

1 Generation and applications of simulated datasets to integrate social  
2 network and demographic analyses

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6

7 **Running headline:** Tools for integrated network-demographic models

8

## 9 Abstract

- 10 1. Social networks are tied to population dynamics; interactions are driven by population  
11 density and demographic structure, while social relationships can be key determinants  
12 of survival and reproductive success. However, difficulties integrating models used in  
13 demography and network analysis have limited research at this interface.
- 14 2. Here we introduce the R package `genNetDem` for simulating integrated network-  
15 demographic datasets. By generating co-capture data with known statistical  
16 relationships it provides functionality for methodological research. We demonstrate its  
17 use with case studies testing how imputation and sampling design influence the  
18 success of adding network traits to conventional Cormack-Jolly-Seber (CJS) models.
- 19 3. We show that incorporating social network effects in CJS models generates qualitatively  
20 accurate results, but with downward-biased parameter estimates when network position  
21 influences survival. Biases are greater when fewer interactions are sampled or fewer  
22 individuals observed in each interaction. However, model performance is robust to  
23 network covariance, network structure and sampling design. Longitudinal imputation of  
24 missing network measures is more stable than cross-sectional imputation but neither  
25 approach successfully generates accurate parameter estimates.
- 26 4. While our results indicate the potential of incorporating social effects within  
27 demographic models, they show that imputing missing network measures alone is  
28 insufficient to accurately estimate social effects on survival. This points to the  
29 importance of incorporating network imputation approaches. `genNetDem` provides a  
30 flexible tool to aid these methodological advancements and help researchers testing  
31 other sampling considerations in social network studies.

32 **Key-words:** co-capture data; hidden Markov model; population dynamics; stochastic block model; survival;

## 33 **Introduction**

34 Network analysis has revolutionised animal social behaviour research by quantifying  
35 how dyadic social interactions and relationships are nested in wider group- and population-  
36 level social structures (J. Krause et al., 2014; Pinter-Wollman et al., 2013). Network studies in  
37 behavioural ecology have often focussed on how the position of an individual within its social  
38 network influences its fitness, either via reproductive success (Formica et al., 2012; Oh &  
39 Badyaev, 2010) or survival (Blumstein et al., 2018; Ellis et al., 2017; Stanton & Mann, 2012).

40 Quantifying direct links between social network position and fitness can help us  
41 understand how selection acts on social behavioural traits. Further, determining how social  
42 behaviour is linked to survival can identify demographic consequences of interactions and  
43 associations (Clements et al., 2022), which can help develop more realistic models for how  
44 social species respond to population declines or environmental change (Snijders et al., 2017).  
45 However, while there is growing interest in linking animal social networks with demography  
46 (Shizuka & Johnson, 2020), there remain many methodological challenges.

47 Currently most studies that link network position and fitness use known fate approaches  
48 such as generalised linear models (e.g. (Blumstein et al., 2018)) or Cox proportional-hazards  
49 models (e.g. (Ellis et al., 2017)). However, application of these approaches is limited in many  
50 wild populations where individuals that are alive are not necessarily detected. In these cases  
51 survival is most commonly estimated using hidden Markov models (HMMs; (McClintock et al.,  
52 2020)) that can simultaneously estimate survival and capture probabilities (Gimenez et al.,  
53 2012; Pradel, 2005). These models also have potential as tools in animal social network  
54 analysis (Clements et al., 2022; Fisher et al., 2017), especially when not all associations are  
55 detected. However, it is challenging to provide universal guidance on the applicability of these

56 approaches given the diversity of animal social systems and sampling designs used to study  
57 them.

58 Here we introduce the R package `genNetDem` to simulate co-capture datasets. We  
59 define a co-capture dataset as one in which a capture-recapture data also provides information  
60 on social structure, such as when individuals are caught or observed in groups (see also (Silk  
61 et al., 2021)). The package generates integrated longitudinal social network and capture-  
62 recapture datasets with known statistical relationships. This provides functionality for  
63 methodological research, power analyses and sampling design. Here we introduce the  
64 package, describe key functions and outline effective workflows. We then provide two case  
65 studies to demonstrate its use. Finally, we identify key next steps in merging social network  
66 and demographic analyses, and discuss the role of `genNetDem` in these.

67

## 68 **genNetDem overview**

69 `genNetDem` is a set of R (R Core Team, 2021) functions that generate longitudinal  
70 social network and/or capture-recapture datasets with known underlying properties.  
71 Functionality can be split into four broad groups: a) population features; b) survival features; c)  
72 social network features; and d) observation features. The package is modular meaning specific  
73 components can be used in isolation or user-generated code can be integrated to extend  
74 functionality to different ecological or social contexts.

75 `genNetDem` is available on GitHub ([https://github.com/matthewsilk/NETDEM/tree/main/  
76 `genNetDem`](https://github.com/matthewsilk/NETDEM/tree/main/genNetDem)). We provide a brief overview of key functions in the main text and more detail in  
77 the Supplementary Materials.

78

## 79 *Population features*

80           The population features provide capability to simulate a population and generate data  
81 about individuals in it. There is then functionality that simulates population dynamics based on  
82 individual survival probabilities (see *survival features*) and stochastic recruitment that  
83 maintains an approximately stable population size when employed.

84           The **population\_generation\_basic()** function generates data for a group-structured  
85 population distributed uniformly in 2D space. The function takes two arguments:  $n$  defines the  
86 population size and  $ng$  the number of groups in the population. When  $n=ng$  individuals are  
87 distributed uniformly across the defined coordinates. When  $n>ng$  groups are distributed  
88 uniformly across the same coordinates with individuals in the same group sharing the same  
89 spatial location. Currently, simulated population size is independent of the extent of the area it  
90 occupies. Therefore, population density will increase with population size and impact spatial  
91 effects on social network structure. This does not represent a problem except when users want  
92 to compare the social structures of populations of different sizes. The **indiv\_info\_gen()** and  
93 **indiv\_info\_add()** functions provide flexibility in generating and updating individual-level trait  
94 data. Variables can be specified as covariates (e.g. size) or categorical factors (e.g. sex), with  
95 further arguments specifying additional features of the variable (e.g. the distribution of a  
96 covariate or the number of levels and level names of a factor).

97           The **timestep\_demographics()** function controls survival and recruitment in the  
98 simulated population. Survival is stochastic based on each individual's survival probability (see  
99 *survival features*). The number of recruits is Poisson distributed ( $\lambda$  calculated from survival  
100 probability). When the population is group-structured individuals can only be recruited into  
101 existing groups.

102

### 103 *Survival features*

104           The **covariates\_survival()** function allows survival probabilities to be calculated for  
105 each individual based on individual traits or the position of an individual within a population  
106 social network (this could be any network provided to the function; the underlying social  
107 networks, simulated interaction network or a separate user-specified network). There is  
108 considerable flexibility in which measures of network position can be included as covariates;  
109 both the function and R package used can be specified within the function, with functionality for  
110 most common packages (e.g. sna: (Butts, 2014); igraph: (Csardi & Nepusz, 2006); tnet:  
111 (Opsahl, 2009)) incorporated. It is also possible to simulate network covariance in survival  
112 whereby closely connected individuals have either more or less similar survival probabilities  
113 than expected by chance. Currently, covariates\_survival() simulates independent (additive)  
114 effects of traits, meaning there is no functionality to capture interactions among variables (e.g.  
115 network position having different effects in males than females). The simpler **basic\_survival()**  
116 function generates population-level survival probabilities in the absence of covariates.

117

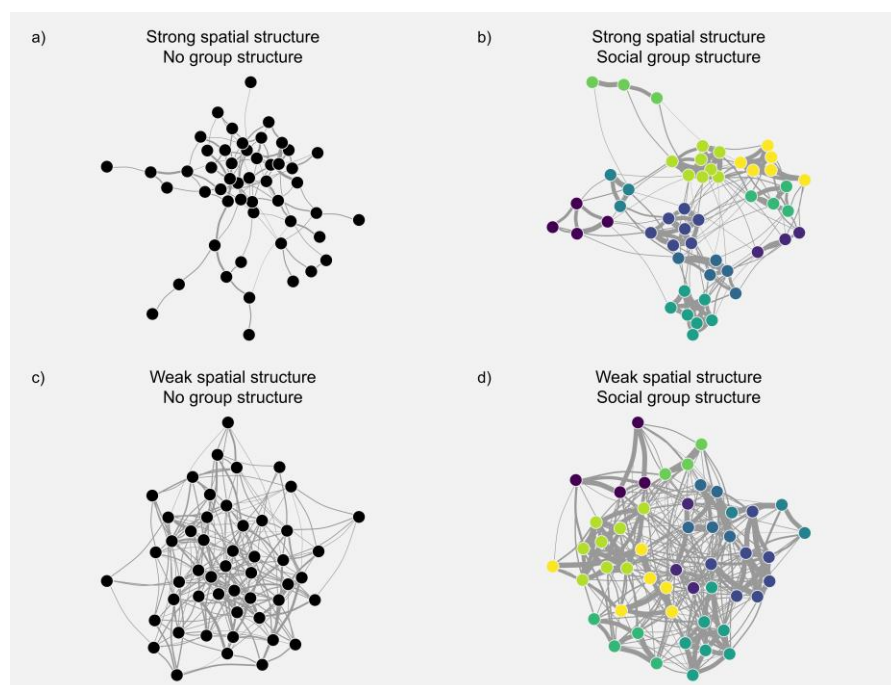
### 118 *Network features*

119           There are two core functionalities of the network features: to generate underlying social  
120 networks for the population; and to generate grouping events (interactions/associations) based  
121 on these networks. There are also two **network\_checker()** functions that quantify and  
122 visualise how well social networks derived from grouping events match the underlying network  
123 used to generate them.

124           The **network\_generation\_covariates()** function generates an underlying network  
125 structure based on group membership (as defined when generating the population), spatial  
126 locations and individual traits. Figure 1 shows examples of networks generated. Current

127 functionality is focussed on how these traits impact the probability of forming social  
128 connections within and between groups separately, thus employing a stochastic block model  
129 (Lee & Wilkinson, 2019). Therefore, genNetDem currently lacks functionality to incorporate  
130 some known social processes such as triadic closure or assortativity (e.g. females being more  
131 closely connected to other females). It is also currently not possible for interaction effects to be  
132 coded directly (e.g. if size effects on connectivity were different for males and females). Edge  
133 probabilities and edge weights are modelled independently to allow variables to explain  
134 variation in one or both of them. There is a simpler **network\_generation\_basic()** function that  
135 uses the same generative model without covariates.

136



137

138 *Figure 1. Examples of the diverse underlying social network structures it is possible to simulate*  
139 *with genNetdem. Here we explore the impact of spatial and social structure on otherwise*  
140 *similar sets of rules for the generation of social relationships. Code to replicate this figure and*  
141 *further explore network possibilities is provided in the Supplementary Materials.*

142

143           When modelling longitudinal network data, the individual social network positions could  
144 vary from relatively stable to highly dynamic (Pinter-Wollman et al., 2013). The  
145 **network\_rewire\_covariates()** function adds newly recruited and removes dead individuals  
146 from the network but also provides functionality to select probabilities that a) an individual  
147 changes some social relationships and b) each social relationship for selected individuals  
148 changes. This allows flexibility in how dynamic simulated networks are.

149           There are two functions that generate grouping events based on underlying network  
150 structure: **interaction\_generation\_simul()** and **interaction\_generation\_seq()**. The difference  
151 between them is that the former divides all individuals in the population into groups (or  
152 isolates) at each time point, while the latter independently samples one group of a defined size  
153 from the population at a time. The former is more widely useful. It uses data on individual IDs,  
154 their underlying social network and a mean group size to divide the populations into groups,  
155 with group membership being stored in a group-by-individual matrix (GBI; see (Farine, 2013)).  
156 The `n_ts` argument defines the number of times this process is repeated (i.e. number of  
157 “**behavioural timesteps**”). Assigning individuals into groups based on the underlying network  
158 can create computational challenges if unconstrained. We use a similar approach to (Evans et  
159 al., 2020), with individuals being added to groups sequentially and the probability of joining  
160 being proportional to the strength of its social relationships with existing members (see  
161 Supplementary Materials). Including a non-zero float argument means it is never impossible to  
162 add an individual to an existing group even in the absence of any social connections. While it  
163 may be tempting to reduce the float to zero this can result in it being impossible (or  
164 computationally challenging) to successfully sample all individuals into groups. However, care  
165 should be taken with particular combinations of group size distributions and underlying network



166 structures that these relaxations do not dominate interaction generation. This can be checked  
167 with the **network\_checker\_simul()** function.

168

### 169 *Observation features*

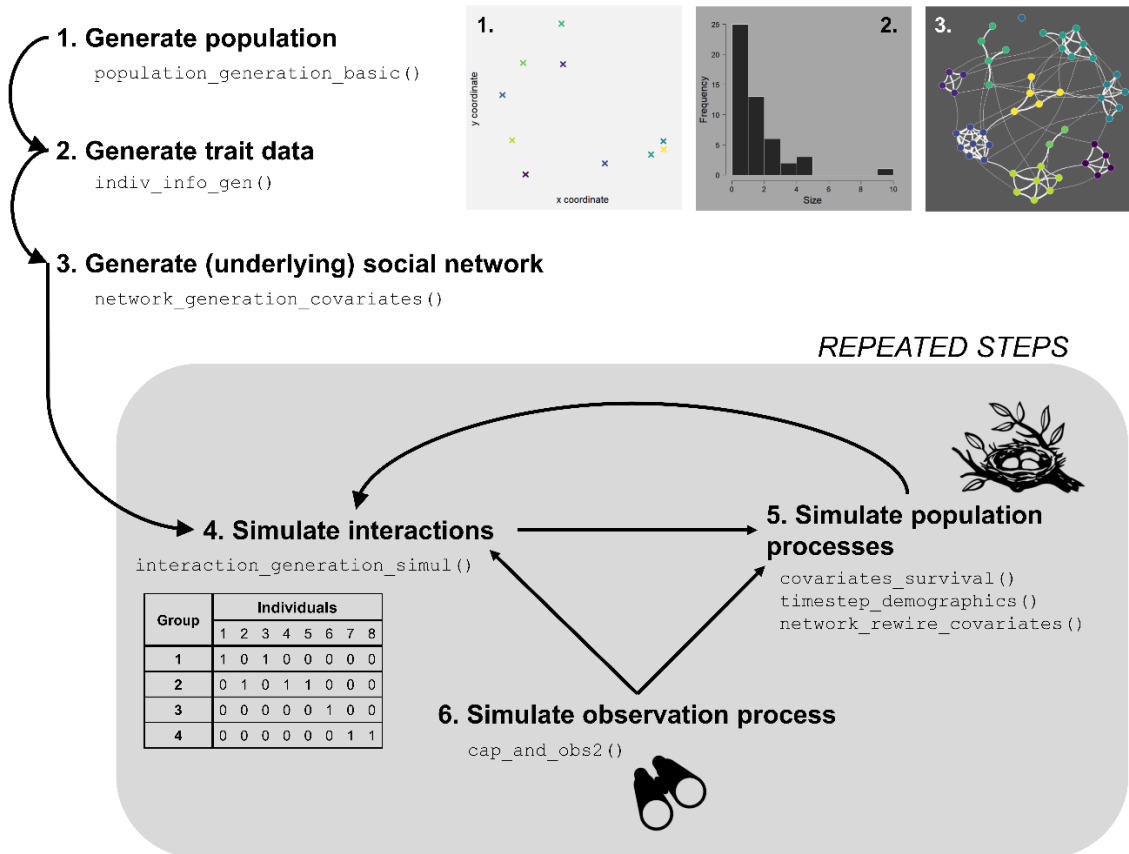
170 The observation features sample the simulated grouping events and generate data for  
171 subsequent analyses. Data is generated in a variety of formats including GBI matrices for  
172 social network analysis and classic capture-recapture formats. There are two **cap\_and\_obs()**  
173 functions that generate an observed network dataset based on the sampling strategy and  
174 design. Of the two **cap\_and\_obs2()** has greater flexibility (see Supplementary Materials).  
175 Inputs include: a) data on true grouping events (the GBI and a vector indicating which  
176 behavioural timestep each group occurred in); b) vectors indicating behavioural timesteps to  
177 be sampled, indicated separately for captures and observations; c) the success of sampling  
178 including both the proportion of groups detected and the proportion of individuals in each  
179 sampled group detected; and d) a vector indicating which (if any) individuals had been  
180 captured previously. The function returns GBIs for captured and observed groups and other  
181 related information. The **cap\_dat\_gen()** function transforms these network datasets into  
182 capture histories for both behavioural timesteps and demographic timesteps.

183

## 184 **Workflows**

185 While genNetDem is designed to be modular so that individual components can be adjusted to  
186 perform a range of tasks, many of the functions fit well within specific workflows. We illustrate  
187 one such common workflow (Fig. 2), but various other applications are demonstrated package  
188 vignettes.

189



190

191 *Figure 2. An example workflow for using genNetDem to simulate integrated network-*  
 192 *demographic datasets. This is a simplified version of the approach used in the case studies*  
 193 *with the grey box capturing a demographic timestep and step 4 involving one or more*  
 194 *behavioural timesteps. Note that multiple set of interactions can be generated prior to the*  
 195 *simulation of population processes if desired. The modular nature of the functions in the*  
 196 *package mean that different parts of this workflow can also be used independently. Further*  
 197 *usage examples are provided in the package vignettes.*

198

## 199 Case Studies

200 We use two complementary case studies to illustrate the use of genNetDem. In the first we test  
 201 how our ability to estimate the relationship between network position and survival depends on  
 202 sampling effort; whether local or global centrality affects survival; and network dynamics. We

203 compare the performance of cross-sectional versus longitudinal imputation of the network  
204 position of non-detected individuals and explore the importance of network covariance in  
205 survival probabilities. In the second we demonstrate how a researcher could use genNetDem  
206 to compare sampling designs. We test how the power to estimate relationship between  
207 network position and survival depends on how sampling effort is distributed through time. Our  
208 simulation asks the question as to whether it is better to concentrate resources into intensively  
209 monitoring more groups in fewer sampling windows or fewer groups in more sampling  
210 windows. We examine whether any differences are impacted by the proportion of individuals  
211 detected in each sampled group and the structure of the underlying social network.

212

### 213 **Methods common to both case studies**

214 In both case studies we use genNetDem to simulate survival and social interactions and then  
215 sample from them to generate capture histories. We fit hidden Markov Models to estimate  
216 survival and capture probabilities using nimble (de Valpine et al., 2017, 2022).

217

### 218 ***Data recorded from simulation runs***

219 We recorded a) the capture-recapture dataset for each demographic timestep; b) the  
220 sampled social network generated from all observed interactions within each demographic  
221 timestep; c) individual survival probabilities for each demographic timestep; and d) information  
222 on true population size and the number of individuals recorded at each demographic timestep.  
223 We estimated the network measure of interest from the sampled social network and scaled it  
224 within each demographic timestep to use as an explanatory variable.

225

### 226 ***Modelling approach***

227 We fitted Cormack-Jolly-Seber models estimating both capture and survival probabilities  
228 (Lebreton et al., 1992) and used Bayesian inference for parameter estimation. We included  
229 explanatory variables of sex and social network measure (either strength or betweenness). In  
230 each model we used weakly informative priors for all parameters (Gaussian distribution with  
231  $\mu=10$  and  $\sigma=10$  for survival-related variables, uniform distribution between 0 and 1 for capture  
232 probability). We used a single Markov chain of 3000 iterations with a burn-in of 500 and a  
233 thinning interval of 5. We confirmed that this number of iterations was typically sufficient for  
234 model convergence and an adequate effective sample size in a subset of simulations.

235

### 236 ***Analysis of simulation results***

237 From each simulation run we calculated the posterior median and standard deviation, the  
238 proportion of the posterior greater than zero, and the 89% HDI. We also calculated a binary  
239 variable indicating whether or not 0 was contained within the 89% HDI. We could then  
240 compare model performance visually and by calculating statistical clarity for positive social  
241 effects on survival as the proportion of simulation runs where 0 fell outside the 89% HDI.

242

## 243 **Case study 1: Performance of basic imputation to estimate social effects**

### 244 **on survival**

#### 245 **Specific Methods**

##### 246 ***Overview of Data generation***

247 We simulated a population of 200 individuals with no underlying group structure. Individual  
248 variation was restricted to a single two-level categorical variable – sex. The underlying social  
249 network had moderate spatial structure.

250 We simulated the behaviour and survival of individuals over 10 demographic timesteps  
251 (over which survival was simulated), each containing five behavioural timesteps (at which  
252 individuals were organised into grouping events). Grouping events had a mean size of two  
253 individuals (many events were dyadic and individuals were frequently alone) to capture a  
254 situation where a species rarely occurs in large aggregations. Survival probability depended on  
255 sex (moderate effect of 0.5 on a logit scale) and position in the social network calculated from  
256 grouping events (see below) with a baseline survival probability of 0.8 in females. We assumed  
257 no recruitment into the population (i.e. the population declined over the simulation).

258 We assumed that all individual in the population were marked or individually-identifiable  
259 prior to the start of the study. Captures and/or observations (which were functionally equivalent  
260 as all individuals were marked) took place in all behavioural timesteps (50 in total). Each group  
261 had either a 25%, 50% or 75% percent chance of being detected (parameter varied between  
262 simulation runs) with the detection probability of an individual in a detected group fixed at 0.9.

263

### 264 ***Simulation structure***

265 In total we generated 3240 simulated datasets, varying five parameters that influenced network  
266 dynamics (one parameter), network effects on survival (three parameters) and sampling (one  
267 parameter).

- 268 • Network dynamics: we varied the probability that an individual's existing connections in  
269 the underlying social network were rewired with values set at 0 (no rewiring), 0.1 and  
270 0.5. The per-edge probability that an individual changed its connections was 0.5.
- 271 • Network effects on survival: a) we varied the network measure that influenced survival  
272 to be either strength (local measure; sum of weighted connections) or betweenness  
273 (global measure; number of shortest paths passing through an individual); b) we varied

274 the effect size to be 0 (no effect), 0.4 (moderate effect) or 0.8 (strong effect); c) we  
275 altered covariance of individual survival within the network to be negative (strongly  
276 connected individuals have more dissimilar survival probabilities), neutral or positive  
277 (strongly connected individuals have more similar survival probabilities).

- 278 • Sampling: we varied the probability of sampling (either capturing or observing) a group  
279 at each behavioural timestep to be 0.25, 0.5 or 0.75.

280 For each combination of parameters (162) we ran 20 replicates.

281

### 282 ***Data recorded***

283 In addition to the four types of data described in the combined methods, we also  
284 recorded the full social network generated from all interactions within each demographic  
285 timestep (including those not observed). We estimated the network measure of interest from  
286 these full networks and scaled them within each demographic timestep as for measures from  
287 partial networks.

288

### 289 ***Model-fitting***

290 From each simulation run we fitted four model versions (see combined methods for  
291 details on model-fitting). The four versions differed in: a) using the measure from the sampled  
292 network and a longitudinal approach for imputing non-observed individuals; b) using the  
293 measure from the sampled network and cross-sectional imputation; c) using the measure from  
294 the full (unobserved) network and longitudinal imputation; and d) using the measure from the  
295 full network and cross-sectional imputation. For cross-sectional imputation missing values  
296 were estimated using the mean and variance of the (scaled) focal network measure from a  
297 given demographic timestep. For longitudinal imputation missing values were estimated using

298 the mean and variance of the focal network measure for each individual where possible and  
299 the overall mean and variance when not.

300

### 301 ***Analysis of simulation results***

302 Prior to the general analysis outlined above, we assessed whether the model had converged  
303 using the posterior median and standard deviation of its estimate for the social effect on  
304 survival. We used k means clustering to identify groups of simulation runs where the model  
305 was unlikely to have converged. We used k=6 clusters and retained 3 out of 6 of these clusters  
306 based on the elbow method and visual inspection of the output (Fig. S1). This method  
307 identified ~2.5% of models had likely not converged.

308 To compare the success of models that used network measures calculated from the  
309 partial versus full network we calculated the earth mover's distance (EMD) of the posterior  
310 distributions (Touzalin et al., 2022) for the parameter of interest from relevant pairs of models.  
311 (i.e. we calculated the EMD for model versions using the full and partial network together with  
312 longitudinal imputation and also the EMD for the model versions using the full and partial  
313 network together with cross-sectional imputation). EMDs provide a measure of overlap of the  
314 posterior distributions.

315

### 316 **Results and Discussion**

317 Overall, we show it is possible to estimate social effects survival from partial networks, albeit  
318 with substantial limitations in power (Fig. 3, Table 1, Table S1). Estimates of social effects on  
319 survival were downward biased meaning that statistical power was limited and only stronger  
320 social effects on survival are likely to be detected. Sampling effort was particularly important  
321 and interacted with how imputation was conducted in determining how well models converged

322 and biases in parameter estimates when they did. Estimates of other parameters were  
323 unaffected.

324 Previous research has demonstrated that network measures from sampled, partial  
325 networks are correlated with those in the full, unobserved network but that these correlations  
326 vary depending on the proportion sampled and network measure calculated (Silk et al., 2015;  
327 Smith & Moody, 2013). Further, the regression slope is rarely 1:1 indicating values for  
328 measures estimated are not perfectly accurate (Silk et al., 2015). This likely explains many of  
329 our results showing the difficulty of detecting social effects on survival in the absence of  
330 network imputation or the use of measures from independently (and better) sampled social  
331 networks).

332

### 333 ***Network variable and covariance structure***

334 When we compared models that used network measures from the full and partial networks we  
335 found downward-biased parameter estimates and reduced statistical clarity of results when  
336 partial network measures were used (Table 1, Fig. 3). These patterns were more striking when  
337 survival was related to a global measure of centrality (betweenness) than a local measure of  
338 centrality (strength). We found that including positive or negative covariance in survival  
339 probabilities related to social network structure had little effect on estimation or power in the  
340 contexts simulated (Fig. S3, Tables S5-6).

341 These results fit well within the literature on how missing individuals impact the  
342 conclusions of social network analysis, with previous studies showing that global estimates of  
343 social centrality (such as betweenness) from partial networks are less well correlated than  
344 measures of local centrality (such as strength) with equivalent measures from the full network  
345 (Silk et al., 2015). While for strength in particular downward-biased parameter estimates in



346 combination with maintained statistical power could also be related to measures of strength  
347 being lower in the smaller, sampled network (Silk et al., 2015), this should be controlled for by  
348 scaling network measures before using them in the model. The lack of a clear effect of network  
349 covariance is somewhat surprising. These results are promising in suggesting that this may  
350 present a more limited issue in this context than often considered (e.g. (Croft et al., 2011;  
351 Farine & Carter, 2020; Silk et al., 2017)). However, the importance of covariance likely  
352 depends substantially on network structure and density, so it would be unwise to generalise  
353 these patterns without further work focussed specifically on this question.

354

### 355 ***Sampling effort, imputation approach and network dynamics***

356 Lower sampling effort was typically associated with both a) reduced likelihood of model  
357 convergence (Table 2, Table S2), and b) downward-biased parameter estimates (Fig. 3).  
358 However, the nature of these relationships depended on the imputation approach selected  
359 (Table 1, Fig. 3), with the performance of different imputation approaches largely independent  
360 of network dynamics (Tables S3-4, Fig. S2).

361 Models were much less likely to converge when sampling effort was low (25% group  
362 capture probability), betweenness centrality from partial networks was used as an explanatory  
363 variable and cross-sectional imputation was used to infer missing values (Table 2, Table S2).  
364 Even when 50% of groups were sampled in these situations there was still a reduction in  
365 convergence rate. Note that this was apparent regardless of whether betweenness centrality  
366 had a positive or no effect on survival probability. Any other changes in the likelihood of model  
367 convergence were of much smaller magnitude, but generally occurred when sampling effort  
368 was low (and measures from partial networks were used).

369 With cross-sectional imputation and use of measures from the full network, estimation of  
370 social effects on survival were largely independent of sampling effort in the contexts examined.  
371 With longitudinal imputation there was some reduction in estimates of the social effect on  
372 survival with low levels of sampling (25% groups sampled). However, both cross-sectional and  
373 longitudinal imputation demonstrated similar relationships between sampling effort and  
374 statistical power (Table 1, Table S1), indicating that posterior distributions had higher variance  
375 when cross-sectional imputation was used.

376 When measures from the partial network were used instead, there was a much more  
377 substantial reduction in both parameter estimates and statistical power apparent even for  
378 higher sampling efforts (Fig. 3, Table 1). Reductions in parameter estimates were more  
379 substantial and remained linear when longitudinal imputation was used, instead flattening out  
380 for cross-sectional imputation so that the difference between 25% and 50% of groups being  
381 sampled was less than the difference between 50% and 75% (Fig. 3). However, similarly to the  
382 pattern for full network measures, this was not reflected in changes to statistical power which  
383 were broadly equivalent for both, indicating a less precise posterior distribution for cross-  
384 sectional imputation. These differences between cross-sectional and longitudinal imputation  
385 changed how EMDs calculated for the differences between posteriors from the full network and  
386 partial network model fits depended on sampling effort (Fig. 4). For cross-sectional imputation  
387 EMDs were highest for low sampling effort ( $p=0.25$ ) while for longitudinal imputation they  
388 peaked at intermediate sampling effort ( $p=0.5$ ). However, in general EMDs were higher for  
389 cross-sectional than longitudinal imputation.

390 Our results show that when social networks are constructed based on the same co-  
391 capture data used to estimate survival, even relatively small drops in sampling effort can lead  
392 to downward biases in parameter estimates and statistical power. While this pattern was

393 especially strong when global measures of centrality such as betweenness explain variation in  
394 survival probability as expected from previous literature (Silk et al., 2015; Smith & Moody,  
395 2013), it was also apparent when strength was associated with survival instead. However, in  
396 this latter case underestimated social effects on survival only caused substantial reductions in  
397 statistical power with very low sampling effort. Consequently, our results fit broadly within the  
398 existing literature where low sampling effort has a greater impact on global measures of  
399 centrality but suggest that missing a high proportion of interaction events leads to wider  
400 problems with subsequent statistical analyses. This was particularly apparent when cross-  
401 sectional imputation was used to estimate missing values for betweenness centrality when  
402 there was a substantial drop-off in how likely models were to converge. Combined with cross-  
403 sectional imputation generating less precise posteriors, this suggests that longitudinal  
404 imputation is a more stable option of the two, although it does lead to greater downward bias in  
405 estimates of social effects on survival. However, neither imputation approach performed well,  
406 highlighting the value of extending network imputation approaches (R. W. Krause et al., 2018,  
407 2020; Young et al., 2020) within capture-recapture models. A good example is provided by  
408 (Clements et al., 2022), who estimate not only the network itself but also the underlying  
409 behaviours that generate the network structure within a Cormack-Jolly-Seber model. While this  
410 was done in the context of a simulation study, and so involved fitting the data-generating  
411 model, it does show the potential of network imputation to improve the accuracy of estimates  
412 of social effects on survival.

413

#### 414 ***Estimates of other parameters***

415 Estimates for other parameter values were unaffected by social effects on survival, use of  
416 measures from full or partial networks or imputation strategy (Figs. S4-6).

417

418

419 *Table 1. Proportion of simulation runs where  $\theta$  falls outside the 89% HDI for different*  
 420 *parameter combinations. M1: partial network - cross-sectional imputation; M2: partial network*  
 421 *- longitudinal imputation; M3: full network - cross-sectional imputation; M4: full network -*  
 422 *longitudinal imputation.*

<b>Network measure</b>	<b>True effect</b>	<b>Model</b>	<b>Group capture probability</b>	<b>Detection rate</b>
Strength	0.4	M1	0.50	0.75
Strength	0.4	M2	0.50	0.72
Strength	0.4	M3	0.50	0.98
Strength	0.4	M4	0.50	0.98
Betweenness	0.4	M1	0.50	0.26
Betweenness	0.4	M2	0.50	0.25
Betweenness	0.4	M3	0.50	0.94
Betweenness	0.4	M4	0.50	0.94
Strength	0.8	M1	0.50	0.99
Strength	0.8	M2	0.50	0.99
Strength	0.8	M3	0.50	1.00
Strength	0.8	M4	0.50	1.00
Betweenness	0.8	M1	0.50	0.70
Betweenness	0.8	M2	0.50	0.68
Betweenness	0.8	M3	0.50	1.00
Betweenness	0.8	M4	0.50	1.00

423

424

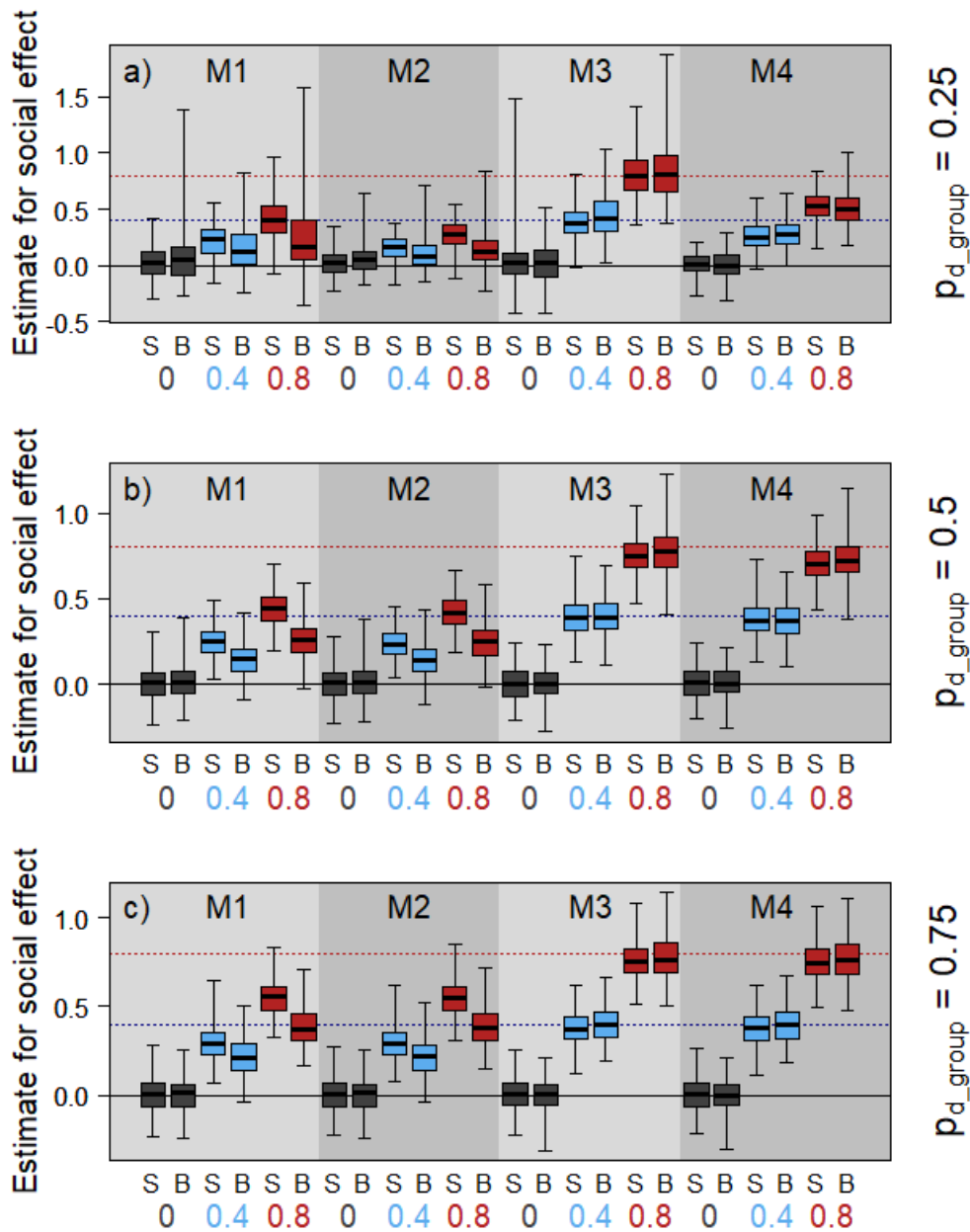
425

426 *Table 2. Convergence rates of models using different imputation approaches for various*  
 427 *parameter combinations. M1: partial network - cross-sectional imputation; M2: partial network*  
 428 *– longitudinal imputation; M3: full network – cross-sectional imputation; M4: full network –*  
 429 *longitudinal imputation.*

<b>Network measure</b>	<b>Model</b>	<b>Group capture probability</b>	<b>Convergence rate</b>
Strength	M1	0.25	0.98
Strength	M1	0.50	0.99
Strength	M1	0.75	1.00
Strength	M2	0.25	1.00
Strength	M2	0.50	1.00
Strength	M2	0.75	1.00
Strength	M3	0.25	1.00
Strength	M3	0.50	1.00
Strength	M3	0.75	1.00
Strength	M4	0.25	1.00
Strength	M4	0.50	1.00
Strength	M4	0.75	1.00
Betweenness	M1	0.25	0.65
Betweenness	M1	0.50	0.94
Betweenness	M1	0.75	0.98
Betweenness	M2	0.25	0.96
Betweenness	M2	0.50	1.00
Betweenness	M2	0.75	1.00
Betweenness	M3	0.25	0.96
Betweenness	M3	0.50	0.97
Betweenness	M3	0.75	0.99
Betweenness	M4	0.25	1.00
Betweenness	M4	0.50	1.00
Betweenness	M4	0.75	1.00

430

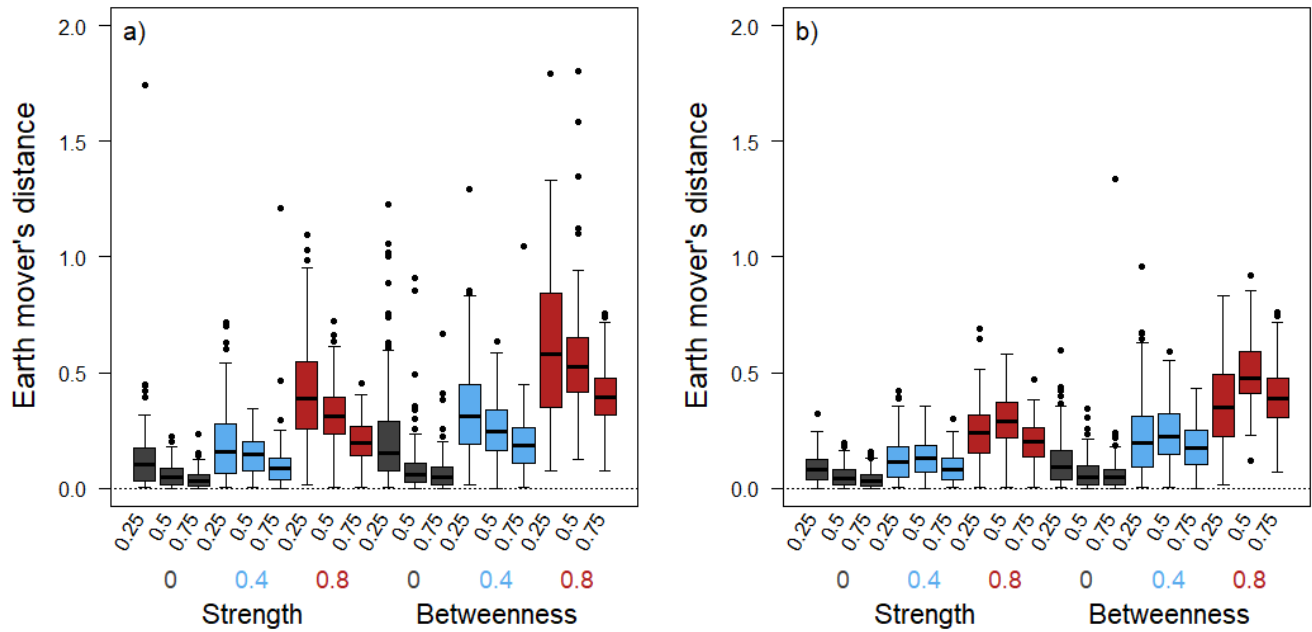
431



433

434 *Figure 3. Distribution of posterior medians for the social effect of survival for different*  
 435 *combinations of model (shaded polygons; M1: partial network - cross-sectional imputation; M2:*  
 436 *partial network – longitudinal imputation; M3: full network – cross-sectional imputation; M4: full*  
 437 *network – longitudinal imputation), network measure (S=Strength; B=Betweenness) and true*  
 438 *effect size (box colour) when a) 25% of groups are sampled, b) 50% of groups are sampled*

439 and c) 75% of groups are sampled. The solid central line represents the median, boxes the  
440 interquartile range and whiskers the full range of values.



442

443 *Figure 4. Earth mover's distances demonstrating the similarity of the posterior distributions for*  
 444 *model estimates of the social effect on survival between the model using full and partial*  
 445 *network measures for a) cross-sectional and b) longitudinal imputation. The solid central line*  
 446 *represents the median, boxes the interquartile range and whiskers extend to 1.5x the*  
 447 *interquartile range. We show the distributions for different combinations of group capture*  
 448 *probability (0.25, 0.5 or 0.75), true simulated effect size (grey for 0, blue for 0.4 and red for 0.8)*  
 449 *and the network measure influencing survival probability (strength vs betweenness). To aid*  
 450 *visualisation we have excluded 19 outlying points with  $EMD > 2$  for panel a).*



451

## 452 **Case study 2: Effective sampling strategies to estimate social effects on** 453 **survival**

### 454 **Specific Methods**

#### 455 ***Overview of Data generation***

456 We simulated a population of 200 individuals with either a) no underlying group structure; or b)  
457 divided into 20 groups. Individual variation in the population was restricted to a single two-level  
458 categorical variable – sex. Underlying network structure depended on parameter choice (see  
459 below).

460 We simulated the behaviour and survival of individuals over 10 demographic timesteps,  
461 each of which contained 20 behavioural timesteps. As previously, grouping events had a mean  
462 size of two individuals. The survival probability of each individual depended on its sex (fixed  
463 effect of 0.5 on a logit scale) and position in the social network calculated from grouping events  
464 (see below) with a baseline survival probability of 0.8 in females. In this case study, there was  
465 recruitment into the population over time (i.e. the population stayed roughly constant over each  
466 simulation). There was a 10% chance that a surviving individual rewired its underlying social  
467 connections and if it did each connection had a 50% chance of changing.

468 The population was initially unmarked. Captures only occurred in the first behavioural  
469 timestep of each demographic timestep with 90% of groups sampled and a 0.9 probability of  
470 individuals in a sampled group being detected. Sampling design and effort for subsequent  
471 observations depended on parameter choice (see below).

472

473 ***Simulation structure***

474 In total we generated 2880 simulated datasets, varying five parameters that influenced network  
475 structure (one parameter), network effects on survival (two parameters) and sampling  
476 effort/design (two parameters).

- 477 • Network structure: we varied underlying network structure so that either a) there was no  
478 group structure and moderate spatial structure driving the probability and weight of  
479 edges; or b) the population was divided into 20 groups with the probability of a within-  
480 group connection of 0.5 and within-group connection weights having a mean of 0.5  
481 (versus a baseline of 0.2 and 0.25 respectively for between-group connections prior to  
482 adjusting for distance effects).
- 483 • Network effects on survival: a) we varied the network measure that influenced survival  
484 to be either strength or betweenness; and b) we varied the effect size to be 0 (no  
485 effect), 0.4 (moderate effect) or 0.8 (strong effect).
- 486 • Sampling: a) we varied sampling design so that the probability of observing a group  
487 within a sampled behavioural timestep covaried with the number of behavioural  
488 timesteps sampled in each demographic timestep resulting in (approximately)  
489 equivalent sampling effort being divided over the full demographic timestep. The  
490 probability of capturing a group was either 0.1, 0.2, 0.4 or 1, with the number of  
491 behavioural timesteps observed being 19, 10, 5 or 2; b) we varied the probability of an  
492 individual being observed in a sampled group to be either 0.5, 0.75 or 1.

493 For each combination of parameters (144) we ran 20 replicates.

494

495 ***Model-fitting***

496 Unlike Case study 1 each Cormack-Jolly-Seber model was conditioned on first capture  
497 (as individuals were not assumed to have been captured previously).

498

499 **Results and Discussion**

500 Overall, survival models performed adequately in detecting social effects on survival (Table 3,  
501 Fig. 5, Tables S7-9, Fig. S7). When we simulated positive effects of network centrality on  
502 survival probabilities model estimates reflected this, although were substantial underestimates,  
503 especially with only moderate social effects on survival. These results support those from Case  
504 Study 1 indicating that it is possible to estimate social effects on survival, but that statistical  
505 power is limited with the presence of non-detected individuals and/or when many interaction  
506 events are unobserved. More encouragingly we show that for two very different social network  
507 structures there is little evidence for strong bias or elevated false positive rates when there is  
508 no social effect on survival.

509

510 ***Network variable***

511 Our statistical models were better able to detect the effect of strength (local centrality measure)  
512 than betweenness (global centrality measure) on survival probabilities. While, the effect size  
513 was underestimated for both measures, this bias was much greater for betweenness centrality  
514 (Fig. 5), and results were more frequently statistically unclear (Table 3). The results here  
515 support those from Case Study 1 and the existing literature (Silk et al., 2015; Smith & Moody,  
516 2013) in highlighting that global measures of network position are more susceptible to  
517 sampling effects than local measures.

518

519 ***Sampling design***

520 There was no clear effect of how groups were sampled within each demographic timestep on  
521 estimates of social effects on survival (Fig. 5). Unsurprisingly, probability of observing  
522 individuals within groups did have some effect, with less downward-biased parameter  
523 estimates and more statistical power when sampling within groups was more complete (Fig.  
524 S7, Tables S7-8), as would be expected.

525 Lower observation success within sampled groups leading to reduced model  
526 performance is unsurprising as it leads to missing edges in the sampled network, reducing its  
527 correlation with the true (unobserved) network. This finding supports related work focussed on  
528 calculating network measures (e.g. (Franks et al., 2010)). (Franks et al., 2010) also tentatively  
529 recommended that more censuses (behavioural timesteps sampled in our case) were  
530 preferable than ensuring a high proportion of interaction events sampled in each census for  
531 calculating weighted measures of centrality. However, we found no clear evidence that this  
532 extended to our survival analysis, where there were only small differences in model  
533 performance and no clear overall trend. It should be noted, however, that the simulation  
534 architecture differed between the two papers.

535

536 ***Social structure***

537 Social structure had a small effect on the ability to detect social effect on survival, with some  
538 differences in statistical power between the two structures investigated. While there were  
539 minimal differences in posterior medians (Fig. 5), results tended to be statistically clearer when  
540 there was no underlying group structure than when the population was divided into 20 groups  
541 (Table 3). Previous studies of sampling in social networks have rarely considered the types of  
542 modular social structures common for group-living animal populations (Silk, 2018). The slight

543 negative impact of this group-structure on our ability to detect social effects on survival  
544 perhaps suggests that the correlation between network measures calculated in the sampled  
545 and full networks is weaker in these types of networks.

546

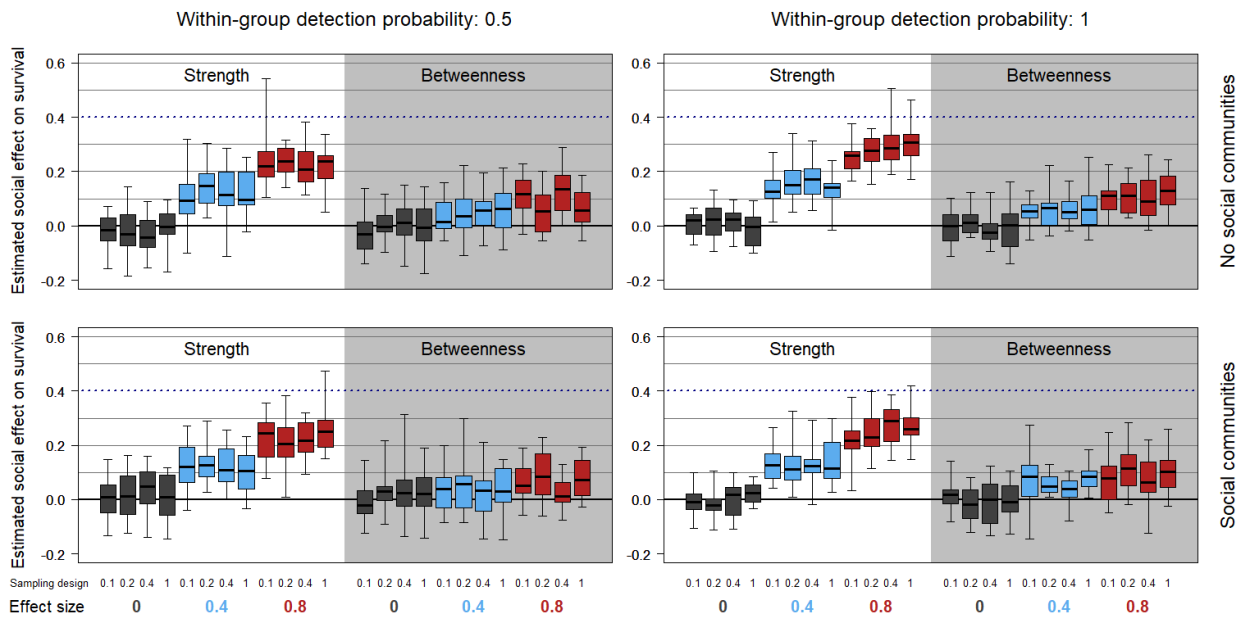
547 ***Estimates of other parameters***

548 Estimates of other parameters were largely unaffected by social effects or sampling design.  
549 Strong social effects on survival were associated with slightly lower estimates of mean survival  
550 probability, but these differences were caused by differences in simulated survival probabilities  
551 rather than model performance (Figs. S8-10). While limited in scope these results provide  
552 evidence that including social effects on survival in demographic models is unlikely to impact  
553 other parameter estimates substantially (see also (Clements et al., 2022)).

554

556 *Table 3. Proportion of simulation runs where  $\theta$  falls outside the 89% HDI for different*  
 557 *parameter combinations with the probability of within-group detection fixed at 1.*

Network measure	True effect	Social structure	Sampling design	Detection rate
Strength	0.4	Communities	0.1	0.45
Strength	0.4	No communities	0.1	0.50
Strength	0.4	Communities	0.2	0.40
Strength	0.4	No communities	0.2	0.70
Strength	0.4	Communities	0.4	0.55
Strength	0.4	No communities	0.4	0.70
Strength	0.4	Communities	1.0	0.50
Strength	0.4	No communities	1.0	0.60
Betweenness	0.4	Communities	0.1	0.20
Betweenness	0.4	No communities	0.1	0.10
Betweenness	0.4	Communities	0.2	0.00
Betweenness	0.4	No communities	0.2	0.10
Betweenness	0.4	Communities	0.4	0.00
Betweenness	0.4	No communities	0.4	0.15
Betweenness	0.4	Communities	1.0	0.10
Betweenness	0.4	No communities	1.0	0.15
Strength	0.8	Communities	0.1	0.95
Strength	0.8	No communities	0.1	1.00
Strength	0.8	Communities	0.2	1.00
Strength	0.8	No communities	0.2	1.00
Strength	0.8	Communities	0.4	0.95
Strength	0.8	No communities	0.4	1.00
Strength	0.8	Communities	1.0	1.00
Strength	0.8	No communities	1.0	1.00
Betweenness	0.8	Communities	0.1	0.15
Betweenness	0.8	No communities	0.1	0.35
Betweenness	0.8	Communities	0.2	0.45
Betweenness	0.8	No communities	0.2	0.45
Betweenness	0.8	Communities	0.4	0.25
Betweenness	0.8	No communities	0.4	0.40
Betweenness	0.8	Communities	1.0	0.40
Betweenness	0.8	No communities	1.0	0.50



560

561 *Figure 5. The impacts of sampling design (within-plot: sets of boxes of the same colour),*  
 562 *within-group detection probability (columns) and social structure (rows) on Cormack-Jolly-*  
 563 *Seber estimates of social effects on survival probability for a range of simulated effect sizes*  
 564 *(colours of boxes). Boxplots show the distribution of posterior medians from multiple simulation*  
 565 *runs with the solid line the median, boxes the interquartile range and whiskers the full range of*  
 566 *values. We illustrate contexts in which a local measure of centrality (strength) and global*  
 567 *measure of centrality (betweenness) are used as explanatory variables. The blue-dotted line*  
 568 *indicates the accurate parameter estimate when the true effect size is 0.4 (the equivalent line*  
 569 *for 0.8 is not illustrated).*

570

571

572

## 574 **Future steps**

575 With the two case studies presented we can only scratch the surface of the potential of  
576 genNetDem as a methodological tool for animal social network analyses. Below we highlight  
577 some logical next steps for methodological studies on this topic, focussing on the integration of  
578 social networks and demography.

579         First, while we demonstrated the capacity for genNetDem to generate diverse social  
580 structures (Fig. 1), this was only a partial focus of our results. Animal social systems vary  
581 widely, and while sampling effects optimal sampling strategies are likely to vary with social  
582 structure (Clements et al., 2022; Silk, 2018; Sunga et al., 2021), this has remained  
583 understudied. Similarly, while we varied network dynamics in our simulations, the network of  
584 every individual was drawn from the same probability distribution. Incorporating greater trait-  
585 based or individual variation in network position would likely influence conclusions drawn about  
586 imputation approaches, for example.

587         Second, it is clear that simple approaches to imputing missing network measures are  
588 only partially successful; while they successfully generate qualitatively correct results,  
589 parameter estimates for social effects on survival are underestimated. While developing more  
590 sophisticated approaches to impute values for network measures may help, exploiting recent  
591 developments in network imputation (R. W. Krause et al., 2018, 2020; Young et al., 2020) are  
592 likely to have the greatest success. The adaptation of these novel approaches for behavioural  
593 ecology, and specifically within this capture-recapture modelling framework is a key challenge.  
594 (Gimenez et al., 2019) applied basic network imputation to study the social structure of  
595 Commerson's dolphin *Cephalorhynchus commersonii*. Similarly, (Clements et al., 2022) et al.  
596 included estimation of network structure within a Cormack-Jolly-Seber model to improve



597 estimation of social effects on survival. However, the latter approach used a rather basic  
598 generative model for the latent network structure that could be improved on or adjusted for  
599 researchers working in different contexts. Consequently, extending these approaches to  
600 incorporate more sophisticated social network models as well as to open populations is a key  
601 priority.

602 Third, to keep our case studies accessible we examined social effects only in Cormack-  
603 Jolly-Seber models to estimate survival probability. (Clements et al., 2022) highlighted the  
604 potential value of incorporating social networks within integrated population models (IPMs),  
605 where different data sources could also be used to inform network structure itself. However,  
606 especially with improvements to imputation of latent network structures, there is also great  
607 potential to incorporate network effects within multi-state models more generally. Given the  
608 central role of social behaviour in mediating interactions between infectious disease dynamics  
609 and demographic processes (Silk et al., 2019; Silk & Fefferman, 2021), extending multistate  
610 models to incorporate social network structure in this way could provide important new insights  
611 into wildlife disease ecology, to provide just one example. genNetDem can provide an ideal  
612 sandbox to refine these models for application to wild systems.

613 Finally, we focus here on dyadic social networks, however many of the social  
614 interactions studied are non-dyadic and may include higher-order interactions (Battiston et al.,  
615 2021; Greening Jr et al., 2015). While there has been limited focus on higher-order  
616 interactions in animal societies (Musciotto et al., 2022), theory suggests they will impact  
617 infectious disease transmission and social contagions (Battiston et al., 2021; Iacopini et al.,  
618 2022; Noonan & Lambiotte, 2021) among other ecological and evolutionary processes.  
619 Therefore, expanding some of the developments here beyond dyadic networks to consider  
620 higher-order effects on survival and imputation of hyperedges (social connections between

621 more than two individuals) will likely represent valuable developments. Because it generates  
622 GBIs that incorporate interactions/associations between more than two individuals genNetDem  
623 is an ideal starting point for methodological research testing higher-order methods in animal  
624 societies.

625

## 626 **Conclusions**

627 We introduce the R package genNetDem as a flexible tool for simulating combined social and  
628 demographic datasets. While we focus on the integration of social network and demographic  
629 models, the modular design of the package allows it to be an equally powerful tool for  
630 generating social network or capture-recapture datasets in their own right. It therefore provides  
631 a general tool for researchers interested in testing key methodological considerations in animal  
632 social network studies, especially as the field moves towards longitudinal analysis. It also helps  
633 researchers wishing to test the power of specific analyses or sampling designs in their own  
634 study systems.

635

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640 discussions related to this work.

641

## 642 **Data/Code availability**

643 Data and code are available at <https://github.com/matthewsilks/NETDEM> and will be archived  
644 on acceptance.

645

## 646 **References**

647 Battiston, F., Amico, E., Barrat, A., Bianconi, G., Ferraz de Arruda, G., Franceschiello, B.,  
648 Iacopini, I., Kéfi, S., Latora, V., & Moreno, Y. (2021). The physics of higher-order  
649 interactions in complex systems. *Nature Physics*, *17*(10), 1093–1098.

650 Blumstein, D.T., Williams, D.M., Lim, A.N., Kroeger, S., & Martin, J.G.A. (2018). Strong social  
651 relationships are associated with decreased longevity in a facultatively social mammal.  
652 *Proc. R. Soc. B*, *285*(1871), 20171934.

653 Butts, C.T. (2014). *Package 'sna.'*

654 Clements, S.J., Zhao, Q., Silk, M.J., Hodgson, D.J., & Weegman, M.D. (2022). Modelling  
655 associations between animal social structure and demography. *Animal Behaviour*, *188*,  
656 51–63.

657 Croft, D.P., Madden, J.R., Franks, D.W., & James, R. (2011). Hypothesis testing in animal  
658 social networks. *Trends in Ecology & Evolution*, *26*(10), 502–507.

659 Csardi, G., & Nepusz, T. (2006). The igraph software package for complex network research.  
660 *InterJournal, Complex Systems*, *1695*(5), 1–9.

661 de Valpine, P., Paciorek, C., Turek, D., Michaud, N., Anderson-Bergman, C., Obermeyer, F.,  
662 Wehrhahn Cortes, C., Rodriguez, A., Temple Lang, D., & Paganin, S. (2022). *{NIMBLE}*:  
663 *{MCMC}*, *Particle Filtering*, and *Programmable Hierarchical Modeling*.  
664 <https://doi.org/10.5281/zenodo.1211190>

665 de Valpine, P., Turek, D., Paciorek, C., Anderson-Bergman, C., Temple Lang, D., & Bodik, R.  
666 (2017). Programming with models: writing statistical algorithms for general model  
667 structures with *{NIMBLE}*. *Journal of Computational and Graphical Statistics*, *26*(2), 403–  
668 413. <https://doi.org/10.1080/10618600.2016.1172487>

669 Ellis, S., Franks, D.W., Natrass, S., Cant, M.A., Weiss, M.N., Giles, D., Balcomb, K.C., & Croft,  
670 D.P. (2017). Mortality risk and social network position in resident killer whales: sex  
671 differences and the importance of resource abundance. *Proc. R. Soc. B*, 284(1865),  
672 20171313.

673 Evans, J., Fisher, D.N., & Silk, M.J. (2020). The performance of permutations and exponential  
674 random graph models when analyzing animal networks. *Behavioral Ecology*, 31(5), 1266-  
675 1276.

676 Farine, D.R. (2013). Animal social network inference and permutations for ecologists in R  
677 using asnipe. *Methods in Ecology and Evolution*, 4(12), 1187–1194.

678 Farine, D.R., & Carter, G.G. (2020). Permutation tests for hypothesis testing with animal social  
679 data: problems and potential solutions. *BioRxiv*.

680 Fisher, D.N., Ilany, A., Silk, M.J., & Tregenza, T. (2017). Analysing animal social network  
681 dynamics: the potential of stochastic actor-oriented models. *Journal of Animal Ecology*,  
682 86(2). <https://doi.org/10.1111/1365-2656.12630>

683 Formica, V.A., Wood, C.W., Larsen, W.B., Butterfield, R.E., Augat, M.E., Hougen, H.Y., &  
684 Brodie, E.D. (2012). Fitness consequences of social network position in a wild population  
685 of forked fungus beetles (*Bolitotherus cornutus*). *Journal of Evolutionary Biology*, 25(1),  
686 130–137.

687 Franks, D.W., Ruxton, G.D., & James, R. (2010). Sampling animal association networks with  
688 the gambit of the group. *Behavioral Ecology and Sociobiology*, 64(3), 493–503.

689 Gimenez, O., Lebreton, J.-D., Gaillard, J.-M., Choquet, R., & Pradel, R. (2012). Estimating  
690 demographic parameters using hidden process dynamic models. *Theoretical Population*  
691 *Biology*, 82(4), 307–316. <https://doi.org/https://doi.org/10.1016/j.tpb.2012.02.001>

692 Gimenez, O., Mansilla, L., Klaich, M.J., Coscarella, M.A., Pedraza, S.N., & Crespo, E.A.

693 (2019). Inferring animal social networks with imperfect detection. *Ecological Modelling*,  
694 401, 69–74.

695 Greening Jr, B.R., Pinter-Wollman, N., & Fefferman, N.H. (2015). Higher-Order Interactions:  
696 Understanding the knowledge capacity of social groups using simplicial sets. *Current*  
697 *Zoology*, 61(1), 114–127. <https://doi.org/10.1093/czoolo/61.1.114>

698 Iacopini, I., Petri, G., Baronchelli, A., & Barrat, A. (2022). Group interactions modulate critical  
699 mass dynamics in social convention. *Communications Physics*, 5(1), 1–10.

700 Krause, J., James, R., Franks, D.W., & Croft, D.P. (2014). *Animal social networks*. Oxford  
701 University Press.

702 Krause, R.W., Huisman, M., & Snijders, T.A.B. (2018). Multiple imputation for longitudinal  
703 network data. *Statistica Applicata-Italian Journal of Applied Statistics*, 1, 33–57.

704 Krause, R.W., Huisman, M., Steglich, C., & Snijders, T. (2020). Missing data in cross-sectional  
705 networks—An extensive comparison of missing data treatment methods. *Social Networks*,  
706 62, 99–112.

707 Lebreton, J.-D., Burnham, K.P., Clobert, J., & Anderson, D.R. (1992). Modeling survival and  
708 testing biological hypotheses using marked animals: a unified approach with case studies.  
709 *Ecological Monographs*, 62(1), 67–118.

710 Lee, C., & Wilkinson, D.J. (2019). A review of stochastic block models and extensions for  
711 graph clustering. *Applied Network Science*, 4(1), 122. [https://doi.org/10.1007/s41109-019-](https://doi.org/10.1007/s41109-019-0232-2)  
712 0232-2

713 McClintock, B.T., Langrock, R., Gimenez, O., Cam, E., Borchers, D.L., Glennie, R., &  
714 Patterson, T.A. (2020). Uncovering ecological state dynamics with hidden Markov models.  
715 *Ecology Letters*, 23(12), 1878–1903.

716 Musciotto, F., Papageorgiou, D., Battiston, F., & Farine, D.R. (2022). Beyond the dyad:

717 uncovering higher-order structure within cohesive animal groups. *BioRxiv*.

718 Noonan, J., & Lambiotte, R. (2021). Dynamics of majority rule on hypergraphs. *Physical*  
719 *Review E*, 104(2), 24316.

720 Oh, K.P., & Badyaev, A.V. (2010). Structure of social networks in a passerine bird:  
721 consequences for sexual selection and the evolution of mating strategies. *The American*  
722 *Naturalist*, 176(3), E80–E89.

723 Opsahl, T. (2009). *Structure and Evolution of Weighted Networks*. University of London  
724 (Queen Mary College). <http://toreopsahl.com/publications/thesis/>

725 Pinter-Wollman, N., Hobson, E.A., Smith, J.E., Edelman, A.J., Shizuka, D., de Silva, S.,  
726 Waters, J.S., Prager, S.D., Sasaki, T., & Wittemyer, G. (2013). The dynamics of animal  
727 social networks: analytical, conceptual, and theoretical advances. *Behavioral Ecology*,  
728 art047.

729 Pradel, R. (2005). Multievent: An Extension of Multistate Capture–Recapture Models to  
730 Uncertain States. *Biometrics*, 61(2), 442–447.  
731 <https://doi.org/https://doi.org/10.1111/j.1541-0420.2005.00318.x>

732 R Core Team. (2021). *R: A Language and Environment for Statistical Computing*.

733 Shizuka, D., & Johnson, A.E. (2020). How demographic processes shape animal social  
734 networks. *Behavioral Ecology*, 31(1), 1–11.

735 Silk, M.J. (2018). The next steps in the study of missing individuals in networks: a comment on  
736 Smith et al. (2017). *Social Networks*, 52. <https://doi.org/10.1016/j.socnet.2017.05.002>

737 Silk, M.J., Croft, D.P., Delahay, R.J., Hodgson, D.J., Weber, N., Boots, M., & McDonald, R.A.  
738 (2017). The application of statistical network models in disease research. *Methods in*  
739 *Ecology and Evolution*, 8(9). <https://doi.org/10.1111/2041-210X.12770>

740 Silk, M.J., & Fefferman, N.H. (2021). The role of social structure and dynamics in the

741 maintenance of endemic disease. *Behavioral Ecology and Sociobiology*, 75(8), 1–16.

742 Silk, M.J., Hodgson, D.J., Rozins, C., Croft, D.P., Delahay, R.J., Boots, M., & McDonald, R.A.

743 (2019). Integrating social behaviour, demography and disease dynamics in network

744 models: applications to disease management in declining wildlife populations.

745 *Philosophical Transactions of the Royal Society B: Biological Sciences*, 374(1781),

746 20180211. <https://doi.org/10.1098/rstb.2018.0211>

747 Silk, M.J., Jackson, A.L., Croft, D.P., Colhoun, K., & Bearhop, S. (2015). The consequences of

748 unidentifiable individuals for the analysis of an animal social network. *Animal Behaviour*,

749 104, 1–11. <https://doi.org/10.1016/j.anbehav.2015.03.005>

750 Silk, M.J., McDonald, R.A., Delahay, R.J., Padfield, D., & Hodgson, D.J. (2021). CMRnet: An r

751 package to derive networks of social interactions and movement from mark–recapture

752 data. *Methods in Ecology and Evolution*, 12(1), 70–75.

753 Smith, J.A., & Moody, J. (2013). Structural effects of network sampling coverage I: Nodes

754 missing at random. *Social Networks*, 35(4), 652–668.

755 Snijders, L., Blumstein, D.T., Stanley, C.R., & Franks, D.W. (2017). Animal social network

756 theory can help wildlife conservation. *Trends in Ecology & Evolution*, 32(8), 567–577.

757 Stanton, M.A., & Mann, J. (2012). Early social networks predict survival in wild bottlenose

758 dolphins. *PloS One*, 7(10), e47508.

759 Sunga, J., Webber, Q.M.R., & Broders, H.G. (2021). Influence of number of individuals and

760 observations per individual on a model of community structure. *Plos One*, 16(6),

761 e0252471.

762 Touzalin, F., Petit, E.J., Cam, E., Stagier, C., Teeling, E.C., & Puechmaille, S.J. (2022). Mark

763 loss can strongly bias demographic rates in multi-state models: a case study with

764 simulated and empirical datasets. *BioRxiv*.

765 Young, J.-G., Cantwell, G.T., & Newman, M.E.J. (2020). Bayesian inference of network  
766 structure from unreliable data. *Journal of Complex Networks*, 8(6), cnaa046.

767