

1 **Telomere heritability and parental age at conception effects**
2 **in a wild avian population**

3 Running title: Heritability of telomere length

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20

21 **Abstract:**

22 Individual variation in telomere length is predictive of health and mortality risk across a range of
23 species. However, the relative influence of environmental and genetic variation on individual
24 telomere length in wild populations remains poorly understood. In previous studies, heritability
25 of telomere length has primarily been calculated using parent-offspring regression, but shared
26 environments can confound such estimates. Furthermore, associations with age and parental age
27 at conception effects are typically not accounted for but can also bias heritability estimates. To
28 control for these confounding variables, quantitative genetic ‘animal models’ can be used.
29 However, the few studies on wild populations using this approach have been restricted by power.
30 Here, we investigated the heritability of telomere length and parental age at conception effects in
31 the Seychelles warbler using 2664 telomere length measures from 1318 birds over 20 years and a
32 multi-generational pedigree. We found a weak negative within-paternal age at conception effect
33 (as fathers aged, their offspring had shorter telomeres) and a weak positive between-maternal age
34 at conception effect (females that survived to older ages had offspring with longer telomeres).
35 While parent–offspring regressions did not detect heritability, animal models provided evidence
36 that heritability of telomere length was low in this population. Environmental and technical
37 variation largely influenced telomere length and would have biased heritability estimates if
38 unaccounted for. Estimating the heritability of telomere length is complex, requiring large
39 sample sizes and accounting for confounding effects in order to improve our understanding of
40 the evolutionary potential of telomere length in the wild.

41 **Keywords:** telomere length, heritability, animal model, paternal age at conception, maternal age
42 at conception, Seychelles warbler

43

44 **Introduction**

45 A complete understanding of the relative impact of genetic and environmental effects on
46 senescence rates requires quantifying individual variation in senescence rates, but this is difficult
47 to achieve, especially in wild populations (van de Pol and Verhulst 2006; Nussey et al. 2008;
48 Charmantier et al. 2014). However, the identification of biomarkers, such as telomeres that
49 reflect an individual's intrinsic state and mortality risk (Wilbourn et al. 2018), have facilitated
50 this (Nakagawa et al. 2004). Telomeres, short repetitive DNA elements that protect the ends of
51 eukaryotic linear chromosomes (Blackburn 1991), shorten with each cell cycle due to the end
52 replication problem (Levy et al. 1992) and other mechanisms including oxidative damage (von
53 Zglinicki 2002). Critically short telomeres can trigger cellular senescence (Harley et al. 1992;
54 Campisi 2005) which may lead to organismal senescence (López-Otín et al. 2013). However,
55 telomeres can also be extended by telomerase (Greider and Blackburn 1989) and alternative
56 lengthening (Cesare and Reddel 2010). Telomere shortening occurs with age in a wide range of
57 species (e.g. Salomons et al. 2009; Aubert et al. 2012). Furthermore, whether causal, or just
58 correlational (Simons 2015; Young 2018), telomere length relative to age positively predicts
59 health (Boonekamp et al. 2013; Blackburn et al. 2015) and survival/lifespan within species
60 (Barrett et al. 2013; Wilbourn et al. 2018). Consequently, telomeres are increasingly used in
61 evolutionary ecology studies as a biomarker of senescence and to measure an individual's
62 physiological response to their environmental experiences (Bize et al. 2009; Bauch et al. 2012;
63 Bebbington et al. 2016; Fairlie et al. 2016).

64
65 A better understanding of the drivers of individual variation in telomere length is important if we
66 are to use them as a biomarker of health and senescence within populations (Dugdale and
67 Richardson 2018). Initial telomere length is inherited from each parent (Delgado et al. 2019). As

68 a mother's gametes are produced prenatally, whereas fathers produce sperm throughout their life,
69 paternal age at conception may impact the telomere lengths of their offspring (Eisenberg and
70 Kuzawa 2018). There is cross-sectional evidence from humans that sperm telomere length is
71 positively correlated with age and older fathers have offspring with longer telomeres (Unryn et
72 al. 2005; Kimura et al. 2008; Aston et al. 2012; Eisenberg et al. 2012; Broer et al. 2013). Such
73 effects may be due to the activity of telomerase in the testes resulting in elongated telomeres with
74 age (Kimura et al. 2008; Aviv and Susser 2013). An alternative (not mutually exclusive)
75 hypothesis is the selective survival or proliferation of germ stem cells with longer telomeres
76 (Kimura et al. 2008; Hjelmborg et al. 2015). However, studies in non-human vertebrates,
77 including a longitudinal study in jackdaws (Bauch et al. 2019), report conflicting results; while
78 some have also found a positive correlation between offspring telomere length and paternal
79 (Eisenberg et al. 2017) or maternal age (Asghar et al. 2015), others have found negative paternal
80 age correlations (Olsson et al. 2011; de Frutos et al. 2016; Criscuolo et al. 2017; Bouwhuis et al.
81 2018; Noguera et al. 2018; Bauch et al. 2019) or no parental age effects (Heidinger et al. 2016;
82 Froy et al. 2017; McLennan et al. 2018; Belmaker et al. 2019; van Lieshout et al. 2020a). Work
83 on a wider range of species, using longitudinal data, is required to identify the drivers of
84 variation in parental age at conception effects.

85

86 In addition to parental age effects, genetic variation also influences the maintenance of telomeres
87 from the first mitotic division, and thus telomere dynamics across an individual's lifetime
88 (Delgado et al. 2019; Eisenberg 2019). Environmental effects and life-history events are also
89 associated with telomere shortening due to the stress they exert on the organism, and such effects
90 will accumulate with age (Hall et al. 2004; Heidinger et al. 2012). The majority of studies
91 looking to quantify the contribution of genetic variation to telomere length are in human
92 populations which typically implicate significant heritability (Dugdale and Richardson 2018). It

93 is, however, difficult to interpret heritability estimates from human studies where processes, such
94 as industrialisation and medical interventions, limit their evolutionary interpretation, and in
95 captive or laboratory populations that exist in controlled environments. Importantly, biologists
96 wanting to understand the ecological and evolutionary significance of telomere variation will be
97 interested in studying telomere heritability in wild populations, which are experiencing natural
98 environmental variation and where natural selection is occurring. To date, very few studies have
99 attempted to separate genetic from environmental contributions to variation in telomere length
100 within wild populations, and most have been restricted in terms of small sample sizes.
101 Consequently, our understanding of the heritability of telomeres in wild populations is limited
102 (Dugdale and Richardson 2018). This is important as the amount of additive genetic variance
103 underlying a trait, such as telomere length, limits the variation that selection can act on and,
104 therefore, a trait's evolutionary potential (Lynch and Walsh 1998).

105

106 In wild populations, telomere length heritability estimates range from 0 to 1 (Dugdale and
107 Richardson 2018). Heritability estimates have, however, been based primarily on parent–
108 offspring regressions, which assume that the similarity between parents and offspring is genetic,
109 when, in fact, relatives often also share environments. These shared environmental effects,
110 including cohort and maternal effects (Asghar et al. 2015; Becker et al. 2015), will artificially
111 inflate heritability estimates (Kruuk and Hadfield 2007; Kruuk et al. 2008). Additionally,
112 telomere length will change throughout life, so measures across the lifetimes of individuals will
113 be the product of inherited telomere length, attrition, and restoration/lengthening (Dugdale and
114 Richardson 2018). Few telomere studies have taken individual age at sampling into account
115 (Reichert et al. 2015), or sampled both offspring and parents at the same age, as both sampling
116 and accurate ageing are difficult in wild populations (but see Becker et al, 2015 and Asghar et al.

117 2015). Furthermore, parental age at conception may also impact the telomere length of their
118 offspring (Eisenberg et al. 2012). Subsequently, it is unclear whether the variation in heritability
119 estimates of telomere length in wild populations reflects true variation, or methodological or
120 analytical differences between studies.

121

122 Quantitative genetic “animal models” offer a strong analytical framework to estimate the relative
123 effects of additive genetic and environmental variation on phenotypic traits (Kruuk 2004).

124 Animal models utilise the relationships in a pedigree to estimate additive genetic variance, thus
125 maximising data and increasing the power to detect heritabilities (Wilson et al. 2010).

126 Additionally, animal models can account for, and estimate the contribution of, other factors
127 known to influence telomeres, to get more accurate estimates of the proportion of phenotypic
128 variance due to additive genetic effects. However, animal models require considerable sample
129 sizes (Wilson et al. 2010). Of the three wild vertebrate studies estimating telomere length
130 heritability using a pedigree-based animal model, two attempted to partition shared maternal
131 environment effects but one did not converge and, therefore, could not separate these effects
132 (Asghar et al. 2015) and the other explained litter variation so was not included (Becker et al.
133 2015). There is a clear need for studies assessing the heritability of telomere length using large
134 datasets with multigenerational pedigrees and animal model approaches (Dugdale and
135 Richardson 2018).

136

137 In this study, we used the long-term individual-based multi-generational data from an isolated
138 population of the cooperatively breeding Seychelles warbler (*Acrocephalus sechellensis*) to
139 investigate additive genetic and environmental variance components underlying telomere length.
140 Telomere length declines with age and adult survival is positively associated with telomere

141 length, independent of age, in this population (Barrett et al. 2013). Telomere loss is greatest in
142 early life in the Seychelles warbler and shows strong cohort effects, and telomere length is
143 positively associated with food abundance (Spurgin et al. 2018). However, telomeres also appear
144 to elongate within individuals in this population (Spurgin et al. 2018). Telomere dynamics have
145 helped reveal the costs of factors such as inbreeding (Bebbington et al. 2016) and social conflict
146 (Bebbington et al. 2017). In the Seychelles warbler, telomere length is therefore an important
147 biomarker of condition and senescence and is impacted by environmental conditions.

148

149 Here, we estimate the heritability of telomere length in the Seychelles warbler using 2664
150 telomere measures from 1318 birds within a 10-generation genetic pedigree. First, we test for
151 parental age at conception effects on offspring telomere length accounting for the age at which
152 offspring were sampled. Next, we estimate the heritability of telomere length using parent–
153 offspring regressions, and investigate how heritability estimates differ between maternal–,
154 paternal– and mid-parent–offspring analyses, when measurements were taken at different ages.
155 We predict that heritability estimates will be higher when sampled at younger ages, since these
156 samples will be closer to the telomere length initially inherited from parents. We then compare
157 heritability estimates from the regressions to those gained from animal models where we control
158 for expected confounding effects. We predict estimates of heritability to be higher in parent–
159 offspring regressions compared to animal models, and higher when we included fewer common
160 environmental effects (due to the upward biasing of heritability as a result of shared
161 environments). Finally, we discuss the broader implications of our results for our understanding
162 of the evolutionary potential of telomeres in this population.

163

164 **Materials and Methods**

165 *Study system*

166 The Seychelles warbler is a small passerine endemic to the Seychelles archipelago (Komdeur et
167 al. 1991). The entire population (ca. 320 adult individuals in 115 territories) on Cousin island (29
168 ha; 04°20'S, 55°40'E) has been monitored extensively since 1985 (Komdeur 1992; Richardson et
169 al. 2007; Hammers et al. 2019; Raj Pant et al. 2019). Seychelles warblers defend year-round
170 territories in which a dominant male and female reside and most clutches contain 1 egg
171 (Komdeur 1994; Richardson et al. 2001). The main breeding season runs from June to
172 September, although some pairs also breed between January and March (Komdeur et al. 1991;
173 Komdeur and Daan 2005). Senescence has been documented in the Seychelles warbler
174 (Hammers et al. 2015) with age-dependent declines in both reproduction and survival (Hammers
175 et al. 2012, 2013). Seychelles warblers are largely insectivorous, and variation in rainfall drives
176 variation in insect abundance (Komdeur and Daan 2005) which was positively associated with
177 telomere length (Spurgin et al. 2018). In addition, the study can compare genetic and social
178 parent effects on telomere variation, due to the presence of extra-group paternity (41% of
179 offspring) and subordinate female cobreeding (11% of offspring) (Raj Pant et al. 2019).

180

181 All protocols were ethically reviewed and approved by the BIO Ethical Review Committee,
182 University of East Anglia, UK, and ratified by the University of Leeds. Each breeding season as
183 many birds as possible are caught using mist nets and all territories monitored for the presence
184 and identity of individually colour-ringed birds. The majority (96%) of individuals have been
185 individually marked with a British Trust for Ornithology ring and unique colour ring
186 combinations (Richardson et al. 2001). Age of unringed birds was estimated using eye colour
187 (Komdeur et al. 1991), and where available lay, hatch or fledge dates. Since 1995, blood samples

188 (ca. 25 ul) have been taken and stored at room temperature in absolute ethanol, thus allowing
189 molecular sexing, parentage assignment (Richardson et al. 2001; Hadfield et al. 2006), pedigree
190 construction (Edwards et al. 2017) and telomere length measurement (Barrett et al. 2013). The
191 population is virtually closed (<0.1% dispersal; (Komdeur et al. 2004)) and extrinsic mortality is
192 low, so birds live long lives (maximum observed lifespan = 18 years). Further, the population is
193 intensively monitored with high annual resighting rates (ca. 0.92 ± 0.02 for birds ≤ 2 years and
194 0.98 ± 0.01 for older birds, Brouwer et al. 2010), so accurate birth and death years are known
195 (Hammers et al. 2015).

196

197 *Telomere data*

198 We used the telomere dataset generated in Spurgin et al. 2018, which included birds caught and
199 blood sampled between 1995 and 2014, when the data were most complete. Relative telomere
200 length (RTL) was estimated using qPCR (Barrett et al. 2013; Bebbington et al. 2016; Spurgin et
201 al. 2018). Our cleaned dataset included 2664 samples from 1318 individuals that passed quality
202 control (Bebbington et al. 2016) and filtering steps (telomere $cq \geq 25$ and cq replicate difference
203 ≥ 0.5 ; GADPH $cq \leq 21$ but ≥ 26 and cq replicate difference ≥ 0.5 ; RTL values ≥ 3). There were no
204 storage time effects on telomere length (Spurgin et al. 2018). To investigate plate variance (by
205 including qPCR plate as a random effect in our statistical models), where samples had replicates
206 across plates ($n=388$), an RTL value for a given blood sample was taken at random.

207

208 *Genetic pedigree*

209 Protocols for genotyping, quality control tests and parentage assignments (*MasterBayes 2.5.2*;
210 (Hadfield et al. 2006)), and pedigree statistics are provided in the supplementary information
211 (Supplementary parentage methods, Figures S1–3 and Tables S1–3). Parentage was assigned at p

212 ≥ 0.8 . The pruned pedigree, calculated using *Pedantics* 1.7 (Morrissey and Wilson 2010),
213 included parentage assignments for individuals born 1992–2014 and contained 1482 informative
214 individuals for telomere length with 1217 maternities and 1268 paternities (Table S3).

215

216 *Statistical analyses*

217

218 *Paternal age at conception effects on offspring telomere length*

219 Statistical analyses were performed in R 3.5.3 (R Core Team 2019). We first investigated
220 associations between parental ages at conception and RTL in offspring using linear mixed effects
221 models with Gaussian error distribution in *lme4* 1.1-21 (Bates et al. 2015). RTL was square root
222 transformed to improve linear mixed model fits, and in each model subset RTL was subsequently
223 z-transformed for comparability of telomere studies (Verhulst 2020). Collinearity between the
224 fixed effects was checked by calculating Variance Inflation Factors (VIF); all VIFs were <3 . We
225 fitted offspring RTL across all ages that offspring were sampled at as the response variable and
226 included offspring sex (factor), offspring age in years (log-transformed for all ages and juvenile
227 model following, Spurgin et al. 2018), parental age at conception (maternal and paternal) and
228 technician identity (factor: 2 levels) as our fixed effects. Random effects included offspring
229 identity, maternal identity, paternal identity, capture season ID and qPCR plate.

230

231 Based on our dataset and model structure we had $\geq 80\%$ statistical power to detect paternal age at
232 conception effect sizes of ≥ 0.02 (Figure S4) using a simulation-based power analysis in the
233 package *simr* 1.0.5 (Green and MacLeod 2016). This was equivalent to a correlation coefficient
234 of 0.059 (following Froy et al. 2017) which is sufficient power to detect paternal age at

235 conception effects of the correlation coefficients previously published (De Meyer et al. 2007;
236 Nordfjäll et al. 2010; Eisenberg et al. 2012, 2017). There was considerable variation in maternal
237 and paternal ages at conception (Figure S5A–B) and a significant but weak correlation between
238 the two ($r=0.12$, $t_{1154}=4.08$, $p<0.001$, Figure S5C) which allowed us to include both variables in
239 the same model. Significance was determined using likelihood ratio tests where the fixed effect
240 of interest was dropped from the full model.

241

242 We also investigated paternal (PAC) and maternal age at conception (MAC) effects where
243 offspring RTL was restricted to the first measurement as a nestling ($n=304$), or all juvenile
244 measures (<1 year old, $n=1137$ measures of 958 offspring). The model structure was identical to
245 the model of all ages, except offspring identity was not included as a random effect for the chick
246 model, and for the juvenile model paternal identity was not included to allow model
247 convergence.

248

249 To investigate whether effects on offspring telomere length were driven by within-parental age
250 rather than between-parental age (selective disappearance) at conception effects, we used within-
251 subject centering (van de Pol and Wright 2009). To the model of RTL across all ages, we first
252 removed PAC and MAC and included mean age at conception per parent (between-parental age
253 effects) and the deviation from the mean age at conception of the parent (within-parental age
254 effects). To test whether the within and between slopes differed from each other, we included
255 parental age at conception (within-individual age effects) and mean parental age at conception
256 (difference between the within and between-individual slopes) in a second model. The
257 significance of mean parental age at conception in this second model indicates that these within
258 and between slopes in the first model are significantly different (van de Pol and Wright 2009).

259

260 Heritability of telomere length

261 We first investigated heritability of telomere length with parent–offspring regressions using a
262 general linear model where offspring RTL was the response variable and parent RTL was a
263 covariate. We used a frequentist approach since no random effects were included. Using *pwr* 1.2-
264 2 (Champely 2018) we had $\geq 80\%$ power to detect correlations ≥ 0.195 and ≥ 0.104 , using the
265 minimum (n=165) and maximum sample sizes (n=585) respectively, at a significance threshold
266 of 0.05 (for all sample sizes see Table S4). We used mother–offspring, father–offspring and mid-
267 parent–offspring regressions of RTL to explore how these affected our heritability estimates as
268 well as investigating maternal/paternal transmission differences or the presence of potential
269 maternal/paternal effects. We also investigated how these heritability estimates changed when
270 we used telomere measures taken at all ages or just juvenile ages (<1 year) for both parents and
271 offspring. For each analysis, a mean RTL measure was taken for each offspring and parent either
272 using RTL across all ages or using only RTL measures taken of the individual when they were a
273 juvenile. To avoid pseudoreplication due to the presence of multiple offspring from the same
274 parent, mean offspring or mid-offspring telomere length was used for each mother, father or
275 parent pair. Hence, a mean of the mean RTL from their offspring was taken. Heritabilities were
276 calculated as twice the slope of maternal or paternal RTL on offspring RTL in mother–offspring
277 or father–offspring regressions, or equal to the slope of midparent RTL on offspring RTL in the
278 mid-parent–offspring regressions (Lynch and Walsh 1998).

279

280 Finally, we investigated heritability of telomere length in quantitative genetic “animal models” in
281 *MCMCglmm* 2.26 (Hadfield 2010). We used a Bayesian approach to provide accurate estimates
282 of our variance components. Our pruned pedigree had $\geq 80\%$ power to detect heritabilities of

283 ≥ 0.17 (Figure S6), determined in *Pedantics* 1.7 (Morrissey and Wilson 2010). These univariate
284 models were fitted with RTL (non-transformed) as the response variable and had an increasingly
285 complex fixed and random effect structure. This allowed us to test for confounds between
286 random effects, and to investigate how the inclusion of random effects affected our estimates of
287 heritability. Model 1 included only individual identity to account for repeated measures (to
288 calculate between-individual variation or ‘repeatability’). In model 2, individual identity was
289 partitioned into additive genetic and permanent environment components using the pruned
290 pedigree. In model 3, we included fixed effects of sex (factor), age (log-transformed following
291 Spurgin et al. 2018) and technician (factor: 2 levels) to investigate how heritability was impacted
292 by the inclusion of fixed effects (following: Wilson 2008; de Villemereuil et al. 2018). In model
293 4 we estimated technical variance by adding qPCR plate ID as a random effect. We subsequently
294 added maternal (model 5) and paternal (model 6) identity, determined from the genetic pedigree,
295 to investigate parental effects underlying telomere length. Maternal effects have previously been
296 observed in other species (Asghar et al. 2015), and maternal inbreeding effects, but not paternal
297 inbreeding effects, on offspring telomere length have been documented in our population
298 (Bebbington et al. 2016). We then added the random effects of season of capture (model 7) and
299 current territory (model 8), to account for spatio-temporal factors associated with telomere length
300 (Spurgin et al. 2018). Finally, we tested for early-life effects of birth season (model 9) to account
301 for long-lasting effects of natal conditions on telomere variation. Although we had information
302 on natal territory, models including natal territory did not converge, and simpler models
303 suggested that natal territory explained no variance in RTL. We used default priors for fixed
304 effects, while for the random effects (except for the residual variance structure which were
305 inverse-Wishart priors, where $V=1$, $n=0.002$) we applied parameter expanded priors (with $V=1$,
306 $\nu=1$, $\alpha.\mu=0$ and $\alpha.V=1000$) as the variance estimates were close to zero (Hadfield
307 2019). We ran our models with a variety of iterations (Models 1-3: 1.2×10^6 iterations, burn-

308 in= 2×10^5 , thinning=500; Models 4-5: 2.4×10^6 iterations, burn-in= 4×10^5 , thinning=1000; Models
309 6-9: 3.6×10^6 iterations, burn-in= 6×10^5 , thinning=1500). To assess convergence of *MCMCglmm*
310 models, we checked: autocorrelation $r < 0.1$, effective sample sizes > 1000 , Heidelberger and
311 Welch's tests were passed and Geweke tests were passed. For estimates of the fixed effects and
312 random effects we took the posterior mode of the posterior distributions. We defined fixed
313 effects as significant if the 95% credible intervals of the posterior modes did not overlap zero.
314 Heritability estimates, and the proportion of phenotypic variance explained by other variance
315 components, were calculated by taking the posterior mode of the ratio of the additive genetic
316 variance to total phenotypic variance for each sample of the posterior distribution.

317

318 To confirm the robustness of our estimates of telomere length heritability, we also ran the final
319 model in a frequentist framework using ASReml-R 3 (Butler et al. 2009) using the same
320 structure as model 9. Significance of random effects was determined by dropping each random
321 effect from a model containing all random effects and performing a likelihood ratio test using
322 twice the absolute difference in log-likelihoods between the two models.

323

324 We then tested whether parental effects were present when using the social (i.e. the dominant
325 breeding pair) rather than the genetic parents, which is possible due to extra-group paternity
326 (41% of offspring) and co-breeding (11% of offspring) (Raj Pant et al. 2019). To do this, we
327 compared the model with genetic parents (model 7) to that where the genetic parents were
328 replaced with social parents (model 10: specifications: 1×10^7 iterations, burn-in= 4×10^6 ,
329 thinning=3000) using model 7's structure, since other random effects in models 8 and 9
330 explained a small proportion of the phenotypic variance (see Results).

331 **Results**

332 Maternal and paternal age at conception were not significantly associated with offspring relative
333 telomere length when using telomere lengths across all ages, or when the dataset was restricted
334 to the first offspring measurements taken as chicks, or when all measurements were taken from
335 juvenile offspring (<1 year old; Figure S7, Table S5). However, when parental age at conception
336 effects were separated into within- versus between-parental age effects for lifelong RTL, there
337 was a significant and negative within-paternal age effect and a significant and positive between-
338 maternal age effect (Table 1, Figure 1). As fathers aged the offspring they produced had
339 progressively shorter telomeres, while females that survived to older ages had offspring with
340 longer telomeres (Figure 1). Within- versus between-parental age slopes were significantly
341 different from each other for both maternal and paternal age at conception (Table 1). However,
342 both the within-paternal and between-maternal age effects on offspring RTL were small (Figure
343 1). There was no difference in lifelong RTL between sexes, but there was a logarithmic
344 association with age and an effect of technician (Table S5).

345

346 Using parent–offspring regression techniques, we found no evidence for mother–offspring,
347 father–offspring or mid-parent–offspring resemblance and hence no heritability of RTL using
348 mean telomere measures across all ages. Further, there was no evidence for parent–offspring
349 resemblance when using just mean juvenile (<1 year old) telomere measures of both offspring
350 and parents (Figure 2, Table S4).

351

352 We estimated heritability with a quantitative genetic animal model using a hierarchical approach
353 (Figure 3). Within-individual repeatability of RTL, the amount of variance due to individual
354 identity, was low across all models and ranged from 0.056 (95% CrI: 0.016-0.092; Table 2:
355 Model 9) to 0.136 (95% CrI: 0.078-0.195; Table S6: Model 1). As repeatability sets the upper
356 limit on standard heritability (when indirect genetic effects are not considered), heritability
357 estimates were also low across all models. RTL heritability was 0.080 (95% CrIs: 0.041-0.144;
358 Table S6: Model 2) in the simplest model and was estimated as 0.031 (95% CrIs: <0.001-0.067)
359 after the inclusion of all fixed and random effects in the final model (Figure 3, Table 2: Model
360 9). We found a small effect of season of capture and moderate qPCR plate effects in the final
361 model (Table 2). There was no evidence for maternal or paternal effects, territory effects or birth
362 season effects (Table 2). If plate variance was not included in the total phenotypic variance, since
363 it represents technical but not biological variance (following de Villemereuil et al. 2018),
364 individual repeatability was 0.077 (95% CrI: 0.028-0.125), heritability was 0.048 (95% CrI:
365 <0.001-0.087), and capture season was 0.036 (95% CrI: 0.018-0.101) in the final model. A
366 frequentist approach using ASReml-R produced similar results: repeatability was 0.057 ± 0.023
367 SE and heritability was low but significant at 0.041 ± 0.018 SE (Table S7). Without plate
368 included, repeatability was 0.074 ± 0.030 SE and heritability was 0.053 ± 0.023 SE.

369

370

371 Parental effects were compared when social parents (dominant breeding pair) or genetic parents
372 (from the pedigree) were included. Maternal and paternal effects were close to zero in both
373 models and did not differ significantly between models based on overlapping 95% credible
374 intervals (Table 3). Heritability estimates were not significantly different based on the 95%
375 credible intervals of the two models (Table 3).

376

377 **Discussion**

378 We found a negative but weak within-paternal age at conception effect and a positive but weak
379 between-maternal age at conception effect on offspring telomere length in the Seychelles
380 warbler, which adds to the growing literature reporting mixed results in wild populations (Asghar
381 et al. 2015; Belmaker et al. 2019; Eisenberg 2019). Simple mother–offspring, father–offspring or
382 mid-parent–offspring regressions did not provide evidence for telomere heritability in this
383 population. However, animal models indicated a low heritability of telomere length, small catch
384 season effects and moderate experimental effects in the form of qPCR plate effects.

385

386 A number of human studies have documented a positive cross-sectional association between
387 paternal age at conception and offspring telomere length (Unryn et al. 2005; Eisenberg et al.
388 2012; Broer et al. 2013). Including just paternal and maternal age at conception in the model, we
389 found no evidence for cross-sectional parental age at conception effects on offspring telomere
390 length in the Seychelles warbler, even with sufficient power to detect paternal age at conception
391 effects of the correlation coefficients previously published (De Meyer et al. 2007; Nordfjäll et al.
392 2010; Eisenberg et al. 2012, 2017). However, using within-subject centering we found weak but
393 significant within-paternal age at conception and between-maternal age at conception effects on
394 offspring telomere length. In contrast to studies in humans (Unryn et al. 2005; Eisenberg et al.

395 2012; Broer et al. 2013), we found that males produced offspring with shorter telomere lengths
396 as they aged, and females that lived longer tended to have offspring with longer telomere
397 lengths. However, both these effects were relatively small, and explained a very small amount of
398 variation in offspring telomere length.

399

400 Despite the consistency in human studies, studies in non-human vertebrate populations are
401 providing mixed evidence of paternal age at conception effects (Eisenberg 2019). While a few
402 have documented positive paternal age at conception effects (Eisenberg et al. 2017; Dupont et al.
403 2018), most find a negative paternal age at conception effect (Olsson et al. 2011; de Frutos et al.
404 2016; Criscuolo et al. 2017; Bouwhuis et al. 2018; Noguera et al. 2018; Bauch et al. 2019).
405 Furthermore, many studies have documented no parental age at conception effects (Heidinger et
406 al. 2016; Froy et al. 2017; McLennan et al. 2018; Belmaker et al. 2019; van Lieshout et al.
407 2020a), while one study found a positive maternal but no paternal age at conception effect
408 (Asghar et al. 2015). In studies investigating telomere heritability, only one has controlled for
409 parental age at conception effects (Asghar et al. 2015). In our study we did not control for
410 parental age at conception effects in our animal model, since overall effects were not significant
411 and within and between parental age effects were small. However, where parental age at
412 conception effects are significant and large in a population, these effects should be controlled for
413 to obtain accurate estimates of the heritability of telomere length (Dugdale and Richardson
414 2018).

415

416 The majority of studies investigating the heritability of telomere length in wild populations have
417 relied on parent–offspring regression techniques, and have found both significant and non-
418 significant heritabilities ranging from 0 to 1 (Dugdale and Richardson 2018). While these could

419 reflect true differences in heritability estimates in different populations they may also be driven
420 by methodological issues. For instance, despite clear relationships between telomere length and
421 age, many studies have measured parents or offspring telomere lengths at different ages, or have
422 controlled for age of sampling in different ways (Dugdale and Richardson 2018). In a king
423 penguin *Aptenodytes patagonicus* study, chicks were measured at 10, 70, 200 and 300 days old,
424 while mothers were sampled during the breeding season in which chicks were hatched. The
425 authors found a positive association between maternal and offspring telomere length, but only
426 when chicks were 10 days old, indicating that age of measurement can impact heritability
427 estimates (Reichert et al. 2015). Therefore, we investigated how our estimates of telomere
428 heritability differed when using parent–offspring regressions with RTL measured across all ages
429 or just as juveniles. We predicted that resemblance would be higher when both parents and
430 offspring were measured as juveniles, since these would be closest to initial telomere length and
431 the accumulation of environmental effects would be lowest (Dugdale and Richardson 2018).
432 However, we found no evidence of heritability of telomeres using either lifelong telomeres
433 measures or just juvenile telomere measures. Our results contrast with studies which typically
434 show significant mother–offspring rather than father–offspring regressions, indicative of
435 heritable and/or maternal effects (either environmental or genetic in origin) underlying telomere
436 variation (Asghar et al. 2015; Becker et al. 2015; Reichert et al. 2015). Our parent–offspring
437 results indicated a very low or non-significant heritability, and/or a lack of parental effects
438 underlying telomere variation in the Seychelles warbler.

439
440 We subsequently investigated heritability using an animal model approach. Telomere length had
441 a low between-individual repeatability (without plate variance: 0.077; 95% CrI: 0.028-0.125) and
442 a low heritability (without plate variance: 0.048; 95% CrI: <0.001-0.087). The three studies in
443 wild populations that have previously estimated heritability using animal models found either no

444 significant heritable variation underlying telomere length variation (Becker et al. 2015), very low
445 heritability estimates (0.011, 95% CrIs: <0.001-0.042, to 0.060, 95% CrIs: 0.023-0.106
446 depending on prior specification, Foley et al. 2020) or a large heritability of 0.48 (95% CIs: 0.24-
447 0.72, Asghar et al. 2015). However, power analyses were not provided, and sample sizes were
448 generally relatively small for animal models ($N \leq 504$; except see Foley et al. 2020). Large
449 samples sizes are particularly needed in order to fully separate additive genetic effects from
450 common environment effects such as maternal effects (Becker et al. 2015). High within-
451 individual variation in telomere measures has been previously documented in other longitudinal
452 studies (Fairlie et al. 2016; Foley et al. 2020) including in the Seychelles warbler (Spurgin et al.
453 2018). Our estimate of repeatability (between-individual variation) was low and sets the upper
454 limit on ordinary narrow-sense heritability (Bijma 2011).

455
456 The low heritability of telomere length in the Seychelles warbler is consistent with individuals
457 with longer telomeres having a higher probability of surviving until the next year, independent of
458 age (Barrett et al. 2013). It is possible that selection for longer telomeres in this population has
459 reduced the genetic variation, and hence the heritability of this trait (Falconer and Mackay 1996).
460 Further, the large contribution of environmental effects to telomere length is supported by the
461 positive association between food availability and telomere length (Spurgin et al. 2018). Our
462 results indicate that environmental variation, beyond the territory and year effects modelled in
463 this study, explains most of the variation in telomere length in the Seychelles warbler, and
464 indicates a low potential for telomere length to respond to selection.

465
466 The previous studies investigating telomere heritability in wild populations using an animal
467 model approach, which separated out some confounding effects, found differing results

468 regarding the contribution of environmental factors to telomere variation. In the white-throated
469 dipper *Cinclus cinclus*, heritability was not significant, but there were strong nest (0.20 ± 0.08
470 SE) and year of birth effects (0.46 ± 0.13 SE) on telomere length variation (Becker et al. 2015).
471 In comparison, in the great reed warbler *Acrocephalus arundinaceus*, high heritability ($0.48 \pm$
472 0.12 SE) and equally large maternal effects (0.47 ± 0.09 SE) appeared to underlie telomere
473 variation (Asghar et al. 2015). In our study, if we do not account for shared environment effects
474 heritability was 0.080 (95% CrI: 0.041-0.144) and 0.048 (95% CrI: <0.001-0.087) after
475 accounting for natal and current environmental effects, technical effects and parental effects.
476 Further, despite the number of environmental factors measured, including cohort, season,
477 territory and parental effects, our final model only provided evidence for small effects of current
478 season on telomere length variation (0.036, 95% CrI: 0.018-0.101). This contrasts with a number
479 of studies which have observed higher telomere loss in poor natal environment cohorts
480 (Boonekamp et al. 2014; Watson et al. 2015), or suggest an impact of cohort or maternal effects
481 on telomere variation (Asghar et al. 2015; Becker et al. 2015; Fairlie et al. 2016). The lack of
482 parental effects in the Seychelles warbler population may have been caused by the high levels of
483 extra-pair paternity or cobreeding by subordinate females (Richardson et al. 2001; Raj Pant et al.
484 2019), which would result in parental care being provided by the social rather than genetic
485 parent. However, including the social rather than genetic parents in the model did not provide
486 evidence for parental effects. Finally, the lack of natal/parental effects may be because these are
487 only apparent early in life and are diluted when looking at lifelong telomere measures. Indeed,
488 previously we have found cohort effects on juvenile telomere length that did not extend to
489 measures beyond the natal year (Spurgin et al. 2018).

490

491 The lack of parental or early life environment effects in our study may also reflect the sampling
492 regime of our population whereby only a small proportion of samples in the dataset were

493 measured as nestlings (12% chicks). Due to the inaccessibility of nests, which may be up to 30
494 metres up in trees, individuals are usually caught as fledglings during their first 3 months when
495 they remain dependent on their parents (Komdeur 1994; Brouwer et al. 2006). While early-life
496 measures of telomeres will be closer to the inherited telomere length, by using a measure of
497 telomere length across the lifetimes of birds we are measuring a product of inheritance, attrition
498 and restoration/lengthening. After birth, telomere attrition occurs rapidly (Hall et al. 2004;
499 Salomons et al. 2009) and telomere length decreases with age quickest in the first few weeks of
500 life in the Seychelles warbler (Spurgin et al. 2018). With more samples from younger or older
501 individuals it would be possible to investigate how different genetic and environmental effects
502 contribute to telomere variation at different time points. Further, we could have tested for genetic
503 correlations between telomere measures in early and late life and investigated the presence of
504 genotype-by-age interactions. However, our power to calculate heritabilities were lower using
505 measures taken only as nestlings (N = 324 measures of 319 birds, power ≥ 0.80 to detect
506 heritabilities of ≥ 0.23) or individuals showing senescent declines in reproduction and survival
507 (>7 years (Hammers et al. 2015), N = 249 measures of 161 birds, power ≥ 0.80 to detect
508 heritabilities of ≥ 0.40). It is likely that the measurement of telomeres in nestlings in previous
509 studies has resulted in higher heritability, or larger maternal or cohort effects due to the use of
510 only very early-life telomere measures (Asghar et al. 2015; Becker et al. 2015). Further studies
511 investigating how additive genetic variance in telomere length changes with age, and
512 investigating genetic correlations between early and late life, are warranted to understand the
513 genetic constraints on relative telomere length (Dugdale and Richardson 2018).

514

515 An important finding from our study is the impact of technical variation on telomere length
516 measurements. Storage time did not affect telomere length (Spurgin et al. 2018). In contrast, the
517 technician handling the qPCR did have an effect on RTL estimates, and we did find considerable

518 plate effects. While the golden sample should standardise samples within a plate to minimise
519 plate variation it is clear that running the golden sample in a few wells is not capturing
520 differences between plates completely resulting in between plate variation. Previous studies
521 estimating heritability of telomere length using animal models have not included experimental
522 effects likely due to small sample sizes. This technical variance has the potential to bias
523 heritabilities if not included in the analyses (Ponzi et al. 2018). Heritabilities can be re-evaluated
524 with the total phenotypic variance excluding any technical variance to reflect true biological
525 variance (de Villemereuil et al. 2018). Measurement error in qPCR studies could come from
526 various factors such as between- and within-plate effects, technician, storage time, changes in
527 reagents and extraction method effects (Eisenberg et al. 2015; Seeker et al. 2016; Reichert et al.
528 2017; van Lieshout et al. 2020*b*). Such measurement error should be incorporated into analyses
529 and reported to prevent it from biasing results (Nettle et al. 2019). Further, future studies should
530 ensure samples on plates reflect multiple years and ages to enable the greatest statistical power to
531 separate variances of technical and biological interest (van Lieshout et al. 2020*b*).

532

533 In conclusion, our results illustrate that heritability of telomere length would not be identified
534 from parent–offspring regression analyses, and only by using more complex quantitative genetic
535 models could a reliable heritability estimate be calculated. In our population, telomere length
536 variation across an individual’s lifetime was largely driven by environmental factors, including a
537 small catch season effect. There was evidence for a negative, but weak, within-paternal age at
538 conception effect and a positive but weak among-maternal age at conception effect. Further work
539 is needed to see how heritability estimates of telomere length, and telomere loss, calculated using
540 the appropriate power and analytical tools, compare across wild populations.

541

542

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837 **Data Accessibility**

838 Data will be deposited in the Dryad Digital Repository upon acceptance.

839

840 **Author Contributions**

841 This study was conceived by H.L.D. and D.S.R. and developed by A.M.S. D.S.R., H.L.D., T.B.

842 and J.K. manage the long-term Seychelles warbler study system. Samples were collected by

843 D.S.R., K.B., H.L.D. and T.B. Molecular work was undertaken by M.V., E.A.F., K.B., L.G.S.

844 and D.S.R. The genetic pedigree was constructed by H.L.D. A.M.S. performed the statistical

845 analyses and wrote the first draft of the manuscript with input from H.L.D. and D.S.R. All

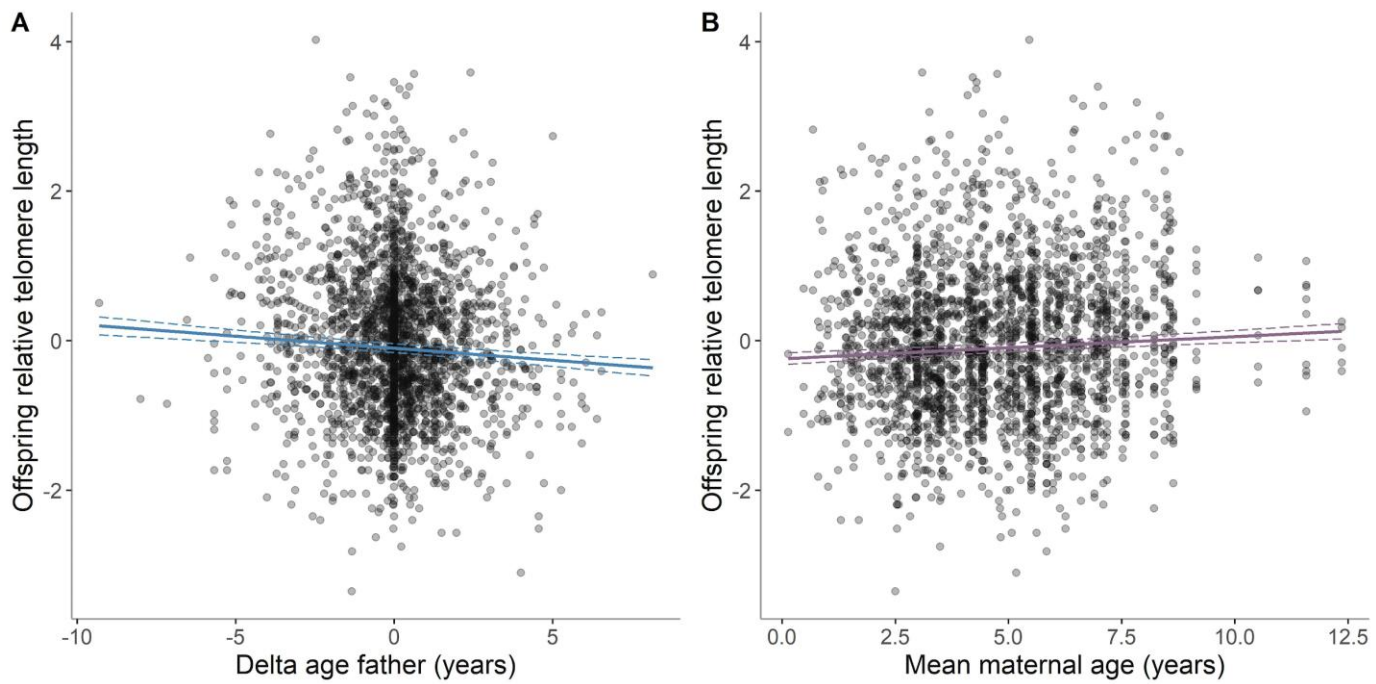
846 authors provided comments on the manuscript and gave final approval for publication.

847

848 **Table 1.** Linear mixed model results investigating between versus within maternal and paternal
849 age at conception (MAC and PAC, respectively) effects on offspring telomere length in the
850 Seychelles warbler using the within-subject centering method (van de Pol and Wright 2009).
851 Associations were investigated in offspring telomere length of all ages (2361 RTL measures of
852 1156 offspring) and included are the estimated effects (estimate), standard errors (SEs), and
853 significance of fixed effects based on a likelihood ratio test (LRT; P-value) where df=1. Relative
854 telomere length was square-root then z-transformed in both models and age was log-transformed.
855 Model 1 investigates within-MAC/PAC effects (deviation from the mean age at conception of
856 the parent: DevMeanMAC/PAC) and between-MAC/PAC (mean age at conception for each
857 parent: meanMAC/PAC) effects. Model 2 investigates whether these within and between slopes
858 are significantly different from each other (mean MAC/PAC representing the difference between
859 the slopes and MAC/PAC which becomes the within-MAC/PAC slope identical to Model 1).
860 P<0.05 are shown in bold.

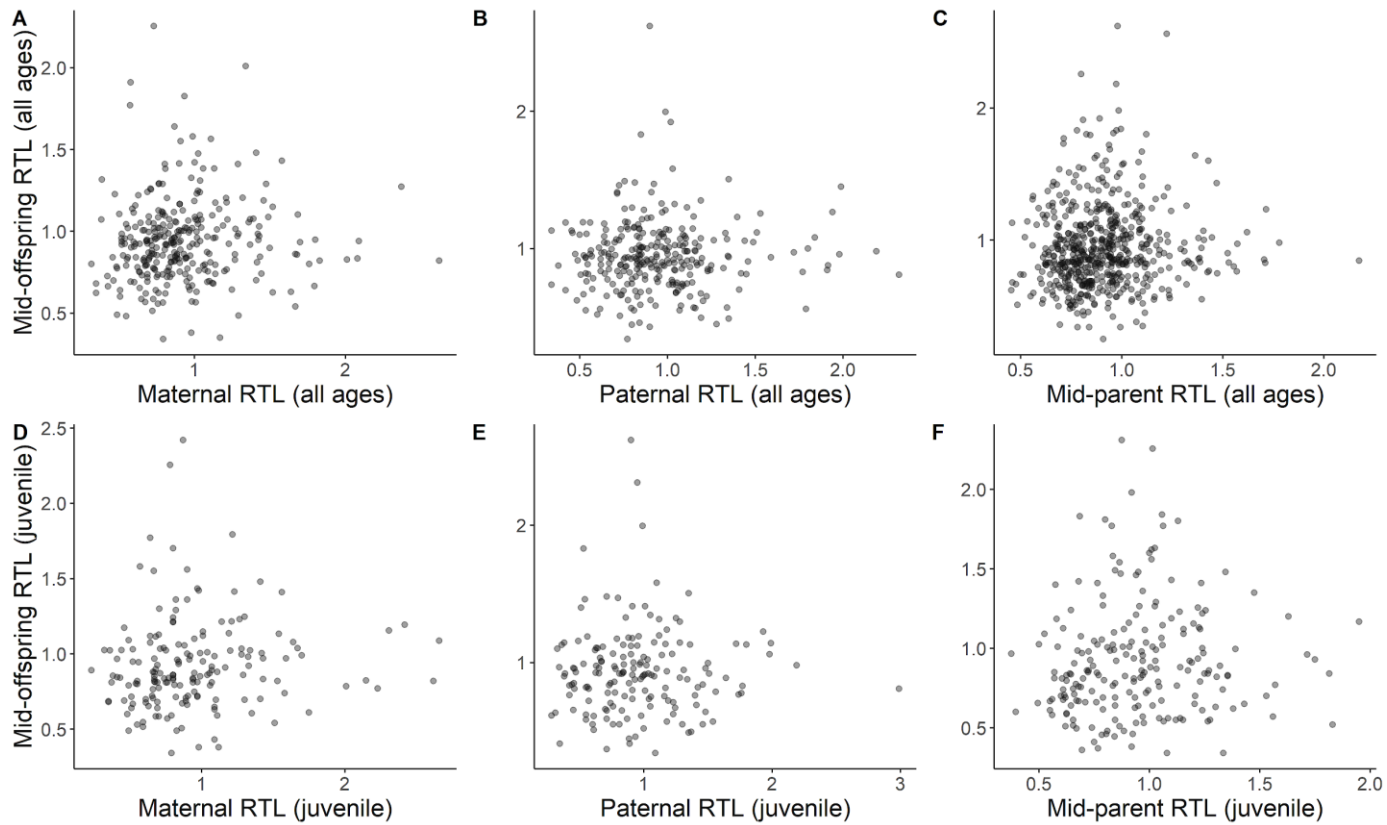
variables	Model 1				Model 2			
	estimate	SE	LRT	P-value	estimate	SE	LRT	P-value
fixed effects								
Intercept	-0.302	0.092			-0.302	0.092		
Log Age (years)	-0.311	0.030	102.250	<0.001	-0.311	0.030	102.250	<0.001
Sex (male)	0.024	0.039	0.374	0.541	0.024	0.039	0.374	0.541
Technician	0.465	0.076	35.954	<0.001	0.465	0.076	35.954	<0.001
MeanMAC	0.030	0.011	7.482	0.006	0.039	0.016	6.002	0.014
DevMeanMAC	-0.010	0.011	0.726	0.394				
MAC					-0.010	0.011	0.726	0.394
MeanPAC	0.001	0.009	0.011	0.918	0.033	0.015	5.102	0.024
DevMeanPAC	-0.032	0.011	8.032	0.005				
PAC					-0.032	0.011	8.032	0.005
random effects								
ID	0.037				0.037			
Mother identity	0.023				0.023			
Father identity	0.005				0.005			
Plate ID	0.186				0.186			
Season ID	0.030				0.030			
Residual	0.653				0.653			

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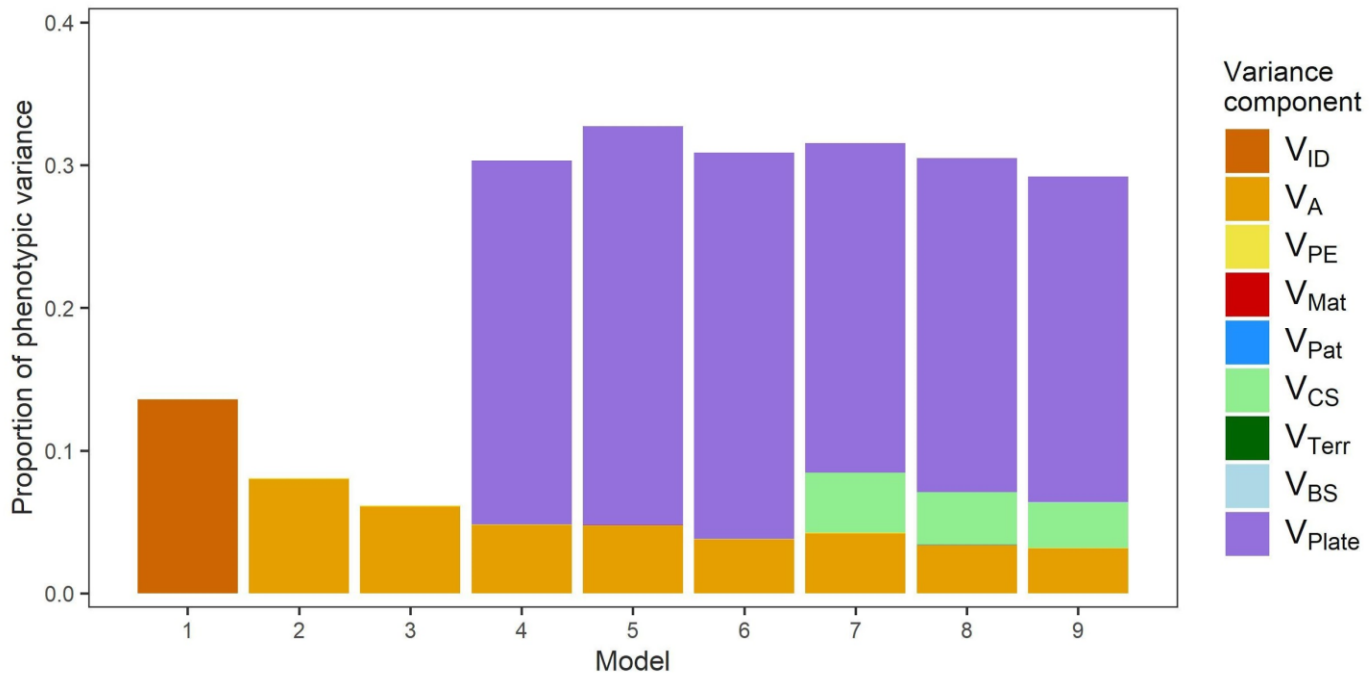
862
 863 **Figure 1.** Scatterplots of raw relative telomere length data from the Seychelles warbler showing
 864 significant negative within-paternal age at conception effects (A) and positive between maternal
 865 age at conception effects (B) on offspring telomere length across all ages (2361 RTL measures of
 866 1156 offspring). Lines indicate mixed model predictions (Model 1, Table 1) using a within-
 867 subject centering method with dashed lines indicating standard errors. Data points are semi-
 868 transparent to show overlapping values.

869



870

871 **Figure 2.** Mid-offspring relative telomere length (RTL) in relation to their mother's (A, D),
 872 father's (B, E) or mid-parent (C, F) RTL in the Seychelles warbler. Data are presented with mean
 873 RTL measures across all ages for both offspring and parents (A-C) and for mean juvenile (<1
 874 year) measures of both offspring and parents (D-F). Where parents had multiple offspring, a
 875 mean of the mean RTL from their offspring was taken. Full model results are provided in Table
 876 S4 and sample sizes are 303 (A), 284 (B), 585 (C), 172 (D), 165 (E) and 210 (F).



877
 878 **Figure 3.** Estimated variance components as proportions of total phenotypic variance in relative
 879 telomere length determined using univariate models in the Seychelles warbler. Models were
 880 fitted additively with increasing random or fixed effects as follows: Model 1 – individual identity
 881 (V_{ID}), 2 – partitioning of V_{ID} into additive genetic (V_A) and permanent environment (V_{PE})
 882 components, 3 – the addition of fixed effects (age, sex, technician), 4 – qPCR plate ID (V_{Plate}), 5
 883 – maternal identity (V_{Mat}), 6 – paternal identity (V_{Pat}), 7 – capture season (V_{CS}), 8 – current
 884 territory (V_{Terr}) and 9 – birth season (V_{BS}). For full model results see Table 2 and S6.

885

886 **Table 2.** Animal model variance component estimates and their associated proportions of the
887 phenotypic variance from a MCMC model of relative telomere length in the Seychelles warbler.
888 Results are from model 9, the model with all variance components and fixed effects estimated
889 (see Methods). Variance components reported are the: additive genetic (V_A), permanent
890 environment (V_{PE}), qPCR plate (V_{Plate}), maternal identity (V_{Mat}), paternal identity (V_{Pat}), capture
891 season (V_{CS}), current territory (V_{CTerr}), birth season (V_{BS}), and residual (V_R) variance. Included
892 are the variance component estimates as the posterior mode along with their lower and upper
893 95% credible intervals (CrI) and the proportion of the total phenotypic variance explained by the
894 term (Prop V_P) with their associated 95% CrI. Significance of fixed effects were determined by
895 whether the 95% CrI did not overlap zero (shown in bold).

Variables	Posterior mode	Lower 95% CrI	Upper 95% CrI	Prop V_P	Lower 95% CrI	Upper 95% CrI
<i>Random effects</i>						
V_A	0.005	<0.001	0.010	0.031	<0.001	0.067
V_{PE}	<0.001	<0.001	0.008	<0.001	<0.001	0.053
V_{Plate}	0.035	0.026	0.045	0.228	0.186	0.287
V_{Mat}	<0.001	<0.001	0.003	<0.001	<0.001	0.022
V_{Pat}	<0.001	<0.001	0.002	<0.001	<0.001	0.011
V_{CS}	0.005	0.002	0.012	0.032	0.013	0.079
V_{Terr}	<0.001	<0.001	0.003	<0.001	<0.001	0.017
V_{BS}	<0.001	<0.001	0.003	<0.001	<0.001	0.021
V_R	0.096	0.090	0.104	0.635	0.588	0.703
<i>Fixed effects</i>						
Intercept	0.864	0.816	0.919			
Sex (male)	0.014	-0.017	0.036	-	-	-
Log Age (years)	-0.117	-0.142	-0.096	-	-	-
Technician	0.199	0.129	0.245	-	-	-

896

897 **Table 3.** Animal model variance component estimates and their associated proportions from a MCMC model of relative telomere length
898 (RTL) in the Seychelles warbler comparing parental effects where genetic parents are included (left, Model 7) and social parents are included
899 (right, Model 10). Variance components reported are the: additive genetic (V_A), permanent environment (V_{PE}), qPCR plate (V_{Plate}), maternal
900 identity (V_{Mat}), paternal identity (V_{Pat}), capture season (V_{CS}), and residual (V_R) variance. Variance component estimates are reported as the
901 posterior mode along with their 95% credible intervals (Lower 95% CrI, Upper 95% CrI) and the proportion of the total phenotypic variance
902 explained by the term (Prop V_P) with their associated 95% credible intervals.

<i>Random effects</i>	Model with genetic parents				Model with social parents							
	Posterior mode	Lower 95% CrI	Upper 95% CrI	Prop V_P	Lower 95% CrI	Upper 95% CrI	Posterior mode	Lower 95% CrI	Upper 95% CrI	Prop V_P	Lower 95% CrI	Upper 95% CrI
V_A	0.006	0.002	0.011	0.042	0.012	0.075	0.007	0.001	0.011	0.042	0.008	0.073
V_{PE}	<0.001	<0.001	0.009	<0.001	<0.001	0.059	<0.001	<0.001	0.009	<0.001	<0.001	0.056
V_{Plate}	0.037	0.028	0.047	0.231	0.193	0.295	0.034	0.028	0.047	0.233	0.193	0.294
V_{Mat}	<0.001	<0.001	0.004	<0.001	<0.001	0.024	<0.001	<0.001	0.003	<0.001	<0.001	0.017
V_{Pat}	<0.001	<0.001	0.002	<0.001	<0.001	0.011	<0.001	<0.001	0.002	<0.001	<0.001	0.016
V_{CS}	0.004	0.002	0.012	0.042	0.013	0.073	0.006	0.002	0.011	0.037	0.011	0.073
V_R	0.097	0.090	0.104	0.634	0.589	0.709	0.096	0.089	0.103	0.648	0.586	0.705

903

904