

Poor nutritional condition promotes high-risk behaviours: A systematic review and meta-analysis

Short Running Title: Condition effects on risky behaviour

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1 **Abstract**

2 Animal behaviour can lead to varying levels of risk, and an individual's physical condition can
3 alter the potential costs and benefits of undertaking risky behaviours. How risk-taking
4 behaviour depends on condition is subject to contrasting hypotheses. The asset protection
5 principle proposes that individuals in better condition should be more risk averse, as they
6 have higher future reproductive potential (i.e. more to lose). The state-dependent safety
7 hypothesis proposes that high-condition individuals that are more likely to survive and
8 maximise the benefits of risky situations may make apparently riskier choices, as their
9 individual risk is in fact lower. We systematically searched for studies that experimentally
10 manipulated animals' nutritional or energetic condition through diet treatments, and
11 subsequently measured risk-taking behaviour in contexts relating to predation, novelty and
12 exploration. Our meta-analysis quantified condition effects on risk-taking behaviour at both
13 the mean and variance level. We preregistered our methods and hypotheses prior to
14 conducting the study. Phylogenetic multilevel meta-analysis revealed that the lower
15 nutritional condition individuals showed on average ca. 26% greater tendency towards risk
16 than high-condition individuals (95% confidence interval: 15% – 38%; n = 126 studies, 1297
17 effect sizes). Meta-regressions revealed several factors influencing the overall effect, such
18 as the experimental context used to measure risk-taking behaviour, and the life-stage when
19 condition was manipulated. Meta-analysis of variance revealed no clear overall effect of
20 condition on behavioural variance (on average ca. 3% decrease in variance in low- vs high-
21 condition groups; 95% confidence interval: -8% – 3%; n = 119 studies, 1235 effect sizes),
22 however, the experimental context was an important factor influencing the strength and
23 direction of the variance effect. Our comprehensive systematic review and meta-analysis
24 provide insights into the roles of state-dependency and plasticity in intraspecific behavioural
25 variation. While heterogeneity among effect sizes was high, our results show that poor
26 nutritional state on average increases risk-taking in ecological contexts involving predation,
27 novelty and exploration.

28 **Keywords:** boldness, exploration, novelty, novel environment, novel object, predation,
29 predator response, animal personality, shoaling, dietary restriction

30 Contents

| | | | |
|----|-------|---|----|
| 31 | I. | Introduction | 2 |
| 32 | II. | Methods | 7 |
| 33 | | (1) Protocol | 7 |
| 34 | | (2) Systematic review and data collection | 8 |
| 35 | | (3) Effect size calculation | 10 |
| 36 | | (4) Data analysis - main effects models | 11 |
| 37 | | (5) Data analysis - hypothesis testing models | 12 |
| 38 | | (6) Data analysis - publication bias tests | 14 |
| 39 | | (7) Data analysis - exploratory models | 15 |
| 40 | III. | Results | 15 |
| 41 | | (1) Main effects models | 16 |
| 42 | | (2) Hypothesis testing models | 16 |
| 43 | | (3) Publication bias tests | 17 |
| 44 | | (4) Exploratory models | 18 |
| 45 | IV. | Discussion | 18 |
| 46 | V. | Conclusions | 24 |
| 47 | VI. | Acknowledgements | 24 |
| 48 | VII. | Authorship | 25 |
| 49 | VIII. | Data Accessibility | 25 |
| 50 | IX. | References | 25 |

51

52 I. Introduction

53 Animals often must gamble with their lives, with behavioural decisions frequently involving
54 trade-offs between resource acquisition, reproduction and survival. Many of those decisions
55 have to be made in face of incomplete information or inherent stochasticity in the outcome.
56 Some behaviours are thus inherently 'risky' (defined as involving high outcome variance),
57 and promise large gains, but also the potential of large losses (Barclay, Mishra, & Sparks,
58 2018). The concept of risk may be applied broadly in animal ecology (e.g. participation in
59 aggressive contests, reproductive investment decisions etc.), and is often used in contexts
60 where the outcome is unpredictable (e.g. responses to novelty, *sensu* boldness; White *et al.*,
61 2013) or contexts with a high relative likelihood of death (e.g. predator responses; Réale *et*
62 *al.*, 2007). When to engage in risky behaviours is an important decision in an individual's life,
63 and thus an important research topic in behavioural ecology. State variables, such as
64 individual condition, can modify the costs and benefits of risk taking (Luttbegg & Sih, 2010).

65 State-dependency of behaviour is an important driver of among-individual variation in
66 behavioural traits (Sih *et al.*, 2015; Niemelä & Dingemanse, 2018; Moiron *et al.* 2019), but its
67 specific relationship to risk taking is subject to unresolved competing hypotheses.

68

69 Individual condition, considered here as variation in nutritional or energetic state, can lead to
70 differences in morphological, behavioural and cognitive traits among individuals

71 (Borcherding & Magnhagen, 2008; Buchanan, Grindstaff, & Pravosudov, 2013; Han &

72 Dingemanse, 2015), which can subsequently affect risk taking in different ways. Animals in

73 high condition might be risk-averse, as these individuals have a lot to lose in terms of future

74 reproductive potential (the 'asset-protection principle'; Ludwig & Rowe, 1990; Clark, 1994),

75 whereas individuals in low condition have more to gain in terms of improved condition,

76 elevated competitiveness, and starvation avoidance, particularly when an individual is

77 relatively close to their starvation threshold (Dall & Johnstone, 2002; Luttbeg & Sih, 2010;

78 also known as the 'needs-based' explanation, Barclay *et al.*, 2018). Contrastingly, the 'state-

79 dependent safety' hypothesis (also known as the 'ability-based' explanation) predicts that

80 individuals may appear to take greater risks where they are better able to survive and

81 maximise the benefits of engaging in risky behaviours, as they individually experience a

82 lower level of risk (Barclay *et al.*, 2018). State-dependent safety might apply if improved

83 condition allows greater investment in physical and/or cognitive capabilities (e.g. increased

84 vigour and/or ability to evade or defend against predation) that reduce the level of risk for the

85 individual (as in Temple, 1987).

86

87 Risk taking can depend on the current and/or past condition of an individual, and physical

88 condition in early life may have a disproportionate effect on risk-taking behaviour. For

89 example, individuals may be developmentally primed to engage in risky behaviours when

90 those behaviours were favoured early in life (Zimmer *et al.*, 2017), and poor early-life

91 environments may drive greater risk taking in adults as a way to compensate for their poor

92 start (Krause & Caspers, 2016). Conversely, a favourable nutritional environment during

93 development in particular can increase investment in traits that improve future survival and
94 fitness, such as defensive or competitive morphologies, or cognitive ability (Luttbeg & Sih,
95 2010; Buchanan *et al.*, 2013). This might allow greater risk taking if those traits provide an
96 advantage in certain risky contexts by altering effective risk levels, if for example high-
97 condition individuals are better protected/less vulnerable than low-condition individuals in the
98 same situation. Theoretical support for any one directional state-effect on risk-taking is
99 mixed, and show that the outcome may depend on environmental conditions, such as overall
100 resource availability or acuteness of the risk factor (Luttbeg & Sih, 2010; Engqvist, Cordes, &
101 Reinhold, 2014). Empirical results are similarly mixed, and thus it remains unknown if there
102 are any generally applicable effects of condition on risk-taking behaviour, or the ecological
103 context in which any one hypothesis applies.

104

105 Regardless of the hypothesis, condition effects on risk taking are often framed as adaptive
106 responses to variation in an individual's future fitness expectations (as in Clark, 1994; Wolf
107 *et al.*, 2007). The key proposition being that decisions to take risks are related to variation in
108 state, where an individual's state includes all intrinsic and extrinsic factors strategically
109 relevant for their fitness (Wolf & Weissing, 2010). State-dependent responses due to
110 nutritional condition may have interactive effects with other state variables, such as life
111 history-differences within- or among-species (McNamara & Houston, 1996). For example,
112 sex is a form of state variation involving differences in reproductive roles, which may alter
113 male and female responses to poor dietary conditions (Han & Dingemanse, 2015). In some
114 cases, males could be more sensitive to condition due to condition-dependent sexual
115 selection, but in other cases, females may be more sensitive to condition since they often
116 bear a disproportionate energetic burden of reproduction (Houslay *et al.*, 2015; English &
117 Uller, 2016). Similarly, interspecific differences in longevity may influence behavioural
118 responses, since long-lived species generally have a larger future reproductive asset and/or
119 more future opportunities to improve their own condition, and thus might be less willing to
120 display risky behaviour (Clark, 1994).

121

122 A subset of ecological contexts where variation in risk-taking behaviour can apply are those
123 involving trade-offs between resource acquisition and (implied or direct) predation risk, which
124 are often used in connection with the concept of 'boldness'. For example, responses to
125 novelty involve inherently high outcome variance, as the potential benefits and dangers of
126 novel situations are unknown to the individual. Furthermore, greater activity or exploration
127 increases the likelihood of both finding new resources or habitat patches, and encountering
128 predators (Réale *et al.*, 2007, Wohlfahrt *et al.*, 2007). Risk taking is therefore often quantified
129 in assays involving the presence of predators directly or via predation cues, which
130 emphasize the risk of mortality (Moschilla, Tomkins, & Simmons, 2018). Furthermore, some
131 studies manipulate the outcome variance of foraging-related behaviour directly (Andrews *et*
132 *al.*, 2018). Studies of risk-taking behaviour across a variety of contexts have shown different
133 responses, for example between predator and novel object experimental setups (Carter *et*
134 *al.*, 2012), or between emergence into a novel environment and startle responses
135 (Beckmann & Biro, 2013). As such, condition effects on behaviour are expected to vary
136 across experimental contexts. For example, state-dependent safety may be more relevant
137 for predator responses, if individuals that are subject to high-condition treatments are able to
138 become less vulnerable to predation.

139

140 Thus far, most studies have focused on mean behavioural effects of condition (i.e. higher or
141 lower levels of risk taking). There has, however, been growing interest in individual-level
142 variation in recent years (Westneat, Wright, & Dingemanse, 2015), and new tools to meta-
143 analyze variances alongside means are revealing that meta-variance effects may be both
144 prevalent and often overlooked (Nakagawa *et al.*, 2015). While a recent meta-analysis of
145 variance has shown diet restriction can increase variation in longevity (Senior *et al.*, 2017),
146 another has shown little evidence of environmental stress (including diet restriction) effects
147 on phenotypic behavioural variance (Sánchez-Tójar *et al.*, 2020). Furthermore, case studies
148 have shown increased within-individual behavioural variation in high-condition animals, via

149 an increased capacity to express behavioural plasticity (Royauté & Dochtermann, 2017;
150 Royauté *et al.*, 2019). Conversely, it is conceivable that extremely poor conditions may lead
151 to the expression of cryptic genetic variation, and thus increased variation in state and
152 behaviour among low-condition individuals. However, if a high-risk strategy is the only viable
153 option for acquiring adequate resources in a poor environment, individuals (including low-
154 condition individuals) may converge on a high-risk phenotype (Han & Dingemanse, 2017a).
155 Overall, condition-dependent effects on the variance in risky behaviours are likely present,
156 but currently are difficult to predict in direction and magnitude.

157

158 We here present a systematic review and meta-analysis of studies that experimentally
159 manipulated individual nutritional or energetic condition through diet quality or quantity
160 treatments, and independently quantified risk-taking behaviours such as exploration, and
161 predation and novelty responses. Specifically, we address six questions, which we
162 preregistered previous to the study (see details below):

- 163 1. Do nutritional condition manipulation treatments have an overall effect on mean risk-
164 taking behaviour? We did not predict a clear non-zero overall effect, but instead
165 expected high heterogeneity among effect sizes resulting from the various contexts in
166 which risk is measured and the multiple mechanisms that may drive condition effects
167 on risk taking.
- 168 2. Is the effect of nutritional condition on mean risk-taking behaviour context-
169 dependent? We expected low-condition treatment groups to show increased risk-
170 taking behaviour in both foraging and feeding contexts (starvation avoidance effect),
171 but reduced risk-taking behaviour in predator-response contexts (state-dependent
172 safety effect). Across the remaining contexts (e.g. novel environment exploration,
173 novel object response), we predicted high-condition treatment groups to show
174 reduced risk-taking behaviour (asset-protection effect).
- 175 3. Does nutritional condition have differential effects on mean risk-taking behaviour in
176 males and females? We did not predict an overall difference between males and

177 females, due to the high heterogeneity in sex-based ecological differentiation across
178 species. However, sex-specific differences in behaviour are widespread, and thus
179 should be quantified.

180 4. Does nutritional condition at different life stages have differential effects on mean
181 risk-taking behaviour? We expected that early-life treatments would have a greater
182 effect on mean risk-taking behaviour than late-life treatments, as early-life treatments
183 may affect mean risk-taking behaviour through both developmental and state-
184 dependent behavioural plasticity.

185 5. Does the life-history of a species determine how nutritional condition affects risk-
186 taking behaviour? We expected that a species' maximum lifespan, a key life-history
187 measure, would influence the condition effect on risk taking. According to the asset
188 protection principle, longer lived species should be less willing to display risky
189 behaviour (Clark 1994).

190 6. Does nutritional condition affect the amount of total variation in risk-taking behaviour
191 within high- and low-condition treatment groups? We did not predict an overall clear
192 variance effect between high- and low-condition experimental groups, however, as
193 for hypotheses 1 and 2, we predicted variance effects to show high heterogeneity
194 and context-dependence.

195 In addition to the hypotheses above, we conducted the following exploratory (i.e. not
196 preregistered) analyses to test for an effect of: (a) manipulation type, e.g. quantity, quality or
197 starvation treatment; (b) manipulation direction, e.g. restriction, enrichment, or combined; (c)
198 manipulation duration relative to maximum longevity; and (d) whether study subjects were
199 reared in the laboratory or the wild.

200

201 **II. Methods**

202 *(1) Protocol*

203 Study protocols (research questions, a priori hypotheses, search methods and planned
204 analyses) were registered prior to data collection to enhance the objectivity of our analyses
205 and conclusions (see preregistration at <https://osf.io/xgrkz/>, Moran *et al.*, 2018). Non-
206 preregistered analyses are hereafter labelled as exploratory. This review was conducted
207 following PRISMA reporting guidelines (Moher *et al.*, 2009, for PRISMA records see
208 Supporting Information S1 and Open Science Framework repository <https://osf.io/3tphj/>).

209

210 (2) *Systematic review and data collection*

211 Database searches were conducted in *Web of Science Core Collection* and *Scopus*, with a
212 search query designed to identify studies involving both diet manipulations (e.g. "*nutrition*",
213 "calori*", "bod* condition*") and risk-taking experiments (e.g. "bold*", "risk*", "novel*",
214 "predat*") within animal behaviour and behavioural ecology (e.g. "personalit*",
215 "temperament*", "behavio* type*", "risk taking behavio*"); for full search strategy see
216 Supporting Information S2).

217

218 We screened records to find original experimental studies that manipulated the condition of
219 animals in independent treatment groups through their diet, via both dietary quantity (i.e.
220 partial restriction, complete deprivation or enrichment) or quality treatments (e.g. protein
221 restriction or enrichment), and including both short term and longer term manipulations up to
222 extended periods of weeks-months. Then we screened for studies that then subjected those
223 animals to behavioural observations in contexts relating to risk (e.g. novel environments,
224 novel object, risk-sensitive foraging, predator response) in independent trials (for inclusion
225 and exclusion decision trees see Supporting Information S1). Our aim was to test for
226 adaptive condition-dependent behavioural responses in non-human animals, therefore we
227 excluded studies using species with compromised genetic diversity and/or evolved adaptive
228 responses (e.g. domesticated animals, laboratory breeds, genetically modified organisms; as
229 per Kelly *et al.*, 2018) as well as studies on humans. Studies manipulating the micronutrient
230 content of diets, or subjecting animals to high fat diets were also excluded as the relationship

231 between these diet manipulations and body condition is not clear and considered beyond the
232 scope of this review. Dietary treatments were excluded as 'non-independent': where the
233 behaviour was measured in the presence of high and low food availability, or dietary
234 treatments such as periods of deprivation were applied within the novel environment (i.e.
235 non-independence of treatments from the behavioural assay); where the dietary treatments
236 were coupled with additional non-dietary factors (non-independence of the diet factor within
237 treatments; e.g. temperature); or, the dietary treatments were applied longitudinally (within
238 individuals) rather than cross-sectionally (i.e. non-independence between high and low
239 treatments).

240

241 Both the title and abstract screening of 5453 records (post-deduplication), and the full-text
242 screening of 641 published papers were conducted by two authors (NPM 100%, AST 25% at
243 both stages) to ensure reliability. Title and abstract screening was done using Rayyan
244 (Ouzzani *et al.*, 2016), from which 626 references were included for full-text screening. The
245 title and abstract screening resulted in 67/1377 (4.9%) conflicted decisions between
246 observers, confirming high inter-screener agreement. All conflicted decisions were resolved
247 collectively by both screeners. A few additional references that were not captured by our
248 search but instead identified from different sources were also included for full-text screening
249 ('non-systematic' records, n = 15). Data from five such papers were included in the final
250 analysis, therefore we conducted a sensitivity analysis to test the potential effects of these
251 additional five references by re-fitting the main effects models without these effect sizes, and
252 results remain very similar (see Supporting Information S3). Full-text screening of 641
253 papers resulted in 5/160 (3.1%) conflicted decisions (i.e. where one screener included a
254 reference, and the other excluded it), that were resolved collectively by both screeners. Full-
255 text screening identified 147 studies meeting the experimental design criteria for inclusion
256 (see <https://osf.io/3tphj/> for full-text screening decision database
257 'CD_FulltextScreeningDatabase.xlsx', and Supporting Information S1 for the PRISMA
258 diagram and the decision tree summarizing the full-text exclusion reasons).

259

260 Data were extracted as comparisons between the *low-condition* groups (i.e. the treatment
261 group for diet restriction treatments, the control group for diet enrichment treatments) and
262 the *high-condition* groups (i.e. the control group for diet restriction treatments, and the
263 treatment group for diet enrichment treatments). Extractions were conducted by NPM with
264 data extracted from figures where necessary using the R package 'metaDigitise' v1.0.0
265 (Pick, Nakagawa, & Noble, 2019b). Data required to calculate effect sizes were (a) group
266 means and (b) estimates of uncertainty (standard error, confidence intervals) or variability
267 (standard deviation, SD) in combination with sample sizes (N) for the behavioural variables
268 of interest. Full or partial extraction of relevant data was possible from the published material
269 of 118 studies (80.2% of all studies included after full-text screening). To recover missing or
270 partially reported data, corresponding authors of 72 studies were contacted via a
271 standardized author correspondence email, such that 395 (29.6%) of 1334 effect sizes in the
272 full final dataset were obtained via author correspondence. Data from 25% of included
273 papers (37 papers) were re-extracted by an independent observer to ensure data reliability.
274 Of 1420 re-extracted values, errors requiring correction were identified in only 6 values
275 (0.4%) affecting only two effect sizes included in the final analyses.

276

277 (3) Effect size calculation

278 We analysed mean effects using the log response ratio of group means ('lnRR'; Hedges,
279 Gurevitch, & Curtis, 1999), instead of Cohen's D or Hedge's g, as lnRR is less sensitive to
280 heteroscedasticity. Variance effects were analyzed using the log coefficient of variation ratio
281 ('lnCVR'), as this effect size, unlike log ratio of variances ('lnVR'), is less sensitive to
282 potential mean-variance correlations (Nakagawa *et al.*, 2015). Both ratios were calculated
283 using *low condition* over *high condition*, such that a positive effect size represents higher risk
284 taking or larger variance in risk taking in low-condition animals, respectively (effect sizes
285 calculated via R package 'metafor' version v2.1-0, Viechtbauer, 2010). To maintain
286 consistent directionality, effect sizes were reversed for a subset of lnRR effect sizes where

287 lower values reflected higher risk behaviours (e.g. 'latency to emerge from a shelter',
288 'distance from a predator' etc.). Since lnCVR directionality is independent of the mean, sign
289 reversals were not required. To assess if our choice of effect sizes affected our conclusions,
290 main effects analyses were also run using alternate effect sizes for mean (standardised
291 mean difference with heteroscedasticity correction 'SMDH'; Bonett, 2009), and variance
292 (lnVR; Nakagawa *et al.*, 2015). Conclusions remained robust (see Supporting Information S4
293 for details).

294

295 *(4) Data analysis - main effects models*

296 Two multilevel intercept-only meta-analytic models were fitted for each effect size, testing for
297 a general effect of condition treatments on risk-taking behaviour at a mean and variance
298 level (using the function 'rma.mv' from the R package 'metafor' v2.1-0, Viechtbauer, 2010).
299 Phylogenetic and non-phylogenetic models were fitted to investigate whether non-
300 independence due to the degree of relatedness between species influenced both the overall
301 effects and their level of uncertainty. Phylogenetic relatedness were estimated based on
302 existing phylogenies and taxonomic information from the Open Tree of Life, and any
303 polytomies were resolved by randomization (Hinchliff *et al.*, 2015; via R package 'rotl' v3.0.7;
304 Michonneau, Brown, & Winter, 2016; for the final phylogenetic tree see Supporting
305 Information S5). Branch lengths were estimated using Grafen's method (Grafen, 1989; via R
306 package 'ape' v5.3; Paradis & Schliep, 2019), and were used to construct a phylogenetic
307 variance-covariance relatedness matrix.

308

309 In addition to phylogeny, we included other random effects in our models to account for non-
310 independence due to the use of the same species across studies (SpeciesID), multiple effect
311 sizes taken from the same study (StudyID), and multiple effect sizes taken from the same
312 experimental group of animals within the same behavioural experiment (ExperimentallID). A
313 unit level random effect (EffectID) was also included to estimate residual heterogeneity. For
314 a subset of effect sizes, an experimental group was compared to multiple treatment groups

315 (i.e. shared-control non-independence). Sampling variances were modeled as variance-
316 covariance matrices that accounted for correlated sampling variances due to the shared
317 group designs, and were constructed following Lajeunesse (2011; for estimation methods
318 see Supporting Information S4).

319

320 A subset of studies used a crossed factorial experimental design by applying an additional
321 treatment factor (e.g. diet x temperature treatments; juvenile x adult dietary treatments etc.).
322 To avoid including variance associated with the additional treatment factor in our analysis,
323 we combined groups across the treatment factor that was not of interest to us (e.g. low
324 condition/low temperature and low condition/high temperature). Groups were combined by
325 calculating marginalised means and SDs (following equations for pooled means and SDs
326 from Pick *et al.*, 2019a).

327

328 For main effects models, we investigated total, residual and random effect specific relative
329 heterogeneity by calculating ' I^2 ' values (Nakagawa & Santos, 2012, via R package
330 v0.0.0.9000 'MetaAidR', Noble, 2019), and estimated absolute heterogeneity ' Q '. For
331 moderator models, we calculated the percentage of heterogeneity explained by the inclusion
332 of moderators in each model ' R^2_{marginal} ' (Nakagawa & Schielzeth 2013, via function 'r2_ml', R
333 package 'oRchard', Nakagawa *et al.* 2020), the residual heterogeneity ' Q_E ', and moderator
334 specific heterogeneity ' Q_M ' (i.e. an omnibus test of coefficients as implemented in the
335 function 'rma.mv' via R package 'metafor' v2.1-0, Viechtbauer, 2010). Where applicable,
336 estimates are presented with 95% confidence intervals in square brackets (hereafter simply
337 referred to as 'confidence interval').

338

339 (5) Data analysis - hypothesis testing models

340 All hypotheses were tested using phylogenetic multilevel meta-regression models for both
341 InRR and InCVR including random effects as above (for detailed descriptions of all
342 moderators used for hypothesis testing models see Supporting Information S6).

343

344 First, we included a categorical moderator (*'RiskContext'*) to test if effects were context-
345 dependent by classifying behavioural variables by both the functional context of the
346 experiment (e.g. assays involving predators or predator cues, novel objects, novel
347 environments etc.; Luttbeg & Sih, 2010) and the specific behavioural measurements (e.g.
348 activity levels, areas explored, willingness to feed and forage, shoaling tendencies etc.; for
349 descriptions of all categories see Supporting Information S6). Second, a categorical
350 moderator (*'Sex'*) tested for differences between male and female experimental groups.
351 Effect sizes were calculated separately for males and females where sufficient data was
352 available, otherwise effect sizes were categorized as mixed (i.e. groups including both
353 sexes), or unknown (i.e. no information about the sex of study subjects). Third, a categorical
354 moderator (*'ManipLifeStage'*) tested for an effect of life-stage at the time of the treatments,
355 with the level of maturity during diet manipulations categorised as juvenile, adult, both (i.e.
356 for treatments spanning both periods), or unknown/mixed. If the paper did not present
357 sufficient information to determine the subject's life-stage, this was inferred from the
358 available information (e.g. age, average length, weight etc.) where possible. If life-stage
359 could not be reasonably inferred or if groups may have included both juvenile and adult
360 individuals, these were classed together as mixed/unknown. Since treatments in juveniles
361 may have been imposed a longer time before behavioural testing (e.g. early-life diet
362 treatments with adult behavioural testing) relative to adult diet treatments, life-stage models
363 also included the time between condition treatment(s) and behavioural experiments relative
364 to the species maximum longevity as a continuous moderator
365 (*'RelativeTimeFromTreatment.C'*). Finally, to assess the role of life-history variation among
366 species, we separately tested for effects of maximum lifespan (*'MaxLongevity.C'*) and the
367 natural logarithm of maximum lifespan (*'InMaxLongevity.C'*) as continuous moderators. Log
368 transformed lifespan was used to better capture the variability in lifespan between species,
369 as estimates for included species were heavily biased towards short lifespans. Lifespan
370 estimates were obtained from online databases (AnAge, *genomics.senescence.info*;

371 FishBase, *fishbase.se*, Animal Diversity Web, *animaldiversity.org*; Longevity Records,
372 *demogr.mpg.de/longevityrecords*). If no estimates were available, *ad hoc* searches for
373 lifespan estimates from primary literature were conducted via *Google Scholar*. Where
374 available, sex-specific and wild/captive-specific longevity estimates were used. Continuous
375 moderators were z-transformed to aid interpretation (Schielzeth, 2010).

376

377 (6) Data analysis - publication bias tests

378 Several meta-regression models were used to assess our InRR dataset for evidence of
379 publication bias (for all included moderators and descriptions see Supporting Information
380 S6).

381

382 First, the precision of each effect was included as a moderator, calculated as the square root
383 of the inverse sampling variance (*'Precision'*, a variant of an Egger's regression based on
384 Nakagawa & Santos, 2012), to test for small-study bias. Next, time-lag bias was tested using
385 the year of publication as a continuous moderator (*'Year.C'*), where a commonly observed
386 trend is a decrease in effect size over time (Jennions & Møler, 2002; Sánchez-Tójar *et al.*,
387 2018). For both the precision and time-lag models, a limited dataset excluding effect sizes
388 obtained through author correspondence was used so that we were specifically testing for
389 effects of publication bias in published material. Finally, using the full dataset, we used a
390 categorical moderator to test whether effect sizes were larger in studies with partial or
391 incomplete reporting of results (*'EffectSizesFromPublication'*, i.e. complete, partial or none;
392 where none refers to studies where all effect sizes had to be obtained via author
393 correspondence). In addition, funnel plots were produced using InRR and precision for a
394 visual assessment of funnel asymmetry (Nakagawa & Santos, 2012; for plots see Supporting
395 Information S7). As there appeared to be some evidence of publication bias, we also
396 calculated fail-safe N to test the robustness of our results (function *'fsn'*, R package *'metafor'*
397 v2.1-0, Viechtbauer, 2010; see Supporting Information S7). Publication bias tests were not
398 conducted for InCVR, as the overwhelming majority of papers were focused on effects at the

399 mean behavioural level, with very few testing for effects on behavioural variance, so we did
400 not expect publication bias on InCVR.

401

402 *(7) Data analysis - exploratory models*

403 Additional exploratory analyses (i.e. not preregistered) were included to test if differences in
404 the experimental designs of included studies influenced the results of both InRR and InCVR
405 (for moderators and descriptions see Supporting Information S6).

406

407 We tested a categorical moderator based on the differing types of diet manipulation included
408 in our analysis (*'ManipType'*). This included quantity (where the amount of food ration/food
409 access differed between groups), starvation (where one group was entirely deprived of food
410 for an extended period), quality (where the nutritional content of food differed between
411 groups) or combined (where both quality and quantity was manipulated in the same
412 treatment group). Since our main models compared low- versus high-condition treatment
413 groups regardless of whether diets corresponded to restriction or supplementation
414 treatments, we also explored potential effects of this by including a categorical moderator
415 (*'ManipDirection'*). This categorised treatments as restriction (where low-condition groups
416 were restricted relative to high condition/control groups), supplementation (where high
417 condition groups were enriched relative to low-condition/control groups), and dual (where
418 both the low-condition group was restricted and the high condition group was enriched from
419 standard conditions). To explore how the duration of diet treatments influenced the outcome,
420 a continuous moderator (*'RelativeManipDuration.C'*) was defined as the time that the
421 treatment was applied as a proportion of the maximum lifespan of the species. Finally, the
422 influence of the source of the study subjects was tested using a categorical moderator
423 (*'WildLabRear'*, wild, laboratory, commercial or mixed).

424

425 **III. Results**

426 (1) *Main effects models*

427 Intercept-only models showed a significant positive effect for lnRR (i.e. log response ratio),
428 with the mean estimate corresponding to a 26% increase in risk-taking behaviour in animals
429 subject to low-condition treatments compared to those from high-condition treatments (non-
430 phylogenetic method: lnRR = 0.23 [0.15 – 0.32], phylogenetic method: lnRR = 0.23 [0.09 –
431 0.38]; Table 1, Figure 1). For lnCVR (i.e. log coefficient of variation ratio), the overall
432 estimate was small, negative and the confidence intervals overlapped zero substantially
433 (lnCVR = -0.03 [-0.09 – 0.03]; Table 1, Figure 1). As phylogeny failed to resolve any
434 heterogeneity in lnCVR, the estimates from the phylogenetic and non-phylogenetic models
435 were identical.

436

437 (2) *Hypothesis testing models*

438 The magnitude of the lnRR was influenced by the experimental context, with the *RiskContext*
439 moderator explaining a large amount of heterogeneity among effect sizes ($R^2_{\text{marginal}} = 8.96\%$;
440 Table 2). Although most context-specific confidence intervals overlapped with zero, all the
441 mean estimates were positive (Table 4). The highest estimates were found for behaviours
442 relating to feeding under predation (lnRR = 0.75 [0.53 – 0.97]), feeding in a novel
443 environment (lnRR = 0.36 [0.20 – 0.52]), and shoaling in a novel environment (lnRR = 0.37
444 [0.06 – 0.67]; Table 4; Fig 2A). The risk context also explained a large amount of
445 heterogeneity in lnCVR ($R^2_{\text{marginal}} = 8.58\%$; Table 3), and the confidence intervals of some
446 context-specific effects did not overlap with zero, including refuge use in a novel
447 environment (lnCVR = 0.18 [0.04 – 0.31]), feeding in a novel environment (lnCVR = -0.16 [-
448 0.25 – -0.07]), and, dispersal/migration decisions (lnCVR = -0.49 [-0.86 – -0.11]; Table 5; Fig
449 2B), showing a reduction in total variance in low- vs. high-condition treatments in those
450 specific risk contexts.

451

452 Sex appeared to have some effect on lnRR (Table 2), but there was little evidence for an
453 effect on lnCVR (Table 3). The lnRR estimates were positive but the confidence intervals

454 slightly overlapped with zero for both females (lnRR = 0.15 [-0.03 – 0.33]) and males (lnRR
455 = 0.12 [-0.06 – 0.30]), while effects were strongest for mixed (lnRR = 0.34 [0.06 – 0.61]) and
456 unknown sex groups (lnRR = 0.29 [0.14 – 0.45]; Fig 2C). Life-stage also influenced lnRR
457 (Table 2), and less clearly also lnCVR (although this model showed a relatively high R^2_{marginal}
458 = 5.28%, Table 3). Life-stage specific estimates for lnRR were lowest and overlapping zero
459 in adult treatments (lnRR = 0.12 [-0.06 – 0.30]), and strongest for treatments that spanned
460 both the juvenile and the adult life stage (lnRR = 0.45 [0.17 – 0.73]; Table 4; Fig 2E). Life-
461 stage effects on lnCVR showed a negative estimate for juvenile treatments (lnCVR = -0.08 [-
462 0.16 – 0.00]), and a positive effect, i.e. an increase in behavioural variance in low-condition
463 treatments, when treatments spanned both the juvenile and the adult life stage (lnCVR =
464 0.18 [0.01 – 0.35]; Table 5; Fig 2F). Untransformed maximum lifespan did not appear to
465 influence lnRR (0.00 [-0.08 – 0.09]). However, log-transformed lifespan showed a positive
466 lnRR effect, with its confidence intervals only slightly overlapping with zero (0.15 [-0.01 –
467 0.30]; Table 2, 4) and did explain a relatively high percentage heterogeneity (R^2_{marginal} =
468 5.73%; Table 2). Neither lifespan estimate appeared to have a clear effect on lnCVR, and
469 both only explained a small amount of heterogeneity (R^2_{marginal} = 0.80% and 0.36%,
470 respectively; Table 3, 5).

471

472 (3) *Publication bias tests*

473 Funnel plots showed some potential evidence of asymmetry (for plots and fail-safe N
474 calculations see Supporting Information S7). The estimated effect of *Precision* on lnRR was
475 negative and the confidence intervals slightly overlapped with zero (-0.002 [-0.005 – 0.000];
476 Table 2, 4), although R^2_{marginal} was low (0.27%; Table 2), showing limited evidence of small-
477 study bias. There was also possible evidence of time-lag bias in published data, with effect
478 sizes appearing to trend slightly downwards over time but the confidence intervals
479 overlapped with zero (-0.05 [-0.14 – 0.05]; Table 2, 4), and R^2_{marginal} was again relatively low
480 (0.82%; Table 2). Last, effects calculated from papers where effect sizes could be partially
481 (lnRR = 0.26 [0.07 – 0.63]) or completely (lnRR = 0.24 [0.09 – 0.40]) calculated from the

482 publicly available material were relatively large (Fig 3), whereas the effect from papers
483 where effect sizes could only be obtained through author correspondence were small and
484 the confidence intervals overlapped with zero ($\ln RR = 0.10 [-0.16 - 0.35]$), however, R^2_{marginal}
485 was relatively low for this moderator (0.64%; Table 2). This difference suggests that non-
486 reported results might be biased towards inconclusive (likely statistically non-significant)
487 results.

488

489 (4) *Exploratory models*

490 There was limited evidence that either the type or direction of diet manipulation influenced
491 $\ln RR$ with all diet types and directional treatments, respectively, showing positive mean
492 estimates, and little heterogeneity explained by either of those moderators ($R^2_{\text{marginal}} = 1.31\%$
493 and 0.08%, respectively; Table 2, 4; Fig 4A, 4C). The effect of the duration of diet treatments
494 on $\ln RR$ was almost zero too (Table 2, 4). There was some heterogeneity explained by the
495 rearing environment of the experimental subjects ($R^2_{\text{marginal}} = 2.17\%$; Table 2, 4), with effect
496 sizes from laboratory reared animals being the smallest ($\ln RR = 0.13 [-0.03 - 0.30]$), and
497 effect sizes from wild reared animals being the largest ($\ln RR = 0.32 [0.16 - 0.48]$; Fig 4E).

498

499 Both the type and direction of diet manipulation did not appear to influence $\ln CVR$
500 substantially, whereas the duration of diet treatments had a small positive effect on
501 behavioural variance (0.05 [0.00 - 0.10]), and explained some heterogeneity ($R^2_{\text{marginal}} =$
502 1.99%; Table 3, 5; Fig 4B, 4D). There was limited evidence that rearing environment
503 influenced $\ln CVR$, with 2.08% of heterogeneity explained by this moderator (Table 3, 5; Fig
504 4F).

505

506 **IV. Discussion**

507 Despite our expectations, we found a convincing directional effect on mean risk-taking
508 behaviour, where animals subject to low condition dietary treatments are more likely to show

509 high-risk behaviour in a range of contexts involving predation and novelty. This condition-
510 dependency may be caused by increased risk aversion in animals from higher-condition
511 treatments, due to their greater reproductive expectations (an interpretation consistent with
512 the asset-protection principle applying to the context of nutritional condition and predation-
513 novelty based risk), or by increased risk preference in animals from low-condition
514 treatments, due to their elevated danger of starvation (a starvation avoidance mechanism;
515 Luttbeg & Sih, 2010). These adaptive interpretations contrast with a recent meta-analysis
516 showing that riskier behavioural types tended to have higher survival in the wild (Moiron,
517 Laskowski, & Niemelä, 2020), which may highlight a distinction between behavioural
518 variation due to personality trait differences and due to state-dependent effects.

519 Nonetheless, our result is consistent with the idea of a trade-off between the potential
520 benefits of high outcome-variance behaviours (e.g. accessing resources) and the potential
521 costs (e.g. predation or starvation), which animals balance based on their current or past
522 nutritional state (Ludwig & Rowe, 1990; Clark, 1994; McNamara & Houston, 1996).

523

524 Although our overall effect was relatively strong, there was high heterogeneity in InRR effect
525 sizes with a large proportion (>20%) related to among-species differences. Variation among
526 species, however, was only minimally related to their shared ancestry, with phylogeny only
527 accounting for a small proportion of heterogeneity (3%). It would be interesting to know if
528 condition-dependence of risk-taking behaviour also applies to humans (Wilson *et al.*, 1994;
529 Gosling, 2008), but the large amount of heterogeneity suggests that any effect might differ in
530 a human context. The high heterogeneity among effect sizes is also evident from the wide
531 prediction intervals estimated, and the substantial heterogeneity among studies and
532 experiments. Since theory predicts that state-dependent effects on risk taking vary in
533 strength and direction with factors such as life history traits (Clark, 1994; McNamara &
534 Houston, 1996) and/or local environmental/ecological conditions (Luttbeg & Sih, 2010), such
535 a pattern of variation among species, studies and experiments was to be expected. Critically,
536 given the high heterogeneity, our overall effect does not preclude the opposite pattern being

537 applicable in certain systems. Also, our findings focus on nutritional state in contexts often
538 involving direct or indirect predation risk, so state-dependent safety may be more directly
539 applicable when considering types of state variables that provide a more direct advantage in
540 reducing predation risk (e.g. defensive traits), or in risk-taking contexts where physical
541 condition provides a clearer advantage (e.g. intraspecific contests).

542

543 The experimental context of risk-taking behaviour was the most explanatory of lnRR
544 moderators, revealing that the effect of condition manipulations in certain contexts was clear
545 and particularly strong, such as those involving feeding. This is consistent with studies
546 showing that the choice of experiment used to measure risk taking is important to the
547 outcome, and that different risk-taking behaviours can show divergent patterns of individual-
548 level variation (e.g. Carter *et al.*, 2012). The concept of a 'risky' behaviour can be applied to
549 a broad range of circumstances, as shown by the range of behavioural variables included
550 here, and 'risk-taking' can refer to a suite of potentially independent behaviours. A risk
551 context that was particularly strongly affected was shoaling behaviour in a novel environment
552 (and, with less certainty, shoaling when exposed to a predator). Whether decisions to
553 venture from a group can be considered a risk-taking behaviour or boldness trait has been
554 disputed, partly due to overlap with sociability traits (Toms, Echevarria, & Jouandot, 2010),
555 but our findings are consistent with these decisions being related to risk taking as a trade-off
556 between resource acquisition and group safety. Contrastingly, the estimated effect was
557 highly uncertain and close to zero for refuge emergence into a novel environment, a
558 commonly used variable to measure bold-exploratory personalities. Studies have shown
559 refuge emergence to be unrelated to within-species variation in other risk-taking behaviours
560 (e.g. startle responses in *Pomacentrus spp.*, Beckmann & Biro, 2013; or novel object tests in
561 *Chlamydogobius eremius*, Moran *et al.*, 2016), such that the relationship between refuge
562 emergence and risk taking remains unclear.

563

564 Sex effects on lnRR did not show evidence of male-female differences, with both male- and
565 female-specific effects being relatively small and similar to each other. It has been
566 suggested that different reproductive roles may lead to sex-specific responses to diet
567 variation (Han & Dingemans, 2015), but there does not appear to be a generalizable
568 direction to this effect. Life-stage effects did show evidence that treatments in juvenile stages
569 had strong and positive effects, while effects in adults were less clear. The effects of life-
570 stage and sex may be interrelated in a way that was not originally anticipated, as the strong
571 effect in unknown sex groups may be related to an overrepresentation of juveniles in that
572 category. Whereas studies where sex was identifiable may have been more likely to involve
573 adult treatments groups, with both sex-specific and adult-specific estimates being smaller.
574 The influence of longevity was ambiguous, but ongoing theoretical support for asset
575 protection to be sensitive to life-history traits (e.g. iteroparous vs. semelparous reproductive
576 strategies; Luttbegg et al., 2020) suggests that a more focused analysis incorporating life-
577 history differences is warranted, particularly in relation to reproductive traits.

578

579 Our exploratory analyses revealed a few key patterns in condition-dependent behavioural
580 responses, and the suitability of our methodology. Modelling studies have suggested there
581 may be non-linearity in state-dependent phenotypic responses in risk-taking behaviour, due
582 to potential factors such as varying correlations between condition and reproductive value
583 (Clark, 1994; McNamara & Houston, 1996; Luttbegg & Sih, 2010). While not directly testing
584 this, evidence of a non-linear effect of condition and risk taking was not detected in the
585 analysis of diet manipulation direction. Effects were similar for each group (i.e. reduced vs.
586 standard condition; standard vs. enriched condition, reduced vs. enriched condition),
587 supporting a more constant directional effect of condition on mean risk taking, and
588 suggesting that our methodology of pooling these designs together for analysis was sound.
589 Similarly, the mean effect estimate was positive across all classes of diet treatment analysed
590 (e.g. quality, quantity etc.), such that pooling these experiments was unlikely to influence
591 results. Finally, wild-reared animals did show the largest effect of treatment on mean risk

592 taking (and also a particularly strong negative effect on behavioural variation), suggesting
593 that these animals might be either more sensitive to imposed dietary manipulations or more
594 responsive to predator-based risk due to past experiences in the wild .

595

596 Contrasting with overall mean effects, support for an overall effect of condition on
597 behavioural variation was limited, with only a small, slightly negative and rather uncertain
598 overall InCVR estimate. This contrasts with the expectation that poor condition may increase
599 phenotypic variability (e.g. by exposing cryptic genetic variation), but agrees with a recent
600 meta-analysis showing that developmental stress does not seem to influence variation in
601 behavioural traits across species (Sánchez-Tójar *et al.*, 2020). Heterogeneity was generally
602 lower in InCVR models relative to InRR ones, which is likely because variance effect sizes
603 are generally associated with larger sampling variances (Sánchez-Tójar *et al.*, 2020).

604 Variance meta-analyses are expected to be more data hungry, although this is unlikely to be
605 the cause of the overall weak InCVR effect found in our study given the large dataset used.

606

607 Variation in behaviour was sensitive to the experimental context of risk-taking behaviour,
608 with variation in both the strength and direction of context-specific effects. In particular,
609 variance in feeding behaviour within novel environments was far lower in low-condition
610 groups, providing some evidence that being highly motivated to feed in this context is an
611 optimum phenotype for individuals in poor energetic state. In contrast, variation in refuge use
612 in a novel environment was higher in low-condition groups, which may be evidence of the
613 opposite (complementary) pattern where high refuge use is a preferred strategy for high
614 condition individuals. Effects of life stage on behavioural variation are consistent with recent
615 empirical evidence suggesting that developmental diet is related to phenotypic plasticity and
616 personality development (see examples in Royauté & Dochtermann 2017; Kelleher *et al.*
617 2019). Buchanan, Grindstaff, & Pravosudov (2013) suggested that poor condition during
618 early life stages may reduce an individual's capacity to express behavioural plasticity. This is
619 potentially consistent with our finding of reduced behavioural variation in groups subject to

620 low-condition treatments as juveniles, while the effect in adults heavily overlapped with zero.
621 We also found that treatments that spanned juvenile and adult life stages (often longer term,
622 chronic diet restriction treatments) had a positive effect on behavioural variation. Similarly,
623 the duration of diet treatments had a positive effect on behavioural variation, consistent with
624 the proposition that extremely poor diet conditions can expose cryptic genetic and
625 phenotypic variation (Han & Dingemans, 2017a). Nonetheless, identifying mechanisms
626 from unpartitioned phenotypic variance remains challenging, as the proposed mechanisms
627 for effects on variability in risk-taking behaviour often apply specifically to among- or within-
628 individual levels (Han & Dingemans, 2015).

629

630 A pertinent question in behavioural ecology is whether phenotypic variation is primarily within
631 or among individuals (Westneat, Wright, & Dingemans, 2015). Any effects on the variance
632 as estimated in our meta-analysis (and more generally in most meta-analysis using InCVR)
633 may arise from either source. Individuals might become more variable in their behaviour in
634 response to some treatment (or some environmental effect) as a form of behavioural bet-
635 hedging or reduce accuracy of performance (i.e. within-individual level). Alternatively,
636 individuals might differ in their average responses to changes in conditions if they have
637 intrinsically different reaction norms (i.e. among-individual level). Only repeated
638 measurements per individual would help to separate the two variance components.
639 However, this type of data is usually not available in the literature (Niemi and Dingemans
640 2018). Future studies should focus on the relative importance of within- vs. among-individual
641 variance in the variance effects identified in our study.

642

643 Considered together, our publication bias analyses suggest there may be some limited
644 influence on the overall results. Time-lag analysis showed that effect sizes might be slightly
645 decreasing over time, while precision analysis showed a small negative effect, both of which
646 can be signs of publication bias toward a positive effect (Jennions & Møller, 2002; Jennions
647 *et al.*, 2013). Moreover, effect sizes obtained from author correspondence where no data

648 could be extracted from published material showed the lowest and most uncertain effect,
649 suggesting preferential publication of positive effects. Intriguingly, publication bias appears to
650 be present even where there are competing hypotheses, with positive effect hypotheses
651 (e.g. the asset protection principle) potentially seemingly preferred. We avoided methods to
652 compensate for bias (e.g. trim and fill) as these can perform poorly in high heterogeneity
653 datasets (Moreno *et al.*, 2009). Instead, we advise caution when interpreting our results, and
654 ecological meta-analyses in general, given the ubiquity of publication bias in the literature.

655

656 **V. Conclusions**

657 (I) The overall evidence of diet and condition-manipulation effects on risk-taking behaviour in
658 the literature is clear, as low-condition individuals appear willing to on average take greater
659 risks in ecological contexts relating to predation risk and novelty.

660 (II) While condition-dependency appears to have broad relevance across the animal
661 kingdom, the strength and certainty of this effect may be somewhat overstated due to
662 publication bias and large heterogeneity among effect sizes.

663 (III) Furthermore, the effect is strongly context-dependent, at both the mean and the variance
664 level, suggesting that the specific ecological (and experimental) factors of any context must
665 be considered when studying risk-taking behaviour.

666 (IV) Overall, there appears to be complex and nuanced effects of diet and condition on
667 behavioural variance warranting further empirical study. Future research should focus on
668 separating among- and within-individual variance effects of individual condition.

669

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683

684 **VII. Authorship**

685 NPM: Conceptualization, Investigation, Methodology, Data curation, Formal analysis, Project
686 administration, Software, Validation, Visualization, Writing - original draft, Writing - review &
687 editing. AST: Conceptualization, Investigation, Methodology, Data collection, Software,
688 Validation, Writing - review & editing. HS: Conceptualization, Funding acquisition, Writing -
689 review & editing. KR: Conceptualization, Funding acquisition, Supervision, Writing - review &
690 editing.

691

692 **VIII. Data Accessibility**

693 All data and code used (including data processing, preparation, analysis and presentation)
694 are available at the Open Science Framework (<https://osf.io/3tphj/>, doi:
695 10.17605/OSF.IO/3TPHJ).

696

697 **IX. References**

698 The asterisk (*) indicates that data from an article was used in the meta-analysis.

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1198 **Tables**

1199 **Table 1:** Main effects models estimates, with random effect specific heterogeneity estimates
 1200 (I^2) expressed as percentages, and Q-test for absolute heterogeneity among effect sizes (Q).
 1201 Square brackets represent 95% confidence intervals. Round brackets represent 95%
 1202 prediction intervals, i.e. the range in which 95% of future or unknown effects are likely to fall.
 1203 Positive log response ratio (lnRR) and log coefficient of variation ratio (lnCVR) effects
 1204 represent higher either risk taking or variance in risk taking in low-condition animals,
 1205 respectively.

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| Effect size | k | Mean effect | I^2 Experiment ID (%) | I^2 Study ID (%) | I^2 Species ID (%) | I^2 Phylogeny (%) | I^2 Effect ID (%) | I^2 Total (%) | Q |
|--------------------------|------|--|-------------------------|-----------------------|-----------------------|---------------------|-----------------------|-----------------------|------------------------|
| InRR (non-phylo) | 1297 | 0.23 [0.15, 0.32] (-0.90, 1.36) | 20.2 [17.1 - 23.3] | 7.8 [6.0 - 9.8] | 24.2 [19.3 - 29.6] | - | 45.8 [41.9 - 49.8] | 98.0 [97.8 - 98.1] | 25864.30 p < 0.0001 |
| InRR (phylo) | 1297 | 0.23 [0.09, 0.38] (-0.91, 1.37) | 19.8 [16.9 - 22.9] | 7.8 [5.9 - 9.8] | 22.0 [17.2- 26.9] | 3.4 [2.5 - 4.4] | 45.3 [41.5 - 49.0] | 98.0 [97.8 - 98.2] | 25864.30 p < 0.0001 |
| InCVR (non-phylo) | 1235 | -0.03 [-0.09,0.03] (-0.78, 0.72) | 11.6 [9.7 - 13.6] | 21.6 [17.4 - 26.2] | 0.0 [0.0 - 0.0] | - | 28.0 [25.8 - 30.3] | 61.2 [58.8 - 63.6] | 2543.32 p < 0.0001 |
| InCVR (phylo) | 1235 | -0.03 [-0.09,0.03] (-0.78, 0.72) | 11.6 [9.7 - 13.5] | 21.6 [17.4 - 26.0] | 0.0 [0.0 - 0.0] | 0.0 [0.0 - 0.0] | 28.0 [25.7 - 30.2] | 61.2 [58.8 - 63.5] | 2543.32 p < 0.0001 |

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1219 **Table 2:** Hypothesis testing, publication bias and exploratory moderators for log response
 1220 ratio (lnRR) models, with Q-test for residual heterogeneity (Q_E), moderator explained
 1221 heterogeneity (Q_M), and the estimated percentage of heterogeneity explained by the
 1222 moderators (R^2_{marginal}). Numbers preceding hypotheses refer to the a priori hypotheses as
 1223 laid out in the introduction.

| Hypothesis (model) | Effect size | <i>k</i> | Moderator(s) | Q_E (residual) p < 0.0001 | Q_M (moderator) p = 0.0001 | R^2_{marginal} (%) |
|--|-------------|----------|---|-----------------------------------|------------------------------------|--------------------------------|
| Hyp. 2. Context-dependency of risk (<i>rr.Full.h2</i>) | lnRR | 1297 | <i>RiskContext</i> | 14657.13 p < 0.0001 | 79.38 *** p < 0.0001 | 8.96 |
| Hyp. 3. Sex difference in risk taking (<i>rr.Full.h3</i>) | lnRR | 1297 | <i>Sex</i> | 24006.28 p < 0.0001 | 15.80 ** p = 0.0033 | 1.97 |
| Hyp. 4. Effects across life stages (<i>rr.Full.h4</i>) | lnRR | 1214 | <i>ManipLifeStage + RelativeTimeFromTreatment.C</i> | 16753.80 p < 0.0001 | 21.20 *** p = 0.0004 | 3.76 |
| Hyp. 5(i). Life-history effects (<i>rr.Full.h5.i</i>) | lnRR | 1214 | <i>MaxLongevity.C</i> | 23933.71 p < 0.0001 | 0.00 p = 0.9651 | 0.00 |
| Hyp. 5(ii). Life-history effects (<i>rr.Full.h5.ii</i>) | lnRR | 1214 | <i>InMaxLongevity.C</i> | 22654.52 p < 0.0001 | 3.46 p = 0.0628 | 5.73 |
| Publication bias 1 (<i>rr.Full.pub1</i>) | lnRR | 908 | <i>Precision</i> | 13245.28 p < 0.0001 | 2.80 p = 0.094 | 0.27 |
| Publication bias 2 (<i>rr.Full.pub2</i>) | lnRR | 908 | <i>Year.C</i> | 21211.43 p < 0.0001 | 0.97 p = 0.3251 | 0.82 |
| Publication bias 3 (<i>rr.Full.pub1</i>) | lnRR | 1297 | <i>EffectSizesFromPublication</i> | 23269.07 p < 0.0001 | 11.58 ** p = 0.009 | 0.64 |
| Exp a. Effect of manipulation type (<i>rr.Full.exp.a</i>) | lnRR | 1297 | <i>ManipType</i> | 22616.48 p < 0.0001 | 8.30 p = 0.0812 | 1.31 |
| Exp b. Effect of manipulation direction (<i>rr.Full.exp.b</i>) | lnRR | 1297 | <i>ManipDirection</i> | 20399.67 p < 0.0001 | 10.41 * p = 0.0154 | 0.08 |
| Exp c. Effect of manipulation duration (<i>rr.Full.exp.c</i>) | lnRR | 1214 | <i>RelativeManipDuration.C</i> | 24024.39 p < 0.0001 | 0.06 p = 0.8007 | 0.02 |
| Exp d. Effect of rearing environment (<i>rr.Full.exp.d</i>) | lnRR | 1297 | <i>WildLabRear</i> | 22799.97 p < 0.0001 | 16.79 ** p = 0.0021 | 2.17 |

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1232 **Table 3:** Hypothesis testing, publication bias and exploratory moderators for log coefficient
 1233 of variation ratio (lnCVR) models, with Q-test for residual heterogeneity (Q_E), moderator
 1234 explained heterogeneity (Q_M), and the estimated percentage of heterogeneity explained by
 1235 the moderators (R^2_{marginal}).

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| Hypothesis (model) | Effect size | k | Moderator(s) | Q_E (residual) $p < 0.0001$ | Q_M (moderator) $p = 0.0002$ | R^2_{marginal} (%) |
|--|-------------|------|---|-------------------------------------|--------------------------------------|--------------------------------|
| Hyp. 2. Context-dependency of risk (cvr.Full.h2) | lnCVR | 1235 | <i>RiskContext</i> | 2450.98 $p < 0.0001$ | 38.4 *** $p = 0.0002$ | 8.58 |
| Hyp. 3. Sex difference in risk taking (cvr.Full.h3) | lnCVR | 1235 | <i>Sex</i> | 2520.50 $p < 0.0001$ | 5.9 $p = 0.2066$ | 3.00 |
| Hyp. 4. Effects across life stages (cvr.Full.h4) | lnCVR | 1153 | <i>ManipLifeStage + RelativeTimeFromTreatment.C</i> | 2158.20 $p < 0.0001$ | 9.5 $p = 0.0908$ | 5.28 |
| Hyp. 5(i). Life-history effects (cvr.Full.h5.i) | lnCVR | 1153 | <i>MaxLongevity.C,</i> | 2185.53 $p < 0.0001$ | 1.41 $p = 0.2348$ | 0.80 |
| Hyp. 5(ii). Life-history effects (cvr.Full.h5.ii) | lnCVR | 1153 | <i>InMaxLongevity.C</i> | 2187.91 $p < 0.0001$ | 0.34 $p = 0.5615$ | 0.36 |
| Exp a. Effect of manipulation type (cvr.Full.exp.a) | lnCVR | 1235 | <i>ManipType</i> | 2535.90 $p < 0.0001$ | 3.1 $p = 0.5406$ | 1.15 |
| Exp b. Effect of manipulation direction (cvr.Full.exp.b) | lnCVR | 1235 | <i>ManipDirection</i> | 2541.40 $p < 0.0001$ | 2.23 $p = 0.5256$ | 0.68 |
| Exp c. Effect of manipulation duration (cvr.Full.exp.c) | lnCVR | 1153 | <i>RelativeManipDuration.C</i> | 2182.57 $p < 0.0001$ | 4.59 * $p = 0.0322$ | 1.99 |
| Exp d. Effect of rearing environment (cvr.Full.exp.d) | lnCVR | 1235 | <i>WildLabRear</i> | 2514.93 $p < 0.0001$ | 4.6 $p = 0.3312$ | 2.08 |

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1247 **Table 4:** Parameter estimates for log response ratio (lnRR) hypothesis testing, publication
1248 bias, and exploratory models, with 95% confidence intervals. *k* shows the number of effect
1249 sizes, and *n_{study}* shows the number of studies. Bold estimates correspond to confidence
1250 intervals that do not overlap zero. Note that models with categorical moderators were fitted
1251 as no-intercept models for ease of interpretation.

| Hypothesis (model) | Moderator(s) | Level | <i>k</i> | <i>n_{study}</i> | Estimate |
|--|-----------------------------------|---|--------------------|--------------------------|--------------------------|
| Hyp. 2. Context-dependency of risk (<i>rr.Full.h2</i>) | <i>RiskContext</i> | <i>novelenvironment_activity</i> | 248 | 46 | 0.09 [-0.06, 0.25] |
| | | <i>novelenvironment_exploration</i> | 153 | 33 | 0.11 [-0.05, 0.28] |
| | | <i>novelenvironment_feeding</i> | 331 | 37 | 0.36 [0.20, 0.52] |
| | | <i>novelenvironment_lightdarktest</i> | 26 | 6 | 0.21 [-0.11, 0.52] |
| | | <i>novelenvironment_refugeemergence</i> | 39 | 7 | 0.03 [-0.23, 0.30] |
| | | <i>novelenvironment_refugeuse</i> | 75 | 16 | 0.23 [0.03, 0.42] |
| | | <i>novelenvironment_shoaling</i> | 29 | 5 | 0.37 [0.06, 0.67] |
| | | <i>novelobject_response</i> | 92 | 11 | 0.19 [-0.04, 0.41] |
| | | <i>predation_feeding</i> | 81 | 14 | 0.75 [0.53, 0.97] |
| | | <i>predation_response</i> | 172 | 34 | 0.19 [0.02, 0.36] |
| | | <i>predation_shoaling</i> | 20 | 4 | 0.28 [-0.04, 0.61] |
| <i>dispersalmigration</i> | 15 | 6 | 0.03 [-0.39, 0.45] | | |
| <i>other</i> | 16 | 5 | 0.23 [-0.16, 0.61] | | |
| Hyp. 3. Sex difference in risk taking (<i>rr.Full.h3</i>) | <i>Sex</i> | <i>female</i> | 421 | 39 | 0.15 [-0.03, 0.33] |
| | | <i>male</i> | 291 | 37 | 0.12 [-0.06, 0.30] |
| | | <i>mixed</i> | 120 | 14 | 0.34 [0.06, 0.61] |
| | | <i>unknown</i> | 465 | 61 | 0.29 [0.13, 0.45] |
| Hyp. 4. Effects across life stages (<i>rr.Full.h4</i>) | <i>ManipLifeStage</i> | <i>adult</i> | 423 | 48 | 0.12 [-0.06, 0.30] |
| | | <i>both</i> | 179 | 8 | 0.45 [0.17, 0.73] |
| | | <i>juvenile</i> | 601 | 66 | 0.30 [0.14, 0.46] |
| | | <i>unknown/mixed</i> | 94 | 11 | 0.40 [0.11, 0.69] |
| | | <i>(covariate)</i> | - | - | 0.01 [-0.03, 0.06] |
| Hyp. 5(i). Life-history effects (<i>rr.Full.h5.i</i>) | <i>MaxLongevity.C</i> | <i>intercept</i> | - | - | 0.26 [0.15, 0.36] |
| | | <i>(covariate)</i> | - | - | 0.00 [-0.08, 0.09] |
| Hyp. 5(ii). Life-history effects (<i>rr.Full.h5.ii</i>) | <i>InMaxLongevity.C</i> | <i>intercept</i> | - | - | 0.22 [0.02, 0.43] |
| | | <i>(covariate)</i> | - | - | 0.15 [-0.01, 0.30] |
| Publication bias 1 (<i>rr.Full.pub1</i>) | <i>Precision</i> | <i>intercept</i> | - | - | 0.28 [0.08, 0.49] |
| | | <i>(covariate)</i> | - | - | 0.00 [-0.01, 0.00] |
| Publication bias 2 (<i>rr.Full.pub2</i>) | <i>Year.C</i> | <i>intercept</i> | - | - | 0.26 [0.07, 0.44] |
| | | <i>(covariate)</i> | - | - | -0.05 [-0.14, 0.05] |
| Publication bias 3 (<i>rr.Full.pub3</i>) | <i>EffectSizesFromPublication</i> | <i>no</i> | 130 | 13 | 0.10 [-0.16, 0.35] |
| | | <i>partial</i> | 360 | 31 | 0.26 [0.07, 0.45] |
| | | <i>yes</i> | 807 | 82 | 0.24 [0.09, 0.40] |
| Exp a. Effect of manipulation type (<i>rr.Full.exp.a</i>) | <i>ManipType</i> | <i>combined</i> | 24 | 4 | 0.27 [-0.07, 0.62] |
| | | <i>quality</i> | 248 | 18 | 0.35 [0.07, 0.63] |
| | | <i>quantity</i> | 390 | 50 | 0.30 [0.07, 0.53] |
| | | <i>starvation</i> | 635 | 59 | 0.19 [-0.04, 0.41] |
| Exp b. Effect of manipulation direction (<i>rr.Full.exp.b</i>) | <i>ManipDirection</i> | <i>dual</i> | 60 | 7 | 0.30 [-0.06, 0.66] |
| | | <i>restrict</i> | 1170 | 112 | 0.23 [0.09, 0.38] |
| | | <i>supplement</i> | 67 | 9 | 0.20 [-0.04, 0.44] |
| Exp c. Effect of manipulation duration (<i>rr.Full.exp.c</i>) | <i>RelativeManipDuration.C</i> | <i>intercept</i> | - | - | 0.25 [0.16, 0.35] |
| | | <i>(covariate)</i> | - | - | -0.01 [-0.07, 0.05] |
| Exp d. Effect of rearing environment (<i>rr.Full.exp.d</i>) | <i>WildLabRear</i> | <i>commercial</i> | 139 | 12 | 0.25 [-0.02, 0.52] |
| | | <i>lab</i> | 711 | 58 | 0.13 [-0.03, 0.3] |
| | | <i>mixed</i> | 15 | 1 | 0.21 [-0.5, 0.93] |
| | | <i>wild</i> | 432 | 57 | 0.32 [0.16, 0.48] |

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1255 **Table 5:** Parameter estimates for log coefficient of variation ratio (lnCVR) hypothesis testing,
1256 and exploratory models, with 95% confidence intervals. *k* shows the number of effect sizes,
1257 and *n_{study}* shows the number of studies. Bold estimates correspond to confidence intervals
1258 that do not overlap zero. Note that models with categorical moderators were fitted as no-
1259 intercept models for ease of interpretation.

| Hypothesis (model) | Moderator(s) | Level | <i>k</i> | <i>n_{study}</i> | Estimate |
|--|--------------------------------|---|--------------------------|--------------------------|-----------------------------|
| Hyp. 2. Context-dependency of risk (cvr.Full.h2) | <i>RiskContext</i> | <i>novelenvironment_activity</i> | 248 | 46 | 0.02 [-0.06, 0.11] |
| | | <i>novelenvironment_exploration</i> | 153 | 33 | -0.05 [-0.15, 0.05] |
| | | <i>novelenvironment_feeding</i> | 312 | 34 | -0.16 [-0.25, -0.07] |
| | | <i>novelenvironment_lightdarktest</i> | 24 | 5 | -0.09 [-0.35, 0.16] |
| | | <i>novelenvironment_refugeemergence</i> | 39 | 7 | 0.04 [-0.18, 0.25] |
| | | <i>novelenvironment_refugeuse</i> | 75 | 16 | 0.18 [0.04, 0.31] |
| | | <i>novelenvironment_shoaling</i> | 29 | 5 | 0.01 [-0.25, 0.26] |
| | | <i>novelobject_response</i> | 88 | 10 | -0.08 [-0.24, 0.08] |
| | | <i>predation_feeding</i> | 61 | 13 | -0.01 [-0.21, 0.18] |
| | | <i>predation_response</i> | 167 | 33 | 0.02 [-0.08, 0.13] |
| | | <i>predation_shoaling</i> | 20 | 4 | 0.01 [-0.24, 0.26] |
| | | <i>dispersalmigration</i> | 13 | 5 | -0.49 [-0.86, -0.11] |
| <i>other</i> | 6 | 3 | 0.59 [0.16, 1.02] | | |
| Hyp. 3. Sex difference in risk taking (cvr.Full.h3) | <i>Sex</i> | <i>female</i> | 401 | 38 | 0.05 [-0.05, 0.16] |
| | | <i>male</i> | 276 | 37 | 0.03 [-0.08, 0.14] |
| | | <i>mixed</i> | 117 | 13 | -0.09 [-0.28, 0.09] |
| | | <i>unknown</i> | 441 | 56 | -0.08 [-0.17, 0.00] |
| Hyp. 4. Effects across life stages (cvr.Full.h4) | <i>ManipLifeStage</i> | <i>adult</i> | 402 | 45 | -0.01 [-0.10, 0.09] |
| | | <i>both</i> | 116 | 7 | 0.18 [0.01, 0.35] |
| | | <i>juvenile</i> | 578 | 63 | -0.08 [-0.16, 0.00] |
| | | <i>unknown/mixed</i> | 89 | 11 | -0.02 [-0.21, 0.17] |
| | | <i>(covariate)</i> | - | - | 0.02 [-0.02, 0.05] |
| Hyp. 5(i). Life-history effects (cvr.Full.h5.i) | <i>MaxLongevity.C</i> | <i>intercept</i> | - | - | -0.03 [-0.09, 0.03] |
| | | <i>(covariate)</i> | - | - | -0.03 [-0.08, 0.02] |
| Hyp. 5(ii). Life-history effects (cvr.Full.h5.ii) | <i>InMaxLongevity.C</i> | <i>intercept</i> | - | - | -0.03 [-0.09, 0.03] |
| | | <i>(covariate)</i> | - | - | -0.02 [-0.09, 0.05] |
| Exp a. Effect of manipulation type (cvr.Full.exp.a) | <i>ManipType</i> | <i>combined</i> | 24 | 4 | 0.07 [-0.21, 0.35] |
| | | <i>quality</i> | 246 | 18 | 0.05 [-0.09, 0.18] |
| | | <i>quantity</i> | 363 | 48 | -0.07 [-0.16, 0.03] |
| | | <i>starvation</i> | 602 | 54 | -0.04 [-0.12, 0.05] |
| Exp b. Effect of manipulation direction (cvr.Full.exp.b) | <i>ManipDirection</i> | <i>dual</i> | 60 | 7 | 0.11 [-0.14, 0.35] |
| | | <i>restrict</i> | 1116 | 106 | -0.04 [-0.10, 0.03] |
| | | <i>supplement</i> | 59 | 8 | -0.06 [-0.27, 0.14] |
| Exp c. Effect of manipulation duration (cvr.Full.exp.c) | <i>RelativeManipDuration.C</i> | <i>intercept</i> | - | - | -0.03 [-0.08, 0.03] |
| | | <i>(covariate)</i> | - | - | 0.05 [0.00, 0.10] |
| Exp d. Effect of rearing environment (cvr.Full.exp.d) | <i>WildLabRear</i> | <i>commercial</i> | 127 | 11 | -0.02 [-0.21, 0.17] |
| | | <i>lab</i> | 679 | 54 | 0.02 [-0.06, 0.11] |
| | | <i>mixed</i> | 15 | 1 | 0.10 [-0.41, 0.62] |
| | | <i>wild</i> | 414 | 55 | -0.09 [-0.18, 0.00] |

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1262 **Figure Legends**

1263 **Fig. 1** Higher mean risk taking in low-condition compare to high-condition animals, but
1264 similar behavioural variation between them. Phylogenetic (black circles) and non-
1265 phylogenetic (white circles) meta-analytic means for log response ratio (lnRR) and log
1266 coefficient of variation ratio (lnCVR) with 95% confidence intervals. The number of effect
1267 sizes used in each model is k .

1268

1269 **Fig. 2** Category-specific estimates for log response ratio (lnRR) and log coefficient of
1270 variation ratio (lnCVR) with meta-regression models testing the effect of (A, B) the
1271 experimental context for risk-taking behaviour; (C,D) sex of study subjects; and (E,F) life-
1272 stage of study subjects during the diet manipulation treatments. lnRR effects are presented
1273 on the left (A, C, D) and lnCVR on the right (B, D, F). The areas of the blue shaded circles
1274 are proportional to the number of effect sizes k used, and bars represent 95% confidence
1275 intervals. A positive effect shows higher risk taking or higher variance in risk taking in low-
1276 condition animals, respectively.

1277

1278 **Fig. 3** Category-specific estimates based on the degree that log response ratio (lnRR)
1279 effect sizes could be extracted from published material. Fully reported effect sizes are from
1280 papers where all effect sizes could be extracted from published material, partially reported
1281 effect sizes are from papers where some effect sizes could be extracted but additional effect
1282 sizes could be obtained from authors (therefore includes effect sizes from published material
1283 and author correspondence), and not reported effect sizes are those that could only be
1284 calculated from data obtained through author correspondence. The areas of the green
1285 shaded circles are proportional to the number of effect sizes k used, and bars represent 95%
1286 confidence intervals. A positive effect shows higher risk taking and higher variance in risk
1287 taking in low-condition animals.

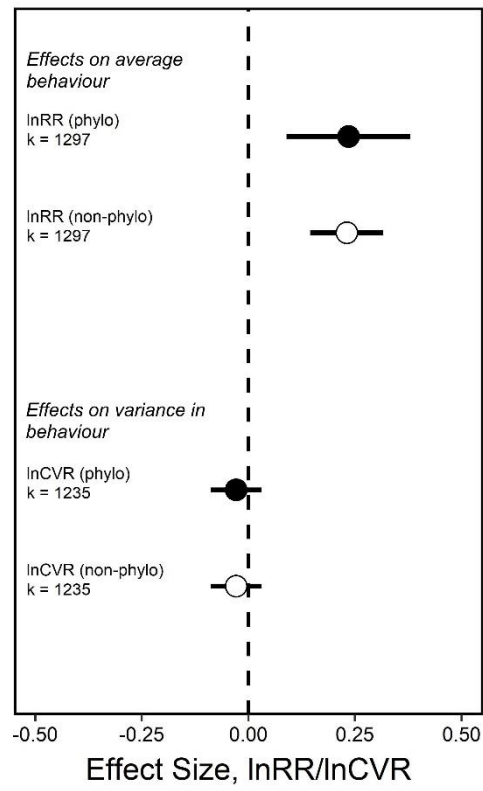
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1289 **Fig 4** Category-specific estimates for log response ratio (lnRR) and log coefficient of
1290 variation ratio (lnCVR) meta-regression models for effect of (A, B) the type of diet
1291 manipulation; (C, D) the direction of the diet manipulation; and (E, F) the rearing
1292 environment of the experimental subjects. lnRR effects are presented on the left (A, C, D)
1293 frames and lnCVR on the right (B, D, F). The areas of the orange shaded circles are
1294 proportional to the number of effect sizes k used, and bars represent 95% confidence
1295 intervals. A positive effect shows higher risk taking and higher variance in risk taking in low-
1296 condition animals, respectively.

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Fig. 1



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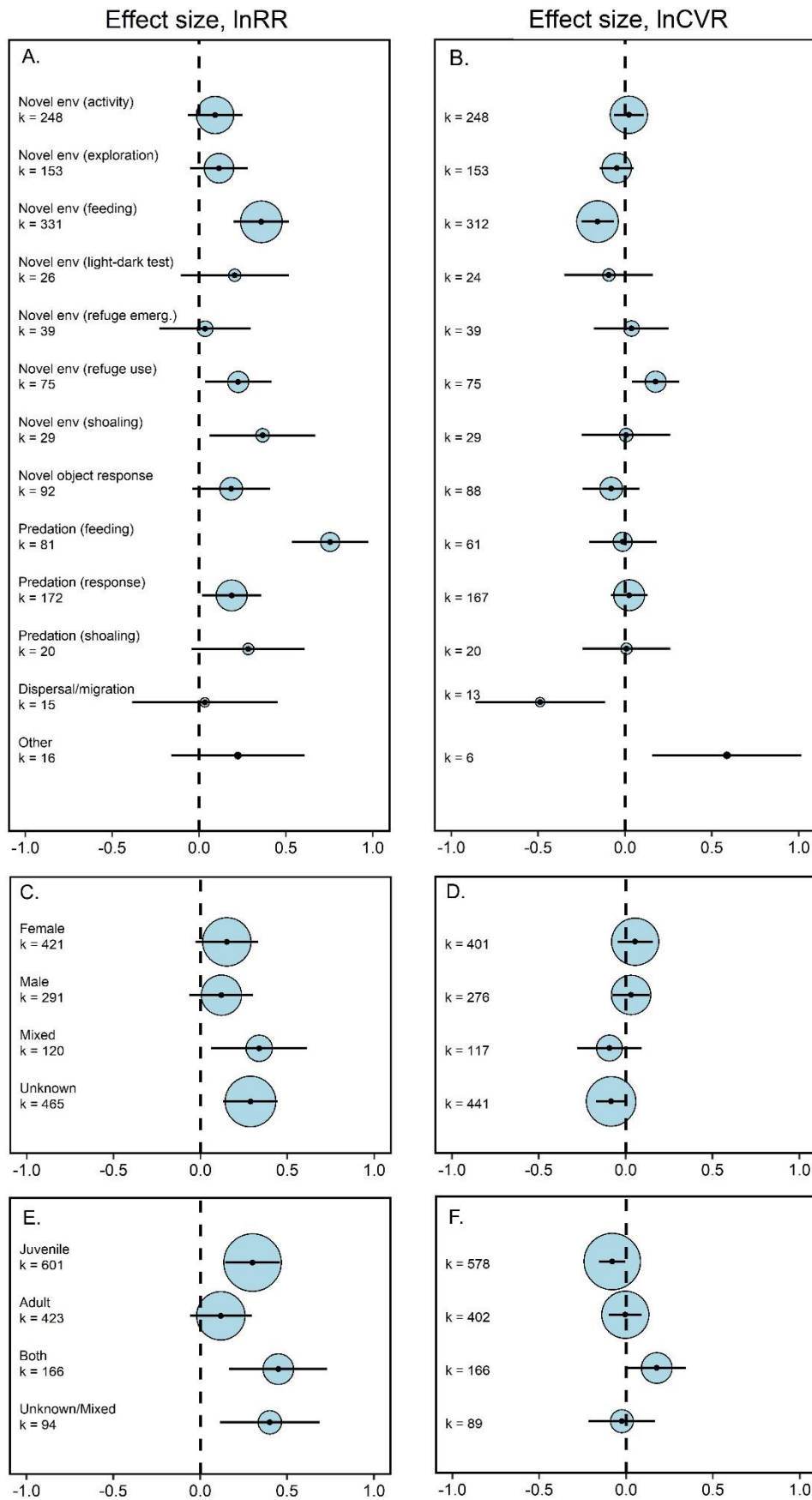
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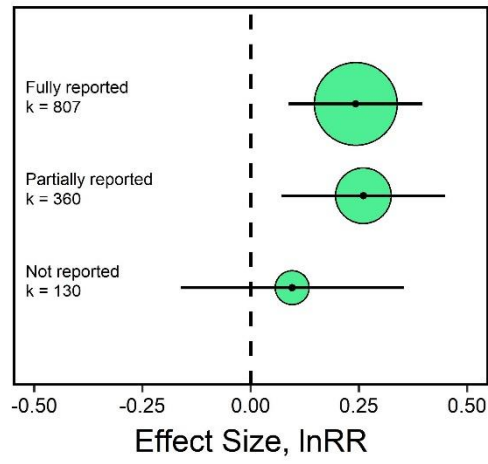
Fig 2



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Fig. 3



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Fig 4.

