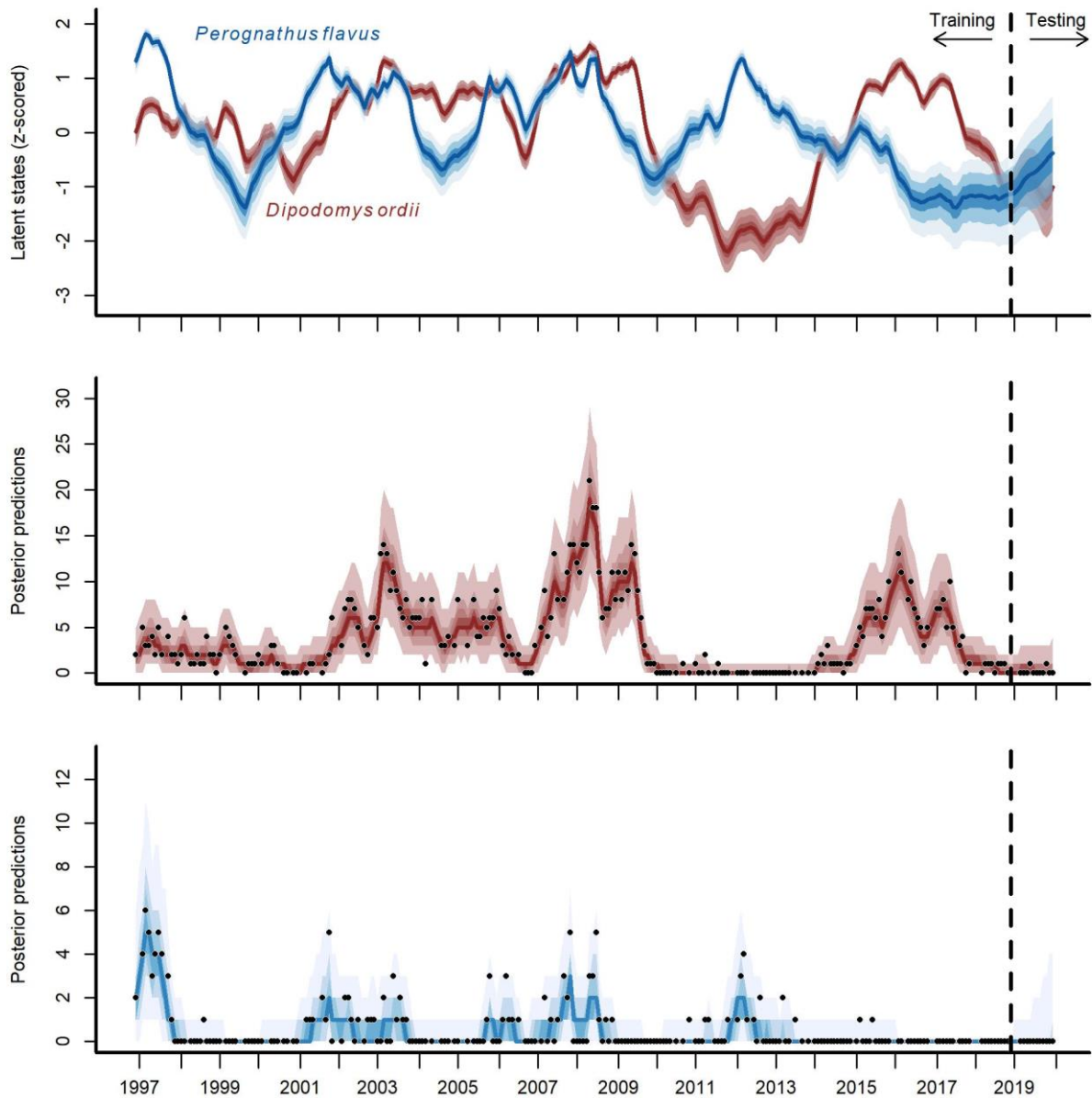
346 following sections, we use simulations to briefly interpret each of the multi-species effects that
 347 allowed the **GAM-VAR** to outperform simpler models.



348

349 **Figure 3:** Cross-validation forecast performances for three of the competing models (we do not show
 350 metrics for the **GAM-AR no pooling** model as they were not clearly distinguishable from the **GAM-**
 351 **AR** metrics). Y-axis shows the log of the weighted variogram score, a scoring rule that penalizes
 352 multivariate forecasts if they are not well calibrated and do not capture inter-series correlations in
 353 observed data (lower scores are preferred). 12-step ahead predictions were evaluated over a
 354 sequence of six evenly spaced origins. Points show individual forecast scores, with lower scores
 355 indicating a better forecast. Lines show Loess smoothed trend lines. Missing points indicate that
 356 sampling did not occur at the time point for that horizon.

357



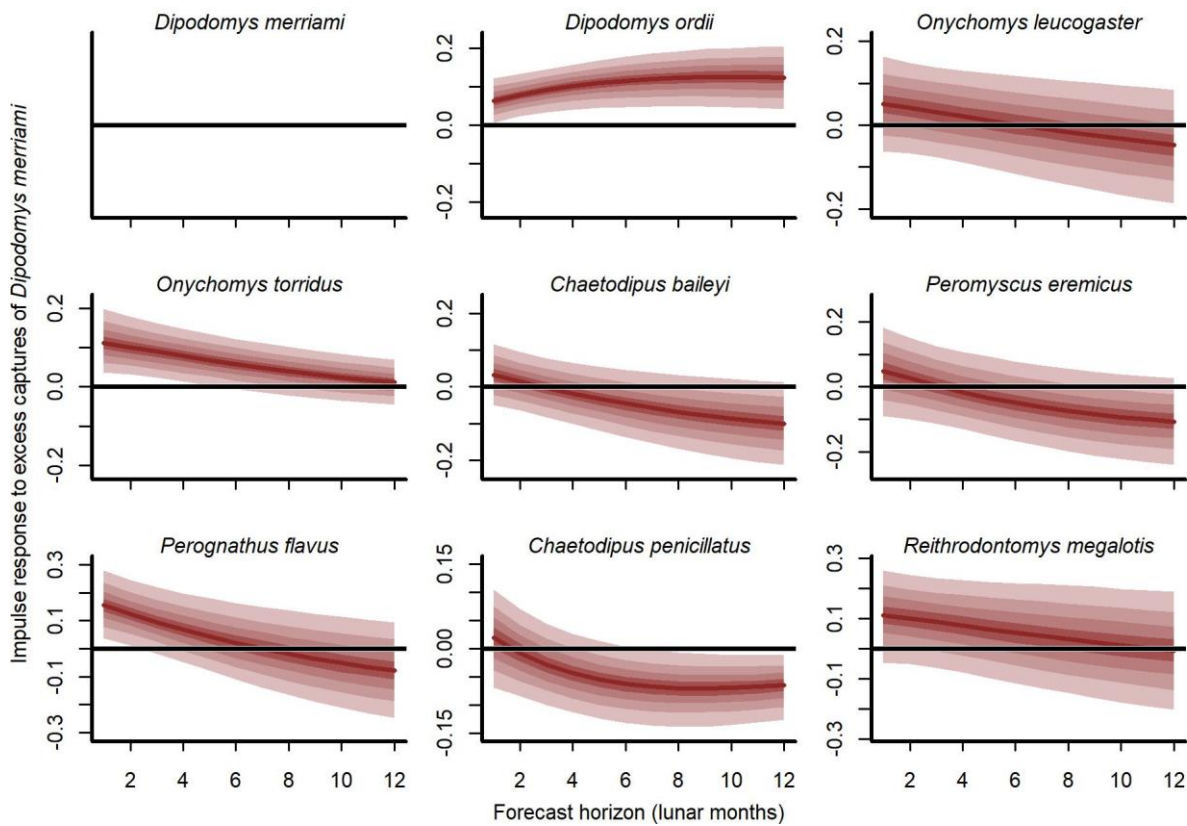
358

359 **Figure 4:** Posterior latent state estimates (top panel) and posterior predictions (bottom two panels)
 360 from the **GAM-VAR** model for Ord’s kangaroo rat (*Dipodomys ordii*; in red) and silky pocket mouse
 361 (*Perognathus flavus*; in blue) for the training and testing periods (demarked by the vertical dashed
 362 line). State estimates were scaled to unit variance for comparisons. Ribbon shading shows posterior
 363 empirical quantiles (90th, 60th, 40th and 20th). Dark lines show posterior medians. Points show
 364 observations.

365

366 **Modeling relationships among species provides unique insights into community dynamics**

367 Our cross-validation metrics strongly favoured the **GAM-VAR** and suggested that the multivariate
368 dynamic component was a particularly important driver of increased performance. Estimates of
369 process error were larger for the benchmarks than the **GAM-VAR** for nearly all species
370 (Supplementary Figure S13), suggesting this model used additional information from multi-species
371 cross-dependencies to produce better predictions. But interpreting this cross-dependence is difficult
372 because VAR effects provide only a marginal view into the complex network of conditional
373 associations. We used impulse response functions (Lütkepohl 1990) to better understand the model.
374 These functions involve simulating an ‘impulse’ in captures for one species and then evaluating how
375 predicted captures for other species changed over the next six months (Figure 5). Following a
376 simulated impulse of three extra captures for Merriam’s kangaroo rat (*D. merriami*), the model
377 expected some initial increases (due to the correlated process errors) followed by declines in
378 captures for most of the other species (Figure 5). The shapes of these declines varied by species.
379 Captures for the two pocket mouse species (*C. baileyi* and *C. penicillatus*) showed more immediate
380 declines, while the two grasshopper mouse species (*O. leucogaster* and *O. torridus*) declined more
381 gradually (Figure 5). In contrast, the other kangaroo rat species (*D. ordii*) was expected to increase
382 following a *D. merriami* pulse (Figure 5). Different effects were expected when changing the focal
383 species (Supplementary Figure S14)



384

385 **Figure 5:** Expected responses to a simulated pulse in captures of Merriam’s kangaroo rat (*D.*
 386 *merriami*). Ribbon plots show how mean captures (μ , on the log scale) are expected to change over
 387 the next six months if three additional *D. merriami* individuals are captured. Ribbons show posterior
 388 empirical quantiles (90th, 60th, 40th and 20th). Dark red lines show posterior medians.

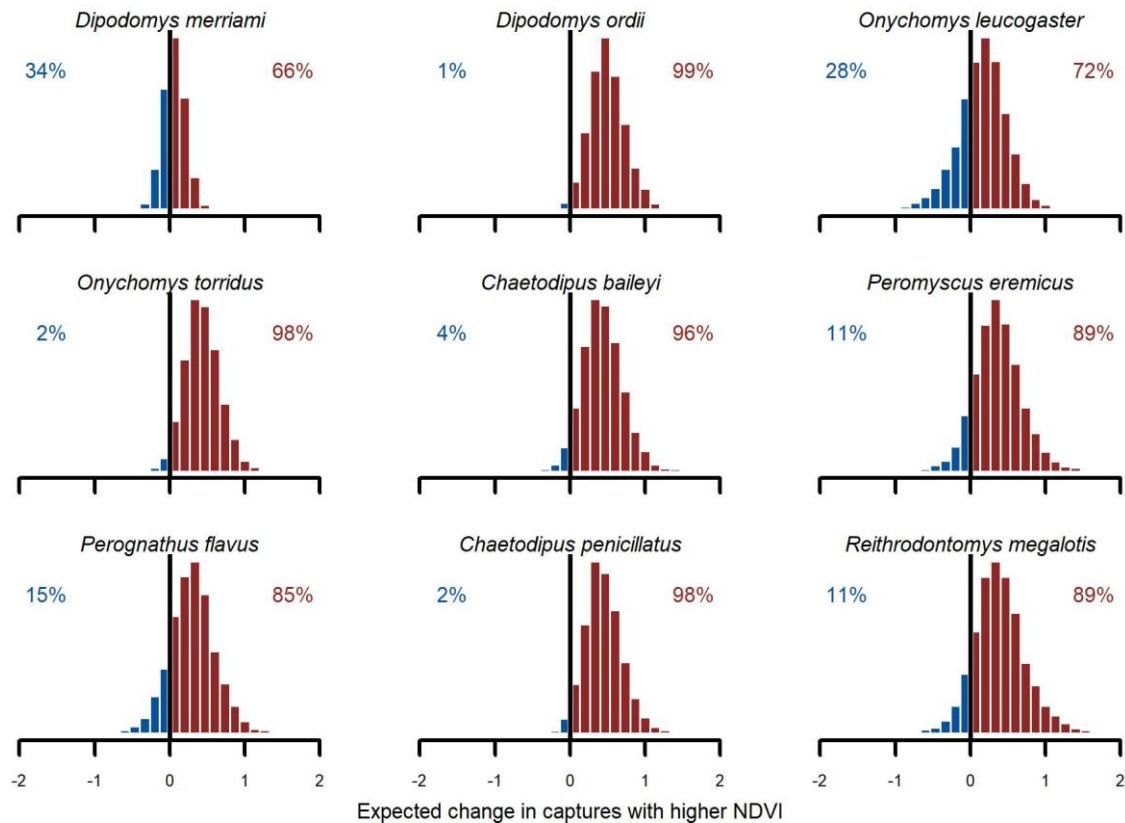
389

390 **Positive NDVI associations and hierarchical minimum temperature effects**

391 We found broad support for positive $NDVI_{MA12}$ associations (Figure 6). Conditional simulations,
 392 which asked how rodents might respond if moved from a relatively dry/brown vegetation state to a
 393 relatively moist/green vegetation state, gave higher probability to increased captures in the
 394 moist/green scenario for all species. But uncertainties about this effect varied. Greatest increases
 395 were expected for Ord's kangaroo rat (*D. ordii*), Western harvest mouse (*R. megalotis*) and cactus
 396 mouse (*Peromyscus eremicus*). The model was less confident about the direction of effect for
 397 Northern grasshopper mouse (*O. leucogaster*) and for one of the most dominant species in the

398 study, Meriam's kangaroo rat (*D. merriami*). For these species, the model expected a increases in
 399 ~70% of simulations and decreases in ~30% (Figure 6). While primary conclusions were generally
 400 similar when using the **GAM-AR no pooling** model, which did not leverage multi-species learning,
 401 the estimates of these contrasts were much more variable (Figure S15).

402



403

404 **Figure 6:** Posterior NDVI contrasts from the hierarchical slopes component of the **GAM-VAR** model.
 405 Histograms illustrate how much the expected number of captures, $exp(\mu)$, would change if the z-
 406 scored NDVI 12-month moving average ($NDVI_{MA12}$) changed from a relatively low value (-0.50) to a
 407 relatively high value (0.50). Numbers in each plot indicate the proportion of probability mass at or
 408 below zero (in blue) vs above zero (in red).

409

410 Interpreting minimum temperature distributed lag effects also required simulation. We visualized
 411 1,000 simulated functions for each species using temperatures from 1997 (Figure S15). There was

412 large uncertainty in function shapes for all species except desert pocket mouse (*C. penicillatus*).
413 Captures for this species were expected to increase from May to October and decrease sharply in
414 winter. For seven of the other eight species, the model generally expected more captures in spring
415 (March – May) and fewer in late summer / autumn (July – October). But the shapes of these
416 responses varied. The two kangaroo rats (*D. merriami* and *D. ordii*) had smoother shapes that
417 decreased gradually from mid-summer to mid-winter. But the model expected *D. ordii* captures to
418 peak slightly later (May as opposed to March for *D. merriami*). The Southern grasshopper mouse (*O.*
419 *torridus*) was the only species that was expected to show higher captures in late autumn / early
420 winter (Figure S16). The five species that relied solely on the global function (*O. leucogaster*, *C.*
421 *baileyi*, *P. eremicus*, *P. flavus* and *R. megalotis*) were expected to show tighter spring peaks and
422 autumn troughs. When simulating from the **GAM-AR no pooling** model, the lack of multi-species
423 learning was immediately obvious. There was not enough information to learn nonlinear distributed
424 lag functions for these five species, with the model instead estimating flat functions centred on zero
425 for all five species (Figure S17).

426

427 **DISCUSSION**

428 Understanding and predicting change in species abundances requires models that capture realistic
429 biotic structure and address data complexities to produce near-term ecological forecasts (Hampton
430 et al. 2013, Holmes et al. 2014). Our results show that incorporating relationships between species
431 to estimate their lagged dependence, and to learn their potentially non-linear associations with
432 environmental drivers, yields more accurate in-sample and out-of-sample predictions. In addition to
433 improved forecasts, incorporating these multi-species complexities provides opportunities for
434 interpretation that are not possible with simpler models. Our dynamic VAR process uncovered biotic
435 structure representing a cascading network of relationships within the community. Captures for all
436 species increased with higher NDVI and responded nonlinearly to temperature change, but the

437 shapes and magnitudes of these responses differed across species. Our results show that models
438 that describe biological complexity, both through nonlinear covariate functions and multi-species
439 dependence, are useful both for generating more accurate near-term forecasts and for asking
440 targeted questions about drivers of ecological change (Ives et al. 2003, Greenville et al. 2016,
441 Ovaskainen et al. 2017, Pedersen et al. 2019).

442

443 **Leveraging relationships between species for ecological forecasting**

444 Our analyses show why models that target multi-species effects in both their environmental
445 responses and their direct biotic interactions should be strongly considered when studying
446 community dynamics. The **GAM-VAR**'s process variance estimates were smaller than those from the
447 benchmarks because it used more information from the data. By learning about the relationships
448 between species the model could better capture both shared responses to environmental factors
449 (e.g., a wet year in the desert is good for most species) and direct temporal effects (e.g., competition
450 for seeds). These relationships between species can allow forecasts for less commonly observed
451 species to borrow strength from more common species, yielding better overall predictions for future
452 population abundances. But like other multivariate autoregressive models (Ives et al. 2003, Holmes
453 et al. 2014, Hannaford et al. 2023) the VAR parameters of the **GAM-VAR** should not be interpreted
454 as a species interaction matrix, because these relationships can result from multiple sources (i.e.,
455 shared environmental responses and direct interactions). While the parameters are not
456 interpretable as direct interactions, this approach does make it possible to gain a more detailed
457 understanding of population dynamics. Conducting simulations from this type of model allows
458 exploring which species have the strongest cascading effects, what changes might we expect if
459 management increases or decreases abundance for target species, and how these effects relate to
460 regime transitions. This approach also makes it possible to partition variance among observation
461 error, environmental responses, and multispecies dependence to guide future efforts to improve

462 ecological forecasting. In our study, forecasts were dominated by uncertainty in the dynamic process
463 model, but using a vector autoregressive process allowed us to dissect this uncertainty in meaningful
464 ways (Lütkepohl 1990, Ives et al. 2003). Simulated responses to sudden impulses in captures were
465 often delayed and nonlinear. Despite the restriction to a VAR of lag of one month, these responses
466 resulted in cascading changes that lasted up to six months.

467

468 **Learning hierarchical nonlinear effects from community data**

469 Our model captured linear, nonlinear, and lagged responses to environmental and climatic
470 covariates that were informed by data from all species at once. We found positive linear associations
471 between capture rates and a 12-month moving average of NDVI. This response was expected
472 because the rodents at Portal depend on plants for food and other resources (Brown and Ernest
473 2002, Ernest et al. 2020) and NDVI measures vegetation greenness in the landscape. Within this
474 overarching community pattern there were interesting patterns of variation in these responses
475 among species. The strongest positive association was shown by Ord's kangaroo rat (*D. ordii*), a
476 species that field evidence suggests consumes and harvests grasses (Kerley et al. 1997). In contrast,
477 Merriam's kangaroo rat (*D. merriami*) showed weaker associations with NDVI. This species has been
478 predicted to increase in prevalence in the study region with more severe droughts, in part due to a
479 preference for more open foraging habitat with less vegetation (Cárdenas et al. 2021).

480 Distributed lag functions determine the best combination of lags for environment covariates
481 but are not commonly used in ecology (but see Ogle et al. 2015, Wells et al. 2016). Our study shows
482 how these effects can be learned hierarchically and provides useful insights into delayed responses
483 to temperature change for rodents at Portal. Most species showed higher captures when minimum
484 temperatures were low 3 – 4 months prior, suggesting increases begin during mid to late spring
485 when resources such as seeds become available. But others, such as Merriam's kangaroo rat and
486 Southern grasshopper mouse, showed increases during cooler months in autumn and winter.

487 Asynchronous phenology, where species show different reproductive timing, is sometimes expected
488 in competitive communities (Carter and Rudolf 2022). Analysis of individual reproductive status in
489 different biotic contexts suggests that some species shift their reproductive timing in the presence of
490 strong competitors in the Portal system (Dumandan et al. 2023). Do these competitive forces play a
491 role in seasonal capture variation over the long-term? Comparing temperature responses on control
492 vs experimental plots would be one interesting way to tackle this question.

493 Interestingly, despite the relatively large number of observations our data contained for
494 each species, estimates of environmental responses were still more precise and informative when
495 using hierarchical models (which use partial pooling) as opposed to a no-pooling model that only
496 considers species' effects in isolation. While hierarchical intercepts and slopes are commonly used in
497 ecological models, there has been less emphasis on hierarchical nonlinear functions (but see
498 Pedersen et al. 2019). Open access to new software that makes it easy to construct and estimate
499 these types of functions, such as the *mvgam* R package that we used here (Clark and Wells 2023),
500 should improve their uptake in ecological forecasting exercises.

501 But despite the power of hierarchical environmental effects to improve predictions, we
502 cannot interpret environmental response estimates as directly causal for several reasons. First, we
503 know NDVI is not a perfect measure of changes in seed production. Second, it is likely that changes
504 to NDVI and minimum temperature are both related to other unmeasured environmental drivers
505 that may also influence rodent abundance. Major storms, the El Niño Southern Oscillation and other
506 factors that influence moisture levels can all influence temperature and vegetation change (Sun and
507 Kafatos 2007). These other drivers could act as unmeasured confounds, biasing estimates in a causal
508 inference framework (McElreath 2020).

509

510 **Future directions**

511 Two additional steps would be useful to fully assess the value of multi-species models for ecological
512 forecasting, both in this system and more broadly as an ecological application. First, a more diverse
513 suite of candidate models could be estimated to determine how forecasts could be combined into
514 an ensemble to provide the best predictions in situations where prediction accuracy is the primary
515 goal (Clark et al. 2022, Powell-Romero et al. 2023). This could be especially useful for detecting
516 changes in the system. For example, **GAM-VAR** gave better forecasts in most cross-validation tests,
517 but its performance was slightly worse than the simpler **GAM-AR** when the training window stopped
518 just prior to a major restructuring of rodent abundances that was taking place in response to a
519 drought. Second, determining which models are best for true forecasting requires evaluating
520 forecasts in the presence of uncertainty in future covariate values. In this study we were hindcasting
521 and therefore used the actual observed environmental measurements for the period reserved for
522 model evaluation. Fortunately, the system is undergoing active forecasting involving a suite of
523 simpler models and leveraging actual forecasts for environmental covariates (White et al. 2019,
524 Simonis et al. 2022). A natural next step for this work is to compare the performance of the **GAM-**
525 **VAR** model to simpler models both using hindcasting with observed covariates and when making
526 true forecasts relying on predictions instead of observations for NDVI and minimum temperature.

527 The Portal Project also provides a unique opportunity to disentangle the combined influence
528 of shared environmental responses and direct species interactions in driving observed relationships
529 between species. The site includes a long-term experimental manipulation where kangaroo rats
530 (*Dipodomys* species) are excluded from some plots. Recent work shows that single species
531 forecasting models for *C. baileyi* do not transfer well between the control plots and this
532 experimental manipulation, likely due to the different competitive environment experience in the
533 absence of the behaviorally dominant kangaroo rats (Dumandan et al. 2023). Multi-species models
534 like the GAM-VAR have the potential to transfer better in situations where one or more species are
535 removed from the system by accurately predicting the response of the other species to this removal.
536 Therefore, a key next step in evaluating the potential strengths of our models is to determine if they

537 can more effectively transfer to make accurate predictions on the plots with the experimentally
538 manipulated species composition. More broadly, we hope that the ability to estimate multi-species
539 dependence and species-level variation in nonlinear environmental responses result in more
540 accurate forecasts, inspire new questions, and lead to an improved understanding of the factors that
541 govern ecological community dynamics.

542

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552

553 **SUPPORTING INFORMATION CAPTIONS**

554 **Figure S1:** Total rodent captures from the Portal Project for the period December 1996 to August
555 2022. Counts represent total captures for nine species across long-term control plots, sampled at
556 each cycle of the lunar moon. Blanks represent missing values.

557 **Figure S2:** Autocorrelation functions of rodent capture time series in the Portal Project. Dashed lines
558 show values beyond which the autocorrelations are considered significantly different from zero.

559 **Figure S3:** Histograms of rodent capture time series in the Portal Project. Counts represent total
560 captures across long-term control plots, sampled at each cycle of the lunar moon.

561 **Figure S4:** Seasonal and Trend decomposition using Loess smoothing (STL) applied to observed
562 minimum temperature time series for the period December 1996 – August 2022. The top panel
563 shows the raw time series. The middle plot shows the estimated long-term trend (calculated using a
564 Loess regression to the de-seasoned time series). The bottom plot shows the time-varying estimate
565 of seasonality (calculated using a Loess regression that smooths across years). STL components were
566 estimated using the *msts()* function in the *forecast* R package (Hyndman and Khandakar 2008).

567 **Figure S5:** Top panel: observed Normalized Difference Vegetation Index (NDVI) time series for the
568 period December 1996 – August 2022, with obvious seasonal fluctuations. Bottom panel: a 12-
569 month moving average that represents smooth, gradual changes in NDVI at the study site.

570 **Figure S6:** Autocorrelation functions of randomized quantile residuals from the **GAM-VAR** model.
571 Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines show
572 posterior medians. Dashed lines show values beyond which the autocorrelations would be
573 considered significantly different from zero in a Frequentist paradigm.

574 **Figure S7:** Normal quantile-quantile plots of randomized quantile residuals from the **GAM-VAR**
575 model. Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark lines show
576 posterior medians.

577 **Figure S8:** Posterior predictions from the **GAM-VAR** model for the training and testing periods
578 (demarcated by the vertical dashed line). Latent state estimates were scaled to unit variance for
579 comparisons. Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark
580 lines show posterior medians. Points show observations.

581 **Figure S9:** Posterior distributions of vector autoregressive coefficients (matrix A). Off-diagonals
582 represent cross-dependencies. For example, the entry in $A[1,2]$ captures the effect of **DO**'s state at
583 time $t - 1$ on the current state estimate for **DM** (at time t). Diagonals (with grey shading) represent
584 autoregressive coefficients (the effect of a species' state at time $t - 1$ on its own state at time t).
585 Colours indicate the proportion of probability mass at or below zero (in blue) vs above zero (in red).
586 **DO**, *Dipodomys merriami*; **DO**, *Dipodomys ordii*; **OL**, *Onychomys leucogaster*; **OT**, *Onychomys*
587 *torridus*; **PB**, *Chaetodipus baileyi*; **PE**, *Peromyscus eremicus*; **PF**, *Perognathus flavus*; **PP**, *Chaetodipus*
588 *penicillatus*; **RM**, *Reithrodontomys megalotis*.

589 **Figure S10:** Posterior distributions for process error correlations (matrix C). Colours indicate the
590 proportion of probability mass at or below zero (in blue) vs above zero (in red). **DO**, *Dipodomys*
591 *merriami*; **DO**, *Dipodomys ordii*; **OL**, *Onychomys leucogaster*; **OT**, *Onychomys torridus*; **PB**,

592 *Chaetodipus baileyi*; **PE**, *Peromyscus eremicus*; **PF**, *Perognathus flavus*; **PP**, *Chaetodipus penicillatus*;
593 **RM**, *Reithrodontomys megalotis*.

594 **Figure S11:** Simulated rodent communities. Using the **GAM-VAR** model's posterior predictive
595 distribution, we simulated communities of 200 individuals at different timepoints to investigate how
596 well the model captured known community transitions. Colours represent different species

597 **Figure S12:** Posterior trend estimates from three competing models for Ord's kangaroo rat
598 (*Dipodomys ordii*; in red) and silky pocket mouse (*Perognathus flavus*; in blue). Trends were scaled to
599 unit variance for comparisons. Ribbon shading shows posterior empirical quantiles (90th, 60th, 40th
600 and 20th). Dark lines show posterior medians.

601 **Figure S13:** Posterior estimates of trend standard deviations from the three competing models.
602 Estimates are the square root of diagonal parameters from the trend covariance matrix (Σ_{VAR}) for
603 the **GAM-VAR** (black), **GAM-AR** (red) and **AR** (blue).

604 **Figure S14:** Expected responses to a pulse in captures of the desert pocket mouse (*Chaetodipus*
605 *penicillatus*). Ribbon plots show how mean captures (μ , on the log scale) are expected to change
606 over the next six months if three additional *C. penicillatus* individuals are captured. Ribbon shading
607 shows posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines show posterior
608 medians.

609 **Figure S15:** Posterior NDVI contrasts from the independent slopes component of the **GAM-AR no**
610 **pooling** model. Histograms illustrate how much the expected number of captures, $exp(\mu)$, would
611 change if the z-scored NDVI 12-month moving average ($NDVI_{MA12}$) changed from a relatively low
612 value (-0.50) to a relatively high value (0.50). Numbers in each plot indicate the proportion of
613 probability mass at or below zero (in blue) vs above zero (in red).

614 **Figure S16:** Conditional distributed lag minimum temperature functions from the hierarchical
615 smooth component of the **GAM-VAR** model, using temperatures observed in 1997. All other effects
616 were ignored. Functions for *O. leucogaster*, *C. baileyi*, *P. eremicus*, *P. flavus* and *R. megalotis* were
617 drawn solely from the global function. Functions for other species were the sum of the global
618 function and a species-specific deviation function. Estimates were scaled to unit variance for
619 comparisons. Ribbons show posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines
620 show posterior medians.

621 **Figure S17:** Conditional distributed lag minimum temperature functions from the independent
622 smooth component of the **GAM-AR no pooling** model, using temperatures observed in 1997. All
623 other effects were ignored. Functions for *O. leucogaster*, *C. baileyi*, *P. eremicus*, *P. flavus* and *R.*

624 *megalotis* were drawn solely from the global function. Functions for other species were the sum of
625 the global function and a species-specific deviation function. Estimates were scaled to unit variance
626 for comparisons. Ribbons show posterior empirical quantiles (90th, 60th, 40th and 20th). Dark red lines
627 show posterior medians.

628

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