

1 Heritabilities and genetic correlations of *Drosophila melanogaster*  
2 locomotory behaviour traits: a high-throughput phenotyping  
3 approach

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19 **Open Research**

20 Code needed to reproduce the analyses and visualisations can be found on GitHub under the following link:  
21 [Code and analyses](#). Data has been deposited on GitHub in the repository: [Data and code repository](#).

22 **Conflict of Interest**

23 We have no conflict of interest to declare.

## Abstract

25 Genetic variance forms the basis for evolutionary inferences as it describes the evolutionary potential  
26 of traits. The major limitation of quantitative genetic studies is achieving sufficient power and sample  
27 sizes to estimate heritabilities with sufficient precision. This issue is especially important in the case  
28 of traits that are inherently susceptible to stochastic, nonbiological variation. Behavioural traits have  
29 long been associated with this group, often yielding quantitative estimates of genetic parameters that  
30 are subject to broad estimation errors, thereby hampering the discovery of genetic variation underlying  
31 such characters. Here, we used a well-established panel of inbred genetic lines of the fruit fly, *Drosophila*  
32 *melanogaster*, to estimate relevant genetic parameters in a range of behavioural traits associated with  
33 mobility and exploration. Using a high-throughput phenotyping approach and automated scoring of  
34 large numbers of individual animals, we provide precise estimates of the quantitative genetic background  
35 behind some basic characters associated with animal behaviour. Fruit flies turn out to harbour significant  
36 genetic variance in traits directly associated with mobility and substantially lower heritabilities of traits  
37 describing the temporal variability of Y-maze movements. Mobility traits also appeared to be only  
38 moderately genetically correlated, except for movement distance vs. variability traits, where we estimated  
39 strongly negative genetic correlations. In general, our results demonstrate the existence of evolutionary  
40 potential in behavioural trait proxies measured by high-throughput methods, additionally hinting at  
41 the potential for sex-specific effects. They also emphasise the growing importance of high-throughput  
42 phenotyping in modern behavioural biology and ecology.

43 **Keywords**— fruit fly, *Drosophila melanogaster*, heritability, genetic correlation, phenomics

# 1 Introduction

The study of behaviour from a genetic and evolutionary perspective presents a rich but demanding field. Behavioural traits differ from most morphological or physiological traits in being more plastic, context-dependent, and often more strongly shaped by both internal states and external environments (Garamszegi and Herczeg, 2012; O’Dea et al., 2022). Despite this, behaviours are now widely recognised as heritable phenotypes subject to natural and sexual selection (Dingemans and Dochtermann, 2013; Dingemans et al., 2004; Johansson et al., 2024). Quantitative genetic parameters such as heritabilities, genetic variances, and genetic correlations are central to understanding the evolutionary potential of behaviour (Walsh and Lynch, 2018). Heritability estimates reveal the genetic contribution to behavioural variation, while genetic correlations inform us about possible trade-offs, constraints, or synergies in the joint evolution of traits. Together, such measures define the scope for behavioural responses to selection in natural and experimental populations.

However, robust estimates of quantitative genetic parameters for behaviour remain elusive. Behaviour is inherently dynamic: it depends on age, sex, circadian rhythms, previous experience, motivational state, and feeding history (Dingemans and Dochtermann, 2013; Thys et al., 2021). These influences reduce repeatability, introduce substantial measurement error, and complicate the separation of genetic from environmental variance components. Traditional behavioural assays exacerbate these problems: they are labour intensive, often rely on direct observation or manual scoring, and are typically limited in scale (Carter et al., 2013; Groothuis and Taborsky, 2015). As a result, studies often lack sufficient replication or sample size to detect the underlying genetic signal, leading to inflated estimates of environmental variance, underestimated heritabilities, and large statistical uncertainty (Kruuk and Hadfield, 2007; Postma, 2014). Moreover, many assays have been criticised for lacking ecological realism, raising questions about how measured behaviours map onto natural histories and fitness consequences (Groothuis and Taborsky, 2015). These challenges help explain why behavioural quantitative genetics has historically produced estimates that are less consistent and less precise than those for morphological or physiological traits (Postma, 2014).

Technological and methodological advances now offer powerful solutions to these technical difficulties. Automated and high-throughput phenotyping systems allow behavioural traits to be measured with unprecedented resolution in large cohorts of individuals (Cleal et al., 2021; Werkhoven et al., 2021). Automated tracking allows for continuous quantification of locomotion and exploration, detecting subtle variations in activity patterns, trajectories, and space use that manual scoring would miss (Werkhoven et al., 2021, 2019; Macartney et al., 2025). Combined with machine learning, such systems can classify behavioural motifs, track state transitions, and reduce observer bias, thereby increasing both precision and repeatability. These advances expand the scale of behavioural datasets, which in turn enables more robust quantitative genetic analyses, including multivariate animal models that partition phenotypic variance into genetic and environmental components and estimate cross-trait genetic correlations (Walsh and Lynch, 2018).

The availability of model organisms further accelerates progress in behavioural quantitative genetics. *Drosophila melanogaster* has long been central in genetics and neurobiology (Anholt and Mackay, 2004), with its short generation time, ease of rearing, and well-annotated genome making it ideal for large-scale quantitative genetic studies (MacKay et al., 2012; Mackay and Anholt, 2024). Many *Drosophila* behaviours, from circadian rhythms (Werkhoven et al., 2021) and learning (Stern et al., 2019) to aggression, courtship, and locomotor activity (Nandy et al., 2016; Kohlmeier et al., 2021; Macartney et al., 2025), show individual variation, fitness consequences, and measurable heritabilities (Harbison et al., 2013; Shorter et al., 2015; Rohde et al., 2017; Ivanov et al., 2015). The *Drosophila* Genetic Reference Panel (DGRP) provides a stable genome-sequenced collection of inbred lines that can be assayed repeatedly and in varying environments. This makes it possible to obtain precise estimates of heritability, dissect polygenic architectures, and quantify genotype-by-environment interactions. Crucially, multivariate analyses of DGRP data also allow the estimation of genetic correlations among behaviours, which can reveal constraints or trade-offs in behavioural evolution (Dingemans and Dochtermann, 2013; Walsh and Lynch, 2018).

In this study, our objective was to estimate the heritabilities and genetic correlations of a set of locomotory and exploratory traits in *D. melanogaster* using the DGRP panel. The target behavioural traits were measured using a high-throughput phenotyping system, which allowed us to obtain precise and repeatable measurements of individual behaviours in an automated way. We focused on several key traits, including overall locomotor activity, Y-maze exploration patterns, and startle response. The choice of traits was already validated in an earlier experimental study (Macartney et al., 2025). Using a large sample size and robust statistical analyses, we provide reliable estimates of genetic parameters and explore the genetic architecture underlying these behaviours.

## 2 Methods

Throughout the Methods section, we use the MeRIT convention (Method Reporting with Initials for Transparency), following Nakagawa et al. (2023).

### 2.1 Fly colony

We used 50 lines from the *Drosophila* Genetic Reference Panel (DGRP). DGRP is a landmark development in the genetic analysis of complex traits in *Drosophila*. It is a collection of more than 200 inbred lines derived from a natural population in Raleigh, North Carolina (MacKay et al., 2012). Each line in the DGRP has been fully sequenced and, because the lines are homozygous, repeated measures of individuals from the same genotype can be used to estimate genetic variance with high precision. The DGRP provides an open resource for genome-wide association studies (GWAS), quantitative trait loci (QTL) mapping, and the estimation of quantitative genetic parameters across a wide range of phenotypes, including behaviour.

Fifty inbred lines from the core set of the DGRP were ordered by SMD in the second half of 2021 from the Bloomington *Drosophila* Stock Centre (Indiana University, Bloomington, IN, USA; see Supplementary Table 1 for the complete list of lines). All lines were shipped to the Charles Perkins Centre (University of Sydney) and their stocks were maintained there under the supervision of ZH. Initially, all lines underwent two rounds of the full reproductive cycle under quarantine. Afterwards, adults from the F3 generation were used in the expansion stage (intensive parallel breeding of flies) to establish the required number of experimental flies in each line (done by ZH, CH, RA and YQL).

DGRP flies were used in a larger study, and the current analysis represents its subset. For the purposes of the large study, 50 lines were grouped in batches of 10 lines; batching was used to create a staggered design, where each group of 10 lines would enter actual assays every few days, allowing phenotyping to be done within the allocated time. Within each line, the flies used in the current study were supplied in groups of 20-30 individuals of each sex. Thus, each batch consisted of 10 lines and two sexes per line, with up to 30 individuals per vial (600 adult flies in total per batch). Subsequent batches entered the assays as freshly emerged adult flies (up to 1 week of age) on the following days of 2022: 23/02, 25/02, 10/03, 15/03, 1/04. Adult flies were placed in acrylic vials with 55-65 mL of standard *D. melanogaster* food medium (by ZH, CH, RA and YQL) and delivered to the School of Biological, Earth and Environmental Sciences, University of New South Wales, Sydney by SMD or commercial shipping service. There, they were transferred to fresh food vials put on plastic flippers (Drososflipper, London, UK) by SMD to facilitate the exchange of used food vials for new ones in later procedures not described here (irrelevant in this study as all fly assayed for the current analysis went through behavioural trials on the third day after arriving at the lab).

Each batch was split by SMD into two ‘sessions’, where half the batch was housed in an incubator where the 12h light cycle started at 9 am, and the other half of the batch was stored in an incubator where the 12h light cycle started at 1 pm. This was so that we could assess all flies from one batch on a single day (that is, in a morning session and an afternoon session) while standardising the circadian rhythm so that all sessions were conducted in the ‘morning’ activity period for the flies and avoided their low activity period of midday. General maintenance of the colony was done by SMD and CV.

### 2.2 Phenotyping

The assay methods are based on the detailed protocols reported elsewhere by Macartney et al. (2022, 2025). Behavioural assays were performed by SMD, ELM, SN, ML, SB, and PP. Briefly, we used high-throughput automated tracking units produced by Zantiks (Zantiks MWP unit, Cambridge, UK; firmware version 02/2021). These units are designed to track small animals, where each unit consists of i) an experimental chamber where the animals are placed (see below for details of the ‘plates’ in which individuals are loaded before being placed in the unit chamber), ii) a camera that tracks the animals, and iii) a computer controlling the units. All units were programmed to maintain 25°C in the chamber.

Using these units, we conducted three different behavioural assays: 1) locomotion tracking in which the overall

140 movement of individuals was measured within a 1 cm deep and round arena, 2) startle response to a light-off stimulus  
141 (also conducted in the same 1 cm deep and round arena as the locomotion assay), and 3) exploration in a Y-maze (see  
142 Macartney et al., (2025) for more details about the arenas). The locomotion tracking assay was run for 30 min (plus  
143 a 10-min habituation period), and the distance travelled by each individual was recorded in three intervals of 10 min  
144 each. The light-off startle response recorded the distance travelled in 1-second intervals following three consecutive  
145 15-ms light-off pulses, allowing us to measure habituation across the pulses. The Y-maze assay recorded ‘trigrams’  
146 of the turn direction flies took while travelling between the maze’s arms (e.g., RRR, LLL RLR, LRL, etc, where R =  
147 right, L = left) (see Macartney et al. (2025)). See Macartney et al. (2025) for a discussion of the biological relevance  
148 and robustness of the assays.

149 For each session (i.e., morning and afternoon sessions per batch), we assayed 300 flies (10 lines/batch, 2 sexes, 5  
150 flies/batch/sex in each of the three assays). In total, 3000 flies were scored; see Results for final sample size after  
151 accounting for deaths and lack of movement detected within arenas. At the beginning of each session, all flies were  
152 anaesthetised by briefly immersing the vials in a bucket of ice. All flies from each line and sex combination were  
153 then tipped into separate Petri dishes. Random flies from each line and sex combination were then selected from  
154 their respective Petri dishes and transferred by aspiration into a 48-well plate (only 45 wells were filled) or into a  
155 Y-maze plate (each session used 3 Y-maze plates, 15 mazes/plate). The flies were then given 10 min at 25°C to  
156 recover and transferred to four different assay units: one for the well-plate with 45 flies and three units for each of the  
157 three Y-maze plates. The unit with the well plate always recorded locomotion first, followed by the light-off startle  
158 response. In particular, each adult fly was used only once in an assay (taking advantage of the inbred property of the  
159 lines, we assumed that all flies from the given line represent the same genotype) and then discarded after a trial.

## 160 2.3 Phenotypic traits

161 The data returned by the units were used to calculate a number of phenotypic descriptors. For the locomotion assay,  
162 the measured trait was the total distance travelled in the arena in each of the 10-minute intervals; the distance was  
163 recorded in raw arbitrary units and then recalculated to millimetres. The startle response assay returned distances  
164 of movement after each light-off stimulus (see the previous subsection). In circular arenas, flies often follow the arena  
165 walls, and thus this behaviour is also referred to as arena circling. Finally, the Y-maze assay returned a series of  
166 left-right turn decisions, together with the timestamps of each turn.

167 To process the Y-maze data, we calculated several maze exploration metrics. First, a logged ratio of the number of  
168 right-turns to the number of left turns was calculated as a measure of movement handedness (the value of zero indicates  
169 lack of turn bias and no handedness; positive values measure increasing levels of right-handedness, negative values -  
170 left-handedness). Secondly, the turn-taking events were assembled into ‘trigrams’ (triplets of consecutive turns, e.g.,  
171 LLL, RLR, RRL, etc.). The counts of specific triplets were used to calculate the numbers of repetitions (occurrences  
172 of three consecutive turns in the same direction, i.e., RRR or LLL) and alternations (turns involve a there-and-back  
173 movement, i.e., RLR or LRL) recorded in a given assay session. Finally, to quantify the variability of maze behaviour,  
174 we calculated two indices. Switchiness recorded the consistency of the turning behaviour (e.g., ...LLLLRRRRR...  
175 represents low switchiness, ...LLRLRLLLLR... - moderate, ...LRLRLRLRLR... - highest possible). Switchiness was  
176 calculated as  $(N_{L \rightarrow R} + N_{R \rightarrow L}) / (2N_R N_L / N)$ , where  $N$  describes the number of all turns,  $N_x$  - the number of x-type  
177 turns and  $N_{x \rightarrow y}$  - the number of x-turns that are followed by y-turn. The second index, clumpiness, expressed the  
178 temporal uniformity of maze exploration. Low value describes individuals consistently exploring the maze without  
179 stopping, large values - flies that explored the maze in bursts of intense exploration separated by periods of no  
180 movement). The clumpiness was calculated using the median absolute deviation (MAD) of the intervals between  
181 turns  $T_i$  (measured in seconds):  $\text{median}[|T_i - \text{median}(T_i)|] / (1800/N)$ .

## 182 2.4 Statistical analysis

183 Data processing and analysis were performed by SMD. Each unit produced a data sheet per assay run. All data  
184 sheets were labelled with a unique ID, which means that we could match each data sheet to individuals coming from  
185 a specific line and sex. All data were cleaned and analysed in the R environment (version 4.2.2) using Visual Studio  
186 Code (Microsoft, USA). Genetic parameters were estimated using ASReml-R v. 4.2 (VSNi, UK) (Butler et al., 2023).  
187 Data processing and parsing was done using `dplyr` (Wickham et al., 2023). Visualisation was done using `ggplot2`

188 ([Wickham, 2016](#)) and `ggcorrplot` ([Kassambara, 2023](#)). Detailed versions and session data can be found in the paper’s  
 189 data supplement.

190 First, we ran univariate general linear mixed models (LMM) for each phenotypic character. The models included  
 191 sex as a fixed effect (females as the reference level). Random terms included: line ID  $L_{ij}$  (genetic effect, its variance  
 192 approximates the broad-sense genetic variance), batch ID  $B_i$ , and assay ID  $A_{ijk}$  (the latter two accounting for  
 193 technical nonindependence introduced through assay clustering). The general form of any univariate model was:

$$y_{ijkl} = \mu + s_{male} + B_i + L_j + A_k + e_{ijkl}, \quad (1)$$

$$L_j \sim \mathcal{N}(0, \sigma_G^2), \quad (2)$$

$$B_i \sim \mathcal{N}(0, \sigma_B^2), \quad (3)$$

$$A_k \sim \mathcal{N}(0, \sigma_A^2), \quad (4)$$

$$e_{ijkl} \sim \mathcal{N}(0, \sigma_e^2). \quad (5)$$

194 In the above models,  $\mu$  is the model intercept,  $s_{male}$  is the fixed effect of sex. All random effects and the residual  
 195 error  $e_{ijkl}$  are assumed to be sampled from relevant normal distributions with zero mean and variance interpretable  
 196 as (conditional) variability estimated between replication units at each level.

197 The random effect of line ID was used to calculate broad-sense heritability (as a fraction of the variance between  
 198 lines in overall variability). The *sensu lato* nature of our heritability estimates is due to the fact that the variation  
 199 between inbred lines cannot be decomposed to separate the dominance variance ([Walsh and Lynch, 2018](#)).

200 Cross-trait correlations were estimated using a multivariate model that allowed for heterogeneous genetic (line) and  
 201 residual variance components between the traits, explicitly modelling the cross-trait genetic covariance matrix  $\mathbf{G}$  and  
 202 residual covariance matrix  $\mathbf{R}$ :

$$y_{ijkl}^{(n)} = \mu^{(n)} + s_{male}^{(n)} + B_i^{(n)} + L_j^{(n)} + A_k^{(n)} + e_{ijkl}^{(n)}, \quad (6)$$

$$\mathbf{L} \sim \mathcal{MVN}(\mathbf{0}, \Sigma_{\mathbf{G}}), \quad \Sigma_{\mathbf{G}} = \mathbf{G}_{n \times n} \otimes \mathbf{I}, \quad (7)$$

$$\mathbf{e} \sim \mathcal{MVN}(\mathbf{0}, \Sigma_{\mathbf{e}}), \quad \Sigma_{\mathbf{e}} = \mathbf{R}_{n \times n} \otimes \mathbf{I}. \quad (8)$$

203 The  $(n)$  indicator indexes the  $n$  phenotypic traits in a model. Both  $\Sigma$  matrices are built using the identity matrices  
 204 of appropriate dimension  $\mathbf{I}$  - i.e., we assume that, e.g., the  $j$ -th line effects for one trait may depend only on the  
 205 effects of this line in other traits.

206 Finally, sex-specific estimates for the analysed traits were derived from models that assumed sex-specific heteroscedas-  
 207 ticity of genetic and residual variances, with the cross-sex genetic correlations estimated rather than fixed at unity:

$$\mathbf{R}_{\times \text{sex}} = \sigma_{e_m}^2 \oplus \sigma_{e_f}^2, \quad (9)$$

208 with  $\sigma_{e_m}^2$  and  $\sigma_{e_f}^2$  being the sex-specific residual variances. The direct product indicates that the cross-sex residual  
209 correlations are not estimated (they are not identifiable).

210 Hypothesis tests for variance components are based on a series of likelihood-ratio tests comparing target models  
211 with their simpler alternatives (i.e., for variance components, models without the focal random term, for genetic  
212 correlations, models with correlations fixed at values expected under  $H_0$ ) (Lynch and Walsh, 1998). The logged  
213 ratios of model likelihoods were assumed to be  $\chi^2$ -distributed with an appropriate number of degrees-of-freedom (Self  
214 and Liang, 1987).

## 215 3 Results

### 216 3.1 Locomotion and Y-maze behaviour

#### 217 3.1.1 Genetic variances

218 All traits studied exhibited low to moderate levels of genetic variance (expressed as variance between genetically  
219 homogeneous lines of the DGRP panel; Fig. 1). Locomotory activity appeared to be the most heritable trait in  
220 the entire set. Overall locomotion activity (measured as arena circling distance covered during the test) harboured  
221 significant genetic variance (likelihood-ratio test (LRT):  $\chi_{df=1}^2 = 203.532$ ,  $p < 0.001$ ; see also Tab. 1). The round  
222 effects (one of the three repetitions recorded for each individual) were marginally important (Tab. 1), highlighting  
223 that the repeatability of the trait was mainly due to genetic effects. Batch and Assay ID effects explained fractions  
224 of variance not distinguishable from zero (Tab. 1).

225 Excluding individuals that did not move throughout the test (but remained alive - see Methods) did not remove  
226 genetic variance; such data still resulted in heritability of  $h^2 = 0.185$  significantly different from zero (see Tab. 1).  
227 Males tended to be more active than females (estimate from data including zeroes  $\pm$ SE:  $b_{M-F} = 0.163 \pm 0.101$ ; data  
228 without zeroes:  $b_{M-F} = 0.226 \pm 0.101$ ) and in the latter case the difference was significant ( $\chi_{df=1}^2 = 5.034$ ,  $p = 0.025$ ).

229 The sexes did not appear to have heterogeneous genetic variances in locomotor activity (males:  $h_m^2 = 0.230 \pm 0.052$ ,  
230 females:  $h_f^2 = 0.228 \pm 0.051$ ), but appeared independent in terms of the underlying genetic effects (low genetic  
231 correlation between sexes:  $r_{g,mf} = 0.271 \pm 0.171$ ; 1-degree-of-freedom LRT for  $H_0 : r_g = 1$ ,  $\chi^2 = 88.278$ ,  $p < 0.001$ ;  
232 for  $H_0 : r_g = 0$ ,  $\chi^2 = 2.288$ ,  $p = 0.13$ ).

233 Y-maze exploration was measured from several perspectives. In all Y-maze behaviour descriptors (except for hand-  
234 edness;  $\chi_{df=1}^2 = 0.000$ ) there was significant genetic variance resulting in heritabilities ranging from low to moderate  
235 (see Tab. 2; LRT p-values: no. of repetitions  $p < 0.001$ , no. of alternations  $p < 0.001$ , switchiness  $p < 0.001$ , MAD  
236 clumpiness  $p < 0.001$ ). Batch effects did not matter (in all cases estimates of variance were not different from zero;  
237 Tab. 2). Assay ID appeared more important in some measured traits (number of alternations and switchiness) indi-  
238 cating that their underlying determinants (e.g., the identity of the person operating the unit, time of day) may have  
239 some impact on measured traits; nonetheless, in the most heritable traits, all other random effects were relatively  
240 negligible.

241 In all traits measured with Y-maze, the sexes did not differ significantly from each other. Nonetheless, in traits that  
242 should be associated with overall locomotor activity, males tended to score higher, in line with the pattern seen in  
243 arena circling behaviour (no. of repetitions:  $b_{M-F} = 0.195 \pm 0.167$ , no. of alternations:  $b_{M-F} = 0.103 \pm 0.124$ ).  
244 The only exception was MAD clumpiness, where males tended to explore the mazes in a “clumpier” way (i.e., more  
245 likely in bursts of activity separated by periods of inactivity, rather than continuously walking between the arms;  
246  $b_{M-F} = -0.340 \pm 0.135$ ).

247 The patterns of sex-specific genetic variance varied between traits. Both the number of repetitions and alternations  
248 were genetically integrated between the sexes (strong genetic correlation, in both cases fixed at  $r_{g,mf} = 1$  by the  
249 model). Interestingly, while the number of repetitions did not exhibit sex-specific genetic variances ( $h_m^2 = 0.142 \pm 0.088$   
250 vs.  $h_f^2 = 0.147 \pm 0.087$ ), heritability of no. of alternations was higher in females ( $h_f^2 = 0.144 \pm 0.088$ ) than in males  
251 (effectively  $h_m^2 = 0$ ). Handedness exhibited no detectable genetic variance (Tab. 2), and this pattern was not different

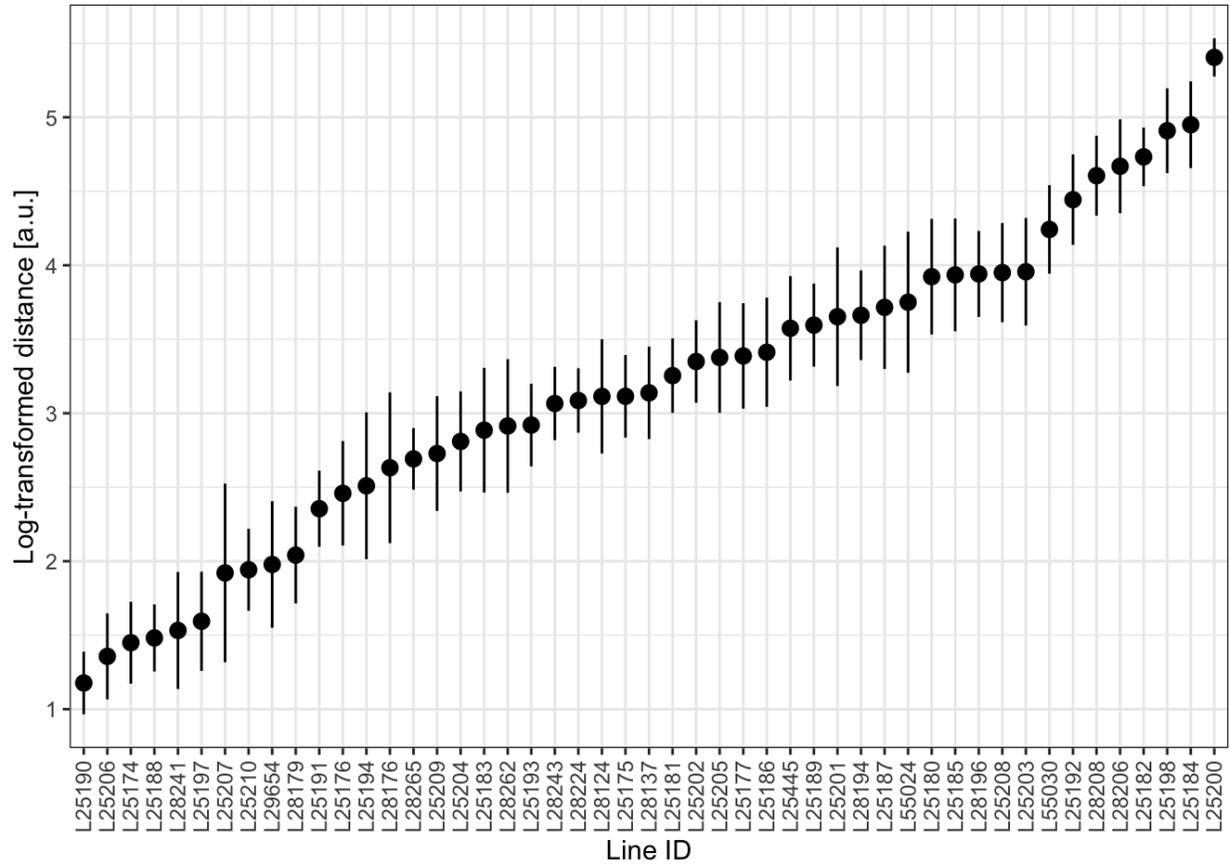


Figure 1: The distribution of average line-specific locomotory scores among the 50 included DGRP lines (transformed with a natural logarithm). The lines are sorted according to the mean trait value. Error bars indicate within-line standard errors; distance was measured in arbitrary units proportional to actual metric distances.

Table 1: Variance components of locomotory activity (arena circling). The "Fraction" column provides proportions of total variance explained by each random effect with its SE (Fraction for line effect is equal to heritability:  $h^2 = \sigma_G^2 / \sigma_{total}^2$ )

Random effect	Variance	SE	Fraction
<i>Locomotion data including zeroes</i>			
Round	0.005	0.013	0.001 ± 0.003
Batch	0.239	0.314	0.053 ± 0.066
Assay ID	0.165	0.121	0.036 ± 0.026
Line	1.011	0.242	0.223 ± 0.046
Residual	3.110	0.125	0.686 ± 0.059
<i>Locomotion data excluding zeroes</i>			
Round	0.006	0.013	0.002 ± 0.003
Batch	0.285	0.310	0.074 ± 0.076
Assay ID	0.128	0.098	0.033 ± 0.025
Line	0.710	0.177	0.185 ± 0.042
Residual	2.701	0.115	0.705 ± 0.065

Table 2: Variance components of Y-maze behaviour. The last column provides proportions of total variance explained by each random effect with its SE. For definitions of variables - see Materials & methods. Estimates without SE (—) have been fixed by the model at the boundary of their parameter space

Random effect	Variance	SE	Fraction
<i>Number of repetitions (log scale)</i>			
Batch	0.360	0.356	0.137 ± 0.119
Assay ID	0.453	0.207	0.172 ± 0.078
Line	0.381	0.173	0.144 ± 0.067
Residual	1.443	0.139	0.547 ± 0.107
<i>Number of alternations (log scale)</i>			
Batch	0.028	0.075	0.022 ± 0.059
Assay ID	0.213	0.096	0.171 ± 0.072
Line	0.198	0.090	0.159 ± 0.07
Residual	0.809	0.078	0.649 ± 0.084
<i>Handedness (log of <math>N_R/N_L</math>)</i>			
Batch	0.000	—	0.000
Assay ID	0.015	0.021	0.026 ± 0.038
Line	0.000	—	0.000
Residual	0.546	0.050	0.974 ± 0.038
<i>Switchiness</i>			
Batch	0.018	0.020	0.125 ± 0.128
Assay ID	0.054	0.017	0.38 ± 0.101
Line	0.013	0.006	0.091 ± 0.045
Residual	0.057	0.006	0.404 ± 0.078
<i>MAD Clumpiness (log scale)</i>			
Batch	0.049	0.102	0.03 ± 0.061
Assay ID	0.139	0.099	0.085 ± 0.06
Line	0.430	0.164	0.262 ± 0.09
Residual	1.022	0.100	0.623 ± 0.09

252 in sex-specific parameters (not shown). MAD clumpiness and switchiness showed no strong sex-specificity on variances  
253 (MAD clumpiness:  $h_m^2 = 0.174 \pm 0.098$  vs.  $h_f^2 = 0.326 \pm 0.105$ ; switchiness:  $h_m^2 = 0.224 \pm 0.098$  vs.  $h_f^2 = 0.102 \pm 0.079$ ).  
254 Inter-sexual  $r_g$  was close to unity in MAD clumpiness, but relatively low in switchiness, indicating weaker cross-sex  
255 genetic integration ( $r_{g,mf} = 0.139 \pm 0.411$ ).

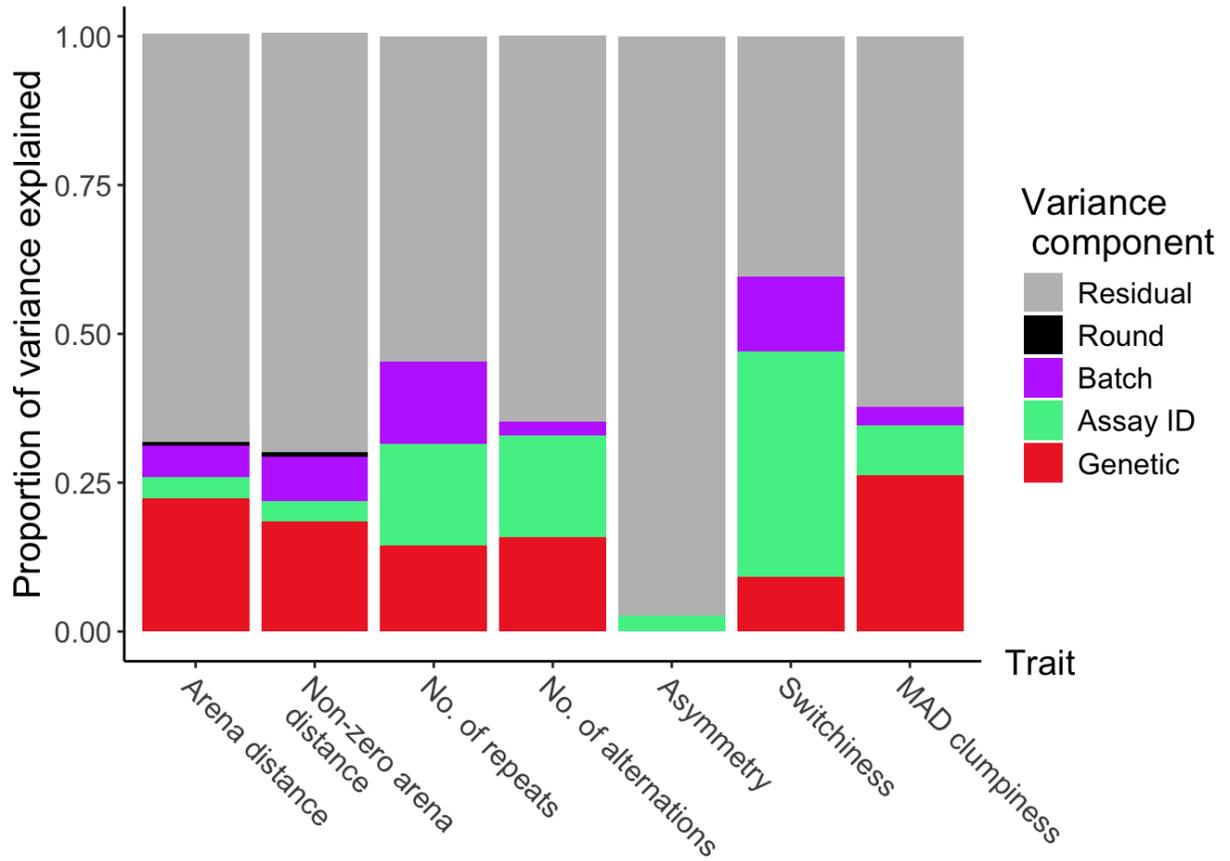


Figure 2: The fraction of overall trait variance explained by the random effects (colour-coded) of the inbred line (genetic), batch (grouping several lines) and assay (grouping multiple individuals in one assay run). The round effect was present only in locomotory data (within each assay, the measurements were grouped into three 10-minute rounds). The genetic (red) effect is estimated as the between-lines variance.

256 **3.1.2 Genetic correlations**

257 Instead of using raw numbers of repeats and alternations (thus avoiding unidentifiable correlations likely close to  
 258 unity), we used a single measure of Y-maze behaviour (proportion of alternations in all trigrams) that can be  
 259 regarded as a measure of working memory. Most traits measured in the locomotion and Y-maze trials were tightly  
 260 genetically correlated - indicating their common underlying genetic underpinnings. In most cases, the cross-trait  
 261 genetic correlations were fixed at the boundary of parameter space ( $|r| = 1$ ; Fig. 3). In a few cases, we observed  
 262 less strict genetic constraints (Fig. 3): relative number of alternations was weakly correlated with total arena circling  
 263 distance, Y-maze clumpiness and Y-maze movement asymmetry (distance vs. fraction of alternations:  $r_g = 0.216 \pm$   
 264  $0.305$ ; clumpiness vs. fraction of alternations:  $r_g = -0.611 \pm 0.176$ ; asymmetry:  $r_g = 0.198 \pm 0.602$ ). More active  
 265 flies also had lower genetic scores for asymmetry in the Y-maze and Y-maze clumpiness ( $r_g = -0.506 \pm 0.512$   
 266 and  $r_g = -0.611 \pm 0.176$ , respectively), but these negative correlations were estimated with errors, making them  
 267 statistically non-significant. In general, the model clearly lacked the power to resolve all genetic correlations.

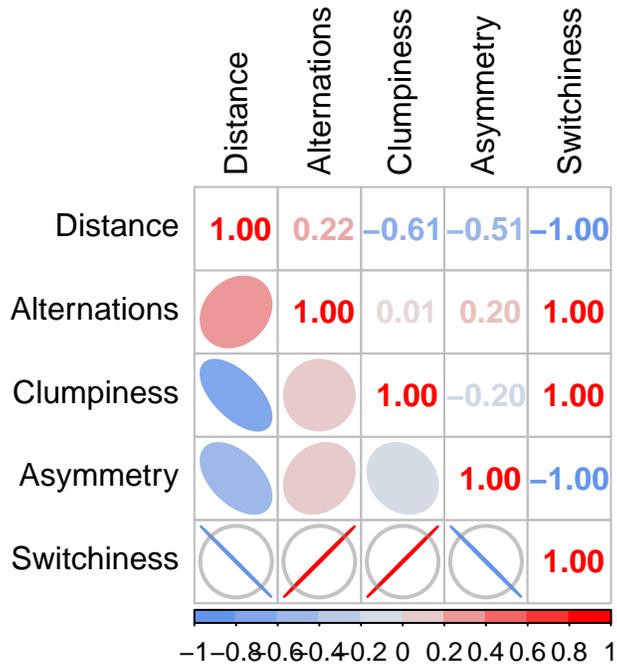


Figure 3: Genetic correlations between traits. Ellipses and their colouration depict the strength of correlation. Correlations close to 1 (or -1) take the form of diagonal lines. Correlations marked with open circles were estimated at the boundary of parameter space (fixed at 1 or -1). Diagonal elements are blank (variances, i.e.,  $r=1$ )

268 **3.2 Startle response**

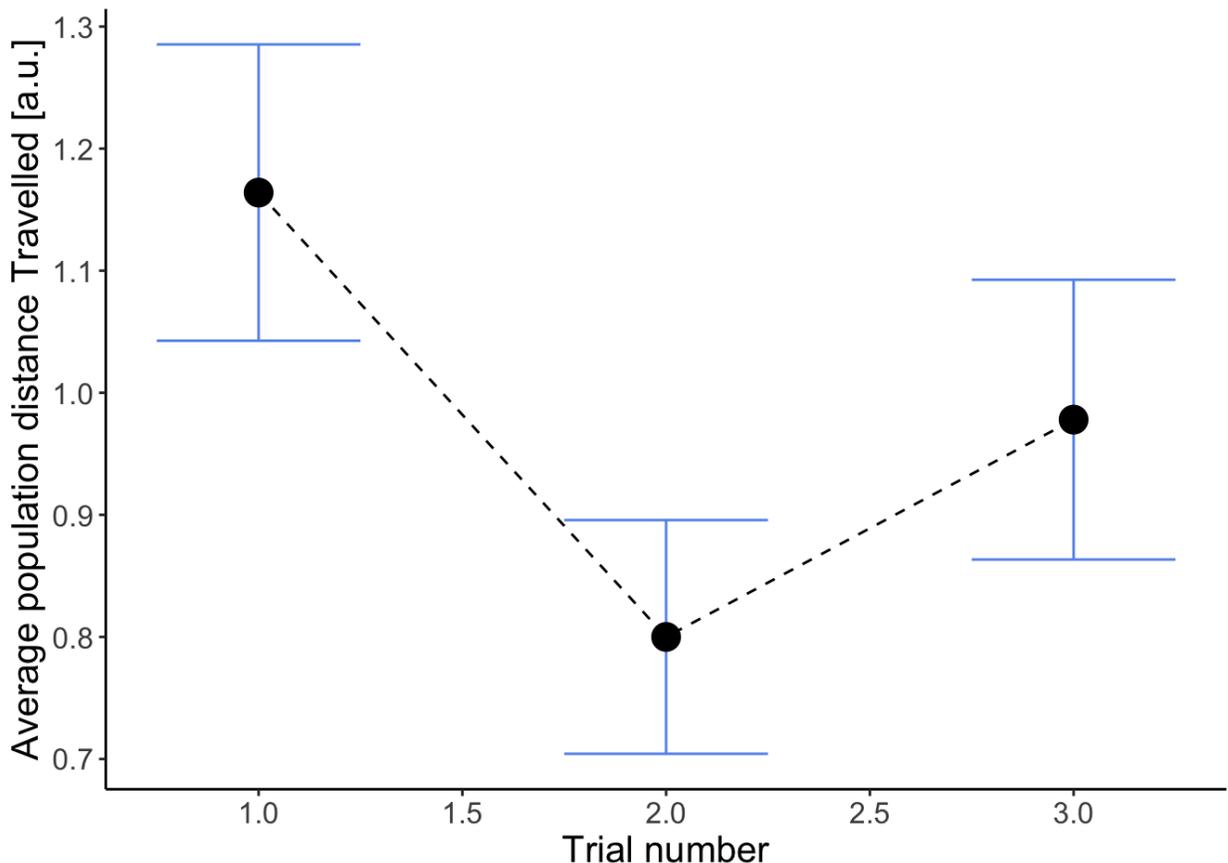


Figure 4: Startle response distance travelled by flies in the Y-maze. Each point represents the mean distance travelled by flies in a given trial. Error bars represent the standard error of the mean.

269 There was a marked average response to the startle stimulus (light-off) between the first and second exposure, and  
 270 no significant difference between the second and third (Fig. 4). Inbred lines differed clearly not only in overall  
 271 activity levels (Fig. S1; see also Fig. 1), but also in the slope and direction of the startle response. However, the  
 272 genetic variation in startle response (line-linked variation in habituation slopes) was not statistically significant (model  
 273 assuming continuous startle number variable:  $\chi^2 = 1.800$ ,  $p = 0.09$ , variance estimate:  $0.046 \pm 0.037$ ; model using  
 274 the contrast convention:  $\chi^2 = 0$ ,  $p = 1.0$ , boundary variance estimate ( $var = 0$ )). Estimation of non-zero genetic  
 275 variation in the startle-response slopes was not possible even after constraining the random regression part to the  
 276 first contrast (time point 1 vs. 2) only (i.e., interval where the response is more likely to be noted due to no prior  
 277 habituation).

278 **3.3 Sex-specific genetic effects**

279 Sex specificity was absent in most genetic parameter estimates, except for the number of Y-maze alternations - in  
 280 this trait, males exhibited no detectable genetic variance, contrary to females ( $h^2 = 0.144$ ; see Tab. 3). Interestingly,  
 281 despite no sex-specific heritabilities, overall locomotor activity was not correlated strongly between the sexes ( $r_{mf} =$   
 282  $0.271$ ), indicating loose cross-sex genetic constraint and crossing intersexual reaction norms.

Table 3: Sex-specific genetic variances and heritabilities (in parentheses) of the analysed traits with relevant SE, together with their cross-sex genetic correlations ( $\pm$ SE). All estimates come from the `Line` ID random effect. Lack of SE indicates estimates at the parameter space boundary (variance close to 0 or correlation close to  $\pm 1$ )

Trait	Female	Male	Correlation
Locomotor activity (log)	0.483 $\pm$ 0.126 (0.228 $\pm$ 0.051)	0.487 $\pm$ 0.126 (0.23 $\pm$ 0.052)	0.271 $\pm$ 0.171
No. of repetitions (log)	0.27 $\pm$ 0.176 (0.147 $\pm$ 0.087)	0.259 $\pm$ 0.172 (0.142 $\pm$ 0.088)	0.999
No. of alternations (log)	0.223 $\pm$ 0.154 (0.144 $\pm$ 0.088)	0.000 0.000	0.999
Handedness (log)	0.000 0.000	0.000 0.000	0.997
Switchiness	0.246 $\pm$ 0.204 (0.102 $\pm$ 0.079)	0.623 $\pm$ 0.321 (0.224 $\pm$ 0.098)	0.139 $\pm$ 0.411
MAD clumpiness (log)	0.604 $\pm$ 0.263 (0.326 $\pm$ 0.105)	0.263 $\pm$ 0.166 (0.174 $\pm$ 0.098)	0.999

## 4 Discussion

Behavioural traits, though more variable and context-dependent than morphological or physiological characters, are crucial phenotypes with direct fitness consequences. Their genetic architecture has remained relatively underexplored compared to other trait classes, due to persistent challenges in consistent phenotyping and the inherently dynamic nature of behaviour (Dingemans and Dochtermann, 2013). In this study, we examined the genetic basis of behavioural variation in *Drosophila melanogaster*, focusing on locomotion, space exploration (via Y-maze-based proxies), and habituation to a light-induced startle response. Our findings emphasise three key areas: heritability of behavioural traits and their evolutionary potential, sex-specific components of genetic variance, and the opportunities and limitations of laboratory-based behavioural assays.

### 4.1 Heritability and Evolutionary Potential

The consistent detection of genetic variance across multiple behaviours indicates measurable evolutionary potential (Lynch and Walsh, 1998). The strongest signals were observed in locomotor activity and exploratory Y-maze behaviours, with heritabilities of  $\sim 20\%$  and  $\sim 15\%$ , respectively. These values are consistent with previous estimates of genetic effects in arthropods, including *D. melanogaster* (15–44%: Jordan et al., 2006, 2007; Noer et al., 2024), *Cydia pomonella* (37%: Schumacher et al., 1997), *Tribolium castaneum* (10–40%: Matsumura et al., 2022; Matsumura and Miyatake, 2015), and *Tetranychus urticae* (24%: Bitume et al., 2011). Importantly, these signals persisted after excluding immobile individuals, suggesting that heritability is not an artefact of extreme phenotypes but reflects underlying genetic architecture. Locomotor traits in particular are known to respond quickly to directional selection (Matsumura et al., 2022), and may be especially important in contexts such as behavioural thermoregulation under rapid climate change (Pottier et al., 2025).

Exploratory behaviour also showed robust line effects (indicative of significant genetic variance). Traits derived from alternation and movement irregularity in the Y-maze retained genetic components despite potential modulation by motivation or physiological state. The fact that traits describing behavioural variability (e.g. switchiness, median absolute deviation (MAD) clumpiness) were heritable highlights the capacity of high-throughput phenotyping to dissect not only mean-level traits but also variance traits (Ayroles et al., 2015). These results add to a growing recognition that even complex cognitive or sensorimotor behaviours, often dismissed as too plastic, can be rigorously quantified and genetically analysed (Werkhoven et al., 2019, 2021; Sauce et al., 2018). Together, they support the view that behaviour harbours substantial, evolvable genetic variation.

### 4.2 Sex Differences in Behavioural Genetics

A central question in behavioural genetics is the degree to which males and females share the underlying genetic architecture (Poissant et al., 2008, 2010). Our results suggest a mixed pattern. For many behaviours, heritabilities

314 were similar across sexes, implying shared architecture. Yet for certain exploratory strategies, we observed sex  
315 differences in the magnitude of genetic variance or cross-sex correlations, consistent with findings in other insects  
316 (Yarwood, 2022; Fuchikawa and Okada, 2013). In line with broader comparative evidence (Wyman and Rowe, 2014),  
317 most of the asymmetries referred to male-biased genetic variance in locomotor traits. However, such biases can be  
318 assay- or population-specific: Videlier et al. (2021) reported a 40% male bias in locomotor variance in *D. melanogaster*,  
319 while in our data, overall activity variance was homogeneous, with sex differences emerging mainly in maze-derived  
320 traits.

321 Low cross-sex genetic correlations in some traits indicate reduced integration and the potential for sex-specific evo-  
322 lutionary trajectories. Such patterns open the door to sexual conflict over shared genetic variation and echo results  
323 on sociality and mating behaviours in flies (Scott et al., 2018). These results resonate with research on behavioural  
324 syndromes and personality, where genetic correlations within individuals may manifest differently between sexes  
325 (Matsumura, 2024). They highlight the role of sexual selection, dimorphism, and potentially sex-specific selection  
326 pressures in shaping behavioural evolution in *Drosophila*.

### 327 4.3 Genetic Correlations and Behavioural Integration

328 Genetic correlations can constrain or facilitate multivariate evolution by linking trait responses. We found a strong  
329 integration between certain behaviours, such as activity level and exploration of the maze (Fig. 3), suggesting a  
330 correlated evolution and a potential structure of behavioural syndromes (Garamszegi and Herczeg, 2012). Similar  
331 coupling has been documented in locomotor traits (Hamlin et al., 2025), although it remains uncertain whether such  
332 patterns extend to natural settings (Matsumura, 2024). At the same time, other traits appeared largely independent,  
333 despite phenotypic similarity, pointing to distinct underlying circuits or motivational bases.

334 Some genetic correlations were difficult to resolve due to limited statistical power—a common issue in multivariate  
335 models when traits differ in measurement reliability or sample distribution. Nonetheless, the evidence for both  
336 integrated and independent axes of variation demonstrates the value of automated phenotyping and multivariate  
337 approaches for uncovering behavioural architecture.

### 338 4.4 Startle Responses and the Challenge of Detecting Transient Behavioural 339 Variation

340 The startle assays provided clear phenotypic responses but limited genetic variance, particularly in habituation. Sim-  
341 ilar patterns were previously observed in outbred *Drosophila* populations (Macartney et al., 2022). The transient  
342 state-dependent nature of startle responses may explain their low repeatability, which complicates genetic analy-  
343 sis. Alternatively, habituation plasticity may be shaped more by developmental or epigenetic mechanisms than by  
344 additive genetic effects (Huang et al., 2020). These challenges illustrate why detecting genetic variance in reactive be-  
345 haviours requires larger datasets, alternative assay designs, or integrative approaches combining genetics with neural  
346 or physiological measurements.

### 347 4.5 Laboratory Environments and the Ecological Validity of Behavioural Assays

348 A recurring concern in behavioural genetics is the ecological validity of laboratory assays. Our paradigms—locomotor  
349 arenas, Y-mazes, and startle assays—were designed for consistency and repeatability and have been validated in  
350 previous work (Fenckova et al., 2019; Cleal et al., 2021; Werkhoven et al., 2021). However, they inevitably simplify  
351 natural challenges, such as predation, foraging, or social interactions. Y-maze exploration may capture novelty  
352 responses but not dispersal ecology, while arena locomotion omits environmental complexity. However, such assays  
353 are generally considered to quantify modular behavioural components that can evolve semi-independently (Berman  
354 et al., 2014; Seidenbecher et al., 2020; Vinicius, 2010; Eberhard, 2018). Our detection of robust heritabilities and  
355 correlations suggests that these modules represent genuine axes of genetic variation. Extending the analyses to  
356 semi-natural or enriched conditions would further clarify their ecological relevance (e.g. Macartney et al., 2025).

357 The trade-off between experimental tractability and ecological realism is well recognised. Automated and high-  
358 throughput assays offer precision and power, yet complementary field-based or semi-natural approaches will be cru-  
359 cial for connecting laboratory phenotypes with natural behaviour. Recent advances in computational ethology (see  
360 [Anderson and Perona, 2014](#); [Branson et al., 2009](#)) provide exciting opportunities to bridge this gap, allowing be-  
361 havioural assays to scale while retaining ecological relevance. Integrating such approaches will strengthen the link  
362 between behavioural genetics, ecology, and evolution (cf. [Kruitwagen et al., 2018](#)).

## 363 4.6 Concluding Remarks

364 Using DGRP in combination with automated phenotyping, we obtained robust estimates of quantitative genetic  
365 parameters and revealed patterns of sex specificity and genetic integration. Our study demonstrates that behavioural  
366 traits in *Drosophila melanogaster*, particularly those related to mobility and exploration, harbour substantial heritable  
367 variation, including general and sex-specific components. These findings confirm that behaviour is not purely plastic  
368 but an evolvable phenotype.

369 Beyond *Drosophila*, these results inform broader questions in behavioural ecology, evolutionary biology, and neuro-  
370 genetics. As behavioural genetics incorporates genome editing, machine learning, and large-scale phenotyping, the  
371 ability to connect genes, circuits, and behaviour within an evolutionary framework becomes increasingly attainable.  
372 Our findings contribute to this larger effort, highlighting both the promise and the complexity of dissecting the genetic  
373 architecture of behaviour.

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634 5 Supplementary materials

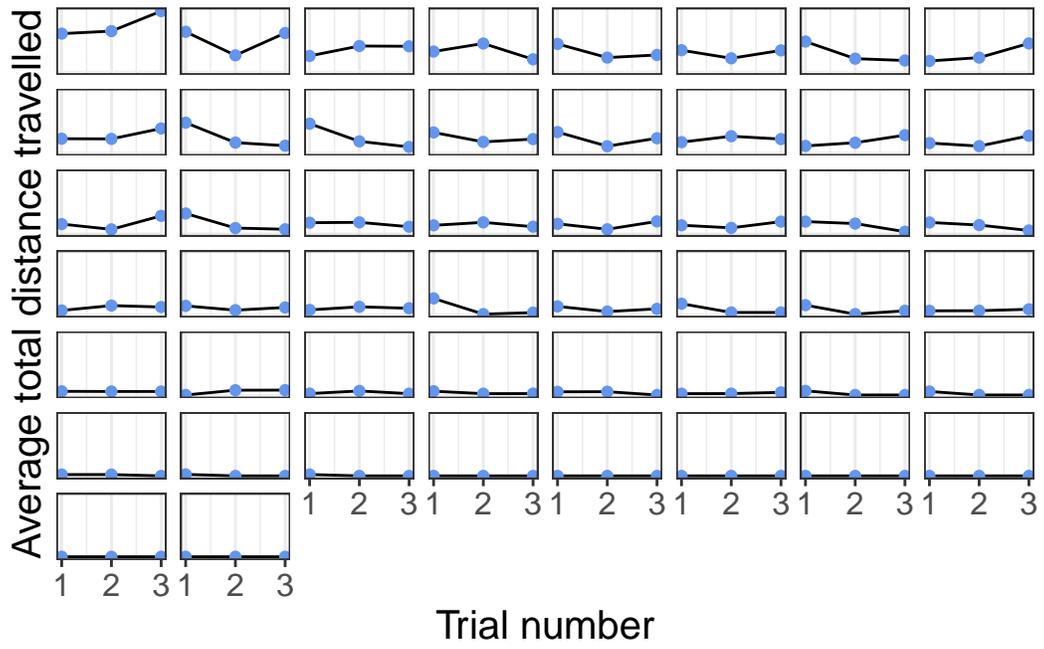


Figure S1: Average startle response distance travelled by flies in each line, sorted by overall level of locomotory activity.