

1 **Title:** Statistical inference for seed mortality and germination with seed bank experiments

2 **Running head:** Statistical inference for seed banks

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6 Last updated: December 21, 2021

7 **Open research statement:** The R scripts, JAGS model code, and Maxima code to reproduce the
8 simulations and analyses in this study have been archived on Zenodo: [https://zenodo.org/record/
9 5794709](https://zenodo.org/record/5794709). The data collected as part of the literature synthesis in this study are available in the
10 repository on Zenodo.

11 **1 Abstract**

12 Plant population ecologists regularly study soil seed banks with seed bag burial and seed addi-
13 tion experiments. These experiments contribute crucial data to demographic models, but we lack
14 standard methods to analyze them. Here, we propose statistical models to estimate seed mortality
15 and germination with observations from these experiments. We develop these models following
16 principles of event history analysis, and analyze their identifiability and statistical properties by
17 algebraic methods and simulation. We demonstrate that seed bag burial, but not seed addition ex-
18 periments, can be used to make inferences about age-dependent mortality and germination. When
19 mortality and germination do not change with seed age, both experiments produce unbiased esti-
20 mates but seed bag burial experiments are more precise. However, seed mortality and germination
21 estimates may be inaccurate when the statistical model that is fit makes incorrect assumptions about
22 the age-dependence of mortality and germination. The statistical models and simulations that we
23 present can be adopted and modified by plant population ecologists to strengthen inferences about
24 seed mortality and germination in the soil seed bank.

25 **Keywords:** seed banks, demography, parameter estimation, identifiability, uncertainty

26 **2 Introduction**

27 Soil seed banks are a crucial part of plant life-history strategies that depend on long-lived stages
28 to persist in variable environments. At the population level, a persistent soil seed bank can buffer
29 populations from temporal variability in reproductive success (Evans et al., 2007), and produce
30 age structure that increases generation time and affects the population growth rate (Kalisz and
31 McPeck, 1992). However, it can be difficult to incorporate seed banks into empirical population
32 models (Menges, 2000; Doak et al., 2002; Nguyen et al., 2019) because seed fates are partially
33 or completely unobservable processes (Rees and Long, 1993). Individual seeds enter the seed
34 bank from seed rain, and eventually leave through death or germination (Simpson et al., 1989).
35 Seeds experience mortality by being consumed or destroyed by predators or pathogens, or through
36 physiological death (Baker, 1989). In the field, seed mortality cannot be directly observed and,
37 because seeds that germinate must have both survived and germinated, seed mortality complicates

38 inferences about germination.

39 Population ecologists measure seed mortality and germination with a range of methods includ-
40 ing experiments (e.g. Kalisz 1991) and natural variability in seed rain and seedling emergence (e.g.
41 Evans et al. 2010). Seed bag burial and seed addition experiments are particularly common: in a
42 literature synthesis, we identified 69 studies from 1991-present that use them to parameterize ma-
43 trix or integral projection models (**Appendix: Literature synthesis**). Ideally, these experiments
44 would be used to obtain accurate estimates for age-dependent seed mortality and germination that
45 are associated with quantified uncertainty (Doak et al., 2002). But it often remains unclear how to
46 translate observations from these experiments into parameter estimates with the desired qualities.
47 Even observations from the same type of experiment are often analyzed in disparate ways. For
48 instance, three recent studies that used seed bag burial experiments each analyzed the observations
49 differently: by regressing seeds in year $t + 1$ on seeds in year t (Kurkjian et al., 2017), fitting an
50 exponential curve (Lommen et al., 2018), or estimating the proportion of surviving seeds (Tanner
51 et al., 2021).

52 In addition to deciding how to estimate seed mortality and germination, population ecologists
53 also have to choose how to represent the seed bank in population models. Because the seed bank
54 cannot be directly observed, these choices are often made with limited information. Studies have
55 evaluated the consequences of incorrectly omitting the seed bank, not including age structure, and
56 obtaining inaccurate or imprecise estimates for seed mortality or germination. Omitting or inac-
57 curately estimating seed mortality or germination can bias estimates for population growth rate,
58 particularly when aboveground rates exhibit high levels of temporal variability (Doak et al., 2002;
59 Nguyen et al., 2019). Age-dependent seed mortality and germination contribute to population age
60 structure, so the decision to represent the seed bank as unstructured or age-structured can affect
61 population growth rates (Kalisz, 1991; Rees and Long, 1993; Doak et al., 2002). In addition, the
62 precision of vital rate estimates, including seed rates, also affects uncertainty in estimates of popu-
63 lation dynamics (e.g. Paniw et al. 2017; Nguyen et al. 2019). Seeds are hard to study and relatively
64 little is often known about them, so authors may omit uncertainty in their estimates of seed related

65 rates and in turn underestimate uncertainty in population growth rate and extinction probability.
66 Challenges associated with estimating seed rates likely contributes to over a third of published
67 matrix models excluding seed banks without justification (Doak et al., 2002; Nguyen et al., 2019).

68 Seed bag burial or seed addition experiments are frequently used to study seed banks in the
69 field. But population ecologists lack a comprehensive statistical approach to using these exper-
70 iments for inferences about seed mortality and germination. We identify three key unanswered
71 questions about seed bag burial and seed addition experiments: (i) When can each experiment be
72 used to obtain estimates for constant versus age-dependent seed mortality and germination? (ii)
73 What is the accuracy and precision of estimates from each experiment? (iii) How are estimates
74 affected by simplifying the age-dependence of seed mortality and germination in statistical mod-
75 els? We answer these questions by describing statistical models for observations from seed bag
76 burial and seed addition experiments and addressing model *identifiability*, the *statistical properties*
77 of estimates for seed mortality and germination, and the consequences of model *misspecification*.

78 First, we describe seed bag burial and seed addition experiments in order to illustrate how ob-
79 servations from these experiments reflect seed fates. We define likelihoods that link observations
80 of surviving seeds and seedlings to estimates of seed mortality and germination. Second, we de-
81 termine the identifiability of the models for each experiment under different assumptions about the
82 age-dependence of mortality and germination. A statistical model is identifiable if it is possible to
83 estimate the parameters in the model from a given set of data. Here, the crux of the issue is that
84 the seed bank experiments produce different observations—seeds and seedlings for seed bag burial
85 experiments, but only seedlings for seed addition experiments. The experiments generate inher-
86 ently different information about seed mortality and germination. To determine if an experiment
87 generates observations that can be used to estimate the desired parameters, we can analyze the
88 identifiability of statistical models. In ecology, similar questions of identifiability arise when an-
89 alyzing observations from presence-only versus presence-absence (Royle et al., 2012; Hastie and
90 Fithian, 2013) or single-visit versus double-visit (Lele et al., 2012; Knappe and Korner-Nievergelt,
91 2015) sampling protocols (reviewed in Stoult 2020). We place model identifiability upstream of

92 the statistical properties of parameter estimates because the latter issues are contingent on having
93 reliable statistical models.

94 Once we determine which statistical models are identifiable for particular experimental obser-
95 vations, we can compare the accuracy and precision of seed mortality and germination estimates
96 from seed bag burial and seed addition experiments. Finally, we assess the consequences of fitting
97 a misspecified model. By this we mean fitting a model that constrains seed mortality or germina-
98 tion to be constant when the observations are produced by age-dependent mortality or germination.
99 In current practice, studies sometimes constrain mortality or germination to reduce the number of
100 parameters in a model (e.g. Leimu and Lehtilä 2006; Burns et al. 2013). The constraint is typically
101 applied without evaluating alternatives, but mortality and germination are likely to depend on seed
102 age in many cases (Lonsdale, 1988; Rees and Long, 1993). But it is not clear if, and how, such
103 constraints change the accuracy of estimates. We demonstrate that incorrect assumptions about the
104 age-dependence of seed rates can produce biased estimates.

105 **3 Developing the statistical models**

106 In the following sections, we characterize seed bag burial and seed addition experiments, and
107 the observations they produce, by way of hypothetical examples (Figure 1). We apply the princi-
108 ples of event history analysis to develop a deterministic model for seed mortality and germination
109 that describes seed fates in seed bank experiments. We then link the observations and determinis-
110 tic processes with probability statements to define a statistical model for observations from each
111 experiment. Throughout, we present general versions of the model to accommodate mortality and
112 germination rates that depend on seed age. At select points, we use the specific case in which
113 mortality and germination rates do not depend on seed age to interpret the general model.

114 **3.1 Observations**

115 We assume that we want to characterize seed mortality and germination for a plant species with
116 a soil seed bank and discrete germination opportunities. For simplicity, we do not compare sites,
117 treatments, or species. The seeds are too small to be followed individually, so we conduct experi-
118 ments with unmarked cohorts of seeds. We consider two possibilities: seed bag burial experiments

119 (Figure 1A-B) and seed addition experiments (Figure 1C-D).

120 In seed bag burial experiments, we add seeds and soil to mesh bags before burying them in
121 the field (0 months in Fig. 1A). Researchers bury seeds in various enclosures (e.g. cages in Kalisz
122 1991, or mesh bags in Quintana-Ascencio et al. 1998) but to be concise we always refer to bags.
123 Bags are recovered from the field after a certain time. Here, we collect bags after germination
124 so that we count intact, ungerminated seeds and germinants (filled circles in Fig. 1A). Sampling
125 tends to be destructive, particularly if intact seeds are tested for viability using a method such as
126 tetrazolium staining. As a result, these studies typically retrieve different subsets of bags for seed
127 and germinant counts at different times (Fig. 1A).

128 We identify each bag by an ID, index i , and the time that it was recovered, index j (columns
129 1-2 in Fig. 1B). We also record time as a variable, t_{ij} (column 3 in Fig. 1B). Each bag has three
130 counts: the number of seeds added to the mesh bags at the start of the experiment, n_{ij} , the number
131 of intact, ungerminated seeds, and the number of germinants, $y_{g,ij}$ (columns 4-6 in Fig. 1B). Here,
132 we assume that all intact seeds are viable (but we discuss combining field experiments and lab
133 viability assays in **Discussion: Extensions**). Finally, we calculate the number of seeds surviving
134 to sampling as the sum of intact, ungerminated seeds and germinants – we assume this is both
135 the number of survivors, y_{ij} , and the number of seeds surviving to just before germination, $n_{g,ij}$
136 (columns 7-8 in Fig. 1B).

137 In seed addition experiments, we lightly bury or sprinkle seeds on the soil surface (0 months
138 in Fig. 1C). Seeds are buried in plots where we do not expect a substantial seed bank, or in pots
139 or trays with seed-free soil. We might also include control plots without seed addition to account
140 for natural seed rain. We survey the plots for seedlings after germination (filled circles in Fig. 1C).
141 Typically, it is not possible to recover intact seeds from the soil but because seedling counts are
142 non-destructive, we can resurvey plots (Fig. 1C).

143 We identify each plot by an ID, index i , and record the time it was surveyed, index j (columns
144 1-2 in Fig. 1D). We also record time as a variable, t_{ij} (column 3 in Fig. 1C). Each plot and survey
145 time has two counts: the number of seeds added to the plot at the start of the experiment, n_{ij} , and

146 the number of seedlings, $y_{g,ij}$ (columns 4 & 6 in Fig. 1C).

147 **3.2 Deterministic model for seed fates**

148 The fate of seeds in the seed bank can be characterized using methods from event history anal-
149 ysis (also called survival or failure time analysis; reviewed in Fox 2001; Landes et al. 2020). By
150 focusing on a seed leaving the seed bank as the event of interest, we can characterize the distri-
151 bution of times at which the event occurs using a set of key functions (Klein and Moeschberger,
152 2003). A survival function describes the probability that a seed remains in the seed bank until time
153 t . The survival function is the term for the probability of the event occurring after time t ; the term
154 applies whether or not the event of interest is death. A probability density function describes the
155 probability that the seed leaves the seed bank at time t . Finally, a hazard function describes the
156 tendency that a seed remaining in the seed bank at time t leaves the seed bank at the next instant
157 in time. The probability density function defines the unconditional probability of events, while the
158 hazard function is associated with the conditional propensity for the event among individuals who
159 have not yet experienced the event (Fox 2001, p. 245). We illustrate the relationship between these
160 functions in **Appendix: Hazards determine the age-structure of the seed bank**.

161 We use these principles to describe how seed loss from the seed bank (the event of interest)
162 depends on mortality and germination. We define hazard functions for each fate. The hazard
163 function for mortality, h_m , is the risk that a seed remaining in the seed bank leaves the seed bank
164 through mortality the next instant. The hazard function for germination, h_g , is the risk that a seed
165 remaining in the seed bank leaves the seed bank through germination the next instant. The hazards
166 thus summarize the “instantaneous risk” (Landes et al., 2020) of mortality or germination. In this
167 paper, we assume that mortality precedes germination, but the principles we describe are flexible
168 and specific equations should be reformulated to correspond to the biology of the study system.

169 We combine the mortality and germination hazards to describe a survival function for the ex-
170 pected probability that seeds remain in the seed bank up to a given time:

$$171 S(t) = \prod_{t_j \leq t} (1 - h_m(t_j)) \times (1 - h_g(t_j)). \quad (1)$$

172
173 Equation 1 is the product of discrete survival functions associated with mortality, $S_m(t_j) = \prod_{t_j \leq t} 1 -$

174 $h_m(t_j)$, and germination, $S_g(t_j) = \prod_{t_j \leq t} 1 - h_g(t_j)$. If the hazards are on an annual timescale, S_m is
 175 the cumulative product of the complement of the mortality probability, up to the number of years
 176 t_j that seeds have been in the soil. Similarly, S_g is the cumulative product of the complement of
 177 the germination probability, up to the number of germination opportunities that seeds have experi-
 178 enced. In terms of the hazards, $h_m(1)$ is a seed's propensity for mortality in the first year and $h_g(1)$
 179 is the propensity for germination of a seed that does not die during the first year. The seeds that
 180 remain in the seed bank past one year do not die with propensity $1 - h_m(1)$ and do not germinate
 181 with propensity $1 - h_g(1)$.

182 We refer to the case with discrete mortality hazards (Equation 1) as a 'non-parametric model,'
 183 and to use it in a statistical model we specify the hazards in terms of probabilities. The mortality
 184 hazard, $h_m(t_j)$, is the probability of mortality during each time interval j , $p_{m,j}$. Specifically, it is
 185 the conditional probability of mortality for seeds that remain in the seed bank. We describe seeds
 186 remaining in the seed bank *after* the period in which they experience mortality, but *before* the
 187 germination opportunity. We assume that after this time interval, seeds have a discrete opportunity
 188 to germinate. The germination hazard, $h_g(t_j)$, is the conditional probability of germination at each
 189 opportunity, $p_{g,j}$, for a seed that remains in the seed bank up to just before germination.

190 With these hazards, the mortality component is defined by $\prod_{j=1}^J 1 - p_{m,j}$. The germination
 191 component is defined by $\prod_{j=1}^J (1 - p_{g,j-1})^{I(j>1)}$, where $I(x)$ is an indicator function equal to 1
 192 if the inequality is true, and 0 if it is not (Metcalf et al., 2009). We use the indicator function
 193 because at the first time j , seeds have not yet experienced a germination opportunity. After the first
 194 germination opportunity, the 'germination history' is defined by the product of past germination
 195 opportunities. The product of the mortality and germination components describe the probability
 196 that seeds remain in the seed bank after j time intervals (e.g. years) as

$$197 \quad f(\mathbf{p}_g, \mathbf{p}_m) = \prod_{j=1}^J \overbrace{(1 - p_{m,j})}^{\text{survival function}} \times \overbrace{(1 - p_{g,j-1})^{I(j>1)}}^{\text{germination history}}. \quad (2)$$

198
 199 The choice of how to represent mortality and germination makes explicit our assumptions about
 200 how those processes operate. The most simple version of the model in Equation 2 is one in which

201 the hazards are constant; neither mortality nor germination probability change with seed age. In
 202 this case, $p_{m,1} = p_{m,2} = \dots = p_{g,J}$ and $p_{g,1} = p_{g,2} = \dots = p_{g,J}$. Mortality and germination are
 203 each described by a single parameter, p_m and p_g .

204 **3.3 Likelihood functions for observations from seed bag burial and seed addition experi-** 205 **ments**

206 To estimate seed mortality and germination, we use probability statements to connect the ob-
 207 servations from field experiments to the deterministic models. We describe likelihood functions
 208 for observations from seed bag burial (Figure 1A-B) and seed addition (Figure 1C-D) experiments.
 209 To illustrate our approach, we assume that mortality and germination do not depend on seed age.
 210 The general structure of the likelihood remains when we relax the assumption of constant hazards
 211 for mortality or germination (Table 1).

212 For the seed bag burial experiment, we construct one likelihood for the observations of ger-
 213 minants and another likelihood for the observations of surviving seeds. First, we use the ob-
 214 servations of germinants to describe a model for the probability of germination, p_g . We as-
 215 sume that the number of seeds that germinate, $y_{g,ij}$, are a binomial sample from the number
 216 of seeds surviving to just before germination, $n_{g,ij}$. Recall that the number of surviving seeds
 217 is the sum of germinants and ungerminated, intact seeds. We estimate the probability of ger-
 218 mination, p_g , for a seed that survives up to just before germination. The likelihood is then
 219 $L(p_g | \mathbf{y}_g) = \prod_{i=1}^I \prod_{j=1}^J \text{binomial}(y_{g,ij} | n_{g,ij}, p_g)$.

220 Next, we use the observations of surviving seeds to describe a survival function for the prod-
 221 uct of germination and mortality hazards. We assume that the number of seeds that survive to a
 222 given time are a binomial sample from the number of seeds that start the experiment in each bag:
 223 $\text{binomial}(y_{ij} | n_{ij}, f(\dots))$. The number of surviving seeds is the sum of germinants and ungermi-
 224 nated, intact seeds. The deterministic model, $f(\dots)$, is the product of the germination history and
 225 the survival function for mortality, and describes the probability of not germinating and not dying
 226 up to the time j . For the case in which mortality and germination do not depend on seed age,
 227 $f(p_m, p_g) = \prod_{k=1}^j (1 - p_m)(1 - p_g)^{I(j>1)}$. The joint likelihood for observations of germinants and

228 surviving seeds is

$$\begin{aligned}
 f(p_m, p_g) &= \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)} \\
 L(p_m, p_g | y_g, y) &= \prod_{i=1}^I \prod_{j=1}^J \left[\text{binomial}(y_{g,ij} | n_{g,ij}, p_g) \text{binomial}(y_{ij} | n_{ij}, f(p_m, p_g)) \right].
 \end{aligned}
 \tag{3}$$

231 Because bags are destructively sampled, the indices for bag ID, i , and recovery time, j , are redun-
 232 dant and the likelihood function will range over unobserved combinations of bag ID and recovery
 233 time (e.g. any bag i at a time j when the bag was not recovered). We retain this notation because it
 234 makes explicit the parallel with the likelihood for observations from seed addition experiments and
 235 because, in practice, we omit the unobserved combinations from the likelihood when implementing
 236 it with statistical software.

237 For the seed addition experiment, we construct a likelihood for the observations of seedlings.
 238 We assume that the number of seedlings is a binomial sample from the number of seeds that start
 239 the experiment: $\text{binomial}(y_{g,ij} | n_{ij}, f(\dots))$. The number of seedlings is the product of mortality
 240 and germination. We describe the combination of those processes with a deterministic model,
 241 $f(\dots)$, that modifies Equation 2 to include germination. Each observation takes place at the time
 242 of germination, rather than after, so that $f(p_m, p_g) = p_g \times \prod_{k=1}^j (1 - p_m)(1 - p_g)^{I(j>1)}$. To account
 243 for germination, the function now includes the probability of germination, p_g , in addition to the
 244 survival function for mortality and the germination history. The likelihood for observations of
 245 seedlings is

$$\begin{aligned}
 f(p_m, p_g) &= p_g \times \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)} \\
 L(p_m, p_g | y_g) &= \prod_{i=1}^I \prod_{j=1}^J \left[\text{binomial}(y_{g,ij} | n_{g,ij}, f(p_m, p_g)) \right].
 \end{aligned}
 \tag{4}$$

248 4 Methods

249 To conduct a comprehensive analysis of statistical models for observations from seed bag burial
 250 and seed addition experiments, we now consider statistical models with four combinations of con-
 251 stant (C) or age-dependent (A) seed mortality and germination. Population models that incorporate
 252 a seed bank typically use one of the following mortality/germination combinations to parameterize

253 seed stages: C/C (e.g. Kurkjian et al. 2017), A/C (e.g. Yates and Ladd 2010), C/A (e.g. Elderd
254 and Miller 2016), and A/A (e.g. Kalisz 1991). We thus consider models to estimate the following
255 cases:

- 256 1. Constant mortality/constant germination (C/C): Mortality, p_m , and germination, p_g , hazards
257 are the same for all seed ages.
- 258 2. Age-dependent mortality/constant germination (A/C): The mortality hazard is a function of
259 seed age, $p_{m,j}$, while the germination hazard is the same for all seed ages, p_g .
- 260 3. Constant mortality/age-dependent germination (C/A): The mortality hazard is the same for
261 all seed ages, p_m , while the germination hazard is a function of seed age, $p_{g,j}$.
- 262 4. Age-dependent mortality/age-dependent germination (A/A): Both mortality, $p_{m,j}$, and ger-
263 mination, $p_{g,j}$, hazards are functions of seed age.

264 For each of these four cases, we study the identifiability of models for seed bag burial and seed ad-
265 dition experiments to determine when each can be used to estimate seed mortality and germination.
266 To directly compare the statistical properties of estimates for seed mortality and germination from
267 seed bag burial and seed addition experiments, we fit a model with constant mortality and constant
268 germination (C/C) to observations from a seed bank with constant mortality and constant germina-
269 tion (C/C). Finally, we study the consequences of model misspecification on parameter estimates.
270 We focus on a special case where observations are generated by a seed bank with age-dependent
271 mortality and constant germination (A/C) but we fit a model with constant mortality and constant
272 germination (C/C).

273 **4.1 Identifiability analysis by the symbolic method**

274 To determine when seed bag burial and seed addition experiments can be used to estimate con-
275 stant or age-dependent seed mortality and germination, we analyze the identifiability of statistical
276 models for the experiments. We study if parameters can be estimated in terms of the structure of
277 the likelihood ('intrinsic identifiability') (Cole 2020). Intrinsic identifiability refers to cases where
278 parameters in a model can be uniquely estimated. For example, models will not be identifiable if
279 different combinations of mortality and germination have the same likelihood for a set of observa-

280 tions. If the model is not identifiable, there are no unique maximum likelihood estimates.

281 To analyze the identifiability of statistical models for different combinations of experiment,
282 hazard, and length of the experiment, we use an algebraic approach called the symbolic method
283 (Catchpole and Morgan 1997; Cole et al. 2010; Cole 2020). With this method, we focus on general
284 issues of experimental design and model structure rather than on specific datasets. We determine
285 the intrinsic identifiability of statistical models for all combinations of experiment (seed bag burial
286 vs. seed addition), hazards (C/C, A/C, C/A, A/A), and length of experiment (1, 2 or 3 years). All
287 the likelihoods that we analyze are shown in Table 1. To apply the symbolic method, we summarize
288 each model by a vector that completely determines the model (an ‘exhaustive summary’). The
289 exhaustive summary is subsequently differentiated with respect to all of the constituent parameters
290 to form a ‘derivative matrix’ (the transpose of the Jacobian). The model is identifiable if the rank of
291 the derivative matrix is equal to the number of parameters in the model; the model is not identifiable
292 if the rank of the derivative matrix is less than the number of parameters. We implement these steps
293 using the computer algebra software Maxima (Maxima, 2014); for detailed methods and scripts,
294 see **Appendix: Identifiability analysis**.

295 **4.2 Simulation experiments**

296 To compare the statistical properties of seed bag burial and seed addition experiments, and
297 study the effect of model misspecification, we conduct numerical experiments in which we fit
298 models to simulated data. To simulate data with the structure of seed bag burial and seed addition
299 experiments (Figure 1), we use the likelihoods corresponding to those observations (Table 1). In
300 practice, we use mortality and germination hazards to calculate the expected probability of a seed
301 remaining in the soil at the end of each year, and its subsequent probability of germinating. We use
302 the expected probability of remaining in the soil to draw a binomial sample of seeds from the initial
303 number of seeds in the bag. We use the probability of germination to draw a binomial sample of
304 germinants from the seeds remaining in the bag. To simulate data with the structure of the seed
305 addition experiment, we retain only the observations of seedlings.

306 Both maximum likelihood and Bayesian methods would be appropriate to fit the models as-

307 sociated with seed bag burial and seed addition experiments. However, we chose to fit Bayesian
308 models to the simulated observations because we can readily estimate the parameters in the joint
309 likelihood. All parameters in our models are probabilities with support $[0, 1]$ on which we place
310 $\text{beta}(1, 1)$ priors. Figure S4 shows the directed acyclic graphs corresponding to the joint and pos-
311 terior distributions for the models. Parameters and sample sizes for simulations are given in the
312 sections that follow. We wrote all simulations and analyzed model output in R version 3.6.2 (R
313 Core Team, 2019). We wrote, fit all models, and sampled posterior distributions using JAGS 4.10
314 with **rjags** (Plummer et al., 2019). For each fit, we ran 3 chains with 3,000 iterations for adap-
315 tation, 5,000 for burn-in, and 5,000 for sampling. For computational efficiency, we thinned the
316 chains and kept every 10th iteration. We used the **MCMCvis** package to work with model output,
317 check chains for convergence, and recover posterior distributions (Youngflesh et al., 2021).

318 **4.2.1 Statistical properties of seed bag burial and seed addition experiments**

319 To compare the statistical properties of estimates from identifiable models, we used a simula-
320 tion experiment in which we fit a model with constant mortality and constant germination (C/C)
321 to observations from a seed bank with constant mortality and constant germination (C/C). We
322 generated data from a 3-year experiment with a sample sizes $n = (5, 10, 15, 20, 25, 30)$ bags or
323 plots each year. Each bag or plot started the experiment with 100 seeds. For each sample size,
324 we simulated 250 replicate datasets for the following combinations of ‘true’ mortality and germi-
325 nation: low mortality/low germination (0.1, 0.1), low mortality/high germination (0.1, 0.5), high
326 mortality/low germination (0.5, 0.1), and high mortality/high germination (0.5, 0.5). We then fit
327 each simulated dataset with two models; one for a seed bag burial experiment and one for a seed
328 addition experiment.

329 To quantify the bias of estimates, we calculated the difference between the posterior modes
330 and the ‘true’ parameters for the probability of mortality or germination. Parameter estimates are
331 unbiased when the difference is 0. To quantify the uncertainty of estimates, we calculated the width
332 of the 95% credible interval for each parameter. For each set of ‘true’ parameters and sample sizes,
333 we estimated the mean difference and width, and quantified 95% confidence intervals for each with

334 a t distribution (Pappalardo et al., 2020). To estimate the coverage of the 95% credible intervals,
335 we calculated the proportion of credible intervals that contain the ‘true’ parameter value. Ideally,
336 a 95% credible interval would contain the ‘true’ parameter value 95% of the time. We calculated
337 confidence intervals for coverage with the Wilson method in the **binom** package (Pappalardo et al.,
338 2020). Finally, we calculated root mean squared error as a measure of the combined effect of bias
339 and uncertainty.

340 **4.2.2 Consequences of model misspecification**

341 To study the consequences of model misspecification, we focused on a special case in which
342 we fit a model with constant mortality and constant germination (C/C) to observations from a seed
343 bank with age-dependent mortality and constant germination (A/C). We generated data from a 3-
344 year experiment with sample sizes of $n = (5, 10, 15, 20, 25, 30)$ bags or plots each year. Each bag
345 or plot started the experiment with 100 seeds. For each sample size, we simulated 250 replicate
346 datasets in which ‘true’ probabilities of mortality in the three years was $p_{m,1} = 0.1$, $p_{m,2} = 0.2$, and
347 $p_{m,3} = 0.3$. The germination rate in all years was $p_g = 0.1$. As before, we fit two models to each
348 simulated dataset; one for a seed bag burial experiment and one for a seed addition experiment. In
349 all cases we fit the C/C model with two parameters, p_m and p_g . Even though we only estimated
350 one parameter for the probability of mortality, we compared properties of the estimate to the age-
351 dependent probability of mortality in each of the three years. For all parameters, we quantified
352 bias, uncertainty, coverage, and root mean squared error.

353 **5 Results**

354 **5.1 Identifiability analysis by the symbolic method**

355 All models for observations from seed bag burial experiments exhibit a deficiency of 0, indi-
356 cating that the models are identifiable (Table 2). In all cases we consider, the models for seed bag
357 burial experiments can be used to estimate parameters for seed mortality and germination. Models
358 for observations from seed addition experiments only show a deficiency of 0 when mortality and
359 germination rates are assumed to be constant, and when more than one year of observations is
360 available (Table 2). In all other cases, models have a deficiency greater than 0, indicating that the

361 models are not identifiable.

362 **5.2 Statistical properties of seed bag burial and seed addition experiments**

363 The C/C models fit to observations from the seed bag burial and seed addition experiments are
364 identifiable when there is more than one year of data (Table 2); here, we analyze simulated data for
365 3-year long experiments. Both experiments produce unbiased estimates of mortality (Fig. 2A-D)
366 and germination (Fig. 2I-L) at large sample sizes. At small sample sizes, seed addition experiments
367 are more likely to produce biased estimates (e.g. Fig. 2A, C). Estimates from seed addition experi-
368 ments display greater uncertainty for all parameter values and sample sizes (Fig. 2E-H, M-P). The
369 difference in uncertainty of estimates between experiments depends on the true probability of mor-
370 tality and germination. Seed mortality estimates show 3-5 times more uncertainty for seed addition
371 experiments when mortality and germination are low, but at most 2 times as much uncertainty when
372 mortality is low but germination is high (Fig. 2E vs. F). For both experiments, coverage is $\sim 95\%$
373 (Fig. S8A-D, I-L), and root-mean squared error decreases with sample size (Fig. S8E-H, M-P).

374 **5.3 Consequences of model misspecification**

375 We fit the C/C model to observations from a simulation in which the probability of seed mor-
376 tality increases over time ($p_{m,1} = 0.1, p_{m,2} = 0.2, p_{m,3} = 0.3$). For both seed bag burial and seed
377 addition experiments, the bias in mortality estimates changes over time (Fig. 3A-C). Both experi-
378 ments progress from overestimating to underestimating mortality, but the magnitude of bias varies
379 depending on the experiment (open vs. filled points). In the first year, seed bag burial experiments
380 exhibit less bias than seed addition experiments; this pattern reverses by the third year. Bias is
381 unaffected by sample size (Fig. 3A-C), but the width of the 95% credible interval decreases with
382 increasing sample size for all parameters and both experiments (Fig. 3E-G). Low accuracy and
383 increased precision at larger sample sizes reduces coverage even when bias does not change (e.g.
384 Fig. 3I-K). The root-mean squared error (RMSE) for mortality is largely determined by the bias of
385 estimates; estimates with a smaller absolute bias also show smaller RMSE (Fig. 3M-O).

386 The ‘true’ probability of germination does not depend on seed age in the simulation, but ger-
387 mination estimates are slightly biased for both seed bag burial and seed addition experiments

388 (Fig. 3D). Although the absolute magnitude of bias is smaller than for mortality estimates, ger-
389 mination is overestimated by 13-20%. The coverage of estimates also decreases with increasing
390 sample size (Fig. 3L), but RMSE is relatively low (Fig. 3P).

391 **6 Discussion**

392 We develop and analyze statistical models for observations from field experiments commonly
393 used to study the soil seed bank. We present the first systematic evaluation and comparison of
394 inferences made with statistical models for seed bag burial and seed addition experiments. We
395 show that seed bag burial experiments can separately estimate mortality and germination even if
396 one, or both, are age-dependent. For seed addition experiments, we demonstrate that seed mortality
397 and germination are only identifiable if both mortality and germination do not change with seed
398 age and with more than one year of observations. In all other cases, it is impossible to separately
399 estimate mortality and germination.

400 To compare the statistical properties of estimates from seed bag burial and seed addition ex-
401 periments, we focus on the identifiable models with constant mortality and constant germination.
402 Estimates from both experiments are unbiased as sample size increases. However, estimates from
403 seed bag burial experiments are more precise for all parameter combinations that we consider.
404 Finally, we evaluate the effect of fitting the wrong model to observations from each experiment.
405 We fit a model with constant mortality and germination rates to observations produced by age-
406 dependent mortality and constant germination. The bias of mortality estimates changes over time,
407 and is exacerbated by increased precision at large sample sizes. Germination estimates are also
408 biased, though to a lesser extent.

409 **6.1 Recommendations for practitioners**

410 We demonstrate how seed bag burial or seed addition experiments can be used to estimate seed
411 mortality and germination. To estimate age-dependent mortality or germination rates in the field,
412 you should conduct a seed bag burial experiment. Even when estimating constant mortality and
413 germination, seed bag burial experiments will produce estimates that are more accurate and precise
414 for a given sample size. Nonetheless, estimates from seed addition experiments will be unbiased

415 when mortality and germination do not change with seed age.

416 We suggest that the best way to adapt our broad-strokes recommendations is to simulate data
417 and fit models to those simulations. Practitioners already likely to know much about many of the
418 key parts of a seed bank experiment. How many seeds could be collected and used for an ex-
419 periment, how many replicates are logistically feasible, and for how long would the experiment
420 run? With these pieces in hand, it is then possible to use plausible values for seed mortality and
421 germination rates to simulate observations. It will not be possible to know the ‘true’ values or their
422 age-dependence, but simulations could explore likely scenarios (e.g. constant vs. increasing mor-
423 tality). Fitting models to these simulations would then make it possible to compare the statistical
424 properties of estimates from seed bag burial versus seed addition experiments. To facilitate this
425 process, we include the code for our study (<https://zenodo.org/record/5794709>); this includes R
426 code to simulate observations, the JAGS code for the models, and the R code to fit the models to
427 observations.

428 Our analysis can also help guide parameter estimation if observations have already been col-
429 lected. Lack of identifiability creates issues for both frequentist and Bayesian statistical methods,
430 which we illustrate in detail in **Appendix: Implications of identifiability for model fitting**. No
431 amount of clever modeling can estimate parameters when they are intrinsically not identifiable.
432 Observations from seed bag burial experiments give you the flexibility to fit models with constant
433 or age-dependent mortality and germination. With observations from seed addition experiments, it
434 is only possible to fit models with constant mortality and germination.

435 Ultimately, the impact of bias or imprecision in estimates of seed mortality or germination on
436 population growth rate depends on the sensitivity of population growth rate to those vital rates. The
437 models and analyses we present will most relevant to researchers working with plant populations in
438 which aboveground vital rates exhibit high temporal variability because these populations are likely
439 sensitive to transitions in the seed bank (Doak et al., 2002; Nguyen et al., 2019). Considering the
440 broader context of the plant life history can help population ecologists determine which fieldwork
441 and modeling approaches are sufficiently accurate and precise for their study system.

442 **6.2 Extensions**

443 Existing studies have used simulations and post-hoc comparisons to explore the consequences
444 of age structure in the seed bank, emphasize how estimates of seed rates interact with temporal
445 variability in aboveground success, and describe the effect of underestimating parameter uncer-
446 tainty (Doak et al., 2002; Paniw et al., 2017; Nguyen et al., 2019). However, these methods do not
447 provide an intuitive way to use observations to test assumptions about seed bank structure and as-
448 sociated parameter uncertainty. For example, the methods do not allow for model checks or model
449 selection, both of which could be used to ask whether the fitted model is consistent with obser-
450 vations. Because accuracy and precision of estimates for seed mortality and germination interact
451 with information about other components of the life cycle, it seems crucial to evaluate the model
452 used to estimate seed mortality or germination separately from the population model.

453 The models we define can accommodate constant and age-dependent seed mortality and germi-
454 nation. In our simulations, we can assess the accuracy of parameter estimates obtained with these
455 models because we picked the values used to generate the data. We lack this luxury for empirical
456 datasets. But standard model checking (e.g. Conn et al. 2018) and model selection (e.g. Hooten
457 and Hobbs 2015; Tredennick et al. 2021) methods could be applied to determine whether the model
458 that we fit to empirical data is consistent with the process that generated the data. Evaluating the
459 performance of these approaches would be a valuable extension to the present work.

460 Studies also describe seed mortality with parametric functions such as exponential models (e.g.
461 Lommen et al. 2018). Analyzing the identifiability and statistical properties of models with contin-
462 uous, parametric descriptions of seed mortality would complement the present study and connect
463 it to the work of Rees and Long (1993), who fit a variety of parametric models for recruitment to
464 observations of seedlings from a seed addition experiment. The authors showed that recruitment is
465 affected by the age-dependence of seed mortality and germination, and that seed banks do not, as a
466 rule, exhibit exponential decay (Rees and Long, 1993). However, they did not separately estimate
467 seed mortality and germination. The models we present could be expanded to include continuous,
468 parametric descriptions for mortality, in which case we would describe the combination of contin-

469 uous mortality and discrete germination hazards with a product integral (Klein and Moeschberger
470 2003, p. 36). In comparison to our non-parametric approach, a parametric description for mor-
471 tality could reduce the number of parameters and facilitate the use of methods from event history
472 analysis (Fox, 2001; Landes et al., 2020).

473 It would also be valuable to combine information from seed bag burial and seed addition exper-
474 iments, and from field experiments with laboratory trials. Studies that have gone to great lengths
475 to carry out both seed bag burial and seed addition experiments (e.g. Liu et al. 2005) have not
476 been able to formally combine observations from those experiments and instead explore a variety
477 of scenarios based on the parameters estimated from each experiment. In addition, a common
478 endpoint for field experiments with seeds is to test intact seeds for viability with lab assays, which
479 may also have uncertainty associated with them. In certain cases, it might be desirable to combine
480 the assays and field experiments to fully account for uncertainty about seed fates.

481 **6.3 Limitations**

482 Event history analysis is developed for and appropriately applied to individual data (Zens and
483 Peart, 2003; Landes et al., 2020), and the models we describe would be completely appropriate if
484 applied to observations of individual seeds. Yet seeds of many plant species are too small for indi-
485 viduals to be tracked in the field. When examining aggregate data—from cohorts, or populations—
486 heterogeneity between subpopulations and change in hazards over time can confound whether pat-
487 terns are the result of changes to hazards or to population structure (Rees and Long, 1993; Zens
488 and Peart, 2003). Our approach is not intended to assess changes to the hazards for individual
489 seeds (unless individual-level data are available) but rather a framework for consistent inferences
490 about seed mortality and germination.

491 To focus on the commonalities between seed bag burial and seed addition experiments, we
492 describe stereotyped versions of each. Not all experiments in the literature exactly follow the
493 schematic we describe; some seed bag burial experiments count intact seeds and estimate germi-
494 nation in another way (e.g. Lommen et al. 2018), or count only seeds at certain times, but both
495 seeds and germinants at other times (e.g. Eckhart et al. 2011). Individual analyses will inevitably

496 have to be tailored to specific data. We sought to explicitly describe the assumptions underlying
497 our statistical models so that they could be readily modified. Investigators will naturally construct
498 models that are appropriate to their system and aims.

499 Other studies have addressed issues of experimental design that could affect observations from
500 seed bag burial or seed addition experiments. For example, high seed densities in mesh bags may
501 promote transmission of pathogenic fungi and increase seed mortality (Van Mourik et al., 2005).
502 Seed bag or seed burial depth may influence mortality and germination rates; for instance, Dille
503 et al. (2017) showed that deeper burial reduced germination, but not mortality, for *Kochia scoparia*
504 seeds. Although beyond the scope of our study, accounting for such considerations is a crucial part
505 of collecting observations that reflect seed mortality in and germination from the soil seed bank.

506 **6.4 Conclusion**

507 Observations from seed bag burial and seed addition experiments are hard-won data, but statis-
508 tical models for observations from these experiments have received little attention to-date. Study-
509 ing these models can help plant population ecologists make the most of existing and future data
510 by identifying potential models to fit, the statistical properties of parameter estimates, and poten-
511 tial bias introduced by making assumptions about age-dependence of mortality and germination.
512 Our analysis contributes to efforts to make richer inferences from the trove of demographic data
513 collected by plant population ecologists.

514 **7 Acknowledgments**

515 We thank W. F. Morris and T. E. X. Miller for discussions about seed bag experiments that
516 helped prompt this research. G. Siegmund was supported by a Graduate Research Fellowship
517 (DGE-1144153) from the U.S. National Science Foundation, and a Presidential Life Science Fel-
518 lowship and Cornell Fellowship from Cornell University. M. A. Geber was supported by the U.S.
519 National Science Foundation (LTREB 1256288 and 1754299).

520 **8 Author contributions**

521 GS and MAG conceived of the ideas in the study. GS developed the statistical model, analyzed
522 identifiability & simulations, and wrote the manuscript with input from MAG.

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Table 1: Likelihoods of models for observations from seed bag burial and seed addition experiments.

Model		
Mortality	Germination	Likelihood
SEED BAG BURIAL EXPERIMENT		
$C(p_m)$	$C(p_g)$	$f(p_m, p_g) = \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g, y) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, p_g) \text{binomial}(y_{ij} n_{ij}, f(p_m, p_g))]]$
$A(p_{m,j})$	$C(p_g)$	$f(p_{m,j}, p_g) = \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g, y) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, p_g) \text{binomial}(y_{ij} n_{ij}, f(p_{m,j}, p_g))]]$
$C(p_m)$	$A(p_{g,j})$	$f(p_m, p_{g,j}) = \prod_{k=1}^j (1 - p_m) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_g y_g, y) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, p_{g,j}) \text{binomial}(y_{ij} n_{ij}, f(p_m, p_{g,j}))]]$
$A(p_{m,j})$	$A(p_{g,j})$	$f(p_{m,j}, p_{g,j}) = \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_g y_g, y) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, p_{g,j}) \text{binomial}(y_{ij} n_{ij}, f(p_{m,j}, p_{g,j}))]]$
SEED ADDITION EXPERIMENT		
$C(p_m)$	$C(p_g)$	$f(p_m, p_g) = p_g \times \prod_{k=1}^j (1 - p_m) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, f(p_m, p_g))]]$
$A(p_{m,j})$	$C(p_g)$	$f(p_{m,j}, p_g) = p_g \times \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_g)^{I(j>1)}$ $L(p_m, p_g y_g) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, f(p_{m,j}, p_g))]]$
$C(p_m)$	$A(p_{g,j})$	$f(p_m, p_{g,j}) = p_{g,j} \times \prod_{k=1}^j (1 - p_m) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_g y_g) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, f(p_m, p_{g,j}))]]$
$A(p_{m,j})$	$A(p_{g,j})$	$f(p_{m,j}, p_{g,j}) = p_{g,j} \times \prod_{k=1}^j (1 - p_{m,j}) \times (1 - p_{g,j})^{I(j>1)}$ $L(p_m, p_g y_g) = \prod_{i=1}^I \prod_{j=1}^J [\text{binomial}(y_{g,ij} n_{g,ij}, f(p_{m,j}, p_{g,j}))]]$

¹ In columns 1 and 2, C is a constant hazard and A is an age-dependent hazard.

² In all likelihoods, $I(x)$ is an indicator function equal to 1 if the inequality is true, and 0 if it is not. As discussed in the main text, the indicator function identifies whether or not seeds have yet experienced a germination opportunity; at the first time point j , they have not.

Table 2: Analysis of intrinsic identifiability for non-parametric models with different assumptions about whether germination and mortality are constant or age-dependent. Each row corresponds to a model in which the germination component is defined in column one and the mortality component is defined in column two. For each model, the columns show the results of the intrinsic identifiability analysis for 1, 2, or 3 years of observations. The analysis identifies the deficiency of the model for a given set of assumptions about the germination and mortality components. The deficiency is calculated as in Cole (2020): the number of parameters in the model minus the rank of the derivative matrix, the latter calculated by the symbolic method. Models with a deficiency of 0 are identifiable; models with a deficiency greater than 0 are not identifiable.

Model		Deficiency		
Mortality component	Germination component	1 year	2 years	3 years
SEED BAG BURIAL EXPERIMENT				
Constant (p_m)	Constant (p_g)	0	0	0
Age-dependent ($p_{m,j}$)	Constant (p_g)	0	0	0
Constant (p_m)	Age-dependent ($p_{g,j}$)	0	0	0
Age-dependent ($p_{m,j}$)	Age-dependent ($p_{g,j}$)	0	0	0
SEED ADDITION EXPERIMENT				
Constant (p_m)	Constant (p_g)	1	0	0
Age-dependent ($p_{m,j}$)	Constant (p_g)	1	1	1
Constant (p_m)	Age-dependent ($p_{g,j}$)	1	1	1
Age-dependent ($p_{m,j}$)	Age-dependent ($p_{g,j}$)	1	2	3

632 **10 Figure captions**

633 **Figure 1.** (A) Schematic of a seed bag burial experiment. Each bag in the experiment is represented
634 by a single line from when the bag is buried at month 0 to when the bag is dug up for sampling
635 (filled circles). The data are organized with indices for bag and sampling time. (B) Data from the
636 seed bag burial experiment. Each row corresponds to a bag and sampling time. (C) Schematic of
637 a seed addition experiment. Each plot in the experiment is represented by a single line from when
638 seeds are added to the plot at month 0 to when plots are censused for seedlings (filled circles). The
639 data are organized with indices for plot and time. (D) Data from the seed addition experiment.
640 Each row corresponds to a plot and sampling time.

641 **Figure 2.** Results of simulation experiment in which we generated observations with constant
642 mortality and germination, and fit a model with constant mortality and germination parameters.
643 (A-D) Bias for estimates of mortality probability, p_m , for different combinations of true mortality
644 and germination probability. (E-H) Width of the 95% credible interval for p_m . (I-L) Bias for esti-
645 mates of germination probability, p_g , for different combinations of true mortality and germination
646 probability. (M-P) Width of the 95% credible interval for p_g . In all panels, error bars represent the
647 95% confidence interval based on a t distribution.

648 **Figure 3.** Results of simulation experiment in which we generated observations with age-dependent
649 mortality and germination, but fit a model with constant mortality and germination. From left to
650 right, columns are analyses of mortality parameters for ages 1, 2, and 3, and germination. (A-D)
651 Bias for estimates of mortality and germination parameters. Error bars represent the 95% confi-
652 dence interval based on a t distribution. (E-H) Width of the 95% credible interval for mortality and
653 germination parameters. Error bars represent the 95% confidence interval based on a t distribution.
654 (I-L) Coverage for mortality and germination parameters. Error bars represent the 95% confidence
655 interval calculated using the Wilson method for binomial proportions. (M-P) Root mean squared
656 error for mortality and germination parameters.

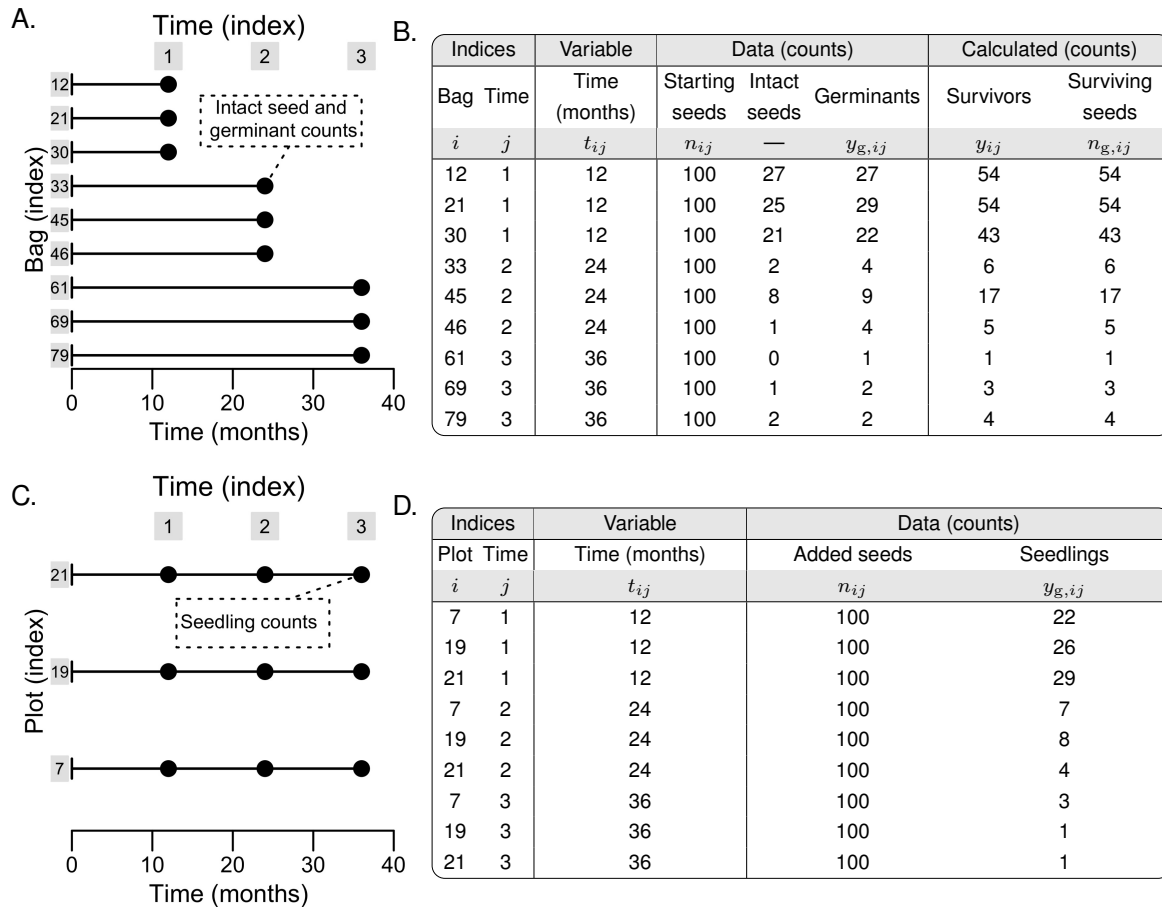


Figure 1: (A) Schematic of a seed bag burial experiment. Each bag in the experiment is represented by a single line from when the bag is buried at month 0 to when the bag is dug up for sampling (filled circles). The data are organized with indices for bag and sampling time. (B) Data from the seed bag burial experiment. Each row corresponds to a bag and sampling time. (C) Schematic of a seed addition experiment. Each plot in the experiment is represented by a single line from when seeds are added to the plot at month 0 to when plots are censused for seedlings (filled circles). The data are organized with indices for plot and time. (D) Data from the seed addition experiment. Each row corresponds to a plot and sampling time.

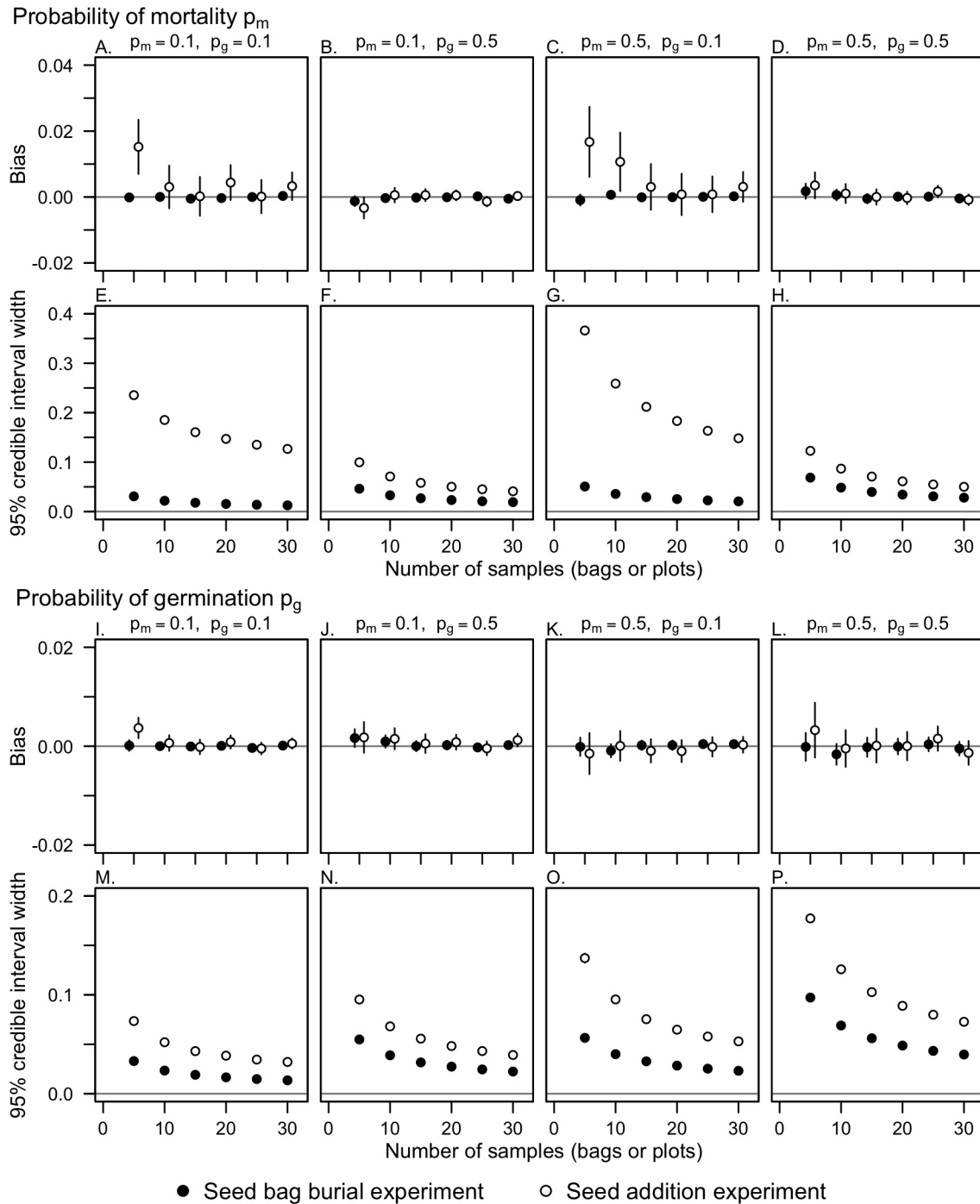


Figure 2: Results of simulation experiment in which we generated observations with constant mortality and germination, and fit a model with constant mortality and germination parameters. (A-D) Bias for estimates of mortality probability, p_m , for different combinations of true mortality and germination probability. (E-H) Width of the 95% credible interval for p_m . (I-L) Bias for estimates of germination probability, p_g , for different combinations of true mortality and germination probability. (M-P) Width of the 95% credible interval for p_g . In all panels, error bars represent the 95% confidence interval based on a t distribution.

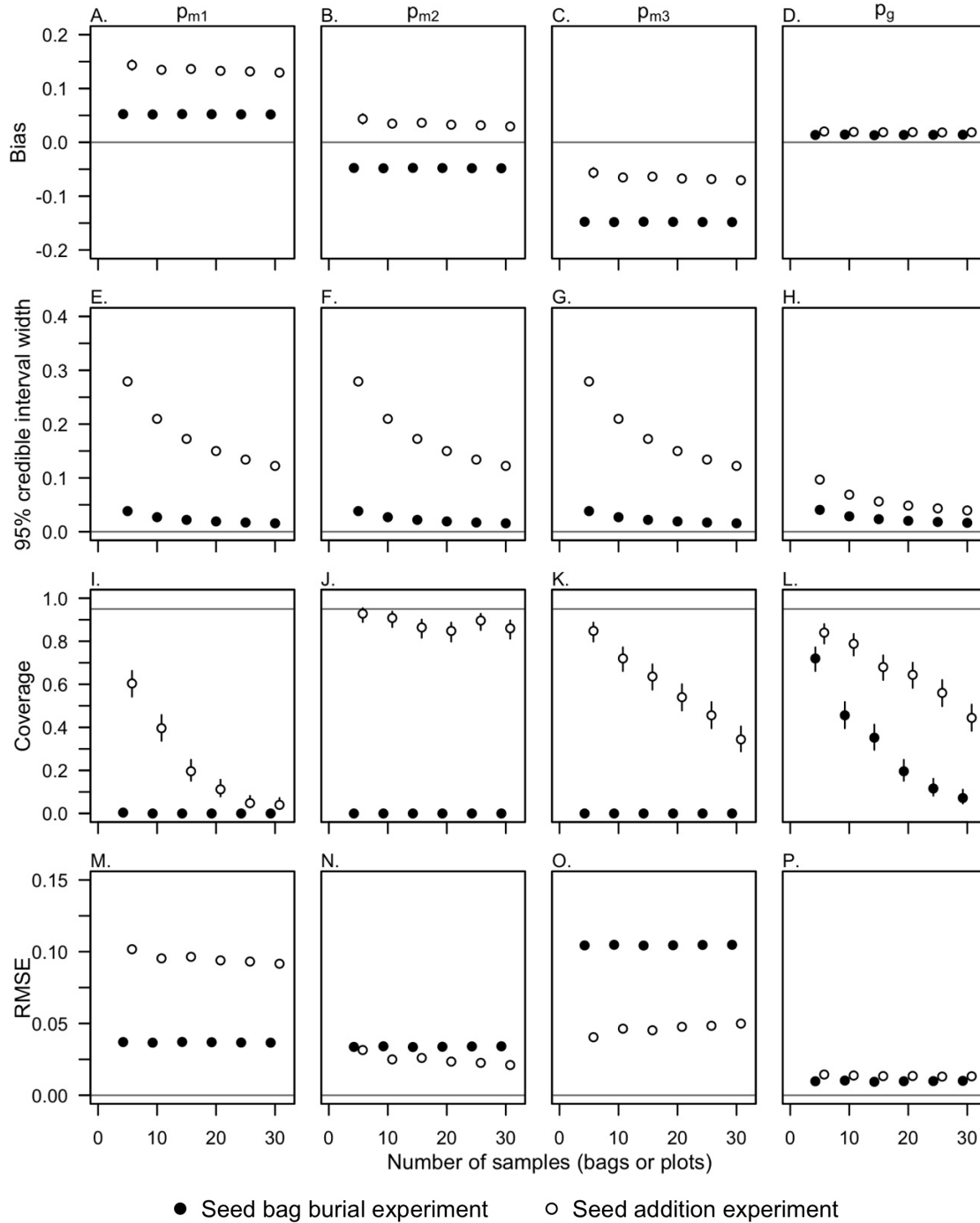


Figure 3: Results of simulation experiment in which we generated observations with age-dependent mortality and germination, but fit a model with constant mortality and germination. From left to right, columns are analyses of mortality parameters for ages 1, 2, and 3, and germination. (A-D) Bias for estimates of mortality and germination parameters. Error bars represent the 95% confidence interval based on a t distribution. (E-H) Width of the 95% credible interval for mortality and germination parameters. Error bars represent the 95% confidence interval based on a t distribution. (I-L) Coverage for mortality and germination parameters. Error bars represent the 95% confidence interval calculated using the Wilson method for binomial proportions. (M-P) Root mean squared error for mortality and germination parameters.