1   **Title:** Telomere length in house sparrows increases in early-life and can be paternally inherited

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Abstract:

Offspring of older parents in many species display decreased longevity, a faster ageing rate and lower fecundity than offspring born to younger parents. Biomarkers, such as telomeres, that tend to shorten as individual age, may provide insight into the mechanisms of parental age effects. Parental age could determine telomere length either through inheritance of shortened telomeres or through indirect effects, such as variation in parental care with parent ages, which in turn might lead to variation in offspring telomere length. There is no current consensus as to the heritability of telomere length, and the direction and extent of parental age effects however. To address this, here we experimentally investigate how parental age is associated with telomere length at two time points in early life in a captive population of house sparrows (*Passer domesticus*). We experimentally separated parental age from sex effects by allowing the parent birds to only mate with young, or old partners. We found that telomere length of the offspring increased between the age of 0.5 and 3 months at the group and individual level, which has been reported previously predominantly in non-avian taxa. We further show that older fathers produced daughters with a greater early-life increase in telomere length, supporting sex-specific inheritance, and or sex-specific non-genetic effects. Overall, our results highlight the need for more studies testing early-life telomere dynamics and sex-specific heritability of telomere length.

Key words: telomere dynamics, ageing, inter-generational effects, z-linked inheritance, transgenerational effects, Lansing effect

Introduction

Parent age at conception is often associated with their offspring's' life-history, with offspring of older parents commonly having reduced reproductive success and longevity (Heidinger et al., 2016; Priest et al., 2002; Schroeder et al., 2015). Moreover, in some species, offspring of older parents experience higher rates of senescence, cellular ageing, and decreased longevity compared to their older siblings (Bouwhuis et al., 2010; Broer et al., 2013; Torres et al., 2011). While some studies do not find such effects (Froy et al., 2017; Unryn et al., 2005), the associations are reported across a wide range of taxa from rotifers (King, 1983) and insects (Priest et al., 2002) to birds and mammals (Bize et al., 2009; Haussmann et al., 2003b), and is termed the Lansing effect (Lansing, 1947).

The relative length of telomeres, the chromosome capping structures consisting of TTAGGG base pair repeats, is associated with biological age and longevity (Heidinger et al., 2012; Mather et al., 2011; Vedder et al., 2021). Telomeres partly function to prevent DNA damage from reactive oxygen species (Aubert and Lansdorp, 2008). The activity levels of telomerase, the RNA-protein complex responsible for ligating TTAGGG repeats, decline rapidly in early life and are tissue specific (Taylor and Delany, 2000). Together this leads to a gradual telomere shortening over an individual's lifetime (Aubert and Lansdorp, 2008; Finkel and Holbrook, 2000), which is why telomere length is often used as a biomarker for biological age (Mather et al., 2011; Zglinicki and Martin-Ruiz, 2005). However, whether there is a direct causal link between telomere length and an individual's age remains unclear (Boonekamp et al., 2013; Simons, 2015).

In birds, telomere loss is fastest in early-life and an initially longer telomere length is associated with longer subsequent lifespans in captive (Reichert et al., 2013; Wilbourn et al., 2018) and wild (Haussmann et al., 2003a; Heidinger et al., 2016; Reed et al., 2008; Richardson et al., 2001; Salomons et al., 2009; Vedder et al. 2021) bird populations. There is evidence for telomere length being heritable in birds (Vedder et al., 2021), and telomere dynamics have been associated with sex-specific parental age and telomere length (Asghar et al., 2015; Horn et al., 2011; Reichert et al., 2015; Salomons et al., 2009). This suggests that indeed, some Lansing-type effects may be inherited via telomere length. However, the direction of the association between telomere length, and maternal and paternal age varies even within bird species (Dugdale and Richardson, 2018). In birds, the offspring of older mothers may have shorter telomeres and a faster attrition rate, especially in early development (Asghar et al., 2015; Salomons et al., 2009). Conversely, negative associations between paternal age and offspring telomere length have been observed in the absence of maternal correlation (Horn et al., 2011).

Between taxa, studies on the heritability of telomere length are conflicting. The heritability of telomere length can be sex-specific and is often larger in the heterogametic sex; suggesting some degree of maternal inheritance in birds (Asghar et al., 2015; Horn et al., 2011; Reichert et al., 2015) and paternal inheritance in humans (Eisenberg et al., 2017; Njajou et al., 2007; Nordfjäll et al., 2009). However, homogametic inheritance of telomere length has also been found in humans (Broer et al., 2013), in some bird species (Bauch et al., 2019; Bouwhuis et al., 2018), and in lizards (Olsson et al., 2011). A sex-specific lack of heritability has also been found in several bird species (Atema et al., 2015; Heidinger et al., 2012; Kucera, 2018). Overall, parental age effects on offspring telomere length, dynamics and heritability are complex, and vary in extent and direction of impact within and between taxa.

Here, we test for sex-specific, age-related parental effects on offspring telomere dynamics in captive house sparrows *Passer domesticus*. By pairing different age categories of parent birds, we experimentally test the hypothesis that offspring of older parents have shorter telomeres and faster telomere attrition rates than offspring from younger parents.

Methods:

Study species and experimental design:

We used captive house sparrows at the Max Planck Institute for Ornithology, Seewiesen, Germany, during the breeding season of 2014. We used 42 pairs of male and female sparrows, which were assigned to four treatments, each with an equal sex ratio and a uniform distribution of ages across both sexes to control for age-assortative mating. We experimentally bred pairs in one of four age combinations: old-female/ old-male (OO, n=8 pairs), old-female/ young-male (OM, n=11 pairs), young-female/ old-male (YO, n=13 pairs), and young-female/ young-male (YY, n=10 pairs). Young birds hatched the preceding summer. Old (O) was defined as sparrows aged 4 years and older, although most individuals were 7 years or older (Males: 8 years = 2, 9 years = 21; Females: 4 years = 1, 7 years= 10, 8 years= 4, and 9 years= 1). The difference in age distribution between females and males corresponded to that observed in the wild, where females live shorter than males(Schroeder et al., 2012). We did not use the middle aged groups because in wild house sparrows, reproductive senescence may start at 3 years for females (Schroeder et al., 2012), or 5 years in males (Hsu et al. 2017). Each treatment group was replicated twice in two separate breeding groups located in separate aviaries. Each replicate aviary contained 15.3 (s.d.= 4.9) males and 14.6±2.4 females of the respective age class. Bird husbandry is described in Girndt et al. (2017).

Each replicate aviary was equipped with one more nest box than breeding pairs to reduce male-male competition for nest boxes. Sparrows were then allowed to naturally display, form pair bonds, choose a mate restricted by the age class present, and raise their young (Girndt et al, 2018). We systematically monitored breeding and identified the parents attending each nest box by observing the individual birds' colour ring combinations.

Blood sample collection:

We took blood samples from chicks 0.5 months after they hatched (n= 75). After fledging, offspring remained in the same aviary as their parents and siblings, and 2.5 months later were blood sampled again (n=59). Blood samples were collected from the brachial vein of offspring using 1mm capillary tubes and stored in 1ml of 96% ethanol. We collected samples of 56 individuals at both 0.5 and 3 months to test for within-individual changes.

DNA extraction and quantification:

Following standard DNA extraction (Richardson et al., 2001), we measured the DNA concentration of the samples using a ThermoScientific NanoDrop8000 Spectrophotometer and standardised the concentration in our samples to 20-30ng/ml to ensure equal amplification of samples during qPCR. Where necessary, samples were diluted with T10E0.1 (10mM Tris-HCl, pH 8.0, 0.1mM EDTA, pH 8.0) or concentrated using a ThermoScientific Savant DNA SpeedVac Concentrator.

Estimation of telomere length:

We used multiplex qPCR to determine relative telomere length. We determined 'T' as the number of telomere repeats and 'S' as the number of control gene repeats. We then used the T/S ratio as a proxy for telomere length. The four DNA primers we used are described in Criscuolo et al. (2009). We used DNA from house sparrows not included in this analysis as standards at five DNA concentrations of 80, 20, 5, 1.25 and 0.31ng/ml, on each plate. We then used these standards to produce a standard curve for all analysed samples. In each well we added 1.5μl of DNA sample, 0.9μl of each primer, 10μl of Sybr®Select Master Mix and 4.9μl ddH$_2$O. We ran each plate with an equal number of 0.5 and 3 months sample pairs from the same individual to account for any potential sample and plate effects when comparing within-individual changes in telomere length. We ran 42 samples, the five standards and a negative (with all components except a DNA sample) in duplicate on each 96-well plate. We ran the qPCR cycling conditions using QuantStudio 12kFlex Software v1.2.2 following the cycle timings given in Cawthon (2009). We analysed the software output to calculate the T/S ratio in each sample (Appendix 1.1). We altered the thresholds for the standard curve of the telomere and GAPDH primers for each plate to optimise amplification efficiencies to between a standard of 95-110. Efficiencies for each plate were between 99.3-99.7 for

129 GAPDH and 99.3-105.8 for telc and telg. The standard curve for each plate had an R^2 of 0.99 and

130 the intra- and inter-plate variation coefficients all met adequate levels (Cawthon, 2009). We also ran

131 a melt curve to examine whether the expected two products were generated in the reaction.

132 Additionally, we checked all plate amplification curves to see if DNA was present in the control, as

133 this would indicate contamination. In all plates DNA was absent, or present only in very low levels

134 in negatives, apart from very late amplification due to primer dimerization. We repeated any sample

135 duplicates that had a standard deviation of >0.05 following thresholding and used the mean T/S

136 ratio of duplicates in our analysis. T/S ratios of offspring at 0.5 months old are referred to as $T/S_{0.5}$

137 and samples at 3 months old T/S_3 in our analyses. We then calculated the difference between the

138 two measurements as $\Delta T/S$. All samples were analysed for telomere length at the same time and had

139 a similar shelf time (Lieshout et al., 2020). All reagents and equipment were produced by Thermo

140 Fisher Scientific, Waltham, Massachusetts, US.

141 ***Ethical Note:***

142 The Government of Upper Bavaria, Germany, approved the care, handling and husbandry of all

143 birds in this study and granted a license for animal experiments to JS (Nr311.5–5682.1/1-2014-

144 024).

145 ***Statistical Analysis:***

146 Next, we tested for a change in telomere length over the 2.5 months period by running a linear

147 mixed-effects model (LMM) with T/S as response variable, time of sampling (0.5 months or 3

148 months) as an explanatory fixed factor, and individual chick ID as a random effect on the intercept.

149 Then we tested whether telomere lengths in offspring were more variable at either 0.5, or 3 months

150 using a two-tailed F-test. Next, we ran two further LMMs with the response variable $T/S_{0.5}$ and

151 T/S_3, respectively. For each model we tested the fixed effects of the paternal and maternal age

152 categories (either 'young' or 'old'). To test for sex-specific parental effects, we included offspring

153 sex as a categorical variable (with 'male' as the reference level) and an interaction of chick sex with

154 parental age in the $T/S_{0.5}$ model. Because not all chicks were sampled at exactly 3 months after

155 hatching (mean= 100.8 days, s.d.= 8.4), we also tested for an effect of the exact age in days of

156 offspring in T/S_3 with a LMM with T/S_3 as the response variable, and 'sample age' as an

157 explanatory covariate. We found that 'sample age' did not have a statistically significant effect on

158 T/S_3 (posterior mode= -0.001, 95% credible interval= -0.01, 0.001, pMCMC=0.809). Still, to

159 account for any potential bias we retained 'sample age' as a fixed effect in the T/S_3 model.

160 As the 0.5 months samples were a mix of newly-, and already-extracted DNA samples, we also

161 tested whether time of extraction had any effect on the calculated T/S ratio as a result of DNA

162 degradation (Madisen et al., 1987) (n samples newly-extracted= 10 out of 75). We fitted a LMM

163 with $T/S_{0.5}$ as the response and the time of extraction as a fixed factor, either 'newly-' or 'already-

164 extracted'. We found no statistically significant difference between newly-, and already-extracted

165 samples (posterior mode= -0.06, 95% credible interval= -0.20, 0.08, pMCMC=0.389).

166 We included the nest box ID and aviary ID in which chicks were born as random effects on the

167 intercept in all models to account for variance between broods and aviaries. We also included the

168 random term of qPCR plate ID in all models to account for between-plate variance on the intercept.

169 All models were run using the Markov chain Monte Carlo (MCMC) method in the R package

170 MCMCglmm v.2.29 (Hadfield, 2010).

171 *Model validation:*

172 As we used a Bayesian modelling approach, we deemed fixed terms to be statistically significant if

173 their 95% credible intervals (95CI) did not span zero, and we also report MCMC-p-values

174 (pMCMC) (Hadfield, 2010). All terms were retained in models irrespective of their statistical

175 significance. We directly assessed model autocorrelation for fixed and random effects to ensure the

176 risk of type I errors was not inflated. We also inspected iteration and density plots to ensure that

177 effects showed equal variation around a constant mode and demonstrated convergence (Gelman and

178 Hill, 2006; Hadfield, 2010). We examined collinearity of fixed effects, as collinearity could distort

179 model results, which did not exceed 0.7 (Dormann et al., 2013). We ran all models for 100,000

180 iterations with a thinning interval of 10 and used default priors. All statistical analyses were carried

181 out in R v.3.6.1 (R Core Team, 2019)

182 *Results:*

183 Unexpectedly, the telomere length for offspring increased within 80% of individuals between 0.5

184 and 3 months of age (n = 45/56) for those where both measurements were available. On average,

185 the difference between $T/S_{0.5}$ and T/S_3 was statistically significantly positive (Figure 1 and Table 1).

186 Further, as chicks aged, they varied more in their telomere lengths; there was greater variance in

187 T/S_3 than in $T/S_{0.5}$ (coefficient of variance (CV)±s.e.: 0.5 months: 0.22±0.02, n=75, 3 months:

188 0.27±0.02, n=59; F-test: F=0.43, p<0.01).

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Table 1: Results from a Bayesian MCMC linear mixed-effects model testing the difference between telomere length in house sparrow chicks at 0.5 and 3 months of age.

Parameter	Estimate	95% confidence intervals	p_{MCMC}
Intercept	**0.93**	**0.84 - 0.99**	**<0.001**
Chick age	**0.19**	**0.12 - 0.26**	**<0.001**
Random effects			
Chick ID	0.00	0.00 - 0.01	
Nest box	0.02	0.01 - 0.04	
Aviary	0.00	0.00 - 0.00	
qPCR plate ID	0.00	0.00 - 0.01	
Residual	0.04	0.02 - 0.05	

Chick age was modelled as a binary variable of either 0.5 months or 3 months, with 0.5 months as a reference level. 0.5 months: n=75 chicks, 3 months: n=59. Estimates shown are posterior modes. Significant effects are shown in bold.
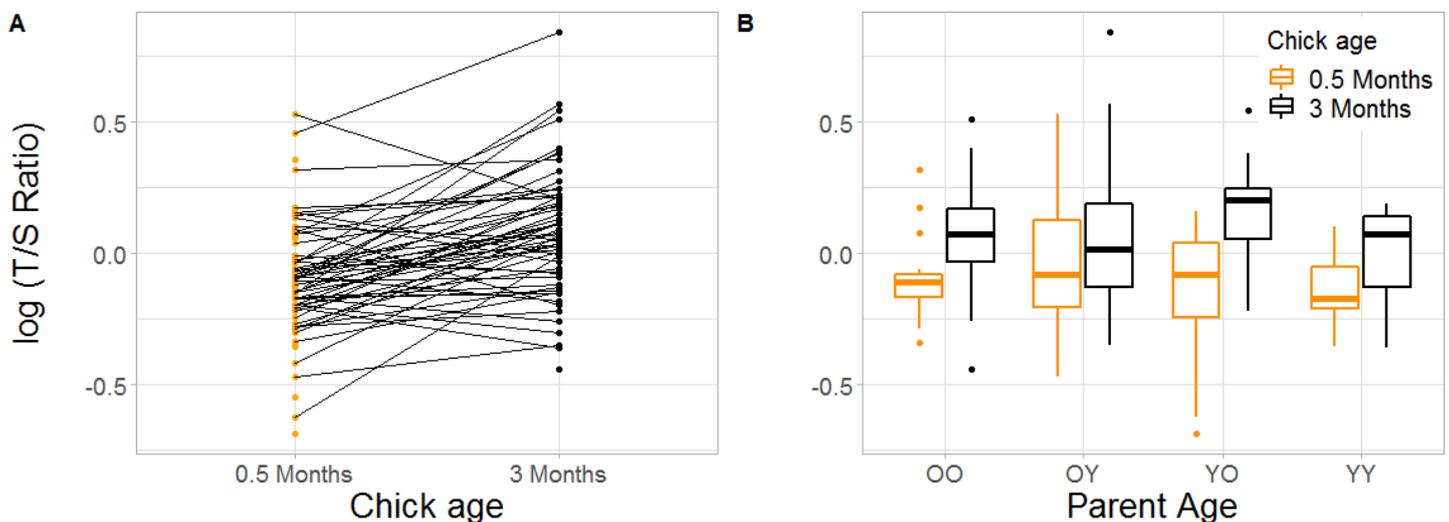
Figure 1: Change in telomere length (log(T/S Ratio)) within house sparrow chicks at 0.5 and 3 months of age. A) Individuals are connected by a line (n offspring with samples at 0.5 months=75, at 3 months=59). B) Boxplots show the mean (central line) and 25th and 75th percentiles (lower and upper box bounds respectively) of the log(T/S Ratio) within age group of the chicks' parents (Y = <2 years old, O > 3 years old for females and >7 years old for males). T/S Ratio is presented on a log scale to aid visualisation. YO = young mothers, old fathers (n=19, 12). OO = both parents old (n=18, 19). OY = old mothers, young fathers (n=17, 18). YY = both parents young (n=15, 10).

We did not find a statistically significant effect of parental age class on $T/S_{0.5}$, which is shortly before sparrows gain independence and fledge from their nest (Table 2). However, the T/S_3 model detected statistically significant effects of paternal age, and the interaction between these two

214 variables. This means that daughters of young fathers had shorter telomeres than daughters of old

215 fathers (Table 2, Fig. 2).

216 **Table 2:** Results from two Bayesian MCMC general linear mixed-effects models with telomere

217 length of house sparrow chicks at age 0.5 months and 3 months as response variables, respectively.

Parameter	$T/S_{0.5}$			T/S_3		
	Estimate	95% CI	p_{MCMC}	Estimate	95% CI	p_{MCMC}
Intercept	**0.97**	**0.84 - 1.10**	**<0.001**	**1.57**	**0.32 – 2.71**	**0.022**
Chick sex	-0.08	-0.22 - 0.11	0.469	-0.23	-0.41 - 0.61	0.121
Maternal age	-0.07	-0.23 - 0.08	0.346	0.06	-0.32 - 0.21	0.711
Maternal age x Chick sex	0.06	-0.15 - 0.25	0.597	0.09	-0.21 - 0.43	0.573
Paternal age	-0.03	-0.20 - 0.10	0.481	**-0.27**	**-0.52 - 0.00**	**0.047**
Paternal age x Chick sex	0.6	-0.07 - 0.34	0.168	**-0.40**	**0.09 - 0.71**	**0.016**
Sample day				0.00	-0.01 - 0.01	0.694
Random effects						
Nest box	0.01	0.00 - 0.03		0.01	0.00 - 0.02	
Aviary	0.00	0.00 - 0.01		0.01	0.00 - 0.01	
qPCR plate ID	0.00	0.00 - 0.01		0.00	0.00 - 0.01	
Residual	0.08	0.03 - 0.11		0.08	0.03 - 0.11	

218 Maternal and paternal age were modelled as a binary variable of either young or old; young was <2 years old, and old

219 was determined as >3 years old for females and >7 years for males). 0.5 months: n=69 chicks, 3 months: n=59. The

220 reference level for parental ages was 'old', and the 'female' was the reference level for chick sex. Estimates shown are

221 posterior modes. Significant effects are shown in bold.

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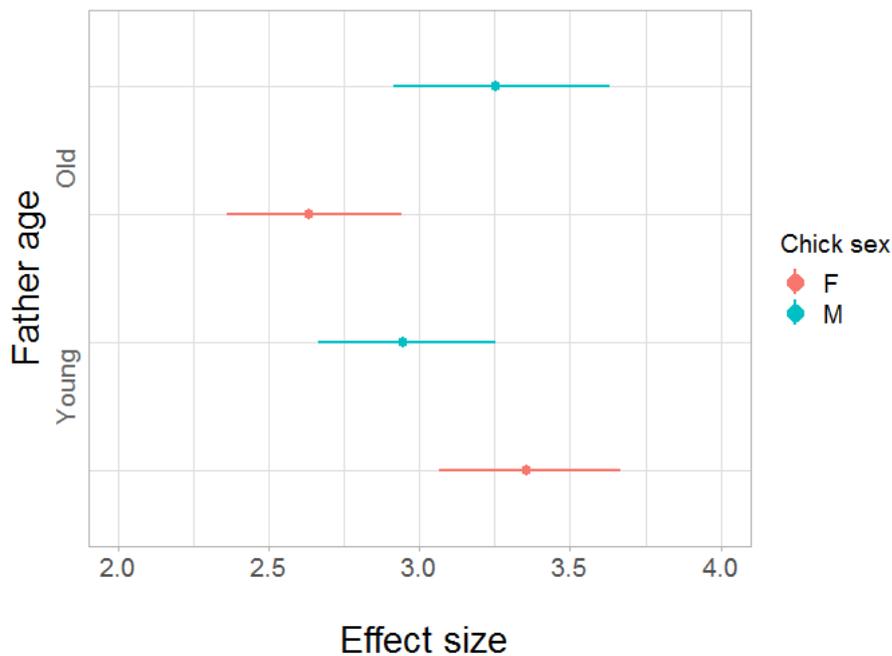
Figure 2: Post-hoc effect size plot from a linear mixed-effects model testing the relationship between T/S_3, father age, and sex of chicks (Table 1). Fathers were assigned an age category of young, 'Y', or old, 'O'. A young father was <2 years old, and an old father was determined as >7 years old. Chick sex is indicated as either female, 'red', or male, 'blue'. The number of offspring in each category; Y, and female=12, male=16, O, and female=16, male=15. Squares represent the model estimates effect sizes of T/S ratio for each paternal age x chick sex, associated lines represent 95% credible intervals.

Discussion:

Individual chick telomere length increased between 0.5 and 3 months of age. This increase disagrees with much of the published literature, which generally find a decrease in telomere length in early-life (Boonekamp et al., 2014; Cerchiara et al., 2017; De Meyer et al., 2007; Hoelzl et al., 2016; Salomons et al., 2009). While a population level increase in telomere length has previously been found in some long-lived bird species (Haussmann et al., 2007; Pauliny et al., 2012), other studies have found that telomeres elongation for a proportion of chicks is more common in smaller, shorter-lived species (Brown et al., 2021; Eisenberg, 2019). For example, a study on jackdaws *Corvus monedula* found that between 5 and 30 days post-hatching, telomere lengths increased for 25% of sampled offspring (Grasman et al., 2011). An increase in early-life telomere length has also been observed in non-avian taxa, including water pythons *Liasis fuscus* (Ujvari and Madsen, 2009) and European badgers *Meles meles* (van Lieshout et al., 2019). A lack of comparable published research exploring a change in telomere length using multiple time points in early life may, in part, explain the surprising nature of our observed increase in telomere length in early-life.

An increase in telomere length can have methodological and/or biological explanations. First, it could be due to DNA in samples degrading over time (Madisen et al., 1987; but see Seutin et al., 1991). Since we used pre-extracted DNA for the majority of 0.5 month samples, we investigated whether differential telomere degradation rates between extracted DNA and blood sample types could be a cause for the observed increase. However, we found no statistically significant difference between the telomere lengths of newly- and already-extracted samples and thus, telomere degradation is an unlikely explanation for our results.

Second, qPCR plates contained both 0.5 and 3 months samples, and between-plate variance was negligible in all our models, highlighting that this element of our methodology had little impact on our results. Overall, we monitored procedural efficiency throughout data collection and did not identify any other potential methodological sources of variation, and so, we are convinced that the increase in telomere length observed in our study has a biological explanation. For example, telomerase activity might have been maintained in the offspring after the first sample was taken. Indeed, two studies have shown that telomerase activity can be maintained up to five weeks post-hatching in zebra finches *Taeniopygia guttata* (Haussmann et al., 2007) and chickens *Gallus gallus* (Taylor and Delany, 2000). Yet, neither of these studies assessed telomerase activity at multiple time points in the same individual's early-life post-hatching, which remains as an interesting future avenue for the field.

While we expected that old parents would produce offspring with shorter telomeres, as found in other short-lived bird species (Bauch et al., 2019; Criscuolo et al., 2017; Sparks et al., 2020), our experimental approach found that old fathers produced daughters with longer telomeres, but only 3 months after hatching, indicating an environmental effect. Similar positive effects of parental age have also previously been found in long and short-lived bird species (Dupont et al., 2018, Asghar et al., 2015; Becker et al., 2015). Positive effects of parental age on offspring telomere length may arise from a potentially improved parental care that older individuals may be able to provide compared to inexperienced, young breeders. Again though, previous studies have found a negative effect of parental age resulting from the poorer condition of these old individuals (Bouwhuis et al., 2018; Criscuolo et al., 2017), or a lack of an effect of parental age on parental care (Nakagawa et al., 2007). Further, in some studies testing parent sex-specific effects, offspring telomere length was found to correlate only with maternal age, and only relatively soon after hatching (ten days: Reichert et al. (2015); nine days: Asghar et al. (2015)). As we found no inflouence of maternal age in our study, an influence of maternal age on offspring telomere length may well have been present, but already diminished below detectable levels 0.5 months after hatching.

While telomere lengths in offspring have been shown to be affected by an offspring's environment (Dugdale and Richardson, 2018; Lieshout et al., 2021), effects of paternal age in birds have been found to be independent of this (Bauch et al., 2019; Boonekamp et al., 2014). As such, overall, there is growing support for at least contributory paternal inheritance of telomere length in some species of birds (Bouwhuis et al., 2018; Olsson et al., 2011; this study). The combined positive effect of having an older father has been theorised to result from an upregulation of telomerase activity in sperm and a subsequent increase in gamete telomere length as males age (De Meyer et al., 2007); as such a positive association of paternal telomere length with age has also been found in humans (Kimura et al., 2008; Unryn et al., 2005). Therefore, a combination of telomerase activity in sperm in fathers, a form of Z-linked inheritance, and potential parental care benefits discussed above may explain the positive effect of increasing father age on offspring telomere lengths, with larger effects seen on daughters compared to sons as observed here.

However, we did not detect an effect of parental age on offspring telomere length at 0.5 months after hatching. Heidinger et al. (2016) similarly found no effect of parental age on offspring telomere length in very early-life at 25 days after hatching in European shags *Phalacrocorax aristotelis*. Further, variation in pre-fledging telomere length may in part be explained by brood-specific additive genetic effects (Voillemot et al., 2012). As such, it may be that at later time points effects of parent age and post-fledging environmental factors appear to be more important than brood-specific effects in determining offspring telomere length. Again, there is a need for more studies investigating the relationship between paternal age and telomere dynamics to detect when and how patterns of telomere dynamics are driven.

In sum, our results indicate that paternal age effects are more influential on offspring telomere length than maternal age effects in our population of house sparrows, with the daughters of older fathers having longer telomeres. Future analyses of telomerase activity levels in both the sperm of adult males and the somatic tissues of offspring would yield further insights into the drivers of parental age effects on offspring telomere dynamics in early-life.

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