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2	Sex-dependent effects of parental age on offspring fitness in a cooperatively breeding bird
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31 Abstract

32 Parental age can have considerable effects on offspring phenotypes and health. However, intergenerational effects may also have longer-term effects on offspring fitness. Few studies 33 have investigated parental age effects on offspring fitness in natural populations while also 34 35 testing for sex- and environment-specific effects. Further, longitudinal parental age effects may be masked by population-level processes such as the selective disappearance of poor 36 quality individuals. Here, we used multi-generational data collected on individually marked 37 Seychelles warblers (Acrocephalus sechellensis) to investigate the impact of maternal and 38 paternal age on offspring lifespan and lifetime reproductive success. We found negative 39 effects of maternal age on female offspring lifespan and lifetime reproductive success which 40 was likely driven by within-mother effects. There was no difference in annual reproductive 41 42 output of females born to older versus younger mothers, suggesting that the differences in 43 offspring lifetime reproductive success are driven by offspring lifespan. In contrast, the lifetime reproductive success of male offspring increased with maternal age, but this was 44 driven by between-mother effects. No within- or between-individual paternal age effects were 45 46 found for female offspring, but fathers that reached old age produced male offspring with higher lifetime reproductive success. We did not find strong evidence for environment-47 dependent parental age effects. Our study provides evidence for parental age effects on the 48 49 lifetime fitness of offspring and shows that such effects can be sex-dependent. These results

add to the growing literature indicating the importance of intergenerational effects on longterm offspring performance and highlights that these effects can be an important driver of
variation in longevity and fitness in the wild.

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54 **Impact summary**

In virtually all animals an individual's health and condition deteriorates with age 55 (senescence), which impacts their survival and the number of offspring produced in later life. 56 Importantly, the quality of offspring produced, as measured through their physiological 57 condition, survival and reproductive success, may also be impacted by their parent's ageing. 58 This may have considerable health and evolutionary implications. However, most of the 59 evidence comes from studies on laboratory animals where artificial conditions exist. Little is 60 61 known about how, if, and why such effects exist in natural populations where a myriad of different stresses and strains act upon individuals. Furthermore, it remains unclear whether 62 63 parental age effects on offspring fitness are sex-specific and affected by environmental 64 conditions. Our study reveals that, in wild-living birds, parental age effects on offspring lifespan and lifetime fitness depend on both the sex of the parent and the sex of the offspring. 65 Cross-sectionally, we found negative effects of maternal age on offspring lifespan and 66 67 lifetime reproductive success of female offspring, which was likely driven by within-maternal age effects. In contrast, there were positive effects of maternal age on the lifetime 68 reproductive success of male offspring. Our study highlights that not including such 69 intergenerational age effects, not separating within- and between-individual effects and not 70 testing for sex-specific effects, may underestimate or obscure important components of 71 72 senescence in wild populations. These intergenerational effects could have important

implications for population dynamics, particularly in age-structured populations, for examplein terms of conservation management and evolutionary dynamics.

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76 Introduction

Early-life environmental conditions can have considerable consequences for later-life 77 individual fitness. Parental effects, whereby an offspring's phenotype is influenced by their 78 79 mother's or father's phenotype, above and beyond the genes they inherit, are widespread in natural populations (Mousseau & Fox 1998) and have important consequences for ecological 80 and evolutionary dynamics (Kirkpatrick & Lande 1989; Wolf et al. 1998). For instance, a 81 recent synthesis of studies found that maternal effects explained half as much phenotypic 82 variance as additive genetic effects across all traits, while for morphological and phenological 83 84 traits maternal effects are comparable to additive genetic effects (Moore et al. 2019). While the impact of parental effects are typically largest in early life (Moore et al. 2019; Gauzere et 85 86 al. 2020), a growing number of studies have indicated that parental effects can have long-term 87 effects on offspring survival and fitness (Arslan et al. 2017; Bock et al. 2019).

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Due to the deterioration of individuals with age in a process known as senescence, parental age is an important component of parental effects with the potential to have large negative consequences on offspring phenotype and performance. For instance, a negative effect of parental age on offspring lifespan has been reported across a range of taxa in both laboratory and field studies ('Lansing effect'; (Plaistow *et al.* 2015; Wylde *et al.* 2019; Monaghan *et al.* 2020), however this effect is not universal (Fox *et al.* 2003; Yilmaz *et al.* 2008; Krishna *et al.* 2012). Age-dependent declines in parental gamete quality (Monaghan & Metcalfe 2019) and

declines in parental care (Hammers et al. 2021) are predicted to cause reduced fitness in 96 97 offspring of elderly parents. Indeed, there is increasing evidence in wild populations that elderly parents produce offspring with lower birth weights, reduced neonatal survival and 98 lower recruitment rates (Hadley et al. 2006; Descamps et al. 2008; Nussey et al. 2009; 99 100 Hoffman et al. 2010; Hayward et al. 2015; Fay et al. 2016; Hammers et al. 2021). However, due to the necessity for long-term longitudinal studies with accurate lifetime reproductive 101 success and survival data, the number of studies that have investigated the impact of parental 102 age on the lifespan and lifetime fitness of offspring in natural populations remains limited 103 (Rödel et al. 2009; Bouwhuis et al. 2010, 2015; Schroeder et al. 2015; Reichert et al. 2019; 104 105 Kroeger et al. 2020). Results from these studies have been mixed, finding for maternal age on 106 offspring lifetime reproductive success: no difference (Bouwhuis et al. 2010), an optimal 107 middle-age effect (Rödel et al. 2009), a lower middle-aged effect (Reichert et al. 2019), or a 108 positive effect (Kroeger et al. 2020). In contrast, other studies have found negative effects of parental age and that older parents produced offspring with shorter lifespans (Bouwhuis et al. 109 2015) and lower lifetime reproductive success (Bouwhuis et al. 2015; Schroeder et al. 2015). 110 111 Further, longitudinal parental age effects may be obscured by population-level processes such as the selective (dis)appearance of poor quality individuals, but these can be separated into 112 within and between-parental age effects using within-subject centering (van de Pol & Wright 113 2009). However, of these studies in natural populations, only two have separated within from 114 115 between parental age effects to account for the effects of selective (dis)appearance (Bouwhuis et al. 2015; Reichert et al. 2019) and only two have investigated both maternal and paternal 116 age effects on offspring of both sexes (Bouwhuis et al. 2015; Schroeder et al. 2015). 117

Parental age studies typically focus on maternal age effects, since these are expected to be 119 120 more important in most species where mothers invest more in offspring than fathers. However, an increasing number are documenting paternal age effects on offspring 121 performance (Bouwhuis et al. 2015; Schroeder et al. 2015; Fay et al. 2016). In addition, 122 123 offspring sex-dependent effects of parental age have been observed (Bouwhuis et al. 2015; Schroeder et al. 2015). These may be due to differences between the sexes in their responses 124 to their early life environment or sex-specific epigenetic inheritance or investment (Bouwhuis 125 et al. 2015). Although the underlying mechanisms are unknown, documented offspring sex-126 specific parental age effects are typically found in offspring of the same sex as the parent 127 128 (Priest et al. 2002; Carnes et al. 2012; Bouwhuis et al. 2015; Lind et al. 2015; Schroeder et al. 129 2015). Further, while early-life environmental conditions affect offspring fitness (Monaghan 130 2008), only one study has investigated how such conditions may interact with parental age 131 (Kroeger et al. 2020). More longitudinal studies investigating both maternal and paternal age effects and testing whether these effects are dependent on offspring sex, or the environment, 132 are needed to fully assess the importance of parental age effects. 133

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Here, we examine the relationships between parental age and offspring lifespan and lifetime 135 136 fitness in the facultatively cooperative-breeding Seychelles warbler (Acrocephalus sechellensis). The entire population on Cousin Island has been the subject of a long-term 137 individual-based monitoring project since 1985. Seychelles warblers live in and defend 138 territories in which a dominant male and female reside. Around half of these territories may 139 140 also contain subordinate individuals, of which approximately one third help with parental care in any given year (Hammers et al. 2019). A combination of high annual resighting rates 141 (Brouwer et al. 2006), virtually no inter-island dispersal (Komdeur et al. 2004) and low 142

extrinsic mortality means that the birds live long lives (up to 19 years) and accurate birth and
death years can be estimated (Hammers *et al.* 2015). High levels of extra-pair paternity and
subordinate female breeding occur (Raj Pant *et al.* 2019; Sparks *et al.* 2021), but a genetically
verified pedigree allows accurate estimation of lifetime reproductive success (Sparks *et al.*2021). Furthermore, the high extra-pair paternity allows us to separate the effects of the age of
the genetic father (e.g. germline deterioration) from that of the social father (deterioration in
parental care).

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Senescence has been well documented in the Seychelles warbler (Hammers et al. 2015) with 151 152 age-dependent declines in both survival and reproduction (Hammers et al. 2012, 2013; Raj Pant et al. 2020). Mothers provide more parental care to the offspring than fathers (Hammers 153 et al. 2019) and the amount of provisioning declines with maternal, but not paternal age 154 (Hammers et al. 2021). Further, the first-year survival of offspring declines with maternal, but 155 156 not paternal age (Hammers et al. 2021). The presence of helpers, however, does compensate 157 for age-related declines in parental care and first year survival with dominant female age 158 (Hammers et al. 2021), providing a means by which social conditions in early life could influence offspring fitness. However, the long-term effects of parental age on offspring 159 160 lifespan and lifetime reproductive success remain unknown.

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In this study, we use a 24-year dataset on individually-marked Seychelles warblers to investigate whether maternal and paternal age are associated with the lifespan and lifetime reproductive success of offspring. We investigate how offspring fitness is associated with maternal age as well as separately testing the effect of the age of the genetic and social father

and for offspring-sex-specific effects on offspring performance. We also investigate whether
parental age effects were dependent on the offspring's early-life environment. We predict that
beneficial early-life environments may counteract the negative impacts of parental age effects
on offspring fitness, whereas negative early-life environments may amplify these effects.
Finally, we separate out parental age effects into within-individual versus between-individual
effects to test for longitudinal versus population-level processes, such as selective
disappearance of poor quality individuals.

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174 Methods

175 *Study system*

The Seychelles warbler is a small insectivorous passerine bird endemic to the Seychelles 176 archipelago. Since 1985, the entire population (ca. 320 adult individuals) on Cousin island 177 (04'20'S, 55'40'E) has been monitored intensively (Komdeur 1992; Richardson et al. 2007; 178 Hammers et al. 2019; Raj Pant et al. 2019). Seychelles warblers live in and defend territories 179 (ca. 115 territories) in which a dominant male and female and any subordinates, if present, 180 reside all year round (Komdeur 1994; Richardson et al. 2001). The dominant breeding pair in 181 each territory is determined by behavioural observations of contact calls and mate guarding 182 (Komdeur 1992; Richardson et al. 2002). The main breeding season runs from June to 183 September and coincides with the peak of insect abundance (Komdeur et al. 1991). However, 184 a smaller proportion of individuals also breed between January and March. On Cousin most 185 clutches contain 1 egg, though clutches of up to 3 eggs can occur (Komdeur et al. 1991; 186 Richardson et al. 2001). While one female can lay more than one egg the presence of multiple 187 eggs in a nest is often indicative of subordinate females breeding (Richardson et al. 2002). 188

Subordinate female cobreeding contributes to 11% of maternities (Raj Pant *et al.* 2019;
Sparks *et al.* 2021), and it is always by subordinate females within the territory, with no
evidence of egg dumping by females from other territories (Richardson *et al.* 2001). Extragroup paternity accounts for 41% of paternities (Raj Pant *et al.* 2019). Only 0.6% of
paternities are gained by (within-group or extra-group) subordinate males (Sparks *et al.*2021).

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Fieldwork is carried out during the breeding seasons, when as many birds as possible are 196 caught using mist nests or captured in the nest. The age of unringed birds is estimated based 197 198 on lay, hatch or fledge dates and/or using eye colour (Komdeur et al. 1991). Since 1995 blood samples have been collected by brachial venepuncture from all birds captured and stored at 199 room temperature in 1 ml of absolute ethanol for molecular sexing and genetic pedigree 200 construction. From 1997 onwards, >96% of individuals in the population have been 201 202 individually marked with a unique combination of a British Trust for Ornithology metal ring 203 and three colour rings (Brouwer et al. 2010). Extrinsic mortality is low due to the absence of predation on adults and individuals have long lifespans (mean = 5.5 years, max= 19 years; 204 (Hammers et al. 2015)). Since the population is virtually closed (Komdeur et al. 2004), and 205 206 annual resighting probability is high (ca. 0.92 ± 0.02 for individuals ≤ 2 years and 0.98 ± 0.01 for older birds, (Brouwer et al. 2010)), individuals can be followed from birth until death 207 (Hammers et al. 2015). 208

209

To investigate the influence that early-life conditions have on parental age effects weinvestigated the effect of natal territory quality, group size and helper presence in the

	offspring's natal territory. We used an index of insect abundance as a measure of territory
213	quality as warblers are almost entirely insectivorous (Komdeur & Daan 2005). Territory
214	quality was calculated using $A*\sum(Cx*Ix)$, where A is the territory size (in hectares), Cx is the
215	amount of foliage cover for tree species x, and Ix is the mean monthly insect density for tree
216	species x per unit leaf area in dm^2 (following (Komdeur 1992; van de Crommenacker <i>et al</i> .
217	2011)). Where territory quality was not available for that field season, we used the mean
218	territory quality of the previous and next field season (Brouwer et al. 2012).
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227 *Statistical analyses*

We included all individuals that were born 1995–2018 with assigned parentage (1643

offspring from 464 mothers and 440 fathers). All statistical analyses were performed in R

v3.6.3 (R Core Team 2019) and all models were run separately for male and female offspring

to allow for differences in lifetime reproductive success and lifespan between the sexes

232 (Figure S1) and to improve interpretability of interactions.

We investigated parental age effects with genetic parents identified using MasterBayes v2.5.2 234 (Hadfield *et al.* 2006) following Sparks et al. (2021). Parentage was assigned where $p \ge 0.8$. 235 The Seychelles warbler pedigree includes parentage assignments for individuals born 1992-236 2018 and contained 1980 individuals with 1708 maternities and 1776 paternities (calculated 237 238 using Pedantics v1.7 (Morrissey & Wilson 2010)). Although 11% of maternities are from subordinate females, there is no difference in the provisioning rates of subordinate or 239 dominant females at nests where a subordinate offspring is present (Table S1). Hence, in this 240 dataset we take the genetic mother as the social mother. In addition to the genetic father, we 241 also included the age of the dominant male in models. Correlations between age of the mother 242 243 and the age of the genetic and/or social father were weak, which allowed them to be included in the same model (maternal age and paternal age: Pearson's correlation; r=0.091, $t_{1641}=3.695$, 244 p<0.001; Figure S2; maternal age and dominant male age: Pearson's correlation; r=0.081, 245 246 t_{1618} =3.288, p<0.001). The correlation between the age of the genetic father and dominant male was moderate (Pearson's correlation; r=0.545, t₁₆₁₈=26.141, p<0.001) which was 247 unsurprising given that 59% of offspring are sired by the dominant male within the territory 248 249 (Raj Pant et al. 2019). We checked for collinearity between fixed effects by calculating 250 Variance Inflation Factors (VIF), as well as overdispersion in all models. To help with model convergence and aid interpretation of the model coefficients, continuous predictor variables 251 were standardised by subtracting the mean and dividing by two standard deviations in arm 252 253 v.1.11-2 (Gelman & Su 2018). Significance of fixed effects was determined by likelihood ratio tests comparing a model with and without the fixed effect of interest. Non-significant 254 255 (i.e. P > 0.05) age² effects and interactions were removed from the models sequentially, in order of least significance, to ensure their inclusion did not affect interpretation of the first-256

order effects. All dropped terms were retested against the base model using likelihood ratiotests.

259

Lifespan was calculated for all individuals who survived until at least one year of age. Since 260 the resighting probability is high, an individual's death year was assigned as the last sighting 261 year. Lifespan analyses were performed using generalised linear mixed models (GLMM) in 262 263 glmmTMB v1.0.2.1 (Brooks et al. 2017) for all individuals who were not translocated as part of a conservation programme (Richardson et al. 2006; Wright et al. 2014) and whose last seen 264 year was before 2019. We included the ages of the genetic mother, genetic father and 265 266 dominant male as fixed effects (linear and squared covariates), in addition to hatch year (linear covariate), the natal group size (linear covariate), presence of helpers in the natal 267 territory (binary factor), natal territory quality (linear covariate), and the presence of siblings 268 in the nest (binary factor). Offspring hatch year was included to account for the fact that 269 270 individuals born more recently in the dataset lived short lives (following (Bouwhuis et al. 271 2015)). Cohort (hatch year), maternal, paternal, and dominant male identity were included as 272 random effects. To calculate a hazards ratio for parental age on offspring survival we also ran these models using Cox mixed effects proportional hazards models in package coxme v2.2-16 273 274 (Therneau 2019) using the same model structure but without standardising variables and without hatch year as a covariate as some individuals in these models were still alive. In these 275 276 models, individuals (130 females, 141 males) that were still alive (i.e. observed in 2019 or 2020), or translocated to other islands (30 females, 44 males) were right-censored. In both 277 278 models, we subsequently tested for environment-dependent parental age effects by testing for two-way interactions of natal territory quality, group size and helper presence with maternal 279 and genetic father and dominant male age. 280

We next tested whether any parental age effects on lifespan translated to differences in reproductive lifespan. We calculated reproductive lifespan as the difference between age at last and first breeding defined by the first and last year an individual had offspring in the genetic pedigree. We included all individuals who survived to one year of age, were not translocated, and whose last seen year was before 2018 (last year of the pedigree). Analyses were performed using a zero-inflated poisson GLMM in glmmTMB with the same model structure as for the lifespan GLMM.

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Lifetime reproductive success (LRS), calculated as the sum of offspring produced who 290 survived to at least one year of age, was calculated for individuals following the same criteria 291 as for reproductive lifespan. Analyses were performed using a zero-inflated poisson GLMM 292 293 in glmmTMB. We included the same fixed and random effects as the lifespan and 294 reproductive lifespan models. To test whether any effects of parental age may be mediated by 295 offspring lifespan, we included offspring lifespan in our models. Similar to the analyses of lifespan, we subsequently tested for environment-dependent parental age effects by including 296 297 interactions between parental age and natal group size, helper presence and natal territory 298 quality.

299

We subsequently investigated whether any parental age effects on offspring lifespan and lifetime reproductive success were driven by within-individual parental age effects rather than between-individual parental age effects, using the within-subject centering method (van de Pol & Wright 2009). We included mean age per parent (between-individual parental age

effects, e.g. testing for selective disappearance of lower-quality individuals) and the deviation 304 from the mean age of the parent (testing for within-individual parental age effects) in the 305 GLMMs of LRS and lifespan. To test whether the within- and between-individual parental 306 age slopes were significantly different from each other, we included age of the parent (within-307 308 parent age effects) and mean age of the parent (difference between the within- and betweenindividual slopes) as predictors in a second model. Using this parameterisation of the model, 309 the significance of mean age indicates that the within and between slopes in the first model 310 are significantly different from each other (van de Pol & Wright 2009). 311

312

313 Finally, in addition to lifespan, parental age effects on offspring lifetime reproductive success may be caused by differences in the annual reproductive output of offspring. Annual 314 reproductive success of offspring was calculated as the number of offspring produced in a 315 year who survived to one year of age. This was calculated for all individuals which had LRS 316 317 data. Analyses were performed using a poisson GLMM in glmmTMB. We included age 318 (linear and squared), lifespan, ages of the mother, father, and dominant male, birth year, 319 territory quality, group size and the presence of helpers (factor) as fixed effects and individual identity and year as random effects. 320

321

322 **Results**

323 Parental age and offspring lifespan

The association between maternal age and offspring lifespan was sex-dependent. Older mothers produced female offspring with shorter lifespans (Table 1 & S2, Figure 1). This effect translated into a 3% increase in mortality risk for female offspring for each year that the

mother was older (Table S2). There was no significant association between maternal age and
lifespan for male offspring (Table 1 & S2). Further, there was no effect of paternal age on
offspring lifespan of either sex (Table 1 & S2). Similar results were found for offspring
reproductive lifespans; older mothers produced female offspring with shorter reproductive
lifespans (Table S3).

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We found no interactions between parental ages and early-life conditions on offspring
lifespan, suggesting no evidence for environment-dependent parental effects (Tables S4 &
S5). However, there was a positive association between natal territory quality and female
offspring lifespan in the Cox proportional hazards model only (Table S2). There was no effect
of helper presence, sibling presence or group size on the lifespan of offspring of either sex.

338

When parental age effects on offspring lifespan were separated into within- versus between-339 individual parental age effects there was a significant negative within-individual maternal age 340 effect on female offspring lifespan (Table 2, Figure 2). As individual mothers aged, the 341 female offspring they produced had shorter lifespans (Table 2, Figure 2). The within-342 individual maternal age slope was significantly more negative than the between-individual 343 344 maternal age slope (Table 2, Table S6). There were no within- or between-individual maternal age effects on male offspring lifespan, and no within- or between-individual paternal age 345 346 effects on offspring lifespan of either sex (Table 2).

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348

349 Parental age and offspring lifetime reproductive success

350 Cross-sectionally, females with older mothers at birth had lower lifetime reproductive success (Table 3, Figure 3A). In contrast, there was a weaker but positive association between 351 352 maternal age and the lifetime reproductive success of male offspring (Table 3, Figure 3B). There was no effect of paternal age, or age of the dominant male, on the lifetime reproductive 353 success of offspring of either sex (Tables 3). There was a strong positive correlation between 354 355 lifespan and lifetime reproductive success (Pearson's correlation; r=0.780, t₉₂₃=37.825, p<0.001; Figure S3). When we added longevity as a covariate to the model, the association 356 between maternal age and female offspring LRS was no longer significant (maternal age 357 β =0.031 ±0.131 SE, $\chi^{2}_{(1)}$ =0.057, p=0.812; lifespan β =1.743 ±0.117 SE, $\chi^{2}_{(1)}$ =204.930, 358 359 p<0.001), which suggests that differences in lifetime reproductive success between female offspring born to different age mothers were driven by differences in longevity rather than by 360 361 differences in annual reproductive output.

362

363 There was no strong evidence for environment-dependent parental age effects (Tables S7). There was a marginally significant paternal age by territory quality and dominant male age by 364 territory quality interaction. It appeared that older males in higher quality territories produced 365 366 female offspring with lower LRS than older males in lower quality territories (Figure S4-7). However, this interaction was driven by a few outliers and the errors around these effects 367 were large (Figure S4-7). In terms of offspring early-life effects, there was a negative effect of 368 helper presence in the natal territory on female offspring lifetime reproductive success (Table 369 3), but no effect on male offspring or effects of group size, sibling presence or territory 370 quality on either sex. 371

373 Parental age effects on offspring LRS were separated into within- versus between-individual parental age effects. There was a significant negative within-individual maternal age effect on 374 375 female offspring lifetime reproductive success (Table 4, Figure 4A). As individual mothers aged, the female offspring they produced had lower lifetime reproductive success (Table 4, 376 Figure 4A). The within- versus between-individual maternal age slopes were not significantly 377 378 different from each other (Table S8). There was also a significant positive between-individual 379 maternal age effect and a marginally significant positive between-individual paternal age effect on male offspring lifetime reproductive success (Table 4, Figure 4B-C). This shows 380 381 that, cross-sectionally, mothers and fathers that survived to, or started to breed at, older ages 382 had male offspring with higher lifetime reproductive success (Table 4, Figure 4B-C). However, there was no difference in the within versus between slopes for either maternal or 383 384 paternal age for male offspring (Table S8).

385

386 Parental age and annual reproductive success

We found no association between maternal, paternal or dominant male age on offspring annual reproductive success (Table S9). This suggests that any parental age effects on offspring lifetime reproductive success are driven by differences in lifespan rather than annual reproductive output. There was a negative effect of helper presence at the nest in which the individual was hatched on the annual reproductive success of female offspring. This suggests that the lower lifetime reproductive success of female offspring from nests with a helper were driven by a reduction in annual reproductive output.

395 **Discussion**

396 We found evidence for sex-dependent parental age effects on offspring fitness in the Seychelles warbler, with negative within-individual maternal age effects for female offspring 397 but positive between-individual maternal age effects for male offspring (summarised in Table 398 5). Female offspring born to older mothers had shorter lifespans and lower lifetime 399 400 reproductive success than females born to younger mothers, which appeared to be driven by 401 within-maternal age effects. Male offspring born to older mothers had higher lifetime reproductive success than males born to younger mothers, which appeared to be driven by 402 between-maternal age effects. Although there were no cross-sectional paternal age effects, 403 404 there was a positive between-paternal age effect on male offspring lifetime reproductive 405 success.

406

407 The mechanisms by which parental age impacts offspring quality or lifespan are not well 408 resolved, especially in wild populations. Direct effects of ageing parents on offspring 409 phenotypes could be caused by age-dependent declines in parental gamete quality (Monaghan & Metcalfe 2019), declines in the quality of parental care (Hammers et al. 2021) or age-410 dependent increases in the costs of inbreeding depression (Charlesworth & Hughes 1996). In 411 412 our study, we tested for effects of the ages of the mother, genetic father and the dominant 413 male on offspring lifespan and lifetime reproductive success. While 11% of maternities are due to subordinate females (Sparks et al. 2021), both dominant and subordinate females 414 415 provide parental care to offspring (Table S1). Consequently, effects of the age of the genetic mother on offspring phenotype may be caused by age-dependent changes in gamete quality, 416 417 incubation of the egg, parental care of the nestling, or epigenetic inheritance. In contrast, 41%

of paternities are sired by an extra group dominant male who provides no parental care to the 418 419 offspring (Raj Pant et al. 2019). This allowed us to compare the effects of the age of the genetic father and the dominant male on offspring phenotype. While age of the genetic father 420 could be associated with changes in gamete quality or epigenetic inheritance, the age of the 421 422 dominant male may be important as they provision and defend the offspring. However, we only found within-individual maternal age effects on offspring, and no paternal age effects, of 423 either the genetic father or dominant male. The greater impact of maternal age on offspring 424 performance is perhaps unsurprising given that only females incubate and females provision 425 more than males (Hammers et al. 2019; van Boheemen et al. 2019). Further, age-dependent 426 427 declines in provisioning rates occur for dominant females but not males, which also suggests that maternal rather than paternal age should be more important for offspring quality in this 428 429 population (Hammers et al. 2021).

430

While negative effects of maternal age on early-life offspring survival have been widely 431 432 reported (Descamps et al. 2008; Nussey et al. 2009; Reid et al. 2010; Torres et al. 2011; 433 Cooper et al. 2020), our study adds to the growing evidence of sex-specific parental age effects on long-term offspring fitness in natural populations (Bouwhuis et al. 2015; Schroeder 434 435 et al. 2015). So far only two studies have investigated both maternal and paternal age effects on the lifespan and lifetime reproductive success of female and male offspring. In both wild 436 house sparrows (Passer domesticus) and common terns (Sterna hirundo) female, but not 437 male, offspring born to older mothers had lower lifetime reproductive success (Bouwhuis et 438 439 al. 2015; Schroeder et al. 2015). While differences in lifespan appear to be driving the lower LRS of female offspring born to older mothers in the Seychelles warbler, in these previous 440 studies differences in annual reproductive success appeared to cause the effect; i.e. there was 441

no association between maternal age and lifespan (Bouwhuis et al. 2015; Schroeder et al. 442 443 2015). In contrast to our positive between-individual maternal age effect observed on the lifetime reproductive success of male offspring, the two previous studies found no maternal 444 445 age effects on males. However, they did find a negative effect of paternal age on male annual 446 reproductive success in the house sparrow and a negative effect of paternal age on male offspring lifespan in the common tern. Both of these effects translated to reduced lifetime 447 reproductive success of male offspring born to older fathers (Bouwhuis et al. 2015; Schroeder 448 et al. 2015). In contrast, our study found a positive between-individual paternal age effect on 449 lifetime reproductive success of male offspring. 450

451

The majority of studies in natural populations have found negative effects of increased 452 maternal or paternal age on offspring lifespan and lifetime reproductive success (Rödel et al. 453 2009; Bouwhuis et al. 2015; Schroeder et al. 2015), or in one case no effect (Bouwhuis et al. 454 455 2010). However, there is also some evidence for positive effects. In yellow-bellied marmots 456 (Marmota flaviventer), daughters born to older mothers had greater annual reproductive 457 success, which translated to higher lifetime reproductive success (Kroeger et al. 2020). The authors suggested that maternal restraint (due to residual lifespan; (Williams 1966; Pianka 458 459 1976)) or constraint (e.g. constrained skills or physiological condition; (Curio 1983)) of younger mothers and increased experience or terminal investment in older mothers may 460 explain this (Kroeger et al. 2020). In a population of semi-captive Asian elephants (Elephas 461 maximus) a complex pattern was observed (Reichert et al. 2019); female offspring survival 462 decreased with increasing maternal age, but daughters from younger and older mothers had 463 higher lifetime reproductive success than daughters from middle-aged mothers. The authors 464 suggest that as middle-aged mothers have the highest annual reproductive rate and working 465

activity in the timber industry, they may invest less in each offspring (Reichert *et al.* 2019). In
our warbler population, the cross-sectional positive effect of maternal age on male lifetime
reproductive success appeared to be driven by between-maternal age effects. While we found
no cross-sectional effect of paternal age, we did find a positive between-paternal age effect on
male offspring lifetime reproductive success. Overall, these results show that mothers and
fathers breeding at older ages have male offspring with higher lifetime reproductive success.

472

Associations between parental age and offspring lifespan could, potentially, be linked to 473 474 telomere dynamics. Telomeres, the protective caps at the ends of chromosomes, tend to 475 shorten with age and shorter telomeres are associated with increased mortality risk (Salomons et al. 2009; Aubert et al. 2012; Wilbourn et al. 2018) including in Seychelles warblers 476 (Barrett et al. 2013; Spurgin et al. 2018). In humans, there is cross-sectional evidence that 477 sperm telomere length is positively correlated with age and older fathers have offspring with 478 479 longer telomeres (Unryn et al. 2005; Kimura et al. 2008; Aston et al. 2012; Eisenberg et al. 480 2012; Broer et al. 2013). However, evidence from natural vertebrate populations is mixed (Eisenberg 2019). In the Seychelles warbler, there is a weak negative within-paternal age at 481 conception effect on offspring telomere length (Sparks et al. 2021). Additionally, there is no 482 483 evidence for within-maternal age at conception effects on offspring telomere length, which might be expected if there is selection for higher-quality oocytes with longer telomeres to be 484 used first (Monaghan et al. 2020). Therefore, the parental age effects on offspring fitness 485 observed in our study are unlikely to be due to parental age at conception effects on offspring 486 487 telomere length.

488

While there is growing evidence of sex-specific parental age effects on offspring lifespan and 489 490 fitness, the mechanisms through which these occur are unknown. Studies have reported different sensitivities between the sexes to poor early-life environments (Jones et al. 2009). 491 For example, in great tits (Parus major) natal environmental conditions impacted the lifespan 492 493 and breeding success of males but not females (Wilkin & Sheldon 2009). This may be because the sexes respond differently to, or require different levels, of early-life investment 494 (Jones et al. 2009; Wilkin & Sheldon 2009). An early comparative study showed that juvenile 495 mortality was proportional to sexual size dimorphism (Clutton-Brock et al. 1985). Further 496 work suggested that sex differences in mortality may be due to sex-linked deleterious alleles 497 498 in the heterogametic sex (Myers 1978), and the effects of these deleterious alleles may be 499 dependent on the environment (e.g. (Marr et al. 2006)). However, a more recent metaanalyses found that sex differences in environmental sensitivities in birds could not be 500 501 explained by sex or size-dependent vulnerabilities alone (Jones et al. 2009). Further explanations could include sex-specific epigenetic inheritance (Hager et al. 2008; Broer et al. 502 2013), age- and sex-specific differences in foraging, offspring provisioning (Weimerskirch 503 504 2018), or differential investment in eggs, which could result in complex sex-specific parental 505 age effects on offspring condition (Bouwhuis et al. 2015).

506

Despite ample evidence that early-life environments can affect offspring performance
(Monaghan 2008), we found no evidence that parental age effects on offspring fitness differed
with early-life conditions. In the warbler system, offspring early-life survival declines with
dominant female age, but helpers may mitigate these declines, suggesting that the early-life
social environment could impact parental age effects (Hammers *et al.* 2021). However, we
found no evidence that helper presence, territory quality or group size, impacted the effects of

parental age on the fitness of offspring that survived their first year of life (i.e. their most 513 514 critical period). Only one study has examined the impact of the environment on long-term parental age effects (Kroeger et al. 2020). In yellow-bellied marmots, there was a positive 515 516 association between maternal age and the lifetime reproductive success of daughters. 517 However, daughters born to older mothers in favourable environments had greater declines in annual reproductive success with age. While contrary to predictions, this was driven by 518 519 daughters of older mothers in harsher environments not living long enough to senesce. Further, natal litter size negatively affected daughters born to older but not younger mothers, 520 indicating an offspring number and quality trade-off for older mothers (Kroeger et al. 2020). 521 522 Although we did not find evidence for environment-dependent parental age effects in our study population, the marmot study indicates that future studies should also investigate 523 524 whether parental age effects are dependent on early life environmental conditions.

525

526 In the Seychelles warbler, provisioning rates and subsequent juvenile survival decline with 527 dominant female, but not male, age (Hammers et al. 2021). Further, female offspring are 528 more likely to forgo dispersal and help as the dominant female ages (Hammers *et al.* 2019). This could explain the negative effect of maternal age on the fitness of female, but not male, 529 530 offspring in our study. However further work is needed to determine the mechanisms by which these sex-specific parental age effects occur and whether being a helper reduces 531 survival. While actuarial and reproductive senescence occur in this warbler population 532 (Hammers et al. 2012, 2013; Raj Pant et al. 2020), our study indicates that parental age 533 534 effects can have longer-term effects on offspring fitness. Not investigating such intergenerational effects, including sex-specific effects, may therefore miss important 535 components of senescence. 536

538 Finally, a negative effect of increasing parental age on offspring lifespan may not necessarily have negative consequences for offspring fitness if the reproductive timing of offspring is 539 540 altered (Monaghan et al. 2020). For instance, in great tits, offspring born to different age mothers had different age-specific reproductive success trajectories, but these equated to the 541 542 same lifetime reproductive success (Bouwhuis et al. 2010). We found no difference in annual 543 reproductive output in relation to parental age. However, females from older mothers had shorter reproductive lifespans and lower lifetime reproductive success than those from 544 younger mothers. In cooperatively breeding species such as the Seychelles warbler, these 545 546 parental age effects may reveal an additional cost to females of becoming a subordinate if it 547 leads to delayed reproduction. Further, older dominant females are more likely to be displaced from their breeding position by a subordinate than are dominant males (Richardson et al. 548 549 2007), which may be adaptive if helping older females to reproduce is less beneficial.

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In conclusion, the existence of complex, often sex-specific, intergenerational effects linked to
parental age, will have consequences for our understanding of the fitness effects of
senescence, and subsequently, evolutionary dynamics. Such effects will also have important
ramifications for animal breeding and the conservation of species, particularly where
populations are small or age-structured and when older individuals are used for breeding or
translocations.

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573 Author Contributions

The idea to test the Lansing effect in this population was conceived by AMS, MH, HLD, DSR,
TB and JK. AMS and MH conceived the specific study design, data selection and methodology.
AMS performed all data analyses with input from MH and HLD. HLD, DSR, JK and TB
manage the long-term Seychelles warbler study system including gaining the relevant funding.
HLD constructed the genetic pedigree. AMS wrote the manuscript with input from MH, HLD
and DSR. All authors gave final approval for publication.

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583 Data Accessibility

- 584 Data will be deposited in the Dryad Digital Repository upon acceptance. All scripts for the
- 585 analysis will be made available at https://github.com/Seychelle-Warbler-Project
- 586
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Figure 1. Survival probability of female Seychelles warbler offspring over time broken down
by age of the mother in the cross-sectional analyses. Maternal age was treated as a continuous

variable in our analyses, but for graphical purposes, here maternal age is grouped into four

age classes 0-3 years, 4-6 years, 7-9 years, and 10+ years.





Figure 2. Within-individual maternal age effects ('Delta age mother') on the lifespan of

female offspring in the Seychelles warbler, using within-subject centering (van de Pol &
Wright 2009). The solid line is the GLMM prediction with mean values for all other

continuous fixed effects in the model and the dashed lines indicate standard errors (Table 2).

Dots show the raw data points. Lifespan values are integers but are jittered to show

overlapping values. Model estimates (Table 2) are similar without the outlier (the offspring

with a lifespan of 20 years) (Table S10).



Figure 3. Cross-sectional maternal age (years) effects on the lifetime reproductive success of female (A) and male (B) offspring (based on

828 genetically assigned offspring) in the Seychelles warbler. Solid lines indicate GLMM predictions with mean values for all other continuous fixed

effects in the model and dashed lines indicate standard errors (Table 3). Dots show the raw data points. LRS values are integers but are jittered to
 show overlapping values.



Figure 4. Within and between genetic parental age effects on the lifetime reproductive success (LRS) of female (A) and male (B-C) offspring in
the Seychelles warbler using within-subject centering (van de Pol & Wright 2009). Lines indicate GLMM predictions with mean values for all
other continuous fixed effects in the model and dashed lines indicate standard errors (Table 4). Dots show the raw data points. LRS values are
integers but are jittered to show overlapping values. Model estimates (Table 4) are similar without the outlier in Figure 4A (Table S11).

Table 1. Generalised linear mixed model results investigating associations between parental age effects on offspring lifespan, for male and
 female offspring separately, in the Seychelles warbler (including all genetically assigned offspring who survived to one year of age). Included are
 the parameter estimates (estimate), their standard errors (SE), and the significance of fixed effects based on a likelihood ratio test (LRT, P) where
 d.f.=1. Significance of dropped squared parental ages, and environment by parental age interactions are shown in Table S4. Significant fixed
 effects are in bold.

	Female offs	pring life	span		Male offs	pring lifes	pan	
	n=467				n=485		-	
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	1.276	0.061			1.196	0.072		
age of mother	-0.242	0.075	10.569	0.001	0.098	0.068	2.098	0.148
age of father	-0.041	0.088	0.217	0.641	-0.062	0.085	0.545	0.460
age of dominant male	0.135	0.082	2.676	0.102	0.125	0.081	2.363	0.124
birth year	-0.584	0.108	22.852	<0.001	-0.556	0.138	13.710	<0.001
territory quality	0.019	0.081	0.054	0.816	0.149	0.093	2.598	0.107
group size	-0.109	0.077	1.994	0.158	-0.075	0.080	0.863	0.353
helper presence (y)	-0.153	0.092	2.787	0.095	-0.128	0.101	1.631	0.202
sibling presence (y)	-0.068	0.103	0.439	0.508	0.135	0.100	1.835	0.176
random effects								
maternal ID	0.085				0.086			
paternal ID	0.163				0.084			
dominant male ID	0.019				0.068			
birth year	0.021				0.068			

Table 2. Generalised linear mixed model results investigating between versus within maternal and paternal age effects on offspring lifespan, in
 each offspring sex separately, in the Seychelles warbler using the within-subject centering method (van de Pol & Wright 2009). Included are the
 estimated effects (estimate), standard errors (SEs), and significance of fixed effects based on a likelihood ratio test (LRT; P-value) where df=1.
 The model investigates within-parental age effects (deviation from the mean age of the parent) and between-parental age effects (mean age for
 each parent). Significant fixed effects are in bold.

	Female offsp	oring lifespa	n		Male offs	pring lifespa	an	
	n=467				n=485			
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	1.284	0.062			1.203	0.070		
between-maternal age	-0.097	0.090	1.140	0.286	0.127	0.095	1.793	0.181
within-maternal age	-0.508	0.121	17.503	<0.001	0.092	0.096	0.920	0.338
between-paternal age	0.004	0.110	0.001	0.973	0.045	0.104	0.189	0.664
within-paternal age	-0.157	0.134	1.361	0.243	-0.198	0.128	2.437	0.119
between-dominant male age	0.113	0.094	1.437	0.231	0.158	0.099	2.570	0.109
within-dominant male age	0.252	0.146	2.974	0.085	0.102	0.131	0.604	0.437
birth year	-0.525	0.112	19.000	<0.001	-0.514	0.138	12.252	<0.001
territory quality	0.022	0.080	0.075	0.784	0.150	0.093	2.645	0.104
group size	-0.104	0.077	1.825	0.177	-0.065	0.081	0.654	0.419
helper presence (y)	-0.153	0.092	2.806	0.094	-0.140	0.101	1.936	0.164
sibling presence (y)	-0.069	0.103	0.446	0.504	0.142	0.100	1.991	0.158
random effects								
maternal ID	0.089				0.088			
paternal ID	0.170				0.081			
dominant male ID	0.004				0.069			
birth year	0.022				0.054			

Table 3. Generalised linear mixed model results investigating associations between parental age effects on offspring lifetime reproductive
 success (LRS) for males and female offspring separately in the Seychelles warbler (including all genetically assigned offspring who survived to
 one year of age). Included are the parameter estimates (estimate), their standard errors (SE), and the significance of fixed effects based on a
 likelihood ratio test (LRT, P) where d.f.=1. Significance of dropped squared parental ages, and environment by parental age interactions are
 shown in Table S7. Significant fixed effects are in bold.

	Female offsp	ring LRS			Male offspri	ng LRS		
	n=441				n=466			
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	0.506	0.128			0.298	0.143		
age of mother	-0.427	0.152	7.817	0.005	0.323	0.140	5.297	0.021
age of father	0.145	0.162	0.798	0.372	0.265	0.174	2.295	0.130
age of dominant male	-0.083	0.158	0.276	0.599	0.023	0.154	0.023	0.879
birth year	-1.082	0.221	25.568	<0.001	-1.169	0.273	17.446	<0.001
territory quality	-0.135	0.143	0.892	0.345	0.245	0.199	1.495	0.222
group size	-0.032	0.151	0.046	0.830	-0.144	0.172	0.708	0.400
helper presence (y)	-0.635	0.203	10.055	0.002	-0.220	0.215	1.069	0.301
sibling presence (y)	-0.078	0.191	0.167	0.683	0.081	0.205	0.156	0.693
random effects								
maternal ID	0.134				0.155			
paternal ID	0.104				0.265			
dominant male ID	0.030				<0.001			
birth year	0.050				0.105			

Table 4. Generalised linear mixed model results investigating between versus within maternal and paternal age effects on offspring lifetime
 reproductive success in each offspring sex separately in the Seychelles warbler, using the within-subject centering method (van de Pol & Wright
 2009). Included are the estimated effects (estimate), standard errors (SEs), and significance of fixed effects based on a likelihood ratio test (LRT;
 P-value) where df=1. The model investigates within-parental age effects (deviation from the mean age of the parent) and between-parental age
 effects (mean age for each parent). Significant fixed effects are in bold.

	Female offsp	ring LRS			Male offsprin	ng LRS		
	n=441				n=466			
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	0.513	0.131			0.345	0.143		
between-maternal age	-0.298	0.166	3.197	0.074	0.442	0.176	6.424	0.011
within-maternal age	-0.835	0.261	9.640	0.002	0.139	0.219	0.402	0.526
between-paternal age	0.046	0.178	0.067	0.796	0.461	0.220	4.383	0.036
within-paternal age	0.291	0.278	1.103	0.294	-0.071	0.277	0.067	0.796
between-dominant male age	-0.140	0.170	0.690	0.406	-0.134	0.192	0.491	0.484
within-dominant male age	0.025	0.301	0.007	0.935	0.324	0.307	1.129	0.288
birth year	-1.076	0.227	23.569	<0.001	-1.119	0.272	16.840	<0.001
territory quality	-0.136	0.142	0.927	0.336	0.219	0.199	1.211	0.271
group size	0.004	0.151	0.001	0.981	-0.097	0.175	0.304	0.581
helper presence (y)	-0.611	0.201	9.336	0.002	-0.279	0.222	1.643	0.200
sibling presence (y)	-0.109	0.189	0.333	0.564	0.070	0.204	0.118	0.732
random effects								
maternal ID	0.100				0.153			
paternal ID	0.090				0.234			
dominant male ID	0.046				<0.001			
birth year	0.064				0.085			

Table 5. Summary of the models investigating parental age effects on offspring lifespan and LRS in the Seychelles warbler. Included are the

867 cross-sectional effects, within-individual (w/n) and between-individual (b/n) parental age effects and whether there was a significant difference

between the within- and between-individual parental age slopes (diff). Positive estimates are indicated by +ve, negative estimates are indicated

by -ve, ns indicates non significant results, and Y and N indicate yes and no, respectively. Significant effects are highlighted in bold.

			Age of	mother			Age of gene	etic father		1	Age of domi	nant male	
		cross-				cross-				cross-			
_		sectional	w/n	b/n	diff	sectional	w/n	b/n	diff	sectional	w/n	b/n	diff
Offspring	Females	-ve	-ve	ns	Y	ns	ns	ns	Ν	ns	ns	ns	Ν
lifespan	Males	ns	ns	ns	Ν	ns	ns	ns	Ν	ns	ns	ns	Ν
Offspring	Females	-ve	-ve	ns	Ν	ns	ns	ns	Ν	ns	ns	ns	Ν
LRS	Males	+ve	ns	+ve	Ν	ns	ns	+ve	Ν	ns	ns	ns	Ν

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872	Supplemental information for:
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874 875	Sex-dependent effects of parental age on offspring fitness in a cooperatively breeding bird
876	
877	Alexandra M. Sparks, Martijn Hammers, Jan Komdeur, Terry Burke,
878	David S. Richardson, Hannah L. Dugdale
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Figure S1. Histograms of the frequency of offspring lifespan (A-B) and lifetime

reproductive success (measured as the number of assigned offspring in the pedigree who
survived to 1 year, C-D) in the Seychelles warbler for females (A,C) and males (B,D)
who survived to one year and had complete lifetime data (A: n=476; B: n=494; C:
n=450; D: n=475).



Figure S2. Scatterplot of raw data showing the correlation between (genetic) maternal and

paternal ages of Seychelles warbler offspring at birth (n=1643 genetically assigned offspring,
464 mothers and 440 fathers).





reproductive success (based on assigned offspring) in the Seychelles warbler for all

individuals who survived to one year and had complete lifetime data (n=925). Lifespan and

LRS are integers but jittered to minimise overplotting.







Figure S5. Interaction plot describing the marginally significant interaction between the age of the genetic father and territory quality on female offspring lifetime reproductive success (LRS) in the Seychelles warbler (Table S7). This multi-panel plot is split into three approximately equalsized groups with the lowest third, middle third and highest third of the territory quality data accordingly. Dots show the raw data points. The lines show the estimated relationship between the age of the genetic father and female offspring LRS within each tercile of the territory quality data with the shaded area indicating the 95% confidence interval. The slopes of each line are based on a few points for older (>10 years) males (N=38) such that there are large error estimates around these values. The interaction plot between the age of the dominant male and territory quality on female offspring LRS showed a similar pattern (Figure S7).



Figure S6. Interaction plot describing the marginally significant interaction between the age 919 of the dominant male and territory quality on female offspring lifetime reproductive success 920 (LRS) in the Seychelles warbler (Table S7). Dots show the raw data points, colour coded by 921 territory quality (lighter = lower territory quality, darker = higher territory quality). Lines 922 show model prediction lines and 95% confidence intervals for the relationship of the age of 923 the dominant male and female offspring LRS on three different levels of territory quality 924 (mean territory quality +/- 1 standard deviation from the mean, see legend). The interaction is 925 based on limited data points (N=34) for older males (>10 years) so there are large error 926 estimates around these values. 927



Figure S7. Interaction plot describing the marginally significant interaction between the age of the dominant male and territory quality on female offspring lifetime reproductive success (LRS) in the Seychelles warbler (Table S7). This multi-panel plot is split into three approximately equalsized groups with the lowest third, middle third and highest third of the territory quality data accordingly. Dots show the raw data points. The lines show the estimated relationship between the age of the dominant male and female offspring LRS within each tercile of the territory quality data with the shaded area indicating the 95% confidence interval. The slopes of each line are based on a few points for older (>10 years) males (N=34) such that there are large error estimates around these values.

Table S1. Generalised linear mixed model (GLMM) results investigating associations 935 between the status of the female and provisioning rate (feed counts), in the Seychelles 936 warbler. The model included nest watch data from nests where at least one of the chicks 937 (brood sizes 1-3) in the nest was from a subordinate female (using the pedigree) and included 938 939 nest watches where there were \geq 90% of birds identified. The response variable was the total 940 number of feed counts of the subordinate and dominant female for each nest watch. To account for differences in length of nest watches (usually 60 minutes), an offset of the nest 941 watch duration (logged) was added. The zero-inflated Poisson GLMM was run in glmmTMB 942 v1.0.2.1 (Brooks et al. 2017). Included are the parameter estimates (estimate), their standard 943 errors (SE), and the significance of fixed effects based on a likelihood ratio test (LRT, P) 944 where d.f.=1. Reference levels for the female status was dominant and for the watch type was 945 feeding. There was no significant difference in provisioning rates between the cobreeding 946 subordinate female and dominant female. N=49 nest watch IDs, each with a provisioning 947 count for the dominant female and subordinate female (N=98 rows). Significant fixed effects 948 are highlighted in bold. 949

	Provisioning r	ate		
variables	estimate	SE	LRT	Р
fixed effects				
intercept	-2.167	0.181		
status of female (subordinate female)	0.076	0.094	0.627	0.428
brood size	0.133	0.107	1.537	0.215
watch type (feeding and incubating)	-0.431	0.120	12.124	<0.001
random effects				
NestWatchID/NestID	0.052			
NestID	< 0.001			
BirdID	0.034			
Observer	< 0.001			
Year	0.073			

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Table S2. Cox proportional hazards mixed model results for parental age effects on offspring lifespan, for males and females separately, in the
 Seychelles warbler. Included are the parameter estimates (estimate) and their standard errors (SE), the exponentiated coefficient (exp(est)) which

represents multiplication effects on the hazard and the significance of fixed effects based on a likelihood ratio test (LRT, P) where d.f.=1.

Significance of dropped squared parental age variables, and the environment by parental age interactions are shown in Table S5. Significant fixed
 effects are in bold.

	oring lifes	span				Male offspri n=670	oring lifespan					
variables	estimate	SE	exp(est)	LRT	Ρ		estimate	SE	exp(est)	LRT	Ρ	
fixed effects												
age of mother	0.034	0.017	1.035	3.871		0.049	0.010	0.018	1.010	0.2791		0.597
age of father	0.001	0.018	1.001	0.013		0.909	0.003	0.020	1.003	0.023		0.880
age of dominant male	0.002	0.019	1.002	0.018		0.895	-0.037	0.020	0.963	3.6529		0.056
territory quality	-0.265	0.122	0.767	4.707		0.030	-0.136	0.144	0.872	0.8909		0.345
group size	0.109	0.069	1.115	2.360		0.125	0.140	0.072	1.150	3.651		0.056
helper presence (y)	0.222	0.130	1.248	2.866		0.090	-0.242	0.139	0.785	3.1135		0.078
sibling presence (y)	-0.011	0.151	0.989	0.006		0.941	-0.161	0.143	0.851	1.2858		0.257
random effects												
maternal ID	<0.001						<0.001					
paternal ID	<0.001						<0.001					
dominant male ID	<0.001						<0.001					
birth year	0.020						0.096					

Table S3. Generalised linear mixed model results investigating associations between parental age effects (in years) on female offspring
 reproductive lifespan in the Seychelles warbler. Included are the parameter estimates (estimate), their standard errors (SE), and the significance
 of fixed effects based on a likelihood ratio test (LRT, P) where d.f.=1. n refers to the number of female offspring with full LRS data. Significant
 fixed effects are highlighted in bold.

	Female offs	pring rep	productive	lifespan
	n=441			
variables	estimate	SE	LRT	Р
fixed effects				
intercept	0.722	0.125		
age of mother	-0.467	0.147	9.932	0.002
age of father	0.118	0.155	0.575	0.448
age of dominant male	-0.049	0.156	0.098	0.755
birth year	-1.353	0.233	32.013	<0.001
territory quality	-0.182	0.138	1.721	0.190
group size	-0.277	0.140	3.978	0.046
helper presence (y)	-0.372	0.195	3.590	0.058
sibling presence (y)	-0.182	0.181	1.002	0.317
random effects				
maternal ID	0.135			
paternal ID	0.109			
dominant male ID	0.147			
birth year	0.067			

Table S4. Significance of quadratic parental age effects (in years) and environment by parental age interactions on female offspring and male

offspring lifespan, in the Seychelles warbler using a GLMM. Significance was determined by a likelihood ratio test (LRT, P) where df=1, by
adding the variable to the minimal model (Table 1).

	Female of	fspring	lifespan	Male offspri	ing lifespan	
variables	LRT	df	Р	LRT	df P)
age of mother ²	0.001	1	0.970	2.523	1	0.11
age of father ²	0.483	1	0.487	0.808	1	0.36
age of dominant male ²	1.547	1	0.214	1.347	1	0.24
territory quality*maternal age	0.407	1	0.524	0.290	1	0.59
territory quality*paternal age	0.052	1	0.820	0.295	1	0.58
territory quality*dominant male age	0.121	1	0.728	0.095	1	0.75
group size*maternal age	2.650	1	0.104	1.164	1	0.28
group size*paternal age	0.126	1	0.722	2.082	1	0.14
group size*dominant male age	1.407	1	0.236	1.563	1	0.21
helper*maternal age	0.002	1	0.968	0.057	1	0.81
helper*paternal age	2.163	1	0.141	0.838	1	0.36
helper*dominant male age	0.326	1	0.568	1.122	1	0.29

Table S5. Significance of quadratic parental age effects (in years) and environment by parental age interactions on female offspring and male
 offspring lifespan, in the Seychelles warbler, in the Cox proportional hazards mixed model. Significance was determined by a likelihood ratio
 test (LRT, P) where df=1, by adding the variable to the minimal model (Table S2).

	Female offspring lifespan		Male offspring lifespan	
variables	LRT	Р	LRT	Р
age of mother ²	0.008	0.929	0.105	0.746
age of father ²	1.954	0.162	0.152	0.697
age of dominant male ²	0.016	0.899	0.878	0.349
territory quality*maternal age	0.044	0.833	0.420	0.517
territory quality*paternal age	0.124	0.724	1.842	0.175
territory quality*dominant male age	0.063	0.802	0.147	0.702
group size*maternal age	0.748	0.387	0.032	0.859
group size*paternal age	1.311	0.252	0.011	0.916
group size*dominant male age	0.010	0.919	0.381	0.537
helper*maternal age	0.007	0.934	0.784	0.376
helper*paternal age	2.042	0.153	0.692	0.406
helper*dominant male age	0.076	0.783	0.020	0.889

Table S6. Generalised linear mixed model results investigating between- versus within- maternal and paternal age effects (in years) on offspring lifespan, in each sex separately, in the Seychelles warbler, using the within-subject centering method (van de Pol and Wright 2009). Included are the estimated effects (estimate), standard errors (SEs), and significance of fixed effects based on a likelihood ratio test (LRT; P-value) where df=1. The model investigates whether the within- and between-parental age slopes from Table 2 are significantly different from each other (mean age represents the difference between the slopes and is significant if the slopes are significantly different from each other). The age component is now equivalent to the within-parental age slope from Table 2. Significant fixed effects are highlighted in bold.

	Female of	fspring life	span		Male offsp	oring lifespa	n	
	n=467				n=485			
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	1.284	0.062			1.204	0.070		
mean maternal age	0.412	0.146	7.939	0.005	0.035	0.134	0.069	0.794
maternal age	-0.508	0.121	17.503	<0.001	0.092	0.096	0.920	0.338
mean paternal age	0.161	0.166	0.928	0.335	0.244	0.154	2.503	0.114
paternal age	-0.157	0.134	1.361	0.243	-0.198	0.128	2.437	0.119
mean dominant male age	-0.139	0.169	0.675	0.411	0.056	0.158	0.126	0.723
dominant male age	0.252	0.146	2.974	0.085	0.102	0.131	0.604	0.437
birth year	-0.525	0.112	19.000	<0.001	-0.514	0.138	12.252	<0.001
territory quality	0.022	0.080	0.075	0.784	0.150	0.093	2.645	0.104
group size	-0.104	0.077	1.825	0.177	-0.065	0.081	0.654	0.419
helper presence (y)	-0.153	0.092	2.806	0.094	-0.140	0.101	1.936	0.164
sibling presence (y)	-0.069	0.103	0.446	0.504	0.142	0.100	1.991	0.158
random effects								
maternal ID	0.089				0.088			
paternal ID	0.170				0.081			
dominant male ID	0.004				0.069			
birth year	0.022				0.054			

Table S7. Significance of quadratic parental age effects and environment by parental age interactions on female offspring and male offspring lifetime reproductive success in the Seychelles warbler. Significance was determined by a likelihood ratio test (LRT, P), df = 1, by adding the variable to the minimal model (Table 3). Significant fixed effects are highlighted in bold.

	Female offspring LRS		Male offspring LRS		
variables	LRT	P	LRT	Р	
age of mother ²	0.646	0.421	0.187	0.66	
age of father ²	0.456	0.500	0.078	0.78	
age of dominant male ²	0.480	0.488	0.072	0.78	
territory quality*maternal age	0.626	0.429	0.082	0.77	
territory quality*paternal age	6.085	0.014	0.009	0.92	
territory quality*dominant male age	4.698	0.030	0.010	0.92	
group size*maternal age	0.951	0.330	0.090	0.76	
group size*paternal age	0.046	0.831	0.714	0.39	
group size*dominant male age	1.773	0.183	0.086	0.77	
helper*maternal age	0.530	0.467	0.612	0.43	
helper*paternal age	0.013	0.911	0.772	0.38	
helper*dominant male age	1.996	0.158	0.001	0.97	

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Table S8. Generalised linear mixed model results investigating between- versus within- maternal and paternal age effects (in years) on offspring lifetime reproductive success, in each sex separately, in the Seychelles warbler using the within-subject centering method (van de Pol and Wright 2009). Included are the estimated effects (estimate), standard errors (SEs), and significance of fixed effects based on a likelihood ratio test (LRT; P-value) where df=1. The model investigates whether the within and between parental age slopes from Table 4 are significantly different from each other (mean age represents the difference between the slopes and is significant if the slopes are significantly different from each other). The age component is now equivalent to the within parental age slope from Table 4. Significant fixed effects are highlighted in bold.

	Female offspri	ng LRS			Male offspring	g LRS		
	n=441 females	-			n=466 males			
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	0.513	0.131			0.345	0.143		
mean maternal age	0.537	0.289	3.344	0.067	0.303	0.274	1.219	0.270
maternal age	-0.835	0.261	9.640	0.002	0.139	0.219	0.402	0.526
mean paternal age	-0.245	0.303	0.659	0.417	0.533	0.345	2.449	0.118
paternal age	0.291	0.278	1.103	0.294	-0.071	0.277	0.067	0.796
mean dominant male age	-0.164	0.322	0.260	0.610	-0.459	0.379	1.493	0.222
dominant male age	0.025	0.301	0.007	0.935	0.324	0.307	1.129	0.288
birth year	-1.076	0.227	23.569	<0.001	-1.119	0.272	16.840	<0.001
territory quality	-0.136	0.142	0.927	0.336	0.219	0.199	1.211	0.271
group size	0.004	0.151	0.001	0.981	-0.097	0.175	0.304	0.581
helper presence (y)	-0.611	0.201	9.336	0.002	-0.279	0.222	1.643	0.200
sibling presence (y)	-0.109	0.189	0.333	0.564	0.070	0.204	0.118	0.732
random effects								
maternal ID	0.100				0.153			
paternal ID	0.090				0.234			
dominant male ID	0.046				<0.001			
birth year	0.064				0.085			

Table S9. Generalised linear mixed model results investigating parental age effects (in years) on offspring annual reproductive success

1001 (measured as the number of offspring in the pedigree who survived to 1 year), in each sex separately, in the Seychelles warbler. Included are the

1002 estimated effects (estimate), standard errors (SEs), and significance of fixed effects based on a likelihood ratio test (LRT; P) where df=1.

1003 Significant fixed effects are highlighted in bold.

	Female offspri	ing ARS			Male offspring	ARS		
	n=2287 observ	vations of 4	ations of 441 IDs n=2341 observations of 466 IDs			66 IDs		
variables	estimate	SE	LRT	Р	estimate	SE	LRT	Р
fixed effects								
intercept	-1.912	0.118			-1.961	0.108		
age	4.983	0.399	-	-	4.870	0.356	-	-
age ²	-4.935	0.435	211.530	<0.001	-4.043	0.331	213.390	<0.001
lifespan	0.392	0.134	8.407	0.004	0.350	0.132	7.017	0.008
age of mother	-0.039	0.107	0.130	0.718	0.117	0.096	1.461	0.227
age of father	-0.015	0.119	0.016	0.900	0.128	0.114	1.256	0.263
age of dominant male	-0.139	0.117	1.416	0.234	0.161	0.113	2.032	0.154
birth year	-0.036	0.207	0.030	0.863	-0.207	0.201	1.062	0.303
territory quality	-0.107	0.114	0.878	0.349	0.129	0.126	1.035	0.309
group size	0.090	0.111	0.634	0.426	-0.193	0.115	2.878	0.090
helper presence (y)	-0.342	0.153	5.109	0.024	0.010	0.147	0.005	0.944
sibling presence (y)	0.043	0.151	0.082	0.775	-0.058	0.136	0.183	0.669
random effects								
ID	0.070				0.066			
Year	0.153				0.120			

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Table S10. Generalised mixed model results investigating between- versus within- maternal and paternal age effects on female offspring lifespan
 in the Seychelles warbler, using the within-subject centering method (van de Pol and Wright 2009). This model excludes the outlier seen in
 Figure 2, but model estimates are very similar with this outlier included (Table 2). Included are the estimated effects (estimate), standard errors
 (SEs), and significance of fixed effects based on a likelihood ratio test (LRT; P-value) where df=1. The model investigates within-parental age
 effects (deviation from the mean age of the parent) and between-parental age effects (mean age for each parent). Significant fixed effects are in
 bold.

	Female offspring lifespan				
	n=466				
variables	estimate	SE	LRT	Р	
fixed effects					
intercept	1.281	0.062			
between-maternal age	-0.104	0.090	1.309	0.253	
within-maternal age	-0.470	0.126	14.046	<0.001	
between-paternal age	0.012	0.110	0.011	0.916	
within-paternal age	-0.162	0.134	1.460	0.227	
between-dominant male age	0.111	0.094	1.383	0.240	
within-dominant male age	0.240	0.147	2.680	0.102	
birth year	-0.522	0.113	18.460	<0.001	
territory quality	0.021	0.081	0.067	0.796	
group size	-0.100	0.077	1.704	0.192	
helper presence (y)	-0.149	0.092	2.678	0.102	
sibling presence (y)	-0.063	0.103	0.375	0.540	
random effects					
maternal ID	0.086				
paternal ID	0.168				
dominant male ID	0.007				
birth year	0.024				

Table S11. Generalised mixed model results investigating between- versus within- maternal and paternal age effects on female offspring lifetime reproductive success in the Seychelles warbler using the within-subject centering method (van de Pol and Wright 2009). This model excludes the outlier seen in Figure 4A, but model estimates are very similar with this outlier included (Table 4). Included are the estimated effects (estimate), standard errors (SEs), and significance of fixed effects based on a likelihood ratio test (LRT; P-value) where df=1. The model investigates within-parental age effects (deviation from the mean age of the parent) and between-parental age effects (mean age for each parent). Significant fixed effects are in bold.

	Female offspring LRS				
	n=440				
variables	estimate	SE	LRT	Р	
fixed effects					
intercept	0.509	0.131			
between-maternal age	-0.296	0.168	3.079	0.079	
within-maternal age	-0.847	0.298	7.695	0.006	
between-paternal age	0.044	0.178	0.061	0.806	
within-paternal age	0.295	0.281	1.109	0.292	
between-dominant male age	-0.141	0.171	0.689	0.407	
within-dominant male age	0.027	0.303	0.008	0.928	
birth year	-1.079	0.228	23.354	<0.001	
territory quality	-0.135	0.142	0.898	0.343	
group size	0.003	0.154	<0.001	0.986	
helper presence (y)	-0.613	0.202	9.347	0.002	
sibling presence (y)	-0.110	0.190	0.338	0.561	
random effects					
maternal ID	0.101				
paternal ID	0.094				
dominant male ID	0.049				