

1 **Title:** Statistical inference for seed mortality and germination with seed bank experiments

2 **Running head:** Statistical inference for seed banks

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7 **Open research statement:** The R scripts, JAGS model code, and Maxima code to reproduce the
8 simulations and analyses in this study have been archived on Zenodo: [https://doi.org/10.5281/
9 zenodo.7317528](https://doi.org/10.5281/zenodo.7317528). The data collected as part of the literature synthesis in this study are available in
10 the repository on Zenodo.

11 **1 Abstract**

12 Plant population ecologists regularly study soil seed banks with seed bag burial and seed addi-
13 tion experiments. These experiments contribute crucial data to demographic models, but we lack
14 standard methods to analyze them. Here, we propose statistical models to estimate seed mortality
15 and germination with observations from these experiments. We develop these models following
16 principles of event history analysis, and analyze their identifiability and statistical properties by
17 algebraic methods and simulation. We demonstrate that seed bag burial, but not seed addition ex-
18 periments, can be used to make inferences about age-dependent mortality and germination. When
19 mortality and germination do not change with seed age, both experiments produce unbiased esti-
20 mates but seed bag burial experiments are more precise. However, seed mortality and germination
21 estimates may be inaccurate when the statistical model that is fit makes incorrect assumptions about
22 the age-dependence of mortality and germination. The statistical models and simulations that we
23 present can be adopted and modified by plant population ecologists to strengthen inferences about
24 seed mortality and germination in the soil seed bank.

25 **Keywords:** seed banks, demography, parameter estimation, identifiability, uncertainty

26 **2 Introduction**

27 Soil seed banks are a crucial part of plant life-history strategies that depend on long-lived stages
28 to persist in variable environments. At the population level, a persistent soil seed bank can buffer
29 populations from temporal variability in reproductive success (Evans et al., 2007), and produce
30 age structure that increases generation time and affects the population growth rate (Kalisz and
31 McPeck, 1992). However, it can be difficult to incorporate seed banks into empirical population
32 models (Menges, 2000; Doak et al., 2002; Nguyen et al., 2019) because seed fates are partially
33 or completely unobservable processes (Rees and Long, 1993). Individual seeds enter the seed
34 bank from seed rain, and eventually leave through death or germination (Simpson et al., 1989).
35 Seeds experience mortality by being consumed or destroyed by predators or pathogens, or through
36 physiological death (Baker, 1989). In the field, seed mortality cannot be directly observed and,
37 because seeds that germinate must have both survived and germinated, seed mortality complicates

38 inferences about germination.

39 Population ecologists studying seed banks are often interested in understanding how seed mor-
40 tality and germination influence population dynamics. Seed mortality and germination can be
41 measured using experiments (e.g., Kalisz 1991) and natural variability in seed rain and seedling
42 emergence (e.g., Evans et al. 2010). In particular, experiments are often used to study seed fates
43 and estimate both seed mortality and germination for population models. Two kinds of experi-
44 ments are most common. Briefly, seed bag burial experiments involve burying seeds in mesh bags
45 and recovering them to count seeds that died or germinated. Seed addition experiments consist of
46 adding seeds to plots and returning to census emerged seedlings. Ideally, these experiments would
47 be used to obtain accurate estimates for age-dependent seed mortality and germination that are
48 associated with quantified uncertainty (Doak et al., 2002). However, there is no standard statistical
49 approach for estimating seed mortality and germination from field experiments, and even observa-
50 tions from the same kind of experiment are often analyzed in disparate ways. For instance, three
51 recent studies that used seed bag burial experiments each analyzed the observations differently: by
52 regressing seeds in year $t + 1$ on seeds in year t (Kurkjian et al., 2017), fitting an exponential curve
53 (Lommen et al., 2018), or estimating the proportion of surviving seeds (Tanner et al., 2021).

54 Decisions about how to estimate seed mortality and germination influence whether and how the
55 seed bank is represented in structured population models. Because the seed bank cannot be directly
56 observed, these choices are often made with limited information. Recent reviews indicate that over
57 a third of published plant population matrix models exclude seed banks without justification (Doak
58 et al., 2002; Nguyen et al., 2019). Omitting the seed bank or inaccurately estimating seed mortality
59 or germination can bias estimates for population growth rate, particularly when aboveground rates
60 exhibit high levels of temporal variability (Doak et al., 2002; Nguyen et al., 2019). Age-dependent
61 seed mortality and germination contribute to population age structure, so the decision to represent
62 the seed bank as unstructured or age-structured can also affect population growth rates (Kalisz,
63 1991; Rees and Long, 1993; Doak et al., 2002). In addition, the precision of vital rate estimates,
64 including seed rates, influences uncertainty in estimates of population dynamics (e.g., Paniw et al.

65 2017; Nguyen et al. 2019). Seeds are hard to study and relatively little is often known about them,
66 so authors may omit uncertainty in their estimates of seed related rates and in turn underestimate
67 uncertainty in population growth rate and extinction probability.

68 Although existing studies address issues related to how seed banks are represented in structured
69 population models, there is no work that examines statistical models for observations from the
70 experiments that are commonly used to study seed banks. We identify three key unanswered
71 questions about seed bag burial and seed addition experiments: (i) When can each experiment be
72 used to obtain estimates for constant versus age-dependent seed mortality and germination? (ii)
73 What is the accuracy and precision of estimates from each experiment? (iii) How are estimates
74 affected by misrepresenting the age-dependence of seed mortality and germination in statistical
75 models? We answer these questions by describing statistical models for observations from seed bag
76 burial and seed addition experiments and addressing model *identifiability*, the *statistical properties*
77 of estimates for seed mortality and germination, and the consequences of model *misspecification*.

78 In seed bag burial and seed addition experiments, observations of surviving seeds and seedlings
79 reflect seed fates. We define likelihoods that link these observations to estimates of seed mortality
80 and germination, and analyze the identifiability of the models under different assumptions about
81 the age-dependence of mortality and germination. Informally, a statistical model is identifiable if
82 it is possible to estimate the parameters in the model from a given set of data. For the models
83 that we analyze, the crux of the issue is that the seed bank experiments produce different observa-
84 tions—seeds and seedlings for seed bag burial experiments, but only seedlings for seed addition
85 experiments. To determine if each experiment generates observations that can be used to esti-
86 mate seed mortality and germination, we analyze the identifiability of statistical models. Once
87 we determine which statistical models are identifiable for particular experimental observations, we
88 compare the accuracy and precision of seed mortality and germination estimates from seed bag
89 burial and seed addition experiments. Finally, we ‘stress-test’ the models by assessing the conse-
90 quences of fitting misspecified models. In current practice, studies may fit models that simplify the
91 age-dependence of mortality or germination by assuming constant mortality or germination (e.g.,

92 Leimu and Lehtilä 2006; Burns et al. 2013). In this case, researchers may want to be parsimonious
93 and reduce the number of parameters in a statistical model. Alternatively, studies may fit models
94 with as many age-dependent parameters as observations permit (e.g. Eckhart et al. 2011; Bricker
95 and Maron 2012). In this case, researchers may want to parameterize an age-structured seed bank
96 for a population model, or may simply want to be cautious and avoid over-simplifying the age
97 structure of the seed bank. These constraints are typically applied without evaluating alternative
98 model structures, and we investigate how such assumptions influence the accuracy and precision
99 of estimates of mortality and germination.

100 **3 Developing the statistical models**

101 In the following sections, we characterize seed bag burial and seed addition experiments, and
102 the observations they produce, by way of idealized examples (Fig. 1). We apply the principles
103 of event history analysis to develop a deterministic model for seed mortality and germination that
104 describes seed fates in seed bank experiments. We then link the observations and deterministic
105 processes with probability statements to define a statistical model for observations from each ex-
106 periment. Throughout, we present general versions of the model to accommodate mortality and
107 germination rates that depend on seed age. At select points, we use the specific case in which
108 mortality and germination rates do not depend on seed age to interpret the general model.

109 We make several choices about how to develop the statistical models that are influenced by how
110 population ecologists use seed bag burial and seed addition experiments. In a literature synthesis,
111 we identified 57 studies conducted from 1991-present that used a total of 69 experiments to pa-
112 rameterize matrix or integral projection models (**Appendix: Literature synthesis**). We used these
113 studies to inform how we constructed our idealized experiments, with the goal of representing the
114 essential attributes of each kind of experiment. The majority of studies (94.7%) built population
115 models with separate seed mortality and germination parameters, and we consider how to estimate
116 both parameters as well. The majority of studies (84.2%) described discrete relationships between
117 seed age and fate, which is also how we represent the relationship in the models we build.

118 3.1 Observations

119 We assume that we want to characterize seed mortality and germination for a plant species with
120 a soil seed bank and discrete germination opportunities. For simplicity, we do not compare sites,
121 treatments, or species. The seeds are too small to be followed individually, so we conduct experi-
122 ments with unmarked cohorts of seeds. We consider two possibilities: seed bag burial experiments
123 (Fig. 1A-B) and seed addition experiments (Fig. 1C-D). We work with idealized versions of these
124 experiments because our study develops and analyzes statistical models.

125 In seed bag burial experiments, we add seeds and soil to mesh bags before burying them in
126 the field (0 months in Fig. 1A). Researchers bury seeds in various enclosures (e.g., cages in Kalisz
127 1991, or mesh bags in Quintana-Ascencio et al. 1998) but to be concise we always refer to bags.
128 Bags are recovered from the field after a certain time. Here, we collect bags after germination
129 so that we count intact, ungerminated seeds and germinants (filled circles in Fig. 1A). Sampling
130 tends to be destructive, particularly if intact seeds are tested for viability using a method such as
131 tetrazolium staining. As a result, these studies typically retrieve different subsets of bags for seed
132 and germinant counts at different times (Fig. 1A).

133 We identify each bag by an ID, index i , and the time that it was recovered, index j (columns
134 1-2 in Fig. 1B). We also record time as a variable, t_{ij} (column 3 in Fig. 1B). Each bag has three
135 counts: the number of seeds added to the mesh bags at the start of the experiment, n_{ij} , the number
136 of intact, ungerminated seeds, and the number of germinants, $y_{g,ij}$ (columns 4-6 in Fig. 1B). Here,
137 we assume that all intact seeds are viable (but we discuss combining field experiments and lab
138 viability assays in **Discussion: Extensions**). Finally, we calculate the number of seeds surviving
139 to sampling as the sum of intact, ungerminated seeds and germinants – we assume this is both
140 the number of survivors, y_{ij} , and the number of seeds surviving to just before germination, $n_{g,ij}$
141 (columns 7-8 in Fig. 1B).

142 In seed addition experiments, we lightly bury or sprinkle seeds on the soil surface (0 months
143 in Fig. 1C). Seeds are buried in plots where we do not expect a substantial seed bank, or in pots
144 or trays with seed-free soil. We might also include control plots without seed addition to account

145 for natural seed rain. We survey the plots for seedlings after germination (filled circles in Fig. 1C).
146 Typically, it is not possible to recover intact seeds from the soil but because seedling counts are
147 non-destructive, we can resurvey plots (Fig. 1C).

148 We identify each plot by an ID, index i , and record the time it was surveyed, index j (columns
149 1-2 in Fig. 1D). We also record time as a variable, t_{ij} (column 3 in Fig. 1C). Each plot and survey
150 time has two counts: the number of seeds added to the plot at the start of the experiment, n_{ij} , and
151 the number of seedlings, $y_{g,ij}$ (columns 4 & 6 in Fig. 1C).

152 Studies that track natural variability in seed rain, n_{ij} , and seedling emergence, $y_{g,ij}$ over time
153 can also produce similar data as seed addition experiments (e.g., Kauffman and Maron 2006; El-
154 derd and Miller 2016). Seed rain is comparable to the number of seeds added to an experimental
155 plot, and seedling surveys are similar to counting seedlings in an experiment. Natural variability
156 can complement or provide an alternative to experimental manipulations, especially for species
157 with limited dispersal or discrete recruitment pulses in which it is possible to link seed rain and
158 seedling emergence. In this manuscript, we describe our statistical framework with reference to
159 experimental observations, but it could also be used to estimate seed mortality and germination
160 from field observations.

161 **3.2 Deterministic model for seed fates**

162 The fate of seeds in the seed bank can be characterized using methods from event history anal-
163 ysis (also called survival or failure time analysis; reviewed in Fox 2001; Landes et al. 2020). By
164 focusing on a seed leaving the seed bank as the event of interest, we can characterize the distri-
165 bution of times at which the event occurs using a set of key functions (Klein and Moeschberger,
166 2003). A survival function describes the probability that a seed remains in the seed bank until time
167 t . The survival function is the term for the probability of the event occurring after time t ; the term
168 applies whether or not the event of interest is death. A probability density function describes the
169 probability that the seed leaves the seed bank at time t . Finally, a hazard function describes the
170 tendency that a seed remaining in the seed bank at time t leaves the seed bank at the next instant
171 in time. The probability density function defines the unconditional probability of events, while the

172 hazard function is associated with the conditional propensity for the event among individuals who
173 have not yet experienced the event (Fox 2001, p. 245). We illustrate the relationship between these
174 functions in **Appendix: Hazards determine the age-structure of the seed bank**.

175 We use these principles to describe how seed loss from the seed bank (the event of interest)
176 depends on mortality and germination. We define hazard functions for each fate. The hazard
177 function for mortality, h_m , is the risk that a seed remaining in the seed bank leaves the seed bank
178 through mortality the next instant. The hazard function for germination, h_g , is the risk that a seed
179 remaining in the seed bank leaves the seed bank through germination the next instant. The hazards
180 thus summarize the “instantaneous risk” (Landes et al., 2020) of mortality or germination. In this
181 paper, we assume that mortality precedes germination, but the principles we describe are flexible
182 and specific equations should be reformulated to correspond to the biology of the study system.

183 We combine the mortality and germination hazards to describe a survival function for the ex-
184 pected probability that seeds remain in the seed bank up to a given time:

$$185 \quad S(t) = \prod_{t_j \leq t} (1 - h_m(t_j)) \times (1 - h_g(t_j)). \quad (1)$$

186
187 Equation 1 is the product of discrete survival functions associated with mortality, $S_m(t_j) = \prod_{t_j \leq t} 1 -$
188 $h_m(t_j)$, and germination, $S_g(t_j) = \prod_{t_j \leq t} 1 - h_g(t_j)$. If the hazards are on an annual timescale, S_m is
189 the cumulative product of the complement of the mortality probability, up to the number of years
190 t_j that seeds have been in the soil. Similarly, S_g is the cumulative product of the complement of
191 the germination probability, up to the number of germination opportunities that seeds have experi-
192 enced. In terms of the hazards, $h_m(1)$ is a seed’s propensity for mortality in the first year and $h_g(1)$
193 is the propensity for germination of a seed that does not die during the first year. The seeds that
194 remain in the seed bank past one year do not die with propensity $1 - h_m(1)$ and do not germinate
195 with propensity $1 - h_g(1)$.

196 To use the survival function (Equation 1) in a statistical model, we specify the hazards in terms
197 of probabilities. The mortality hazard, $h_m(t_j)$, is the probability of mortality during each time in-
198 terval j , $p_{m,j}$. Specifically, it is the conditional probability of mortality for seeds that remain in
199 the seed bank. We describe seeds remaining in the seed bank *after* the period in which they expe-

200 rience mortality, but *before* the germination opportunity. We assume that after this time interval,
 201 seeds have a discrete opportunity to germinate. The germination hazard, $h_g(t_j)$, is the conditional
 202 probability of germination at each opportunity, $p_{g,j}$, for a seed that remains in the seed bank up to
 203 just before germination.

204 With these hazards, the mortality component is defined by $\prod_{j=1}^J 1 - p_{m,j}$. The germination
 205 component is defined by $\prod_{j=1}^J (1 - p_{g,j-1})^{I(j>1)}$, where $I(x)$ is an indicator function equal to 1
 206 if the inequality is true, and 0 if it is not (Metcalf et al., 2009). We use the indicator function
 207 because at the first time j , seeds have not yet experienced a germination opportunity. After the first
 208 germination opportunity, the ‘germination history’ is defined by the product of past germination
 209 opportunities. The product of the mortality and germination components describe the probability
 210 that seeds remain in the seed bank after j time intervals (e.g., years) as

$$211 \quad f(\mathbf{p}_g, \mathbf{p}_m) = \prod_{j=1}^J \overbrace{(1 - p_{m,j})}^{\text{survival function}} \times \overbrace{(1 - p_{g,j-1})^{I(j>1)}}^{\text{germination history}}. \quad (2)$$

212

213 The choice of how to represent mortality and germination makes explicit our assumptions about
 214 how those processes operate. The most simple version of the model in Equation 2 is one in which
 215 the hazards are constant; neither mortality nor germination probability change with seed age. In
 216 this case, $p_{m,1} = p_{m,2} = \dots = p_{m,J}$ and $p_{g,1} = p_{g,2} = \dots = p_{g,J}$. Mortality and germination are
 217 each described by a single parameter, p_m and p_g .

218 **3.3 Likelihood functions for observations from seed bag burial and seed addition experi-** 219 **ments**

220 To estimate seed mortality and germination, we use probability statements to connect the ob-
 221 servations from field experiments to the deterministic models. We describe likelihood functions
 222 for observations from seed bag burial (Fig. 1A-B) and seed addition (Fig. 1C-D) experiments. To
 223 illustrate our approach, we assume that mortality and germination do not depend on seed age.
 224 The general structure of the likelihood remains similar when we relax the assumption of constant
 225 hazards for mortality or germination (Table 1).

226 For the seed bag burial experiment, we construct one likelihood for the observations of ger-

227 minants and another likelihood for the observations of surviving seeds. First, we use the ob-
 228 servations of germinants to describe a model for the probability of germination, p_g . We as-
 229 sume that the number of seeds that germinate, $y_{g,ij}$, is a binomial sample from the number of
 230 seeds surviving to just before germination, $n_{g,ij}$. Recall that the number of surviving seeds is the
 231 sum of germinants and ungerminated, intact seeds. We estimate the probability of germination,
 232 p_g , for a seed that survives up to just before germination. The likelihood is then $L(p_g|\mathbf{y}_g) =$
 233 $\prod_{i=1}^I [\prod_{j=1}^{J_i} \text{binomial}(y_{g,ij}|n_{g,ij}, p_g)]$.

234 Next, we use the observations of surviving seeds to describe a survival function for the prod-
 235 uct of germination and mortality hazards. We assume that the number of seeds that survive to a
 236 given time is a binomial sample from the number of seeds that start the experiment in each bag:
 237 $\text{binomial}(y_{ij}|n_{ij}, f(\dots))$. The number of surviving seeds is the sum of germinants and ungermi-
 238 nated, intact seeds. The deterministic model, $f(\dots)$, is the product of the germination history and
 239 the survival function for mortality, and describes the probability of not germinating and not dying
 240 up to the time j . For the case in which mortality and germination do not depend on seed age,
 241 $f(p_m, p_g) = \prod_{k=1}^j (1 - p_m)(1 - p_g)^{I(j>1)}$. The joint likelihood for observations of germinants and
 242 surviving seeds is

$$243 \quad f(p_m, p_g) = \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)} \quad (3)$$

$$244 \quad L(p_m, p_g|y_g, y) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} [\text{binomial}(y_{g,ij}|n_{g,ij}, p_g) \text{binomial}(y_{ij}|n_{ij}, f(p_m, p_g))] \right]$$

245 In seed bag burial experiments, bags are destructively sampled so the indices for bag ID, i , and
 246 recovery time, j , are redundant (Fig. 1A). We write the likelihood function so that the index for
 247 the time the bag was recovered, j , is nested within the index for bag, i . We adopt this notation to
 248 avoid unobserved combinations of bag ID and recovery time (e.g., any bag i at a time j when the
 249 bag was not recovered). Using this notation also makes explicit the parallel with the likelihood for
 250 observations from seed addition experiments, for which there are multiple observation times j per
 251 bag i (Fig. 1B).

252 For the seed addition experiment, we construct a likelihood for the observations of seedlings.

253 We assume that the number of seedlings is a binomial sample from the number of seeds that start
 254 the experiment: $\text{binomial}(y_{g,ij}|n_{ij}, f(\dots))$. The number of seedlings is the product of mortality
 255 and germination. We describe the combination of those processes with a deterministic model,
 256 $f(\dots)$, that modifies Equation 2 to include germination. Each observation takes place at the time
 257 of germination, rather than after, so that $f(p_m, p_g) = p_g \times \prod_{k=1}^j (1 - p_m)(1 - p_g)^{I(j>1)}$. To account
 258 for germination, the function now includes the probability of germination, p_g , in addition to the
 259 survival function for mortality and the germination history. The likelihood for observations of
 260 seedlings is

$$\begin{aligned}
 f(p_m, p_g) &= p_g \times \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)} \\
 L(p_m, p_g|y_g) &= \prod_{i=1}^I \left[\prod_{j=1}^{J_i} [\text{binomial}(y_{g,ij}|n_{g,ij}, f(p_m, p_g))] \right].
 \end{aligned}
 \tag{4}$$

263 4 Methods

264 We now conduct a comprehensive analysis of statistical models for seed bag burial and seed
 265 addition experiments. First, we determine whether the models can be used to estimate different
 266 combinations of constant (C) or age-dependent (A) seed mortality and germination. Population
 267 models that incorporate a seed bank typically parameterize seed mortality and germination with
 268 one of the following combinations of mortality/germination: C/C (e.g., Kurkjian et al. 2017), A/C
 269 (e.g., Yates and Ladd 2010), C/A (e.g., Elderd and Miller 2016), and A/A (e.g., Kalisz 1991). We
 270 thus analyze models for the following cases:

- 271 1. Constant mortality/constant germination (C/C): Mortality, p_m , and germination, p_g , hazards
 272 are the same for all seed ages.
- 273 2. Age-dependent mortality/constant germination (A/C): The mortality hazard is a function of
 274 seed age, $p_{m,j}$, while the germination hazard is the same for all seed ages, p_g .
- 275 3. Constant mortality/age-dependent germination (C/A): The mortality hazard is the same for
 276 all seed ages, p_m , while the germination hazard is a function of seed age, $p_{g,j}$.
- 277 4. Age-dependent mortality/age-dependent germination (A/A): Both mortality, $p_{m,j}$, and ger-
 278 mination, $p_{g,j}$, hazards are functions of seed age.

279 For each of these four cases, we study the identifiability of models for seed bag burial and seed ad-
280 dition experiments to determine when each can be used to estimate seed mortality and germination.
281 To directly compare the statistical properties of estimates for seed mortality and germination from
282 seed bag burial and seed addition experiments, we fit a model with constant mortality and constant
283 germination (C/C) to observations from a seed bank with constant mortality and constant germina-
284 tion (C/C). Finally, we study the consequences of model misspecification on parameter estimates.
285 We stress-test the models with two cases in which the true age structure of the seed bank and the
286 age-dependence of parameters in the statistical model are mismatched. First, we fit a model with
287 age-dependent mortality and constant germination (A/C) to observations from a seed bank with
288 constant mortality and germination (C/C). Second, we we fit a model with constant mortality and
289 constant germination (C/C) to observations from a seed bank with age-dependent mortality and
290 constant germination (A/C).

291 **4.1 Identifiability analysis by the symbolic method**

292 To determine when seed bag burial and seed addition experiments can be used to estimate con-
293 stant or age-dependent seed mortality and germination, we analyze the identifiability of statistical
294 models for the experiments. We study if parameters can be estimated in terms of the structure of
295 the likelihood (‘intrinsic identifiability’) (Cole 2020). Intrinsic identifiability refers to cases where
296 parameters in a model can be uniquely estimated. For example, models will not be identifiable if
297 different combinations of mortality and germination have the same likelihood for a set of observa-
298 tions. If the model is not identifiable, there are no unique maximum likelihood estimates regardless
299 of the quantity of data that is available.

300 To analyze the identifiability of statistical models for different combinations of experiment,
301 hazard, and length of the experiment, we use an algebraic approach called the symbolic method
302 (Catchpole and Morgan 1997; Cole et al. 2010; Cole 2020). With this method, we focus on general
303 issues of experimental design and model structure rather than on specific datasets. We determine
304 the intrinsic identifiability of statistical models for all combinations of experiment (seed bag burial
305 vs. seed addition), hazards (C/C, A/C, C/A, A/A), and length of experiment (1, 2 or 3 years). All

306 the likelihoods that we analyze are shown in Table 1. To apply the symbolic method, we summarize
307 each model by a vector that completely determines the model (an ‘exhaustive summary’). The
308 exhaustive summary is simply the likelihood associated with each observation. The exhaustive
309 summary is subsequently differentiated with respect to all of the constituent parameters to form
310 a ‘derivative matrix’ (the transpose of the Jacobian). The model is identifiable if the rank of the
311 derivative matrix is equal to the number of parameters in the model; the model is not identifiable if
312 the rank of the derivative matrix is less than the number of parameters. We implement these steps
313 using the computer algebra software Maxima (Maxima, 2014); for detailed methods and scripts,
314 see **Appendix: Identifiability analysis**.

315 **4.2 Simulation experiments**

316 To compare the statistical properties of seed bag burial and seed addition experiments, and
317 study the effect of model misspecification, we conduct numerical experiments in which we fit
318 models to simulated data. To simulate data with the structure of seed bag burial and seed addition
319 experiments (Fig. 1), we use the likelihoods corresponding to those observations (Table 1). In
320 practice, we use mortality and germination hazards to calculate the expected probability of a seed
321 remaining in the soil at the end of each year, and its subsequent probability of germinating. We use
322 the expected probability of remaining in the soil to draw a binomial sample of seeds from the initial
323 number of seeds in the bag. We use the probability of germination to draw a binomial sample of
324 germinants from the seeds remaining in the bag. To simulate data with the structure of the seed
325 addition experiment, we retain only the observations of seedlings.

326 Both maximum likelihood and Bayesian methods would be appropriate to fit the models asso-
327 ciated with seed bag burial and seed addition experiments. We chose to fit Bayesian models to the
328 simulated observations because we can readily estimate the parameters in the joint likelihood. All
329 parameters in our models are probabilities with support $[0, 1]$ on which we place $\text{beta}(1, 1)$ priors;
330 this is equivalent to a uniform prior. Fig. S4 shows the directed acyclic graphs corresponding to
331 the joint and posterior distributions for the models. Parameters and details of simulations are given
332 in the sections that follow. We wrote simulations and analyzed model output in R version 3.6.2 (R

333 Core Team, 2019). We wrote all models and sampled posterior distributions using JAGS 4.10 with
334 **rjags** (Plummer et al., 2019). For each fit, we ran 3 chains with 3,000 iterations for adaptation,
335 5,000 for burn-in, and 5,000 for sampling. For computational efficiency, we thinned the chains
336 and kept every 10th iteration. We used the **MCMCvis** package to work with model output, check
337 chains for convergence, and recover posterior distributions (Youngflesh et al., 2021).

338 **4.2.1 Statistical properties of seed bag burial and seed addition experiments**

339 To compare the statistical properties of estimates from identifiable models, we used a simula-
340 tion experiment in which we fit a model with constant mortality and constant germination (C/C)
341 to observations from a seed bank with constant mortality and constant germination (C/C). We gen-
342 erated data from a 3-year experiment with $n = (5, 10, 15, 20, 25, 30)$ bags or plots each year. Each
343 bag or plot started the experiment with 100 seeds. For each number of bags or plots, we simulated
344 250 replicate datasets for four combinations of ‘true’ mortality and germination: low mortality/low
345 germination (0.1, 0.1), low mortality/high germination (0.1, 0.5), high mortality/low germination
346 (0.5, 0.1), and high mortality/high germination (0.5, 0.5). We fit each simulated dataset with two
347 models; one for a seed bag burial experiment and one for a seed addition experiment.

348 To quantify the bias of estimates, we calculated the difference between the posterior modes
349 and ‘true’ parameters for the probability of mortality or germination. Estimates are unbiased when
350 the difference is 0. To quantify the uncertainty of parameter estimates, we calculated the width
351 of the 95% credible interval. For each set of ‘true’ parameters and number of bags or plots, we
352 estimated the mean difference and width, and quantified 95% confidence intervals for each with
353 a t distribution (Pappalardo et al., 2020). To estimate the coverage of the 95% credible intervals,
354 we calculated the proportion of credible intervals that contain the ‘true’ parameter value. Ideally,
355 a 95% credible interval would contain the ‘true’ parameter value 95% of the time. We obtained
356 confidence intervals for coverage with the Wilson method in the **binom** package (Pappalardo et al.,
357 2020), and calculated root mean squared error as a measure of the combined effect of bias and
358 uncertainty.

359 **4.2.2 Consequences of model misspecification**

360 To study the consequences of model misspecification, we studied two cases with a mismatch
361 between the true age structure of the seed bank and the age-dependence of the statistical model.
362 For these analyses, we were interested in the interaction between the data-generating process and
363 the statistical model. We thus use simulations with a large number of bags or plots ($n = 30$).

364 First, we fit a model with age-dependent mortality and constant germination (A/C) to observa-
365 tions from a seed bank with constant mortality and constant germination (C/C). In this analysis, we
366 examined the effect of fitting a model that has more parameters than the process that generated the
367 observations (i.e., the model is overspecified). We used the observations that we simulated from
368 a seed bank with constant mortality and germination in the previous section. We then fit a model
369 with age-dependent mortality and constant germination, the A/C model, to these observations. The
370 model has four parameters: one mortality parameter each for one-, two-, and three-year old seeds,
371 $p_{m,1}$, $p_{m,2}$, and $p_{m,3}$, and one parameter for germination, p_g . For all parameters, we quantified
372 bias, uncertainty, coverage, and root mean squared error.

373 Second, we fit a model with constant mortality and constant germination (C/C) to observations
374 from a seed bank with age-dependent mortality and constant germination (A/C). In this analysis,
375 we studied the effect of fitting a model that has fewer parameters than the process that generated the
376 data (i.e., the model is underspecified). We generated data from a 3-year experiment with $n = 30$
377 bags or plots each year, and each bag or plot started the experiment with 100 seeds. The ‘true’
378 probability of mortality increased over time, so that $p_{m,1} = 0.1$, $p_{m,2} = 0.2$, and $p_{m,3} = 0.3$. The
379 germination rate in all years was constant. To examine the influence of germination on statistical
380 properties, we varied p_g from 0.1 to 0.5 across simulations. For each true germination probability,
381 we simulated 250 replicate datasets. As before, we fit two models to each simulated dataset; one
382 for a seed bag burial experiment and one for a seed addition experiment. Even though we only
383 estimated one parameter for the probability of mortality, we compared properties of the estimate
384 to the age-dependent probability of mortality in each of the three years. For all parameters, we
385 quantified bias, uncertainty, coverage, and root mean squared error.

386 **5 Results**

387 **5.1 Identifiability analysis by the symbolic method**

388 All models for observations from seed bag burial experiments exhibit a deficiency of 0, indi-
389 cating that the models are identifiable (Table 2). In all cases we consider, the models for seed bag
390 burial experiments can be used to estimate parameters for seed mortality and germination. Mod-
391 els for observations from seed addition experiments only show a deficiency of 0 when mortality
392 and germination rates are assumed to be constant, and when more than one year of observations
393 is available (Table 2). In all other cases, models have a deficiency greater than 0, indicating that
394 the models are not identifiable. Identifiability is directly related to the structure of a model's like-
395 lihood (Table 1). Models for a seed bag burial experiment contain the product of likelihoods for
396 observations of intact seeds and germinants, each of which provides separate information on seed
397 mortality and germination (Fig. S4A). However, models for a seed addition experiment only con-
398 sist of a single likelihood for observations of seedlings, in which seed mortality and germination
399 always appear as a product (Fig. S4B). In this case, the model is identifiable only when there are
400 as many, or more, years of seedling observations as there are parameters in the model.

401 **5.2 Statistical properties of seed bag burial and seed addition experiments**

402 The C/C models fit to observations from the seed bag burial and seed addition experiments are
403 identifiable when there is more than one year of data (Table 2); here, we analyze simulated data
404 for 3-year experiments. Both experiments produce unbiased estimates of mortality (Fig. 2A-D)
405 and germination (Fig. 2I-L) with large numbers of bags or plots. With small numbers of bags
406 or plots, seed addition experiments are more likely to produce biased estimates for mortality (e.g.,
407 Fig. 2A, C). Estimates from seed addition experiments display greater uncertainty in all simulations
408 (Fig. 2E-H, M-P). The difference in uncertainty of estimates between experiments depends on the
409 true probability of mortality and germination. Seed mortality estimates show seven to nine times
410 more uncertainty for seed addition experiments when germination is low, but roughly twice as
411 much uncertainty when germination is high (Fig. 2E-H). Germination estimates from seed addition
412 experiments displayed a similar pattern, though the overall difference in uncertainty was smaller.

413 Estimates from seed addition experiments are 2.3 times as uncertain when when mortality is low,
414 and 1.8 times as uncertain when germination is high. Fig. 2M-P).

415 For both experiments, coverage is $\sim 95\%$ (Fig. S8A-D, I-L), and root-mean squared error de-
416 creases with the number of bags or plots (Fig. S8E-H, M-P). For seed addition experiments, esti-
417 mates of seed mortality show the greatest error when germination is low (Fig. S8E, G). The joint
418 posterior distribution for mortality and germination is more positively correlated when germina-
419 tion is low; put another way, the model structure makes it challenging to determine whether small
420 numbers of observed seedlings are due to low germination or high mortality.

421 **5.3 Consequences of model misspecification**

422 Fitting the A/C model to observations from a seed bank with constant mortality and germina-
423 tion has a strong influence on parameter estimates for seed addition experiments. In particular,
424 estimates of first year seed mortality from seed addition experiments are biased and associated
425 with high uncertainty when the true probability of germination is low, and when both true prob-
426 abilities of mortality and germination are high (Fig. S9A-H). In addition, parameter estimates for
427 seed mortality in subsequent years also exhibit higher uncertainty (Fig. S9E-H). In contrast, param-
428 eter estimates for seed bag burial experiments are unbiased (Fig. S9A-D). However, uncertainty of
429 second and third year mortality estimates is greater when the mismatched A/C model, rather than
430 the correct C/C model, is fit to the data (compare filled points in Fig. 2E-H and Fig. S9E-H). In-
431 creased bias and uncertainty are especially pronounced in seed addition experiments because the
432 likelihood produces strong positive correlations in the joint posterior distribution of first year seed
433 mortality and germination (Fig. S6L). Especially for low germination rates, this correlation can
434 introduce bias and uncertainty into parameter estimates. For seed bag burial experiments, corre-
435 lations between second and third year seed mortality (Fig. S6I) similarly contribute to increased
436 uncertainty in those estimates.

437 We fit the C/C model to observations from simulations in which the probability of seed mortal-
438 ity increases over time ($p_{m,1} = 0.1, p_{m,2} = 0.2, p_{m,3} = 0.3$), for a range of true germination prob-
439 abilities. Note that when we analyze this case, we are comparing a single estimated parameter to

440 three true mortality parameters. For both seed bag burial and seed addition experiments, seed age
441 and the true probability of germination interact to determine the direction and magnitude of bias
442 in mortality (Fig. 3A-C). Both experiments progress from overestimating to underestimating mor-
443 tality. When the true probability of germination is low, mortality is overestimated more strongly
444 (greater positive bias) for young than old seeds, but mortality is underestimated to a lesser extent
445 (less negative bias) for old seeds. In the first year, seed bag burial experiments exhibit less bias
446 than seed addition experiments; this pattern reverses by the third year. The relationship between
447 true germination probability and parameter uncertainty is the same across all mortality parame-
448 ters (Fig. 3E-G). Uncertainty decreases with germination rate for seed addition experiments, but
449 increases slightly with germination rate for seed bag burial experiments. For most of the scenarios
450 we considered, low accuracy of parameter estimates translates into low coverage (Fig. 3I-K). The
451 single exception is that year two mortality exhibited nominal coverage (roughly 95%) at intermedi-
452 ate germination rates. The root-mean squared error (RMSE) for mortality is largely determined by
453 the bias of estimates; estimates with a smaller absolute bias also show smaller RMSE (Fig. 3M-O).

454 The ‘true’ probability of germination does not depend on seed age in the simulation, but ger-
455 mination estimates are slightly biased for both seed bag burial and seed addition experiments
456 (Fig. 3D). For both experiments, the bias of germination estimates also increases with the true
457 probability of germination. Although the absolute magnitude of bias is smaller than for mortality
458 estimates, germination is overestimated from 6-19% across all scenarios. The greater uncertainty
459 of germination estimates from seed addition experiments (Fig. 3H) translates into higher coverage
460 (Fig. 3L). However, coverage for both experiments is far below the desired level of 95%. As with
461 mortality, RMSE is largely a function of the bias of parameter estimates (Fig. 3P).

462 **6 Discussion**

463 We develop and analyze statistical models for observations from field experiments commonly
464 used to study the soil seed bank. We present the first systematic evaluation and comparison of
465 inferences made with statistical models for seed bag burial and seed addition experiments. We
466 show that seed bag burial experiments can separately estimate mortality and germination even if

467 one, or both, are age-dependent. For seed addition experiments, we demonstrate that seed mortality
468 and germination are only identifiable if both mortality and germination do not change with seed
469 age and with more than one year of observations. In all other cases, it is impossible to separately
470 estimate mortality and germination.

471 To compare the statistical properties of estimates from seed bag burial and seed addition ex-
472 periments, we focus on identifiable models with constant mortality and constant germination. We
473 place model identifiability upstream of the statistical properties of parameter estimates because the
474 latter issues are contingent on having reliable statistical models. Estimates from both experiments
475 are unbiased as the number of bags or plots increases. However, estimates from seed bag burial
476 experiments are more precise for all parameter combinations that we consider.

477 In practice, researchers may fit models that describe seed mortality and germination with more
478 or fewer parameters than necessary. We thus evaluate two scenarios in which we fit the wrong
479 model to observations. In one case, we fit a model with age-dependent mortality and constant
480 germination rates to observations produced by constant mortality and constant germination. Fitting
481 the more complicated model makes the parameter estimates more sensitive to the structure of the
482 model. The effect is strongest for seed addition experiments when germination rates are low, and
483 introduces bias and uncertainty into estimates of mortality. In a second case, we fit a model with
484 constant mortality and germination rates to observations produced by age-dependent mortality and
485 constant germination. The bias of mortality estimates changes over time, and is exacerbated by
486 increased precision at higher germination rates. Germination estimates are also biased, though to
487 a lesser extent.

488 **6.1 Recommendations for practitioners**

489 We demonstrate how seed bag burial or seed addition experiments can be used to estimate
490 seed mortality and germination. From a statistical perspective, seed bag burial experiments have
491 several useful properties. When estimating constant mortality and germination, seed bag burial
492 experiments will produce estimates that are more accurate and precise for a given number of bags
493 or plots.

494 We suggest that the best way to adapt our general recommendations is to simulate data and
495 fit models to those simulations. Practitioners likely know much about many of the key parts of
496 a seed bank experiment. How many seeds could be collected and used for an experiment, how
497 many replicates are logistically feasible, and for how long would the experiment run? With these
498 pieces in hand, it is possible to use plausible values for seed mortality and germination rates to
499 simulate observations. It will not be possible to know the ‘true’ values or their age-dependence, but
500 simulations could explore likely scenarios (e.g., constant vs. increasing mortality). Fitting models
501 to these simulations would make it possible to compare the statistical properties of estimates from
502 seed bag burial versus seed addition experiments. To facilitate this process, we include the code for
503 our study (<https://doi.org/10.5281/zenodo.7317528>); this includes R code to simulate observations,
504 the JAGS code for the models, and the R code to fit the models to observations.

505 Our analysis can also help guide parameter estimation if observations have already been col-
506 lected. Lack of identifiability creates issues for both frequentist and Bayesian statistical methods,
507 which we illustrate in detail in **Appendix: Implications of identifiability for model fitting**. No
508 amount of clever modeling can estimate parameters when they are intrinsically not identifiable.
509 Observations from seed bag burial experiments give you the flexibility to fit models with constant
510 or age-dependent mortality and germination. With observations from seed addition experiments,
511 only models with constant mortality and germination are identifiable. Our analysis of model mis-
512 specification indicates that fitting a model with more age-dependence than necessary introduces
513 noise into estimates from seed addition experiments but not seed bag burial experiments. In con-
514 trast, fitting a model with constant mortality or germination to observations from an age-structured
515 seed bank produces biased estimates for both types of experiments.

516 Ultimately, the impact of bias or imprecision in estimates of seed mortality or germination on
517 population growth rate depends on the sensitivity of population growth rate to those vital rates. The
518 models and analyses we present will be most relevant to researchers working with plant popula-
519 tions in which aboveground vital rates exhibit high temporal variability because these populations
520 are likely sensitive to transitions in the seed bank (Doak et al., 2002; Nguyen et al., 2019). Consid-

521 ering the broader context of the plant life history can help population ecologists determine which
522 fieldwork and modeling approaches are sufficiently accurate and precise for their study system.

523 **6.2 Extensions**

524 Existing studies have used simulations and post-hoc comparisons to explore the consequences
525 of age structure in the seed bank, emphasize how estimates of seed rates interact with temporal
526 variability in aboveground success, and describe the effect of underestimating parameter uncer-
527 tainty (Doak et al., 2002; Paniw et al., 2017; Nguyen et al., 2019). However, these methods do not
528 provide an intuitive way to use observations to test assumptions about seed bank structure and as-
529 sociated parameter uncertainty. For example, the methods do not allow for model checks or model
530 selection, both of which could be used to ask whether the fitted model is consistent with obser-
531 vations. Because accuracy and precision of estimates for seed mortality and germination interact
532 with information about other components of the life cycle, it seems crucial to evaluate the model
533 used to estimate seed mortality or germination separately from the population model.

534 Although estimates from seed addition experiments will be unbiased when mortality and ger-
535 mination do not change with seed age, researchers will generally not know whether this is the case
536 before conducting an experiment. In our simulations, we can assess the accuracy of parameter es-
537 timates obtained with these models because we picked the values used to generate the data. While
538 we lack this luxury for empirical datasets, it is possible that standard model checking (e.g., Conn
539 et al. 2018) and model selection (e.g., Hooten and Hobbs 2015; Tredennick et al. 2021) methods
540 could help determine whether the fitted model is consistent with the process that generated the
541 data. For example, it may be possible to use model selection to determine whether a model with
542 constant mortality and germination is a good fit to data from a seed addition experiment. However,
543 further research is required to determine the effectiveness of such approaches.

544 The relationship between seed age and fate can also be described with continuous functions
545 such as exponential models (e.g., Lommen et al. 2018). Population ecologists have not typically
546 used this approach to analyze observations from seed bank experiments (**Appendix: Literature**
547 **synthesis**; 12.3% fit models with continuous functions). However, continuous descriptions of seed

548 fate could reduce the number of model parameters in cases where estimating the age-dependence of
549 mortality or germination at many ages is of interest (Fox, 2001; Landes et al., 2020). Rees and Long
550 (1993) studied seed bank dynamics by fitting continuous models for recruitment to observations
551 of seedlings from a seed addition experiment. The authors used these models to demonstrate that
552 recruitment is affected by the age-dependence of seed mortality and germination and that seed
553 banks do not, as a rule, exhibit exponential decay (Rees and Long, 1993). However, they did not
554 separately estimate seed mortality and germination, which are processes that population ecologists
555 are often interested in obtaining separate information on (**Appendix: Literature synthesis**; 3.5%
556 of studies built population models in which mortality and germination were combined).

557 Our study focuses on statistical issues associated with estimating seed mortality and germi-
558 nation from field experiments, and suggests that seed bag burial experiments present statistical
559 advantages for jointly estimating seed mortality and germination. While we do not evaluate how
560 experimental design affects observations from seed bag burial or seed addition experiments, it is
561 crucial to collect observations that reflect natural levels of seed mortality and germination from
562 the soil seed bank. For example, high seed densities in mesh bags may promote transmission of
563 pathogenic fungi and increase seed mortality (Van Mourik et al., 2005), and seed burial depth can
564 influence mortality and germination rates (Dille et al., 2017). However, researchers can aim to
565 control how much burying seeds in bags alters natural biological processes. Fungal infection can
566 be minimized by mixing seeds with sand or soil to reduce seed-to-seed contact and decrease seed
567 densities (Van Mourik et al., 2005). Appropriate burial depths could be evaluated by conducting
568 pilot experiments (Hernandez et al., 2020) or burying seeds at multiple depths (Philippi, 1993).

569 Ultimately, how researchers choose to study the seed bank will likely depend on multiple fac-
570 tors including issues related to experimental design, statistical properties of estimates, and the
571 research question being addressed. A strategy for balancing these competing considerations may
572 be to combine information from different types of experiments (e.g., Liu et al. 2005; Bricker and
573 Maron 2012), or from experiments and field surveys (e.g., García 2003; Adams et al. 2005). Ap-
574 proaches that formally integrate data from multiple sources are an active area of research in ecol-

575 ogy (Zipkin et al., 2021), and applying these ideas to seed bank studies could help researchers
576 effectively use existing datasets and optimize the collection of future data.

577 **6.3 Limitations**

578 Event history analysis is developed for and appropriately applied to individual data (Zens and
579 Peart, 2003; Landes et al., 2020), and the models we describe would be completely appropriate if
580 applied to observations of individual seeds. Yet seeds of many plant species are too small for indi-
581 viduals to be tracked in the field. When examining aggregate data—from cohorts, or populations—
582 heterogeneity between subpopulations and change in hazards over time can confound whether pat-
583 terns are the result of changes to hazards or to population structure (Rees and Long, 1993; Zens
584 and Peart, 2003). Our approach is not intended to assess changes to the hazards for individual
585 seeds (unless individual-level data are available) but rather a framework for consistent inferences
586 about seed mortality and germination.

587 To focus on the commonalities between seed bag burial and seed addition experiments, we
588 describe stereotyped versions of each. Not all experiments in the literature exactly follow the
589 schematic we describe; some seed bag burial experiments count intact seeds and estimate germi-
590 nation in another way (e.g., Lommen et al. 2018), or count only seeds at certain times, but both
591 seeds and germinants at other times (e.g., Eckhart et al. 2011). Individual analyses will inevitably
592 have to be tailored to specific data. We sought to explicitly describe the assumptions underlying
593 our statistical models so that they could be readily modified. Investigators will naturally construct
594 models that are appropriate to their system and aims.

595 **6.4 Conclusion**

596 Observations from seed bag burial and seed addition experiments are hard-won data, but statis-
597 tical models for observations from these experiments have received little attention to-date. Study-
598 ing these models can help plant population ecologists make the most of existing and future data
599 by identifying potential models to fit, the statistical properties of parameter estimates, and poten-
600 tial bias introduced by making assumptions about age-dependence of mortality and germination.
601 Our analysis contributes to efforts to make richer inferences from the trove of demographic data

602 collected by plant population ecologists.

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611 **8 Author contributions**

612 GS and MAG conceived of the ideas in the study. GS developed the statistical model, analyzed
613 identifiability & simulations, and wrote the manuscript with input from MAG.

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Table 1: Likelihoods of models for observations from seed bag burial and seed addition experiments.

Model		
Mortality	Germination	Likelihood
SEED BAG BURIAL EXPERIMENT		
$C(p_m)$	$C(p_g)$	$f(p_m, p_g) = \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g, y) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} [\text{binomial}(y_{g,ij} n_{g,ij}, p_g) \text{binomial}(y_{ij} n_{ij}, f(p_m, p_g))] \right]$
$A(p_{m,j})$	$C(p_g)$	$f(p_{m,j}, p_g) = \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g, y) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} [\text{binomial}(y_{g,ij} n_{g,ij}, p_g) \text{binomial}(y_{ij} n_{ij}, f(p_{m,j}, p_g))] \right]$
$C(p_m)$	$A(p_{g,j})$	$f(p_m, p_{g,j}) = \prod_{k=1}^j (1 - p_m) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_{g,j} y_g, y) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} [\text{binomial}(y_{g,ij} n_{g,ij}, p_{g,j}) \text{binomial}(y_{ij} n_{ij}, f(p_m, p_{g,j}))] \right]$
$A(p_{m,j})$	$A(p_{g,j})$	$f(p_{m,j}, p_{g,j}) = \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_{g,j} y_g, y) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} [\text{binomial}(y_{g,ij} n_{g,ij}, p_{g,j}) \text{binomial}(y_{ij} n_{ij}, f(p_{m,j}, p_{g,j}))] \right]$
SEED ADDITION EXPERIMENT		
$C(p_m)$	$C(p_g)$	$f(p_m, p_g) = p_g \times \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} \text{binomial}(y_{g,ij} n_{g,ij}, f(p_m, p_g)) \right]$
$A(p_{m,j})$	$C(p_g)$	$f(p_{m,j}, p_g) = p_g \times \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} \text{binomial}(y_{g,ij} n_{g,ij}, f(p_{m,j}, p_g)) \right]$
$C(p_m)$	$A(p_{g,j})$	$f(p_m, p_{g,j}) = p_{g,j} \times \prod_{k=1}^j (1 - p_m) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_{g,j} y_g) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} \text{binomial}(y_{g,ij} n_{g,ij}, f(p_m, p_{g,j})) \right]$
$A(p_{m,j})$	$A(p_{g,j})$	$f(p_{m,j}, p_{g,j}) = p_{g,j} \times \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_{g,j} y_g) = \prod_{i=1}^I \left[\prod_{j=1}^{J_i} \text{binomial}(y_{g,ij} n_{g,ij}, f(p_{m,j}, p_{g,j})) \right]$

¹ In columns 1 and 2, C is a constant hazard and A is an age-dependent hazard.

² In all likelihoods, $I(x)$ is an indicator function equal to 1 if the inequality is true, and 0 if it is not. As discussed in the main text, the indicator function identifies whether or not seeds have yet experienced a germination opportunity; at the first time point j , they have not.

Table 2: Analysis of intrinsic identifiability for models with different assumptions about whether germination and mortality are constant or age-dependent. Each row corresponds to a model in which the germination component is defined in column one and the mortality component is defined in column two. For each model, the columns show the results of the intrinsic identifiability analysis for 1, 2, or 3 years of observations. The analysis identifies the deficiency of the model for a given set of assumptions about the germination and mortality components. The deficiency is calculated as in Cole (2020): the number of parameters in the model minus the rank of the derivative matrix, the latter calculated by the symbolic method. Models with a deficiency of 0 are identifiable; models with a deficiency greater than 0 are not identifiable.

Model		Deficiency		
Mortality component	Germination component	1 year	2 years	3 years
SEED BAG BURIAL EXPERIMENT				
Constant (p_m)	Constant (p_g)	0	0	0
Age-dependent ($p_{m,j}$)	Constant (p_g)	0	0	0
Constant (p_m)	Age-dependent ($p_{g,j}$)	0	0	0
Age-dependent ($p_{m,j}$)	Age-dependent ($p_{g,j}$)	0	0	0
SEED ADDITION EXPERIMENT				
Constant (p_m)	Constant (p_g)	1	0	0
Age-dependent ($p_{m,j}$)	Constant (p_g)	1	1	1
Constant (p_m)	Age-dependent ($p_{g,j}$)	1	1	1
Age-dependent ($p_{m,j}$)	Age-dependent ($p_{g,j}$)	1	2	3

728 **10 Figure captions**

729 **Figure 1.** (A) Schematic of a seed bag burial experiment. Each bag in the experiment is represented
730 by a single line from when the bag is buried at month 0 to when the bag is dug up for sampling
731 (filled circles). The data are organized with indices for bag and sampling time. (B) Data from the
732 seed bag burial experiment. Each row corresponds to a bag and sampling time. (C) Schematic of
733 a seed addition experiment. Each plot in the experiment is represented by a single line from when
734 seeds are added to the plot at month 0 to when plots are censused for seedlings (filled circles). The
735 data are organized with indices for plot and time. (D) Data from the seed addition experiment.
736 Each row corresponds to a plot and sampling time.

737 **Figure 2.** Results of simulation experiment in which we generated observations with constant
738 mortality and germination, and fit a model with constant mortality and germination. (A-D) Bias for
739 estimates of mortality, p_m , for different combinations of true mortality and germination probability.
740 (E-H) Width of the 95% credible interval for p_m . (I-L) Bias for estimates of germination, p_g , for
741 different combinations of true mortality and germination probability. (M-P) Width of the 95%
742 credible interval for p_g . In all panels, error bars represent the 95% confidence interval based on a
743 t distribution.

744 **Figure 3.** Results of simulation experiment in which we generated observations with age-dependent
745 mortality and germination, and fit a model with constant mortality and germination. The true prob-
746 ability of mortality increases over time: $p_{m,1} = 0.1$, $p_{m,2} = 0.2$, and $p_{m,3} = 0.3$. From left to right,
747 columns are analyses of mortality for ages 1, 2, and 3, and germination. Each panel shows the
748 statistical properties for parameters from simulations with true probabilities of germination from
749 0.1 to 0.5. The number of bags or plots is always $n = 30$. (A-D) Bias for estimates of mortality
750 and germination. Error bars represent the 95% confidence interval based on a t distribution. (E-H)
751 Width of the 95% credible interval for mortality and germination. Error bars represent the 95%
752 confidence interval based on a t distribution. (I-L) Coverage for mortality and germination. Er-
753 ror bars represent the 95% confidence interval calculated using the Wilson method for binomial
754 proportions. (M-P) Root mean squared error for mortality and germination.

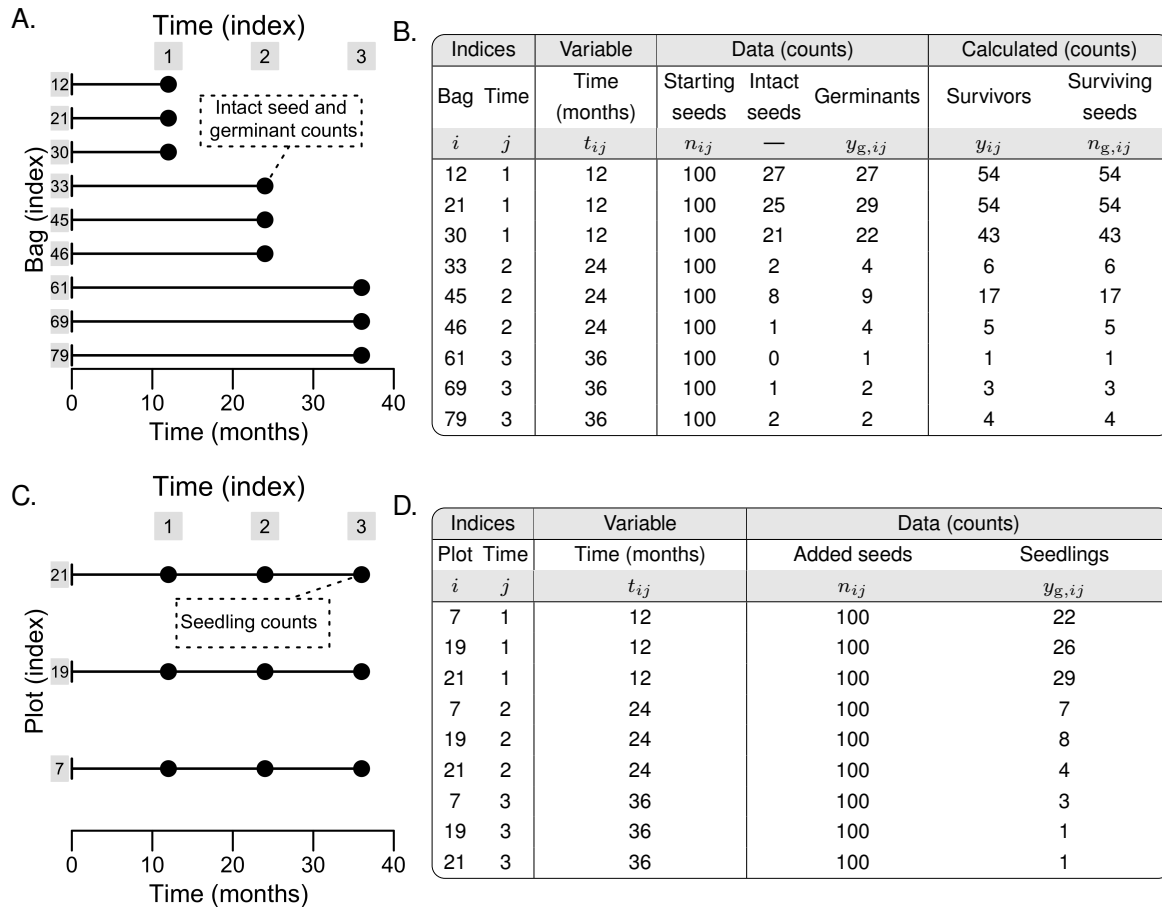


Figure 1: (A) Schematic of a seed bag burial experiment. Each bag in the experiment is represented by a single line from when the bag is buried at month 0 to when the bag is dug up for sampling (filled circles). The data are organized with indices for bag and sampling time. (B) Data from the seed bag burial experiment. Each row corresponds to a bag and sampling time. (C) Schematic of a seed addition experiment. Each plot in the experiment is represented by a single line from when seeds are added to the plot at month 0 to when plots are censused for seedlings (filled circles). The data are organized with indices for plot and time. (D) Data from the seed addition experiment. Each row corresponds to a plot and sampling time.

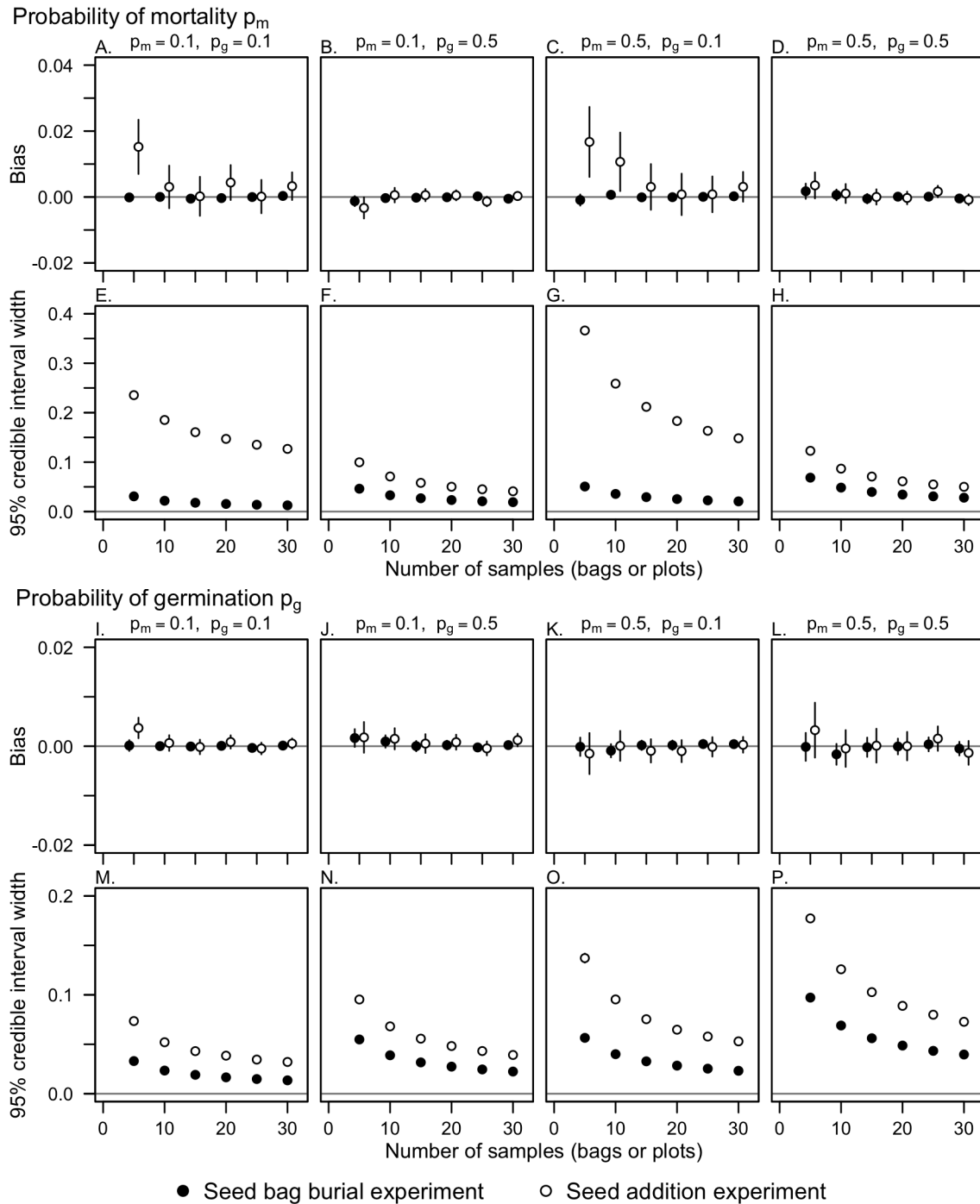


Figure 2: Results of simulation experiment in which we generated observations with constant mortality and germination, and fit a model with constant mortality and constant germination parameters. (A-D) Bias for estimates of mortality probability, p_m , for different combinations of true mortality and germination probability. (E-H) Width of the 95% credible interval for p_m . (I-L) Bias for estimates of germination probability, p_g , for different combinations of true mortality and germination probability. (M-P) Width of the 95% credible interval for p_g . In all panels, error bars represent the 95% confidence interval based on a t distribution.

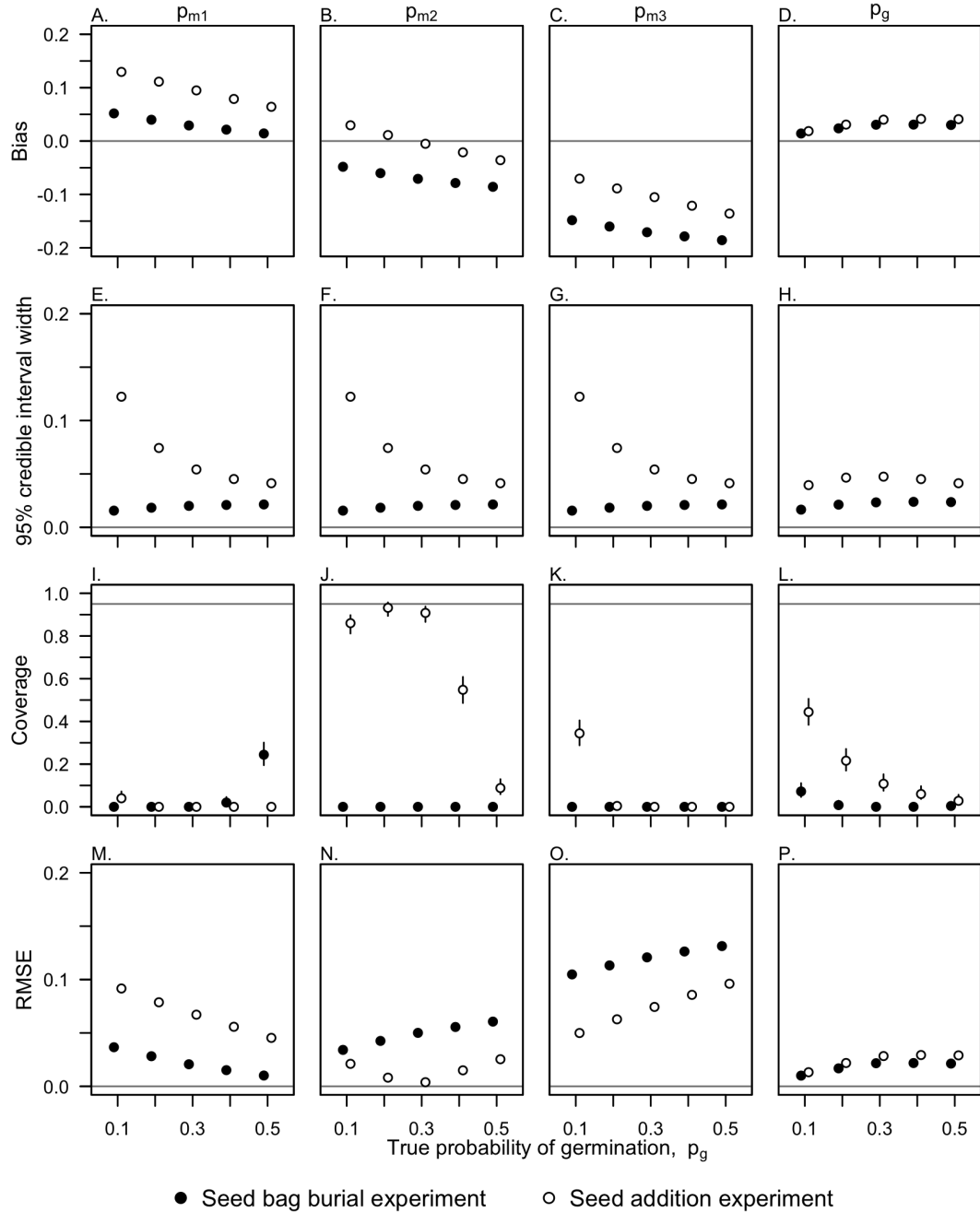


Figure 3: Results of simulation experiment in which we generated observations with age-dependent mortality and germination, and fit a model with constant mortality and germination. The true probability of mortality increases over time: $p_{m,1} = 0.1$, $p_{m,2} = 0.2$, and $p_{m,3} = 0.3$. From left to right, columns are analyses of mortality for ages 1, 2, and 3, and germination. Each panel shows the statistical properties for parameters from simulations with true probabilities of germination from 0.1 to 0.5. The number of bags or plots is always $n = 30$. (A-D) Bias for estimates of mortality and germination. Error bars represent the 95% confidence interval based on a t distribution. (E-H) Width of the 95% credible interval for mortality and germination. Error bars represent the 95% confidence interval based on a t distribution. (I-L) Coverage for mortality and germination. Error bars represent the 95% confidence interval calculated using the Wilson method for binomial proportions. (M-P) Root mean squared error for mortality and germination.